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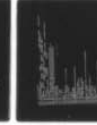
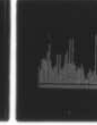
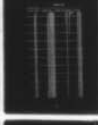
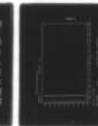
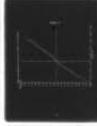
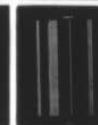
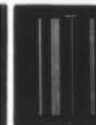
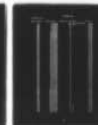
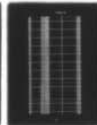
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NAVAL POSTGRADUATE SCHOOL

Monterey, California



THESIS

APL HISTOGRAM, DENSITY ESTIMATION
AND
PROBABILITY PLOTTING ROUTINES

by

Dennis Roy Hutchinson

December 1976

Thesis Advisor:

P. A. W. Lewis

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APL HISTOGRAM, DENSITY ESTIMATION
AND
PROBABILITY PLOTTING ROUTINES

by

Dennis Roy Hutchinson
Captain, United States Army
B.S., United States Military Academy, 1969

Submitted in partial fulfillment of the
requirements for the degree of

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December 1976

Author

Dennis R. Hutchinson

Approved by:

Peter A. W. Lewis
Thesis Advisor

Richard W. Butterworth
Second Reader

J. K. Hartman (Acting)
Chairman, Department of Operations Research

David A. Schradz
Dean of Information and Policy Sciences

ABSTRACT

This paper introduces several data analysis routines that were designed for interactive use with APL (A Programming Language) and placed in the APL user library at the Naval Postgraduate School. Specifically, histograms, density estimation and probability plotting routines are both explained in detail and demonstrated with actual data. In addition, applications and limitations on each of the routines are explored. And, the combined routines give the general user an extensive tool to analyze either discrete or continuous data.

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Professor Richard W. Butterworth contributed greatly with his extensive knowledge of all aspects of APL. His cooperation and willingness to assist resulted in the efficient and extensive use of all current features available for APL at the Naval Postgraduate School.

I. INTRODUCTION

The Naval Postgraduate School acquired APL (A Programming Language) from IBM in 1974. Since that time more and more students and faculty have become familiar with the extensive and efficient capabilities of APL and have been putting these features to good use. With the acquisition of APL came several extensive library routines that are both well documented and varied in scope. However, on close examination of these library routines it was found that statistics and data analysis were areas where some additions would be particularly useful.

Because of the efficiency and ease of APL in manipulating vectors, matrices and arrays, it is ideal for use in the area of data analysis. After a complete and thorough screening of the existing APL library routines pertaining to data analysis, it was found that by adding six additional data analysis routines to the present library, the Naval Postgraduate School could enhance its present APL capability and provide the student and general user with a more varied and flexible tool for analyzing data.

To this end the purpose of this thesis will be (1) to completely describe the six data analysis routines added to the APL library, (2) to explain the features and capabilities of each of the routines and (3) to demonstrate the use of each of the routines with "real world data".

The data to be used in this paper has come from two different sources. The first source of data was from tests performed jointly by IBM Germany and the German Public Telephone Network on errors in transmission of binary data on telephone lines (Lewis & Cox, 1966). From this source two sets of data are used and each data set contains the times between errors in binary bits transmitted over telephone lines. The first data set contains 672 elements (times-between-errors: actually number of bits between errors) and will hereby be referred to as "telephone data 1". The second data set contains 736 elements and will be referred to as "telephone data 2". The second source of data was obtained from percent overrun or underrun on selected military contracts during the year 1950 (Dixon, 1973). This data set contains 22 elements and will be referred to as "cost overrun data".

II. HISTOGRAM ROUTINE

A. DESCRIPTION

The first routine to be presented is the histogram routine which is used for estimating from given data the probability density function $f(x)$ of a continuous random variable. The current APL library has several small histogram routines that are general in nature but lack the overall detail necessary for good data analysis. For this reason HIST (histogram routine) was created. HIST represents the adaptation and modification of the fortran library version of HISTG/F, which was developed at N.P.S. by D. R. Robinson under the guidance of Professor P.A.W. Lewis. By modifying and adapting HISTG/F to APL the power and efficiency of the APL language could be put to full use.

A complete description of how HIST operates is contained in the variable HISTHOW. If the users APL workspace is properly loaded (see section IX.B. for workspace loading procedures) all that is necessary is to type HISTHOW. The user then receives the following printed response on the terminal:

HISTHOW

SYNTAX HIST

HIST ALLOWS YOU TO INTERACTIVELY OBTAIN A HISTOGRAM OF YOUR DATA ALONG WITH A SET OF BASIC DESCRIPTIVE STATISTICS. IN ADDITION, HIST HAS THE FOLLOWING CAPABILITIES WHICH ALLOW YOU:

- (1) THE OPTION OF A TITLE FOR YOUR HISTOGRAM
- (2) THE OPTION OF DISPLAYING A SMOOTHED EMPIRICAL DENSITY FUNCTION OVER THE HISTOGRAM
- (3) THE OPTION OF SCALING AND SELECTING THE NUMBER OF CELLS FOR YOUR HISTOGRAM
- (4) THE OPTION OF SELECTING AN INTERVAL AND PERFORMING A HISTOGRAM ON ALL THE DATA POINTS OR CONDITIONALLY SELECTING AN INTERVAL IN THE RANGE OF THE DATA.
- (5) THE OPTION OF HAVING YOUR OUTPUT APPEAR ON THE OFFLINE PRINTER OR ON YOUR TERMINAL

WHEN YOU TYPE HIST YOU WILL BE ASKED TO DO THE FOLLOWING:

- (1) ENTER YOUR DATA IN VECTOR FORM - YOU CAN TYPE YOUR DATA IN SINGLY OR YOU CAN TYPE THE NAME OF A VARIABLE THAT HAS YOUR DATA IN IT. YOU MUST ENSURE THAT YOU HAVE AT LEAST 10 DATA POINTS IN YOUR VECTOR AND THAT THERE IS SOME DIFFERENCES IN THE DATA POINTS (MAX SIZE OF INTEGER VECTOR IS APPROX. 2500 , MAX SIZE OF REAL VECTOR IS 2000). AFTER YOU HAVE ENTERED YOUR DATA YOU WILL BE ASKED
- (2) IF YOU DESIRE A SMOOTHED EMPIRICAL DENSITY FUNCTION OR NOT. THE EMPIRICAL DENSITY FUNCTION WHEN PLOTTED GIVES ESSENTIALLY A MORE EXACT PICTURE OF THE DATA THAN DOES THE HISTOGRAM ALONE, ALTHOUGH THIS FEATURE IS SLIGHTLY BLURRED BY THE PRECISION WHICH CAN BE OBTAINED WITH THE APL BALL (THE APL FINE PLOT IS NOT PRESENTLY AVAIL-ABLE ON THE NPS SYSTEM). THE SMOOTHED EMPIRICAL DENSITY IS DEFINED BY THE RELATION (LEWIS, LIU, ROBINSON, AND ROSENBLATT, 1975; ROSENBLATT, 1956)

$$\bar{F}(Z) = \frac{1}{N} \sum_{I=1}^N \frac{W((X - Z) \div B(N))}{N \times B(N)}$$

WHERE N IS THE NUMBER OF DATA POINTS, B(N) IS A BAND-WIDTH FUNCTION,

$$B(N) = \text{RANGE} \div \text{SQRT}(N)$$

AND W IS A WEIGHT FUNCTION,

$$W(Z) = \begin{cases} 0 & \text{IF } |Z| > 1 \\ 1 - |Z| & \text{OTHERWISE} \end{cases}$$

- $\bar{F}(z)$ IS COMPUTED FOR VALUES OF z BETWEEN THE MAXIMUM AND THE MINIMUM OF THE SAMPLE AND PLOTTED OVER THE HISTOGRAM USING THE SYMBOL -F-. THE RELATIVE FREQUENCY MARKS ON THE LEFT OF THE OUTPUT REFER TO THE HISTOGRAM, AND NOT TO THE DENSITY FUNCTION. AFTER THIS QUERY YOU WILL BE ASKED
- (3) IF YOU DESIRE TO TITLE YOUR HISTOGRAM. IF YOU ELECT TO TITLE YOUR HISTOGRAM, SIMPLY TYPE YOUR TITLE, ENSURING THAT YOUR TITLE IS MORE THAN ONE CHARACTER IN LENGTH. IF NO TITLE IS DESIRED JUST HIT THE CARRIAGE RETURN. AFTER THE TITLE QUERY YOU WILL BE ASKED
 - (4) IF YOU WANT TO SET YOUR OWN SCALE AND THE NUMBER OF CELLS. YOUR RESPONSE MUST BE A VECTOR OF 3 ELEMENTS THE FIRST ELEMENT IS THE NUMBER OF CELLS YOU DESIRE, THIS MUST BE AN INTEGER BETWEEN 10 AND 28, THE SECOND ELEMENT IS THE LEFT SCALE POINT AND THE THIRD ELEMENT IS THE RIGHT SCALE POINT (HIST DOES NOT REQUIRE THAT YOUR INTERVAL BE DIVISIBLE BY THE NUMBER OF CELLS). IF YOU WANT HIST TO AUTOMATICALLY SCALE AND PICK THE CELLS YOU SHOULD TYPE THE VECTOR 0 0 0. AFTER YOU HAVE SELECTED YOUR SCALING TECHNIQUE YOU WILL BE ASKED
 - (5) IF YOU WANT DATA POINTS NOT INSIDE THE SCALE LIMITS INCLUDED IN THE HISTOGRAM ROUTINE. MOST HISTOGRAMS LUMP DATA POINTS THAT FALL OUTSIDE THE SCALE LIMITS IN THE END CELLS. HOWEVER, HIST GIVES YOU THE OPTION OF INCLUDING THEM OR EXCLUDING THEM, I.E. OF OBTAINING A HISTOGRAM FOR THE CONDITIONAL DENSITY. AFTER YOUR RESPONSE TO THIS QUERY YOU WILL BE ASKED
 - (6) IF YOU WANT YOUR OUTPUT TO APPEAR ON THE OFFLINE PRINTER OR ON YOUR TERMINAL. IF YOU SELECT THE OFFLINE PRINTER THE NEXT RESPONSE YOU WILL RECEIVE ON YOUR TERMINAL IS - HISTOGRAM SENT TO PRINTER -. THIS RESPONSE WILL TAKE SEVERAL SECONDS AND AFTER IT IS RECEIVED YOUR TERMINAL IS FREE FOR FURTHER USE. HOWEVER, IF YOU ELECTED TO HAVE YOUR HISTOGRAM PRINTED ON YOUR TERMINAL THE PRINTING WOULD BEGIN IN JUST A FEW SECONDS BUT WOULD TAKE BETWEEN 5 AND 10 MINUTES TO COMPLETE.

THE FOLLOWING BASIC DESCRIPTIVE STATISTICS ARE COMPUTED AND PRINTED OUT BY HIST.

MEAN, MEDIAN, TRIMEAN, MIDMEAN, MODE
GEOMETRIC AND HARMONIC MEANS (POSITIVE SAMPLES ONLY)
VARIANCE, STANDARD DEVIATION, COEFFICIENT OF VARIATION,
RANGE AND MIDSREAD
THIRD AND FOURTH CENTRAL MOMENTS, COEFFICIENTS OF SKEW-
NESS AND KURTOSIS
MAXIMUM, MINIMUM AND 5 SAMPLE QUANTILES

IN ADDITION, THE MEAN IS DISPLAYED ON THE HISTOGRAM BY A VERTICAL COLUMN OF -M- AND THE QUARTILES BY COLUMNS OF DOTS.

INTERPRETING THE OUTPUT

THE DEFINITIONS OF THE BASIC STATISTICS COMPUTED BY HIST ARE LISTED BELOW. PAGE NUMBER REFERENCES ARE TO THE CRC STANDARD MATH TABLES, 19TH EDITION (1971).

- MEAN AVERAGE OF THE SAMPLE (P 554).
- MEDIAN MID-VALUE OF THE SAMPLE, IF THERE ARE AN ODD NUMBER OF SAMPLE POINTS, OR THE AVERAGE OF THE TWO MIDDLE VALUES FOR AN EVEN NUMBER OF POINTS (P 555)
- SAMPLE QUARTILES THE $Q(1)=.25$, $Q(2)=.50$, AND $Q(3)=.75$ POPULATION QUARTILES ARE THE SOLUTION TO THE EQUATION $PROB (X \leq X(Q(I))) = Q(I)$ $I=1,2,3$. THE SAMPLE QUARTILES, WHICH ESTIMATE THE POPULATION QUARTILES ARE, THE J TH ORDERED VALUE IN THE SAMPLE, WHERE $J = [Q(I) \times N] + 1$. WHERE $N =$ SAMPLE SIZE.
- TRIMEAN $0.25 \times (Q(1) + 2Q(2) + Q(3))$, WHERE THE $Q(I)$ ARE THE QUARTILES.
- MIDMEAN THE AVERAGE OF ALL THE SAMPLE VALUES BETWEEN THE UPPER AND LOWER QUARTILES.
- MODE THE DATA POINT THAT OCCURS MOST OFTEN (IF ALL THE DATA POINTS ARE DIFFERENT OR IF THERE ARE MORE THAN 300 DATA POINTS THE MODE WILL NOT BE PRINTED. IF TWO OR MORE MODES OCCUR HIST WILL PRINT THE FIRST MODE.)
- MIDRANGE AVERAGE OF THE MAXIMUM AND MINIMUM.
- GEOMETRIC (P 554).
MEAN
- HARMONIC (P 555).
MEAN
- VARIANCE (P 557). UNBIASED ESTIMATORS FOR VARIANCE AND STANDARD DEVIATION ARE USED.
- STANDARD (P 557).
DEVIATION

COEFFICIENT OF VARIATION = STANDARD DEVIATION + |MEAN| WHEN THE MEAN IS LESS THAN 1E-30, THE COEFFICIENT OF VARIATION IS SET TO ZERO.

MEAN (P 556). THE AVERAGE OF THE SUM OF THE ABSOLUTE DEVIATION DIFFERENCES BETWEEN THE SAMPLE VALUES AND THE MEDIAN.

RANGE MAXIMUM - MINIMUM (P 557).

MIDSPREAD $Q(3) - Q(1)$, ALSO CALLED THE INTERQUARTILE DISTANCE.

M3 THIRD CENTRAL MOMENT. UNBIASED ESTIMATOR IS USED. (P 558)

M4 FOURTH CENTRAL MOMENT. UNBIASED ESTIMATOR IS USED. (P 558)

COEFFICIENT OF SKEWNESS $M3 + (STD DEV)*3$

COEFFICIENT OF KURTOSIS $(M4 + (STD DEV)*4) - 3$

BETA1 BIASED ESTIMATE OF THIRD CENTRAL MOMENT. CAN BE USED IN TESTING FOR NORMALITY. (BIOMETRIKA TABLES FOR STATISTICIANS, 1966).

BETA2 BIASED ESTIMATE OF FOURTH CENTRAL MOMENT. (BIOMETRIKA TABLES FOR STATISTICIANS, 1966).

MAXIMUM LARGEST SAMPLE VALUE.

MINIMUM SMALLEST SAMPLE VALUE.

SAMPLE THE α -QUANTILE, $X(\alpha)$, IS THE SOLUTION TO THE EQ. QUANTILES PROBABILITY $(X \leq X(\alpha)) = \alpha$.

With this complete description the general user should be able to take full advantage of HIST and put to use all its options.

B. USAGE WITH TELEPHONE DATA 1 AND TELEPHONE DATA 2, OFFLINE, ALL DATA, ECDF, AND TITLE

HIST was now used on two sets of data. Both telephone data 1 and telephone data 2 were first used with the offline printer demonstrating the title option, the empirical density function option and using the conditional option with any data points outside the designated interval being lumped into the end cells. When HIST was typed the following responses to each of the queries were entered.

HIST

ENTER DATA IN VECTOR FORM

□:

TEL DAT1

IF YOU ALSO WANT A SMOOTHED EMPIRICAL DENSITY FUNCTION ENTER A 1 . IF YOU DO NOT WANT IT ENTER A 0 .

□:

1

IF YOU WANT TO TITLE YOUR HISTOGRAM TYPE YOUR TITLE. IF YOU DO NOT WANT A TITLE JUST HIT THE CARRIAGE RETURN.

TELEPHONE DATA 1

IF YOU WANT TO SET THE NUMBER OF CELLS AND THE SCALE ENTER FIRST THE NUMBER OF CELLS (AN INTEGER BETWEEN 10 AND 28) FOLLOWED BY A SPACE AND THEN YOUR LEFT SCALE POINT FOLLOWED BY A SPACE AND THEN YOUR RIGHT SCALE POINT. HOWEVER, IF YOU WANT HIST TO AUTOMATICALLY SCALE ENTER 0 0 0 .

□:

23 0 20000

GIVEN THAT YOU HAVE SET YOUR OWN SCALE, TO INCLUDE DATA POINTS THAT MIGHT BE OUTSIDE YOUR SCALE LIMITS IN THE END CELLS, TYPE 1 . IF YOU DESIGNATED AUTOSCALE ALSO, TYPE 1 . IF HOWEVER, YOU DO NOT WANT THE DATA OUTSIDE THE SCALE LIMITS INCLUDED IN THE HISTOGRAM, TYPE 0 .

□:

1

IF YOU WANT YOUR OUTPUT TO APPEAR ON THE OFFLINE PRINTER, TYPE 1 . IF YOU WANT YOUR OUTPUT TO APPEAR ON YOUR TERMINAL, TYPE 0 . (NOTE IF YOU TYPED 0 BE SURE YOUR TERMINALS CARRIAGE PAGE SETTING IS ON THE MAXIMUM WIDTH)

□:

1

HISTOGRAM SENT TO PRINTER

Note that telephone data 1 was contained in the variable TELDAT1 and that the number of cells chosen was 28 with the left scale point being 0 and the right scale point being 20,000.

After the response - HISTOGRAM SENT TO PRINTER - was received. HIST was again typed under identical conditions and telephone data 2 was entered through the variable TELDAT2.

HIST

ENTER DATA IN VECTOR FORM

□:

TELDAT2

IF YOU ALSO WANT A SMOOTHED EMPIRICAL DENSITY FUNCTION ENTER A 1 . IF YOU DO NOT WANT IT ENTER A 0 .

□:

1

IF YOU WANT TO TITLE YOUR HISTOGRAM TYPE YOUR TITLE. IF YOU DO NOT WANT A TITLE JUST HIT THE CARRIAGE RETURN.

TELEPHONE DATA 2

IF YOU WANT TO SET THE NUMBER OF CELLS AND THE SCALE ENTER FIRST THE NUMBER OF CELLS (AN INTEGER BETWEEN 10 AND 28) FOLLOWED BY A SPACE AND THEN YOUR LEFT SCALE POINT FOLLOWED BY A SPACE AND THEN YOUR RIGHT SCALE POINT. HOWEVER, IF YOU WANT HIST TO AUTOMATICALLY SCALE ENTER 0 0 0 .

□:

28 0 20000

GIVEN THAT YOU HAVE SET YOUR OWN SCALE, TO INCLUDE DATA POINTS THAT MIGHT BE OUTSIDE YOUR SCALE LIMITS IN THE END CELLS, TYPE 1 . IF YOU DESIGNATED AUTOSCALE ALSO, TYPE 1 . IF HOWEVER, YOU DO NOT WANT THE DATA OUTSIDE THE SCALE LIMITS INCLUDED IN THE HISTOGRAM, TYPE 0 .

□:

1

IF YOU WANT YOUR OUTPUT TO APPEAR ON THE OFFLINE PRINTER, TYPE 1 . IF YOU WANT YOUR OUTPUT TO APPEAR ON YOUR TERMINAL, TYPE 0 . (NOTE IF YOU TYPED 0 BE SURE YOUR TERMINALS CARRIAGE PAGE SETTING IS ON THE MAXIMUM WIDTH)

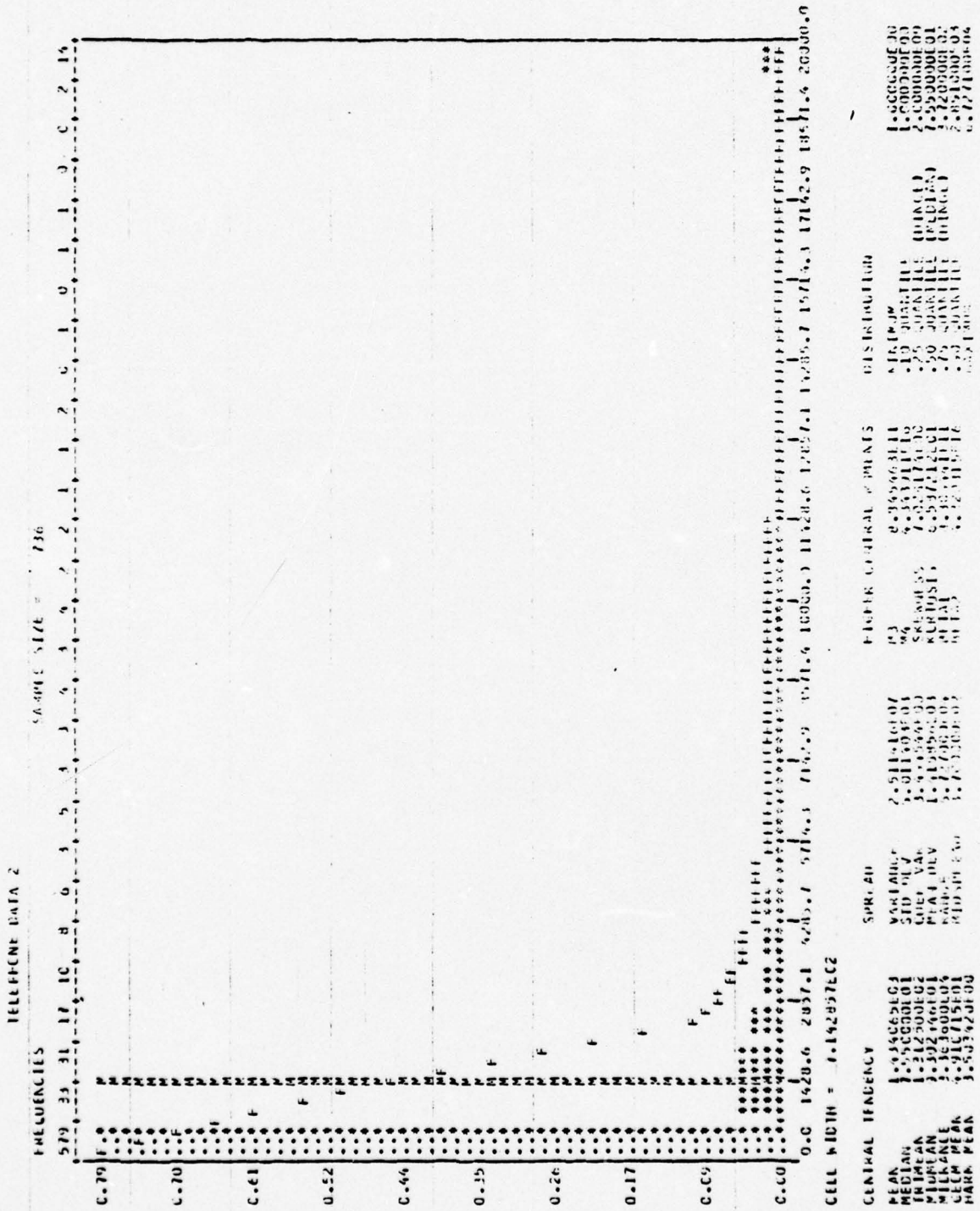
□:

1

HISTOGRAM SENT TO PRINTER

Now by looking at figure 1 (output for telephone data 1) and figure 2 (output from telephone data 2) the similarities and differences in the histograms can be compared. Without getting into specifics, the empirical density function plot seems to indicate that both sets of data are similar. However, the one time-between-errors dominate the data; a more detailed discussion of this data and its analysis is given in Section VIII.

FIGURE 2



C. USAGE WITH TELEPHONE DATA 1 AND TELEPHONE DATA 2, ON LINE; CONDITIONAL DATA BETWEEN 2 AND 140, ECDF, AND TITLE

Because both sets of data contain:

- (1) a large number of elements,
- (2) a large number of times-between-error equal to 1 (this becomes more apparent when HISTLIST is described), and
- (3) the range of the data sets is so extensive,

it would appear that the conditional option available on HIST could be used to see if the two data sets are in fact similar over a smaller interval. This in fact was done using the on line printer option, the empirical density function option, the title option and the conditional option with any data points outside the designated interval excluded from the histogram.

HIST
ENTER DATA IN VECTOR FORM

□: TELDAT1

IF YOU ALSO WANT A SMOOTHED EMPIRICAL DENSITY FUNCTION ENTER
A 1 . IF YOU DO NOT WANT IT ENTER A 0 .

□: 1

IF YOU WANT TO TITLE YOUR HISTOGRAM TYPE YOUR TITLE.
IF YOU DO NOT WANT A TITLE JUST HIT THE CARRIAGE RETURN.

TELEPHONE DATA 1 BETWEEN 2 AND 140

IF YOU WANT TO SET THE NUMBER OF CELLS AND THE SCALE ENTER
FIRST THE NUMBER OF CELLS (AN INTEGER BETWEEN 10 AND 28)
FOLLOWED BY A SPACE AND THEN YOUR LEFT SCALE POINT FOLLOWED
BY A SPACE AND THEN YOUR RIGHT SCALE POINT. HOWEVER, IF YOU
WANT HIST TO AUTOMATICALLY SCALE ENTER 0 0 0 .

□: 28 2 140

GIVEN THAT YOU HAVE SET YOUR OWN SCALE, TO INCLUDE DATA POINTS THAT MIGHT BE OUTSIDE YOUR SCALE LIMITS IN THE END CELLS, TYPE 1 . IF YOU DESIGNATED AUTOSCALE ALSO, TYPE 1 . IF HOWEVER, YOU DO NOT WANT THE DATA OUTSIDE THE SCALE LIMITS INCLUDED IN THE HISTOGRAM, TYPE 0 .

□:

0

IF YOU WANT YOUR OUTPUT TO APPEAR ON THE OFFLINE PRINTER, TYPE 1 . IF YOU WANT YOUR OUTPUT TO APPEAR ON YOUR TERMINAL, TYPE 0 . (NOTE IF YOU TYPED 0 BE SURE YOUR TERMINALS CARRIAGE PAGE SETTING IS ON THE MAXIMUM WIDTH)

□:

0

Note that the same variable TELDAT1 is used but this time the interval was between 2 and 140. Also, the HISTOGRAM SENT TO PRINTER - was not typed because the on-line printer (terminal) option was employed.

After the output for telephone data 1 was printed HIST was again typed and telephone data 2 was entered under identical conditions.

HIST

ENTER DATA IN VECTOR FORM

□:

TELDAT2

IF YOU ALSO WANT A SMOOTHED EMPIRICAL DENSITY FUNCTION ENTER A 1 . IF YOU DO NOT WANT IT ENTER A 0 .

□:

1

IF YOU WANT TO TITLE YOUR HISTOGRAM TYPE YOUR TITLE. IF YOU DO NOT WANT A TITLE JUST HIT THE CARRIAGE RETURN.

TELEPHONE DATA 2 BETWEEN 2 AND 140

IF YOU WANT TO SET THE NUMBER OF CELLS AND THE SCALE ENTER FIRST THE NUMBER OF CELLS (AN INTEGER BETWEEN 10 AND 28) FOLLOWED BY A SPACE AND THEN YOUR LEFT SCALE POINT FOLLOWED BY A SPACE AND THEN YOUR RIGHT SCALE POINT. HOWEVER, IF YOU WANT HIST TO AUTOMATICALLY SCALE ENTER 0 0 0 .

□:

28 2 140

GIVEN THAT YOU HAVE SET YOUR OWN SCALE, TO INCLUDE DATA POINTS THAT MIGHT BE OUTSIDE YOUR SCALE LIMITS IN THE END CELLS, TYPE 1 . IF YOU DESIGNATED AUTOSCALE ALSO, TYPE 1 . IF HOWEVER, YOU DO NOT WANT THE DATA OUTSIDE THE SCALE LIMITS INCLUDED IN THE HISTOGRAM, TYPE 0 .
□:

0

IF YOU WANT YOUR OUTPUT TO APPEAR ON THE OFFLINE PRINTER, TYPE 1 . IF YOU WANT YOUR OUTPUT TO APPEAR ON YOUR TERMINAL, TYPE 0 . (NOTE IF YOU TYPED 0 BE SURE YOUR TERMINALS CARRIAGE PAGE SETTING IS ON THE MAXIMUM WIDTH)
□:

0

Figure 3 (output from telephone data 1 between 2 and 140) and figure 4 (output from telephone data 2 between 2 and 140) now appear quite different in shape based on the empirical density function plot. This is, again, because of the extensive range of the data (85,993 for telephone data 1 and 67,271 for telephone data 2) and the large number of times-between-error equal to one. Both sets of data are actually discrete, only occurring at multiples of 1, but as an initial analysis the data sets were treated as continuous. Thus, by employing the conditional option available on HIST differences in the two sets of data become quite apparent whereas before, the differences were not so easily detected.

FIGURE 3

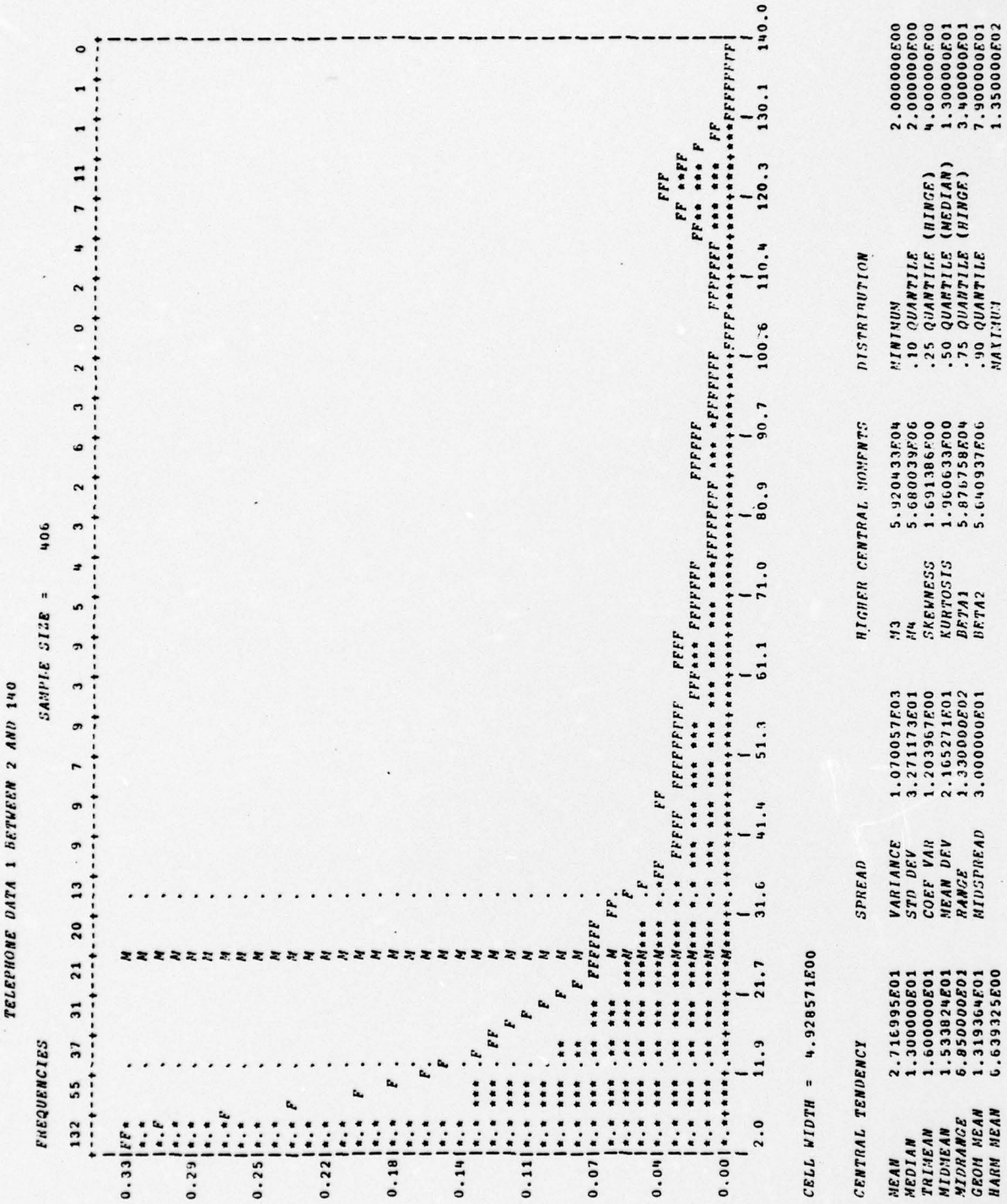
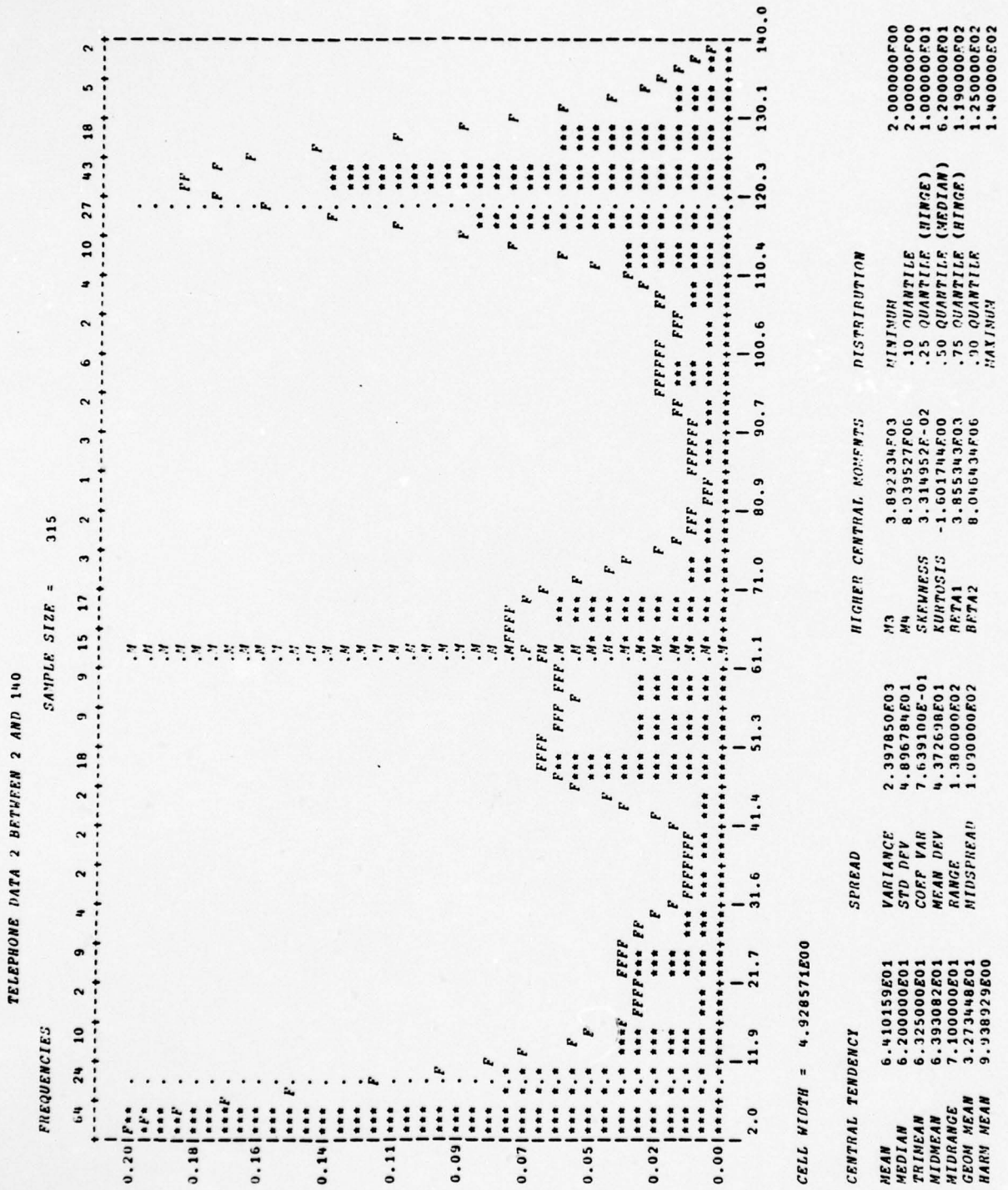


FIGURE 4



III. LISTING ROUTINE

A. DESCRIPTION

The second routine presented is a listing routine. APL has a function that will automatically sort the data and print the results. However, the unique feature of HISTLIST (listing routine) is that it takes advantage of like occurrences in the data and prints the ordered data ascendingly in a compressed form. This becomes highly useful when listing a large number of data points that contain multiple occurrences. It is also a tool for finding multiplicities in supposedly continuous data, and a probability function estimating routine for data which is known to be discrete.

A complete description of how HISTLIST operates is contained in the variable HISTLISTHOW. When the user types HISTLISTHOW the following response is printed on the terminal:

HISTLISTHOW

SYNTAX HISTLIST

HISTLIST IS A HIGHLY CONVENIENT WAY TO LIST YOUR DATA. HISTLIST TAKES YOUR DATA, ORDERS IT AND COMPRESSES IT. FOR EXAMPLE, IF THREE DATA POINTS WERE ALL THE SAME VALUE HISTLIST WOULD JUST PRINT THE VALUE ONCE AND THEN PRINT THE NUMBER OF OCCURENCES OF THAT VALUE. HISTLIST WILL ALSO PRINT THE SERIAL NUMBER OF THE DATA, THE PERCENTAGE THIS SAMPLE VALUE IS TO THE WHOLE SAMPLE, AND A SMALL HISTOGRAM (STARS) SHOWING RELATIVE PERCENTAGES. EXAMPLE: 6 4 4 3 4

HISTLIST

SER. NUM.	ORDERED DATA	NUMBER OF OCCURENCES	PER CENT
1	3	1 ****	.20
2	4	3 *****	.60
5	6	1 ****	.20

HISTLIST IS IDEALLY SUITED FOR A LARGE SAMPLE THAT COULD POSSIBLY HAVE A LOT OF LIKE OCCURENCES. HISTLIST FURTHER HAS THE ADVANTAGE OF BEING USED WITH EITHER THE OFFLINE PRINTER OR THE USERS TERMINAL.

B. USAGE WITH TELEPHONE DATA 1 AND TELEPHONE DATA 2 OFFLINE
 HISTLIST was used with the title option and offline printer option on both telephone data 1 and telephone data 2. When HISTLIST was typed the following responses to each of the queries were entered.

HISTLIST
 HISTLIST PRINTS THE SERIAL NUMBER OF THE COMPRESSED DATA, THE ORDERED DATA COMPRESSED, AND THE NUMBER OF LIKE OCCURENCES. ENTER YOUR DATA IN VECTOR FORM.

□:
 TELDAT1

IF YOU WANT TO TITLE YOUR DATA TYPE YOUR TITLE. IF YOU DO NOT WANT A TITLE JUST HIT THE CARRIAGE RETURN.

TELEPHONE DATA 1

IF YOU WANT YOUR OUTPUT TO APPEAR ON THE OFFLINE PRINTER TYPE 1 . IF YOU WANT YOUR OUTPUT TO APPEAR ON YOUR TERMINAL TYPE 0 .

□:
 1
 HISTLIST SENT TO PRINTER

After the response - HISTLIST SENT TO PRINTER - was received HISTLIST was again typed and telephone data 2 was entered.

HISTLIST

HISTLIST PRINTS THE SERIAL NUMBER OF THE COMPRESSED DATA, THE ORDERED DATA COMPRESSED, AND THE NUMBER OF LIKE OCCURENCES. ENTER YOUR DATA IN VECTOP FORM.

□:

TEL DAT 2

IF YOU WANT TO TITLE YOUR DATA TYPE YOUR TITLE. IF YOU DO NOT WANT A TITLE JUST HIT THE CARRIAGE RETURN.

TELEPHONE DATA 2

IF YOU WANT YOUR OUTPUT TO APPEAR ON THE OFFLINE PRINTER TYPE 1 . IF YOU WANT YOUR OUTPUT TO APPEAR ON YOUR TERMINAL TYPE 0 .

□:

1

HISTLIST SENT TO PRINTER

Looking at figure 5 (output with telephone data 1) and figure 6 (output with telephone data 2) the listings of the two data sets can be compared. It can be seen that both telephone data 1 and telephone data 2 contain a large number of multiple occurrences of the number one and the number two. In fact 19% of telephone data 1 is the number one and 24% of telephone data 2 is the number one. Also, telephone data 2 has many more multiple occurrences in the 120 to 130 range than telephone data 1. This was quickly apparent when one looked at the stars to the right of the ordered data.

FIGURE 5A

TELEPHONE DATA 1

SERIAL NUMBER	ORDERED DATA	NUMBER OF OCCURENCES	PER CENT
1	000000	128	0.150
2	000000	128	0.150
3	000000	128	0.150
4	000000	128	0.150
5	000000	128	0.150
6	000000	128	0.150
7	000000	128	0.150
8	000000	128	0.150
9	000000	128	0.150
10	000000	128	0.150
11	000000	128	0.150
12	000000	128	0.150
13	000000	128	0.150
14	000000	128	0.150
15	000000	128	0.150
16	000000	128	0.150
17	000000	128	0.150
18	000000	128	0.150
19	000000	128	0.150
20	000000	128	0.150
21	000000	128	0.150
22	000000	128	0.150
23	000000	128	0.150
24	000000	128	0.150
25	000000	128	0.150
26	000000	128	0.150
27	000000	128	0.150
28	000000	128	0.150
29	000000	128	0.150
30	000000	128	0.150
31	000000	128	0.150
32	000000	128	0.150
33	000000	128	0.150
34	000000	128	0.150
35	000000	128	0.150
36	000000	128	0.150
37	000000	128	0.150
38	000000	128	0.150
39	000000	128	0.150
40	000000	128	0.150
41	000000	128	0.150
42	000000	128	0.150
43	000000	128	0.150
44	000000	128	0.150
45	000000	128	0.150
46	000000	128	0.150
47	000000	128	0.150
48	000000	128	0.150
49	000000	128	0.150
50	000000	128	0.150
51	000000	128	0.150
52	000000	128	0.150
53	000000	128	0.150
54	000000	128	0.150
55	000000	128	0.150
56	000000	128	0.150
57	000000	128	0.150
58	000000	128	0.150
59	000000	128	0.150
60	000000	128	0.150
61	000000	128	0.150
62	000000	128	0.150
63	000000	128	0.150
64	000000	128	0.150
65	000000	128	0.150
66	000000	128	0.150
67	000000	128	0.150
68	000000	128	0.150
69	000000	128	0.150
70	000000	128	0.150
71	000000	128	0.150
72	000000	128	0.150
73	000000	128	0.150
74	000000	128	0.150
75	000000	128	0.150
76	000000	128	0.150
77	000000	128	0.150
78	000000	128	0.150
79	000000	128	0.150
80	000000	128	0.150
81	000000	128	0.150
82	000000	128	0.150
83	000000	128	0.150
84	000000	128	0.150
85	000000	128	0.150
86	000000	128	0.150
87	000000	128	0.150
88	000000	128	0.150
89	000000	128	0.150
90	000000	128	0.150
91	000000	128	0.150
92	000000	128	0.150
93	000000	128	0.150
94	000000	128	0.150
95	000000	128	0.150
96	000000	128	0.150
97	000000	128	0.150
98	000000	128	0.150
99	000000	128	0.150
100	000000	128	0.150
101	000000	128	0.150
102	000000	128	0.150
103	000000	128	0.150
104	000000	128	0.150
105	000000	128	0.150
106	000000	128	0.150
107	000000	128	0.150
108	000000	128	0.150
109	000000	128	0.150
110	000000	128	0.150
111	000000	128	0.150
112	000000	128	0.150
113	000000	128	0.150
114	000000	128	0.150
115	000000	128	0.150
116	000000	128	0.150
117	000000	128	0.150
118	000000	128	0.150
119	000000	128	0.150

FIGURE 5C

6331	6208	.CCCCC	00000
6332	6209	.CCCCC	00000
6333	6210	.CCCCC	00000
6334	6211	.CCCCC	00000
6335	6212	.CCCCC	00000
6336	6213	.CCCCC	00000
6337	6214	.CCCCC	00000
6338	6215	.CCCCC	00000
6339	6216	.CCCCC	00000
6440	6217	.CCCCC	00000
6441	6218	.CCCCC	00000
6442	6219	.CCCCC	00000
6443	6220	.CCCCC	00000
6444	6221	.CCCCC	00000
6445	6222	.CCCCC	00000
6446	6223	.CCCCC	00000
6447	6224	.CCCCC	00000
6448	6225	.CCCCC	00000
6449	6226	.CCCCC	00000
6450	6227	.CCCCC	00000
6451	6228	.CCCCC	00000
6452	6229	.CCCCC	00000
6453	6230	.CCCCC	00000
6454	6231	.CCCCC	00000
6455	6232	.CCCCC	00000
6456	6233	.CCCCC	00000
6457	6234	.CCCCC	00000
6458	6235	.CCCCC	00000
6459	6236	.CCCCC	00000
6460	6237	.CCCCC	00000
6461	6238	.CCCCC	00000
6462	6239	.CCCCC	00000
6463	6240	.CCCCC	00000
6464	6241	.CCCCC	00000
6465	6242	.CCCCC	00000
6466	6243	.CCCCC	00000
6467	6244	.CCCCC	00000
6468	6245	.CCCCC	00000
6469	6246	.CCCCC	00000
6470	6247	.CCCCC	00000
6471	6248	.CCCCC	00000
6472	6249	.CCCCC	00000
6473	6250	.CCCCC	00000
6474	6251	.CCCCC	00000
6475	6252	.CCCCC	00000
6476	6253	.CCCCC	00000
6477	6254	.CCCCC	00000
6478	6255	.CCCCC	00000
6479	6256	.CCCCC	00000
6480	6257	.CCCCC	00000
6481	6258	.CCCCC	00000
6482	6259	.CCCCC	00000
6483	6260	.CCCCC	00000
6484	6261	.CCCCC	00000
6485	6262	.CCCCC	00000
6486	6263	.CCCCC	00000
6487	6264	.CCCCC	00000
6488	6265	.CCCCC	00000
6489	6266	.CCCCC	00000
6490	6267	.CCCCC	00000
6491	6268	.CCCCC	00000
6492	6269	.CCCCC	00000
6493	6270	.CCCCC	00000
6494	6271	.CCCCC	00000
6495	6272	.CCCCC	00000
6496	6273	.CCCCC	00000
6497	6274	.CCCCC	00000
6498	6275	.CCCCC	00000
6499	6276	.CCCCC	00000
6500	6277	.CCCCC	00000
6501	6278	.CCCCC	00000
6502	6279	.CCCCC	00000
6503	6280	.CCCCC	00000
6504	6281	.CCCCC	00000
6505	6282	.CCCCC	00000
6506	6283	.CCCCC	00000
6507	6284	.CCCCC	00000
6508	6285	.CCCCC	00000
6509	6286	.CCCCC	00000
6510	6287	.CCCCC	00000
6511	6288	.CCCCC	00000
6512	6289	.CCCCC	00000
6513	6290	.CCCCC	00000
6514	6291	.CCCCC	00000
6515	6292	.CCCCC	00000
6516	6293	.CCCCC	00000
6517	6294	.CCCCC	00000
6518	6295	.CCCCC	00000
6519	6296	.CCCCC	00000
6520	6297	.CCCCC	00000
6521	6298	.CCCCC	00000
6522	6299	.CCCCC	00000
6523	6300	.CCCCC	00000
6524	6301	.CCCCC	00000
6525	6302	.CCCCC	00000
6526	6303	.CCCCC	00000
6527	6304	.CCCCC	00000
6528	6305	.CCCCC	00000
6529	6306	.CCCCC	00000
6530	6307	.CCCCC	00000
6531	6308	.CCCCC	00000
6532	6309	.CCCCC	00000
6533	6310	.CCCCC	00000
6534	6311	.CCCCC	00000
6535	6312	.CCCCC	00000
6536	6313	.CCCCC	00000
6537	6314	.CCCCC	00000
6538	6315	.CCCCC	00000
6539	6316	.CCCCC	00000
6540	6317	.CCCCC	00000
6541	6318	.CCCCC	00000
6542	6319	.CCCCC	00000
6543	6320	.CCCCC	00000
6544	6321	.CCCCC	00000
6545	6322	.CCCCC	00000
6546	6323	.CCCCC	00000
6547	6324	.CCCCC	00000
6548	6325	.CCCCC	00000
6549	6326	.CCCCC	00000
6550	6327	.CCCCC	00000
6551	6328	.CCCCC	00000
6552	6329	.CCCCC	00000
6553	6330	.CCCCC	00000

FIGURE 6A

TELEPHONE DATA 2

SERIAL NUMBER	ORDERED DATA	NUMBER OF OCCURENCES	PER CENT
1	000000	178	0.242
179	000000	36	0.049
2	000000	11	0.015
3	000000	6	0.008
4	000000	6	0.008
5	000000	6	0.008
6	000000	6	0.008
7	000000	6	0.008
8	000000	6	0.008
9	000000	6	0.008
10	000000	6	0.008
11	000000	6	0.008
12	000000	6	0.008
13	000000	6	0.008
14	000000	6	0.008
15	000000	6	0.008
16	000000	6	0.008
17	000000	6	0.008
18	000000	6	0.008
19	000000	6	0.008
20	000000	6	0.008
21	000000	6	0.008
22	000000	6	0.008
23	000000	6	0.008
24	000000	6	0.008
25	000000	6	0.008
26	000000	6	0.008
27	000000	6	0.008
28	000000	6	0.008
29	000000	6	0.008
30	000000	6	0.008
31	000000	6	0.008
32	000000	6	0.008
33	000000	6	0.008
34	000000	6	0.008
35	000000	6	0.008
36	000000	6	0.008
37	000000	6	0.008
38	000000	6	0.008
39	000000	6	0.008
40	000000	6	0.008
41	000000	6	0.008
42	000000	6	0.008
43	000000	6	0.008
44	000000	6	0.008
45	000000	6	0.008
46	000000	6	0.008
47	000000	6	0.008
48	000000	6	0.008
49	000000	6	0.008
50	000000	6	0.008
51	000000	6	0.008
52	000000	6	0.008
53	000000	6	0.008
54	000000	6	0.008
55	000000	6	0.008
56	000000	6	0.008
57	000000	6	0.008
58	000000	6	0.008
59	000000	6	0.008
60	000000	6	0.008
61	000000	6	0.008
62	000000	6	0.008
63	000000	6	0.008
64	000000	6	0.008
65	000000	6	0.008
66	000000	6	0.008
67	000000	6	0.008
68	000000	6	0.008
69	000000	6	0.008
70	000000	6	0.008
71	000000	6	0.008
72	000000	6	0.008
73	000000	6	0.008
74	000000	6	0.008
75	000000	6	0.008
76	000000	6	0.008
77	000000	6	0.008
78	000000	6	0.008
79	000000	6	0.008
80	000000	6	0.008
81	000000	6	0.008
82	000000	6	0.008
83	000000	6	0.008
84	000000	6	0.008
85	000000	6	0.008
86	000000	6	0.008
87	000000	6	0.008
88	000000	6	0.008
89	000000	6	0.008
90	000000	6	0.008
91	000000	6	0.008
92	000000	6	0.008
93	000000	6	0.008
94	000000	6	0.008
95	000000	6	0.008
96	000000	6	0.008
97	000000	6	0.008
98	000000	6	0.008
99	000000	6	0.008
100	000000	6	0.008
101	000000	6	0.008
102	000000	6	0.008
103	000000	6	0.008
104	000000	6	0.008
105	000000	6	0.008
106	000000	6	0.008
107	000000	6	0.008
108	000000	6	0.008
109	000000	6	0.008
110	000000	6	0.008
111	000000	6	0.008
112	000000	6	0.008
113	000000	6	0.008
114	000000	6	0.008
115	000000	6	0.008
116	000000	6	0.008
117	000000	6	0.008
118	000000	6	0.008
119	000000	6	0.008
120	000000	6	0.008
121	000000	6	0.008
122	000000	6	0.008
123	000000	6	0.008
124	000000	6	0.008
125	000000	6	0.008
126	000000	6	0.008

FIGURE 6D

692	6385	.000000	1	0.001
693	6055	.000000	1	0.001
694	6816	.000000	1	0.001
695	6821	.000000	1	0.001
696	7307	.000000	1	0.001
697	7329	.000000	1	0.001
698	7146	.000000	1	0.001
699	7927	.000000	1	0.001
700	8039	.000000	1	0.001
701	8053	.000000	1	0.001
702	8153	.000000	1	0.001
703	8147	.000000	1	0.001
704	9206	.000000	1	0.001
705	9256	.000000	1	0.001
706	9317	.000000	1	0.001
707	9341	.000000	1	0.001
708	9362	.000000	1	0.001
709	9365	.000000	1	0.001
710	10020	.000000	1	0.001
711	10376	.000000	1	0.001
712	1037	.000000	1	0.001
713	10518	.000000	1	0.001
714	12042	.000000	1	0.001
715	12793	.000000	1	0.001
716	13012	.000000	1	0.001
717	13175	.000000	1	0.001
718	14692	.000000	1	0.001
719	15225	.000000	1	0.001
720	16377	.000000	1	0.001
721	18335	.000000	1	0.001
722	19243	.000000	1	0.001
723	19846	.000000	1	0.001
724	19849	.000000	1	0.001
725	19851	.000000	1	0.001
726	20840	.000000	1	0.001
727	21440	.000000	1	0.001
728	24273	.000000	1	0.001
729	24770	.000000	1	0.001
730	26278	.000000	1	0.001
731	27238	.000000	1	0.001
732	29133	.000000	1	0.001
733	29513	.000000	1	0.001
734	39607	.000000	1	0.001
735	53122	.000000	1	0.001
736	57271	.000000	1	0.001

In addition, HISTLIST saved on printing time and paper. By printing the data in compressed form HISTLIST saved printing 448 lines (6 additional pages) in the case of telephone data 1 and 419 lines (5 additional pages) in the case of telephone data 2. Thus, HISTLIST not only gives the user more information than an ordered listing of the data, but also is cost effective in terms of printing time and paper used. Finally, note that it is not possible to look at the data in as much detail with routine HIST as with HISTLIST. If the data is continuous and there are no multiplicities, then HISTLIST gives only this information and an ordered listing of the data. The shape of the density function can best be seen (estimated) in using routine HIST.

IV. SECTIONING ROUTINE

A. DESCRIPTION

The third routine presented is the sectioning routine, HISTS. HISTS (sectioning routine) gives a way of assessing the variability of estimates of descriptive statistics from sample data. It is essential that the data be in random order.

The basic idea is as follows: Assume we have m independent observations y_1, y_2, \dots, y_m of a random variable Y . The usual estimate of its mean value $\mu = E(Y)$ is the sample mean \bar{y} , where $\bar{y} = \frac{1}{m} \sum_{i=1}^m y_i$. Now \bar{y} is the least-squares estimate of μ , and therefore unbiased with variance $\text{var}(\bar{y}) = \sigma^2/m$, where $\sigma^2 = \text{var}(y)$. Of course σ^2 is unknown, but we can estimate it from the data with the sample variance

$$s^2 = \frac{1}{m-1} \sum_{i=1}^m (y_i - \bar{y})^2$$

and then estimate the variance of the estimate \bar{y} of μ as

$$\widetilde{\text{var}}(\bar{y}) = \frac{s^2}{m} = \frac{1}{m(m-1)} \sum_{i=1}^m (y_i - \bar{y})^2$$

This is the basis for the sectioning routine: here the y_i are estimates of descriptive statistics from the m sections of the data and \bar{y} is the average of the statistics

from each section. Estimates are assumed independent because the original data is assumed to be independent.

A complete description of how HISTS operates is contained in the variable HISTSHOW. When the user types HISTSHOW the following response is printed on the terminal:

HISTSHOW

SYNTAX HISTS

HISTS ALLOWS YOU TO INTERACTIVELY SECTION YOUR DATA AND ASSESS THE VARIABILITY IN EACH OF THE DESCRIPTIVE STATISTICS BY USING THE SECTIONED SAMPLE DATA.

WHEN YOU TYPE HISTS YOU WILL BE ASKED TO DESIGNATE THE NUMBER OF SECTIONS YOU DESIRE. HISTS WILL THEN TAKE THE UNORDERED DATA AND DIVIDE THE DATA INTO THE NUMBER OF SECTIONS YOU INDICATE DISCARDING ANY DATA POINTS LEFT OVER. FOR EXAMPLE, IF YOU HAVE 301 DATA POINTS AND YOU SELECT 10 SECTIONS HISTS WILL PLACE THE FIRST 30 DATA POINTS IN THE FIRST SECTION, THE SECOND 30 DATA POINTS IN THE SECOND SECTION AND SO ON UNTIL THE LAST DATA POINT IS OMITTED. YOU WILL NOW HAVE 10 SECTIONS WITH 30 DATA POINTS PER SECTION.

HISTS WOULD NOW PRINT THE FOLLOWING STATISTICS ON EACH OF THE SECTIONS: MEAN, MEDIAN, VARIANCE, STD DEV, COEF VAR, SKEWNESS, KURTOSIS, MINIMUM AND MAXIMUM. IN ADDITION, THE ABOVE STATISTICS WOULD BE PRINTED FOR THE UNSECTIONED DATA TO ALLOW FOR COMPARISONS.

FINALLY, HISTS WILL PRINT (1) THE MEAN OF THE SECTIONED DATA STATISTICS. FOR EXAMPLE, THE MEAN FOR SKEWNESS WOULD BE EACH SECTION VALUE FOR SKEWNESS SUMMED UP AND DIVIDED BY THE NUMBER OF SECTIONS. (2) THE VARIANCE AND STD DEV OF THE SECTIONED DATA STATISTICS. AND, (3) THE STD DEV DIVIDED BY THE SQUARE ROOT OF THE NUMBER OF SECTIONS, WHICH ESTIMATES THE STANDARD DEVIATION OF THE STATISTICS.

AS A RESULT, HISTS WILL GIVE YOU AN UNBIASED ESTIMATE OF THE VARIANCE OF THE SAMPLE MEAN, MEDIAN, VARIANCE, STD DEV, COEF VAR, SKEWNESS AND KURTOSIS FROM USING THE SAMPLE VARIANCE OF THE SECTIONED DATA. WITH THIS RESULT, CONFIDENCE INTERVALS CAN ALSO BE OBTAINED FOR EACH OF THE ABOVE STATISTICS, IF THE ESTIMATES FROM THE SECTIONS ARE NORMALLY DISTRIBUTED. HISTS IS BEST SUITED FOR LARGE AND MODERATE SIZED SAMPLES; FOR SMALL SAMPLES JACKKNIFING SHOULD BE CONSIDERED.

B. USAGE WITH TELEPHONE DATA 1

HISTS was now used on telephone data 1 to assess the variability in the mean, median, variance, standard deviation, coefficient of variation, skewness and kurtosis. When HISTS was typed the following responses were entered (see figure 7).

The 672 data points of telephone data 1 were broken down into 16 sections with 42 data points per section. Because of this breakdown no data points were discarded.

The unsectioned statistics printed can be compared with the values printed by HIST (figure 1) and are in fact the same. Providing that the estimates are normally distributed (this can be checked with the normal plots, described later), confidence intervals for each of the statistics (mean, median, variance, standard deviation, coefficient of variation, skewness and kurtosis) based on the t-statistic can be obtained in the following manner

$$\bar{y}_n \pm \frac{s_{\bar{y}_n}}{\sqrt{m}} t_{(1-\frac{1}{2}\alpha), (m-1)}$$

Here \bar{y}_n is the mean of the sectioned data statistics (obtained from column one under summary for sectioned data); $\frac{s_{\bar{y}_n}}{\sqrt{m}}$ is the standard deviation of the sectioned data statistic divided by the square root of the number of sections (obtained from column four under summary for sectioned data); m is the number sections chosen; and, $t_{(1-\frac{1}{2}\alpha), (m-1)}$ is the $1-\frac{1}{2}\alpha$ quantile of the t-distribution with m-1 degrees of freedom.

FIGURE 7

HISTS
 TYPE THE NUMBER OF SECTIONS YOU DESIRE (INTEGER
 BETWEEN 2 AND 28) BE SURE TO PICK YOUR NUMBER OF
 SECTIONS SO AS TO MINIMIZE THE NUMBER OF DATA
 POINTS THAT WILL HAVE TO BE DISCARDED. (HISTS
 PLACES THE DATA INTO THE EQUAL NUMBER OF SECTIONS
 YOU INDICATE DISCARDING ANY DATA LEFT OVER)

16

ENTER YOUR DATA TO BE SECTIONED IN VECTOR FORM

11:

TELDATA1

SECTION	MEAN	MEDIAN	VARIANCE	STD DEV	COEF VAR	SKEWNESS	KURTOSIS	MINIHUN	MAXIHUN
1	1.0526F03	8.5000F00	3.4598E07	5.8820E03	5.5879F00	6.3484E00	3.7831E01	1.0000F00	3.8003F04
2	3.2133E03	1.4500E01	1.8494E08	1.3599E04	4.2320E00	5.8106E00	3.3017E01	1.0000F00	8.5993F04
3	1.7662E03	1.4500E01	4.7383E07	6.5103E03	3.6860E00	4.2806E00	1.7836E01	1.0000F00	3.5644F04
4	6.0669E02	1.1000F01	5.3412E06	2.3111E03	3.8094F00	4.3148E00	1.6486E01	1.0000F00	1.1280F04
5	1.5639F03	5.0500F01	2.1924E07	4.6824E03	2.9941F00	4.2209E00	1.6565E01	1.0000F00	2.6443F04
6	2.5343E03	5.7000F01	4.0337E07	6.3511E03	2.5061F00	3.1573E00	9.5654E00	1.0000F00	3.0974F04
7	2.6778E03	2.2000F01	7.2756E07	8.5297E03	3.1853F00	4.1587E00	1.7579E01	1.0000F00	4.7120F04
8	9.8881E02	1.8500E01	3.8282E07	6.1873E03	6.2573E00	6.4801E00	3.8995E01	1.0000F00	4.0131E04
9	1.5176E03	2.2000F01	2.0792E07	4.5599E03	3.0046E00	2.9866E00	6.8551F00	1.0000F00	1.7174E04
10	2.7682E03	1.4000F01	1.0906E08	1.0443E04	3.7726E00	4.8932E00	2.4134E01	1.0000F00	6.1710F04
11	1.9250E03	1.4000F01	5.9852E07	7.7364E03	4.0173E00	5.4734E00	2.9059E01	1.0000F00	4.7592E04
12	8.1955E02	4.9500F01	8.2895E07	2.8791E03	3.5131E00	4.5999E00	1.9785E01	1.0000F00	1.5868F04
13	2.1201E03	4.0000F00	1.2224E08	1.1056E04	5.2150E00	5.9400E00	3.3861E01	1.0000F00	6.9775F04
14	2.3062E02	1.1500F01	3.3035E05	5.7476E02	2.4923E00	3.4695E00	1.2010E01	1.0000F00	2.9620F03
15	4.3752E02	7.0000E00	5.7201E06	2.3917E03	5.4664F00	6.3983E00	3.8289E01	1.0000F00	1.5504F04
16	5.4838E02	6.5000F00	1.1340E07	3.3675F03	6.1408E00	6.4765E00	3.8964E01	1.0000F00	2.1848F04
UNSECTIONED	1.5482E03	1.4000E01	4.8362E07	6.9543E03	4.4918E00	7.1531E00	6.2608F01	1.0000F00	8.5993F04

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SUMMARY FOR SECTIONED DATA

	MEAN	VARIANCE	STD DEV	STD1(SECS)*.5
MEAN	1.5482E03	8.6488E05	9.2999E02	2.3250E02
MEDIAN	2.0313E01	2.8023E02	1.6740E01	4.1850E00
VARIANCE	4.8637E07	2.6217E15	5.1203E07	1.2801E07
STD DEV	6.0664E03	1.2625E07	3.5532E03	8.8830E02
COEF VAR	4.1175E00	1.5503E00	1.2451E00	3.1128E-01
SKEWNESS	4.3343E00	1.4701E00	1.2125E00	3.0312E-01
KURTOSIS	2.4552E01	1.2448E02	1.1173E01	2.7933E00

C. INTERPRETATION OF RESULTS

As an example, a confidence interval for the coefficient of variation was obtained in the following manner. The mean value of the coefficient of variation for the 16 sections is 4.1175 (column 1). The standard deviation divided by the square root of 16 is .31128 (column 4). Using $\alpha = .05$, the t value with 15 degrees of freedom is 2.131. Thus, the 95% confidence interval for the coefficient of variation for telephone data 1 is $4.1175 \pm (.31128)(2.131)$ which is [3.454, 4.781]. Confidence intervals on the six other statistics could be obtained in the same fashion.

Again note that the use of the variance estimate from the sectioned data to give confidence intervals is based on the assumption that the estimates from the sections are independent and normally distributed. The normality will depend on the number of observations in each section, which should be kept large to induce normality. This requirement conflicts with the need to make the number of sections large to reduce the variability in the estimate of the variance of the statistics.

Another problem is that if the number of observations in each section is small, the estimates may be severely biased. This effect can be seen in figure 7: note that all of the 16 estimates of skewness from the sections are smaller than the estimate 7.1531 from the unsectioned data.

V. JACKKNIFE ROUTINE

A. DESCRIPTION

The fourth routine presented is the jackknife routine. HISTJACK (jackknife routine) is another way of assessing the variability in the estimates from sample data, and also of reducing bias in estimates of the descriptive statistics.

The jackknife procedure, like the previous sectioning method, is based on the assumption that an independent and identically distributed random sample x_1, x_2, \dots, x_n have come from a population with an unknown distribution function $F_X(x)$. If we divide the sample into r groups, with each group containing the same number of elements, we can obtain estimates $\tilde{\theta}$ of the descriptive statistics, which we denote generically as θ , in the same manner as previously done with the sectioning method. The difference here is that the descriptive statistics are computed with the j^{th} group deleted $j=1, 2, \dots, r$. We then let $\tilde{\theta}_{(j)}$ be the result or the descriptive statistic estimate computed with the j^{th} subgroup omitted, and $\tilde{\theta}_{a11}$ is the corresponding result or descriptive statistic estimated from the entire sample (no group omitted). The jackknife pseudo-values are then computed in the following way:

$$\tilde{\theta}_{*j} = (r)(\tilde{\theta}_{a11}) - (r-1)(\tilde{\theta}_{(j)}) \quad j = 1, 2, \dots, r$$

Then we define the jackknifed estimator to be:

$$\tilde{\theta}_* = \frac{1}{r} \sum_{j=1}^r \tilde{\theta}_{*j}$$

The pseudo-values can be used to obtain variance estimates for $\tilde{\theta}_*$, and to set approximate confidence limits, using Student's t. The idea is that the pseudo-values will be approximately independent and possibly normally distributed. The jackknifed estimator $\tilde{\theta}_*$ is a sample average so we form an estimate s_*^2 of its variance given by the following relationship (Miller, 1974):

$$s^2 = \frac{\sum \tilde{\theta}_{*j}^2 - \frac{1}{r} (\sum \tilde{\theta}_{*j})^2}{r-1}$$

$$s_*^2 = s^2/r$$

This procedure is particularly useful if the number n of data points is small, but it must be used with care. Note, that the estimator $\tilde{\theta}_*$ is designed to eliminate a $1/n$ bias term in the estimator $\tilde{\theta}$.

A complete description of how HISTJACK operates is contained in the variable HISTJACKHOW. When the user types HISTJACKHOW the following response is printed on the terminal.

HISTJACKHOW

SYNTAX HISTJACK

HISTJACK ALLOWS YOU TO INTERACTIVELY JACKKNIFE YOUR DATA AND ASSESS THE VARIABILITY IN EACH OF THE STATISTICAL ESTIMATES BY USING THE SAMPLE DATA.

WHEN YOU TYPE HISTJACK YOU WILL BE ASKED TO DESIGNATE THE NUMBER OF GROUPS YOU DESIRE. HISTJACK WILL TAKE THE UNORDERED DATA AND DIVIDE THE DATA INTO THE NUMBER OF GROUPS YOU INDICATE DISCARDING ANY DATA POINTS LEFT OVER. FOR EXAMPLE, IF YOU HAVE 22 DATA POINTS AND YOU SELECT 7 GROUPS HISTJACK WILL PLACE THE FIRST 3 DATA POINTS IN GROUP 1, THE SECOND 3 DATA POINTS IN GROUP 2, AND SO ON UNTIL THE LAST DATA POINT IS OMITTED. YOU WOULD NOW HAVE 7 GROUPS WITH 3 DATA POINTS PER GROUP. IF YOU HAD ELECTED TO DO A COMPLETE JACKKNIFE, THAT IS TYPED 22, YOU WOULD NOW HAVE 22 GROUPS WITH 1 DATA POINT OMITTED PER GROUP.

HISTJACK WOULD NOW PERFORM STATISTICAL COMPUTATIONS USING THE JACKKNIFE PROCEDURE. THAT IS, BY OMITTING ONE GROUP AT A TIME, STARTING WITH THE FIRST GROUP, HISTJACK WOULD PRINT THE FOLLOWING STATISTICS: MEAN, MEDIAN, VARIANCE, STD DEV, COEF VAR, SKEWNESS, KURTOSIS, MINIMUM AND MAXIMUM. IN ADDITION, THE ABOVE STATISTICS WOULD BE PRINTED FOR THE UNGROUPED DATA TO ALLOW FOR COMPARISONS. (NOTE, THE COLUMNS GIVE THE STATISTIC ESTIMATED FROM ALL THE DATA WITH ONE GROUP MISSING, AND NOT THE PSEUDO-VALUES)

FINALLY, HISTJACK WILL PRINT (1) THE JACKKNIFE ESTIMATE (2) THE SAMPLE VARIANCE OF THE PSEUDO-VALUES DERIVED IN THE JACKKNIFE ESTIMATE (3) AND, THE ESTIMATED STD DEV OF THE JACKKNIFE ESTIMATE DIVIDED BY THE SQUARE ROOT OF THE NUMBER OF GROUPS.

AS A RESULT, HISTJACK WILL GIVE YOU AN ESTIMATE OF THE VARIANCE OF THE SAMPLE MEAN, MEDIAN, VARIANCE, STD DEV, COEF VAR, SKEWNESS AND KURTOSIS USING THE SAMPLE VARIANCE OF THE JACKKNIFED DATA. WITH THIS RESULT, CONFIDENCE INTERVALS CAN BE OBTAINED FOR EACH OF THE ABOVE STATISTICS, AGAIN ASSUMING THAT THE PSEUDO-VALUES ARE APPROXIMATELY INDEPENDENT AND NORMALLY DISTRIBUTED. HISTJACK IS BEST SUITED FOR SMALL SAMPLES.

B. USAGE WITH TELEPHONE DATA 1

HISTJACK was now used on telephone data 1 to assess the variability in the mean, median, variance, standard deviation, coefficient of variation, skewness and kurtosis. When HISTJACK was typed the following responses were entered. (see figure 8)

The 672 data points were broken down into 16 groups with 42 data points per group. Again, because of this breakdown no data points were discarded.

The ungrouped statistics printed are again the same values that were printed by HIST (figure 1). Using the jackknife method, confidence intervals for each of the statistics (mean, median, variance, standard deviation, coefficient of variation, skewness and kurtosis) can be obtained in the following manner;

$$\tilde{\theta}_* \pm (s_*) t_{(1-\frac{1}{2}\alpha), (r-1)} .$$

Here $\tilde{\theta}_*$ is the jackknife estimate of the sample data (obtained from column one under summary for jackknifed data); s_* is the jackknife estimate of the standard deviation divided by the square root of the number of groups (obtained from column four under summary for jackknifed data); r is the number of groups chosen; and, $t_{(1-\frac{1}{2}\alpha), (r-1)}$ is the $1-\frac{1}{2}\alpha$ quantile of the t-distribution with $r-1$ degrees of freedom. The basis for these assertions about the confidence intervals using the jackknifing technique is asymptotic and great care must be taken in using them.

HISTJACK
 TYPE THE NUMBER OF GROUPS YOU DESIRE (INTEGER
 BETWEEN 2 AND 50.) BE SURE TO PICK YOUR NUMBER
 OF GROUPS SO AS TO MINIMIZE THE NUMBER OF DATA
 POINTS THAT WILL HAVE TO BE DISCARDED. (HISTJACK
 PLACES THE DATA INTO THE EQUAL NUMBER OF GROUPS
 YOU INDICATE DISCARDING ANY DATA LEFT OVER)
 []:

16

ENTER YOUR DATA TO BE JACKKNIFED IN VECTOR FORM

{}:

TELDAT1

GROUP	MEAN	MEDIAN	VARIANCE	STD DEV	COEF VAR	SKEWNESS	KURTOSIS	MINIMUM	MAXIMUM
1	1.5813E03	1.5000E01	4.9318E07	7.0227E03	4.4412E00	7.1746E00	6.3025E01	1.0000E00	8.5993E04
2	1.4372E03	1.4000E01	3.9338E07	6.2720E03	4.3641E00	6.5220E00	5.0690E01	1.0000E00	6.9775E04
3	1.5337E03	1.4000E01	4.8825E07	6.9875E03	4.5560E00	7.3093E00	6.4762E01	1.0000E00	8.5993E04
4	1.6110E03	1.4000E01	5.1180E07	7.1540E03	4.4408E00	6.9781E00	5.9257E01	1.0000E00	8.5993E04
5	1.5472E03	1.3500E01	5.0162E07	7.0825E03	4.5777E00	7.1494E00	6.1827E01	1.0000E00	8.5993E04
6	1.4825E03	1.3000E01	4.8893E07	6.9923E03	4.7166E00	7.3563E00	6.5160E01	1.0000E00	8.5993E04
7	1.4729E03	1.3000E01	4.6758E07	6.8800E03	4.6425E00	7.5081E00	6.8311E01	1.0000E00	8.5993E04
8	1.5855E03	1.4000E01	4.9073E07	7.0032E03	4.4183E00	7.1850E00	6.3371E01	1.0000E00	8.5993E04
9	1.5503E03	1.3500E01	5.0236E07	7.0877E03	4.5720E00	7.1592E00	6.1773E01	1.0000E00	8.5993E04
10	1.4669E03	1.4000E01	4.4376E07	6.6615E03	4.5413E00	7.4572E00	6.9618E01	1.0000E00	8.5993E04
11	1.5230E03	1.4000E01	4.7680E07	6.9050E03	4.5337E00	7.3200E00	6.5949E01	1.0000E00	8.5993E04
12	1.5968E03	1.3000E01	5.3013E07	7.1423E03	4.4729E00	7.0992E00	5.9701E01	1.0000E00	8.5993E04
13	1.5101E03	1.6000E01	4.3600E07	6.6030E03	4.3726E00	7.1493E00	6.5032E01	1.0000E00	8.5993E04
14	1.6361E03	1.4000E01	5.1446E07	7.1726E03	4.3841E00	6.9200E00	5.8526E01	1.0000E00	8.5993E04
15	1.6223E03	1.5000E01	5.1130E07	7.1506E03	4.4078E00	6.9784E00	5.9319E01	1.0000E00	8.5993E04
16	1.6149E03	1.5000E01	5.0781E07	7.1261E03	4.4128E00	7.0291E00	6.0129E01	1.0000E00	8.5993E04
UNGROUPED	1.5482E03	1.4000E01	4.8362E07	6.9543E03	4.4918E00	7.1531E00	6.2608E01	1.0000E00	8.5993E04

FIGURE 8

SUMMARY FOR JACKKNIFED DATA

	JACKKNIFE ESTIMATE	VARIANCE	(VARIABLES)*.5 JACKKNIFE ESTIMATE OF STD DEV OF MEAN OF PSEUDO-VALUES
MEAN	1.5482E03	8.6488E05	2.3250E02
MEDIAN	1.3063E01	1.5656E02	3.1281E00
VARIANCE	4.8344E07	2.5453E15	1.2613E07
STD DEV	7.0154E03	1.3879E07	9.3135E02
COEF VAR	4.5053E00	2.4262E00	3.8940E-01
SKEWNESS	7.3732E00	1.2963E01	9.0012E-01
KURTOSIS	6.7077E01	4.6806E03	1.7110E01

C. INTERPRETATION OF RESULTS

To compare the confidence interval obtained for the coefficient of variation using the sectioning routine with that obtained using the jackknife routine the following was done. The jackknife estimate of the coefficient of variation for the 16 groups is 4.5053 (column 1). The jackknife estimate of the standard deviation divided by the square root of 16 is .3894. Using $\alpha = .05$, the t value with 15 degrees of freedom is 2.131. Thus, the 95% confidence interval for the coefficient of variation for telephone data 1 is $4.5053 \pm (.3894)(2.131)$ which is [3.676, 5.335]. This compares with the confidence interval of [3.454, 4.781] using the sectioning routine described in section IV. Likewise, confidence intervals on the remaining six statistics could be obtained in a similar manner. Note that the values obtained for the skewness coefficient from the sections are now not evidently biased; of the 16 values, 7 have values below the value 7.1531 for all the data.

D. USAGE WITH COST OVERRUN DATA

To demonstrate how the complete jackknife could be used and why it is better to use when possible, the following was done. The 22 data points of the cost overrun data were used with the jackknife routine (HISTJACK). When HISTJACK was typed the data was entered in the variable YROVR and 22 was typed as the number of groups. By typing 22, which is the same as the number of data points, a complete jackknife was done.

Looking at the output from the complete jackknife (figure 9), the cost overrun data can be studied. One can note that by using the complete jackknife the mean, median, and variance of the jackknife estimate (column one under summary for jackknifed data) are the same value as the ungrouped mean, median and variance. But, also note that the coefficient of variation is less than zero which can happen when using the jackknife technique.

FIGURE 9

HISTJACK
 TYPE THE NUMBER OF GROUPS YOU DESIRE (INTEGER
 BETWEEN 2 AND 50) BE SURE TO PICK YOUR NUMBER
 OF GROUPS SO AS TO MINIMIZE THE NUMBER OF DATA
 POINTS THAT WILL HAVE TO BE DISCARDED. HISTJACK
 PLACES THE DATA INTO THE EQUAL NUMBER OF GROUPS
 YOU INDICATE DISCARDING ANY DATA LEFT OVER)

U:

ENTER YOUR DATA TO BE JACKKNIFE IN VECTOR FORM

Y:ROVR

GROUP	MEAN	MEDIAN	VARIANCE	STD DEV	COFF VAR	SKENNESS	KURTOSIS	MINIMUM	MAXIMUM
1	1.0524E00	-1.6000E00	1.0228E02	1.0113E01	9.6101E00	7.7349E-01	7.3991E-02	-1.3600E01	2.5300E01
2	1.2048E00	-1.2000E00	1.0189E02	1.0094E01	8.3784E00	7.2904E-01	5.1529E-02	-1.3600E01	2.5300E01
3	1.3190E00	-1.2000E00	1.0089E02	1.0044E01	7.6149E00	7.0897E-01	7.9345E-02	-1.3600E01	2.5300E01
4	1.7714E00	-1.2000E00	9.1014E01	9.5401E00	5.3856E00	8.7130E-01	3.2139E-01	-1.3000E01	2.5300E01
5	9.9048E-01	-1.6000E00	1.0213E02	1.0106E01	1.0203E01	7.9519E-01	1.0417E-01	-1.3600E01	2.5300E01
6	7.6190E-01	-1.6000E00	1.0066E02	1.0003E01	8.8022E-01	8.8022E-01	3.1867E-01	-1.3600E01	2.5300E01
7	1.2286E00	-1.2000E00	1.0173E02	1.0086E01	8.2096E00	7.2372E-01	5.4384E-02	-1.3600E01	2.5300E01
8	1.4381E00	-1.2000E00	9.9207E01	9.9603E00	6.3260E00	7.0632E-01	1.4054E-01	-1.3600E01	2.5300E01
9	1.5286E00	-1.2000E00	9.7491E01	9.8738E00	6.4595E00	7.2120E-01	2.0046E-01	-1.3600E01	2.5300E01
10	1.2000E00	-1.2000E00	1.0192E02	1.0095E01	8.4128E00	7.3016E-01	5.1155E-02	-1.3600E01	2.5300E01
11	1.0190E00	-1.6000E00	1.0222E02	1.0111E01	9.9216E00	7.8499E-01	8.8707E-02	-1.3600E01	2.5300E01
12	1.7429E00	-1.2000E00	9.1918E01	9.5874E00	5.5009E00	8.4272E-01	3.1785E-01	-1.3600E01	2.5300E01
13	1.6000E00	-1.2000E00	9.5869E01	9.7913E00	6.1195E00	7.4622E-01	2.4899E-01	-1.3600E01	2.5300E01
14	1.2857E00	-1.2000E00	1.0124E02	1.0062E01	7.8260E00	7.1331E-01	6.7672E-02	-1.3600E01	2.5300E01
15	6.0476E-01	-1.6000E00	9.7232E01	9.8607E00	1.6305E01	9.3010E-01	5.4286E-01	-1.3600E01	2.5300E01
16	1.9048E-01	-1.6000E00	8.4311E01	9.1821E00	4.0206E01	8.9026E-01	9.9254E-01	-1.3600E01	2.5300E01
17	6.8571E-01	-1.6000E00	9.8831E01	9.9414E00	1.4498E01	9.0629E-01	4.2160E-01	-1.3600E01	2.5300E01
18	-8.0952E-02	-1.6000E00	7.1546E01	8.4585E00	1.0449E02	4.6734E-01	-2.0707E-01	-1.3600E01	1.9600E01
19	1.1810E00	-1.6000E00	1.0202E02	1.0101E01	8.5529E00	7.3487E-01	5.0328E-02	-1.3600E01	2.5300E01
20	9.6190E-01	-1.6000E00	1.0201E02	1.0100E01	1.0500E01	8.0563E-01	1.2225E-01	-1.3600E01	2.5300E01
21	1.3333E00	-1.2000E00	1.0072E02	1.0036E01	7.5270E00	7.0754E-01	8.5165E-02	-1.3600E01	2.5300E01
22	5.8095E-01	-1.6000E00	9.6705E01	9.8339E00	1.0927E01	9.3606E-01	5.8008E-01	-1.3600E01	2.5300E01
UNGROUPED	1.0727E00	-1.4000E00	9.7420E01	9.8702E00	9.7010E00	7.0191E-01	2.2400E-01	-1.3600E01	2.5300E01

SUMMARY FOR JACKKNIFE DATA

(VAR(GROUPS))^*.5
 JACKKNIFE ESTIMATE OF STD DEV
 OF MEAN OF PSEUDO-VALUES

MEAN	1.0727E00	9.7420E01
MEDIAN	-1.4000E00	1.8480E01
VARIANCE	9.7420E01	2.3840E04
STD DEV	1.0026E01	6.7477E01
COFF VAR	-1.2279E02	2.0893E05
SKENNESS	8.7458E-01	4.8496E00
KURTOSIS	4.3524E-01	2.7843E01
		2.1043E00
		9.1652E-01
		3.2919E01
		1.7513E00
		9.7452E01
		4.6951E-01
		1.1250E00

VI. EXPONENTIAL PLOTTING ROUTINE

A. DESCRIPTION

The fifth routine presented is an exponential plotting routine. Routine EXPONP is a way of plotting the data to see if it "fits" an exponential distribution, and also to give some indication of what alternative distributions could be used if the exponential hypothesis is rejected.

A complete description of how EXPONP operates is contained in the variable EXPONPHOW. When the user types EXPONPHOW the following response is printed on the terminal.

EXPONPHOW

SYNTAX EXPONP

EXPONP ORDERS THE DATA X(I) AND COMPUTES THE EMPIRICAL LOG SURVIVER FUNCTION FOR THE DATA. THAT IS,

$$\begin{array}{c} \backslash / \\ X \\ / \backslash \end{array} (I) \quad \text{VS} \quad \begin{array}{c} | \quad | \quad | \\ | \quad | \quad | \\ | _ _ | \quad | \end{array} \quad \begin{array}{c} / \\ | \\ \backslash \end{array} \quad \begin{array}{c} 1 \\ - \\ N+1 \end{array} \quad \begin{array}{c} I \\ \text{-----} \\ N+1 \end{array} \quad \begin{array}{c} \backslash \\ | \\ / \end{array}$$

THE ORDERED DATA IS PLOTTED AGAINST THE LOG SURVIVER FUNCTION TO SEE IF THERE IS A LINEAR FIT. EXPONP ALSO ALLOWS YOU TO TITLE YOUR PLOT.

B. USAGE WITH TELEPHONE DATA 1

EXPONP was used with telephone data 1 to see if the data plotted as a relative straight line. When EXPONP was typed the following responses were entered.

EXPONP

EXPONP ORDERS THE DATA YOU GIVE AND COMPUTES THE EMPIRICAL LOG SURVIVER FUNCTION FOR THE DATA. A PLOT OF THE LOG SURVIVER FUNCTION FOR THE DATA IS THEN PRINTED TO SEE IF THERE IS A LINEAR FIT.

IF YOU WANT TO TITLE YOUR PLOT TYPE YOUR TITLE. IF YOU DO NOT WANT A TITLE JUST HIT THE CARRIAGE RETURN.

TELEPHONE DATA 1

ENTER YOUR DATA IN VECTOR FORM

□:

TEL DAT1

Looking at figure 10 (plot of telephone data 1 using EXPONP), it was found that the data did not plot linearly from the origin, but that the data did appear somewhat linear in the tail (5,000 to 90,000 range).

C. USAGE WITH RANDOM GENERATED EXPONENTIALLY DISTRIBUTED SAMPLE WITH MEAN SAME AS TELEPHONE DATA 1

As a comparison, EXPONP was used with an exponentially generated random sample with the same mean as telephone data 1 (figure 11). As expected, this plot is, within limits of sample fluctuations, linear from the origin and in fact, what telephone data 1 would have looked like if the data was truly exponential. The quantization because of the coarseness of the APL type-ball is evident in this plot. The sample size is 672 , but not all these points can be plotted separately.

FIGURE 10

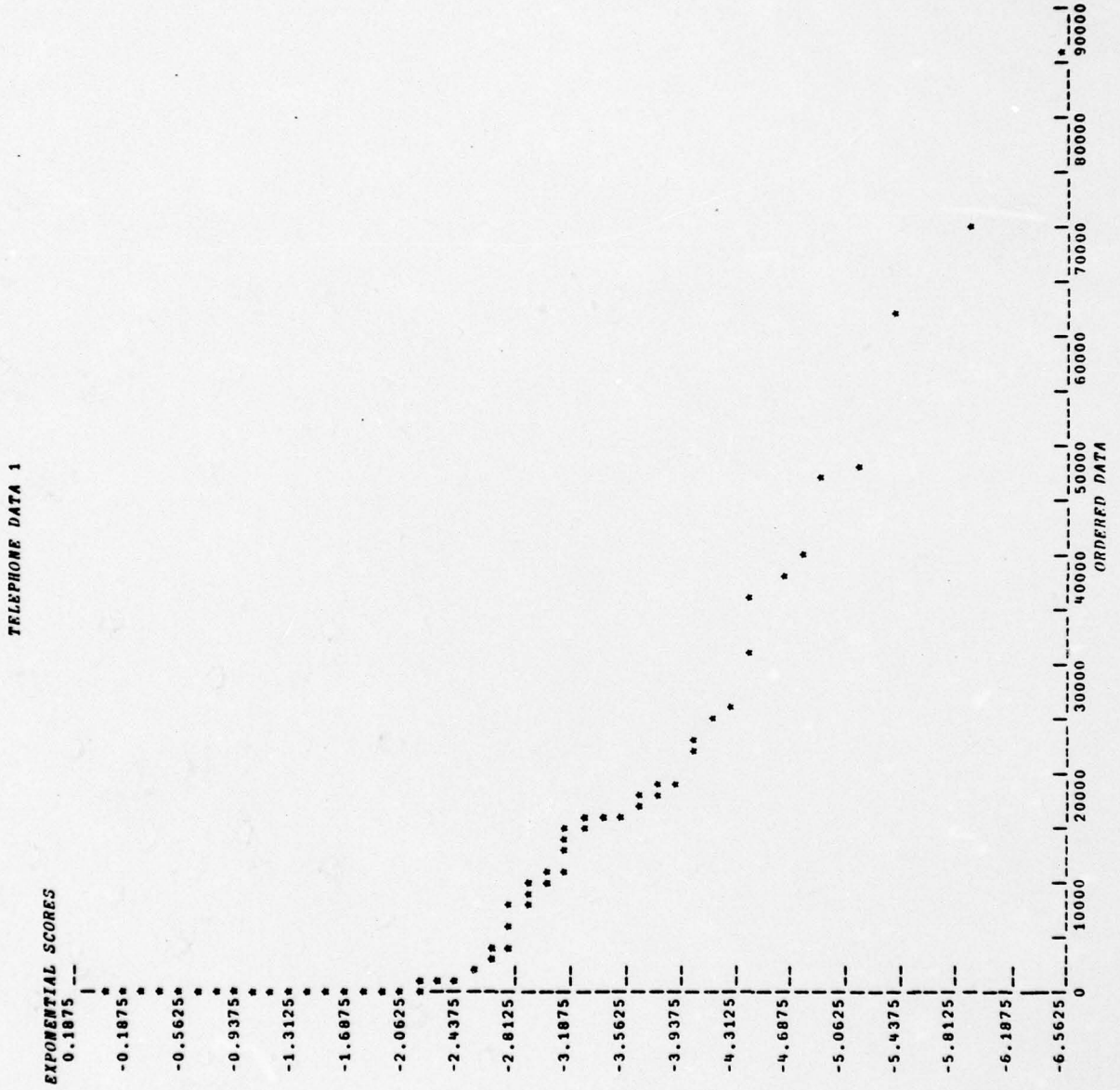
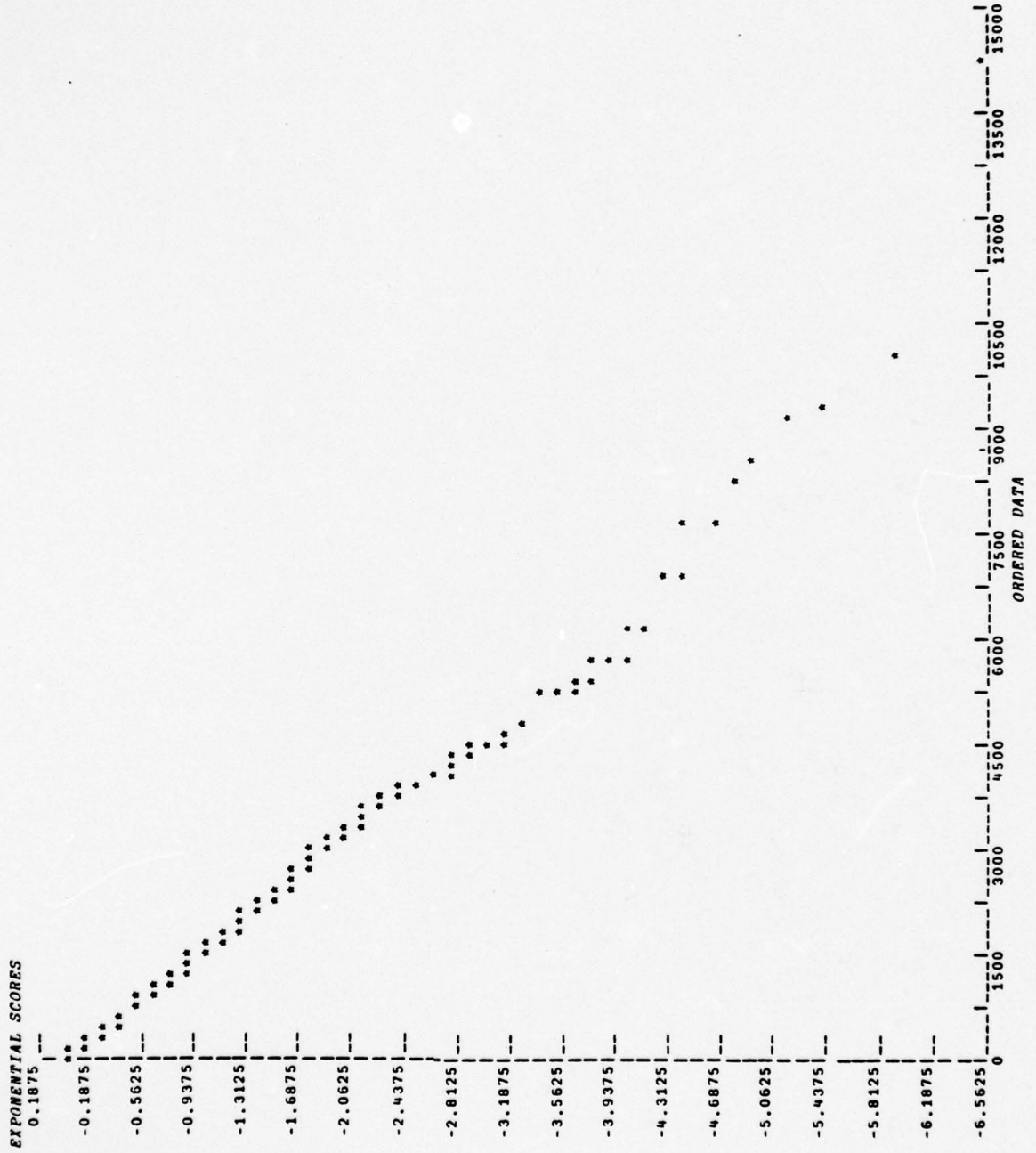


FIGURE 11

EXPONENTIALLY GENERATED RANDOM SAMPLE WITH MEAN THE SAME AS TELEPHONE DATA 1



VII. NORMAL PLOTTING ROUTINE

A. DESCRIPTION

The final routine presented is a normal plotting routine. Routine `NORMP` is a way of plotting the data to see if it "fits" a normal distribution. In particular one might want to look at estimates of descriptive statistics obtained from sections and groups in routines `HISTS` and `HISTJACK`.

A complete description of how `NORMP` operates is contained in the variable `NORMPHOW`. When the user types `NORMPHOW` the following response is printed on the terminal.

`NORMPHOW`

`SYNTAX NORMP`

`NORMP` ORDERS THE DATA $X(I)$ AND COMPUTES THE INVERSE OF THE UNIT NORMAL CUMULATIVE DISTRIBUTION. THAT IS,

$$\begin{array}{c} \backslash / \\ X \\ / \backslash (I) \end{array} \quad \text{VS} \quad \begin{array}{c} T-1 / I \backslash \\ \Phi | \text{-----} | \\ 1 \backslash N+1 / \end{array}$$

THE ORDERED DATA IS PLOTTED AGAINST THE INVERSE OF THE UNIT NORMAL CUMULATIVE DISTRIBUTION TO SEE IF THERE IS A LINEAR FIT. `NORMP` ALSO ALLOWS YOU TO CONVIENTLY TITLE YOUR PLOT.

B. USAGE WITH COST OVERRUN DATA

NORMP was used with the cost overrun data to see if the data plotted as a relative straight line. When NORMP was typed the following responses were entered.

```
NORMP
NORMP ORDERS THE DATA YOU GIVE AND COMPUTES THE
INVERSE OF THE UNIT NORMAL CUMULATIVE DISTRIBUTION
FOR THE DATA. A PLOT OF THE INVERSE OF THE
UNIT NORMAL CUMULATIVE DISTRIBUTION VS THE ORDERED
DATA IS THEN PRINTED TO SEE IF THERE IS A
LINEAR FIT.
```

```
IF YOU WANT TO TITLE YOUR PLOT TYPE YOUR TITLE.
IF YOU DO NOT WANT A TITLE JUST HIT THE CARRIAGE
RETURN.
```

```
COST OVERRUNS
```

```
ENTER YOUR DATA IN VECTOR FORM
```

```
□:
    YROVR
```

Note that the cost overrun data was contained in the variable YROVR . Looking at figure 12 (plot of cost overrun data using NORMP), it was found that the data did in fact plot fairly linear through the range -14 to 26 (formal tests are available; see Wilk & Gnanadesikan, 1968).

C. USAGE WITH NORMAL SAMPLE GENERATED WITH MEAN AND VARIANCE THE SAME AS COST OVERRUN DATA

As a comparison, NORMP was used with a normal sample with the same mean and variance as the cost overrun data (figure 13). As expected, this plot is very linear. But again, this plot is not that much different from that of figure 12, which gives credence to the fact that the cost overrun data might in fact be normally distributed.

FIGURE 12

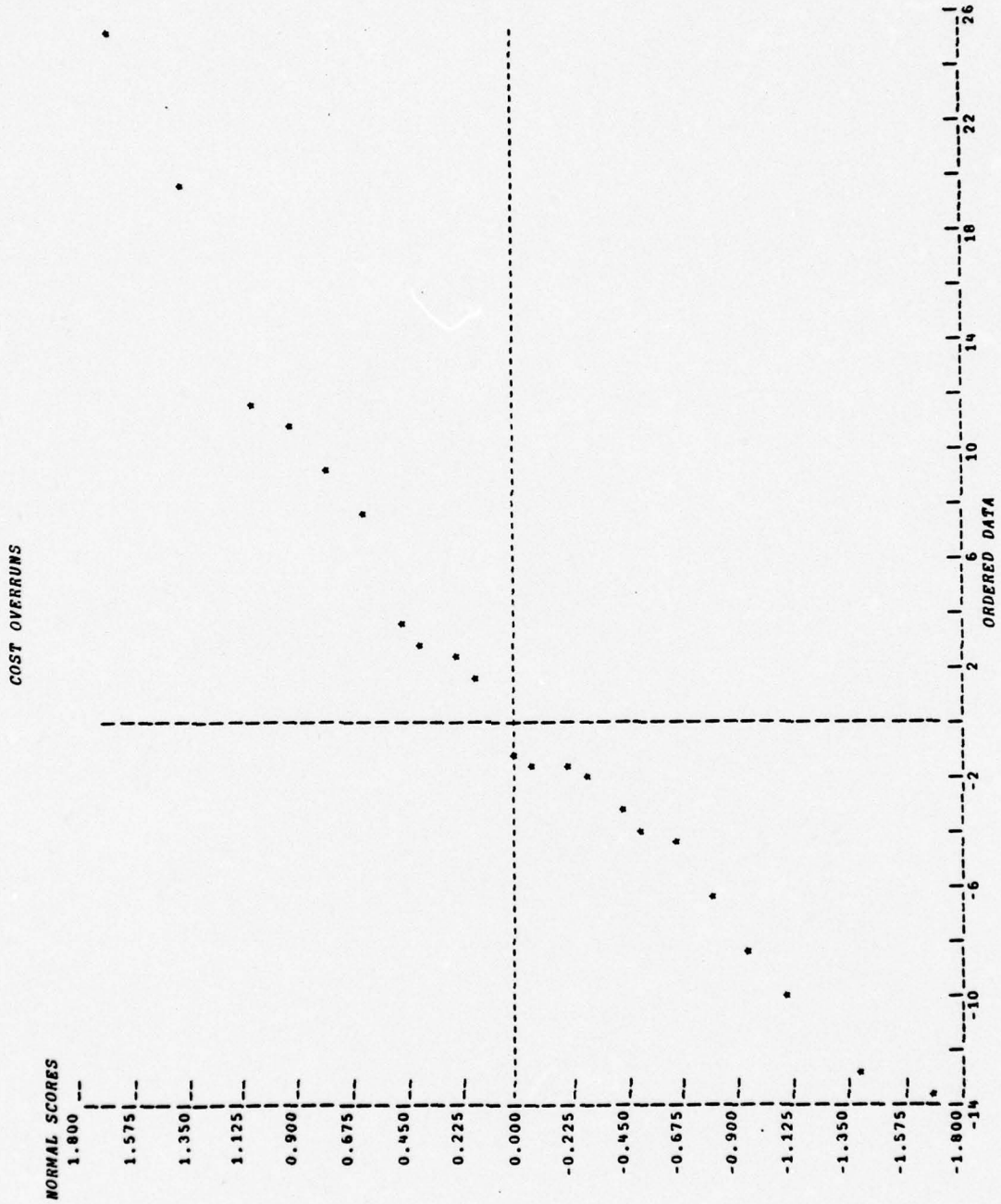
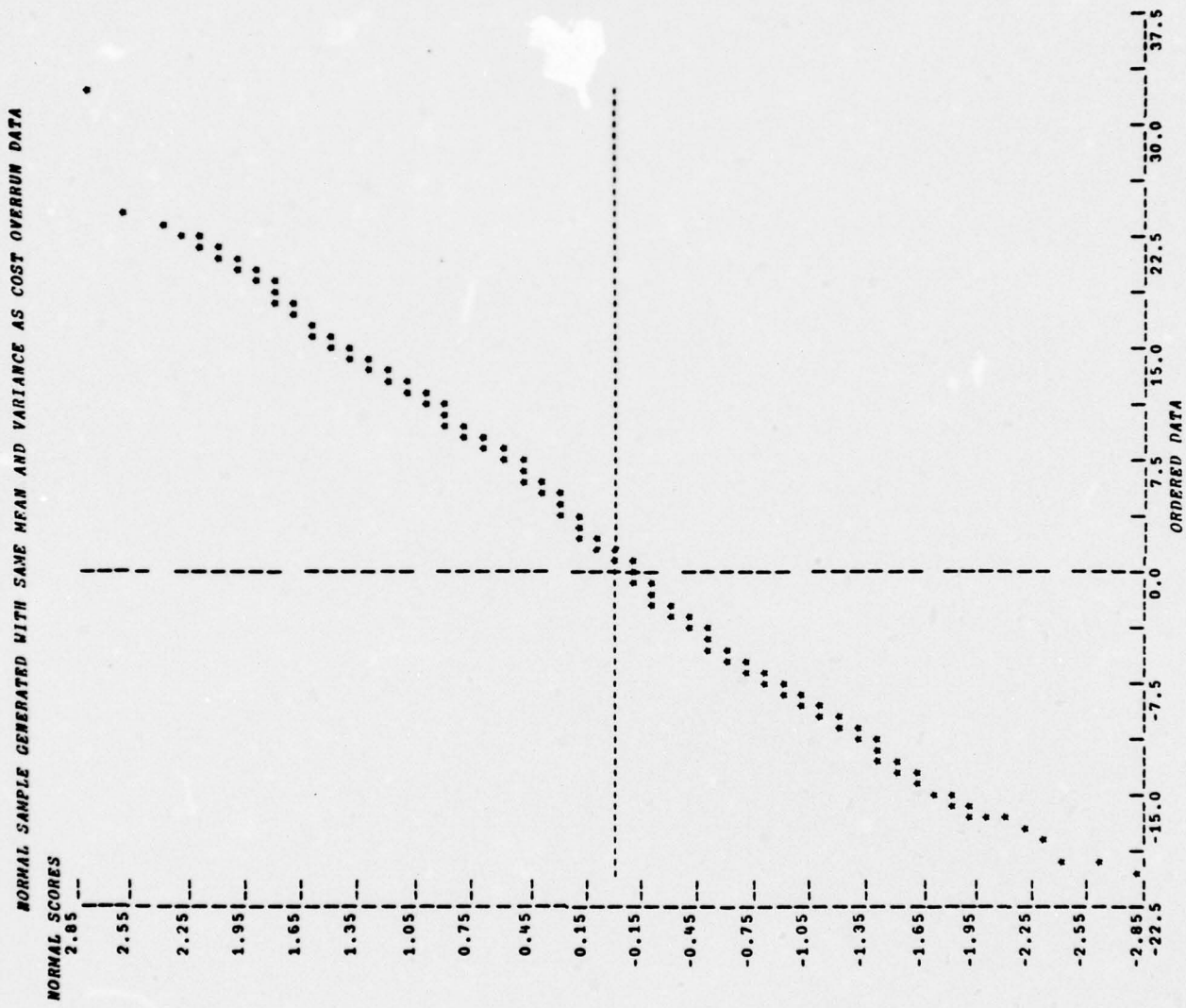


FIGURE 13

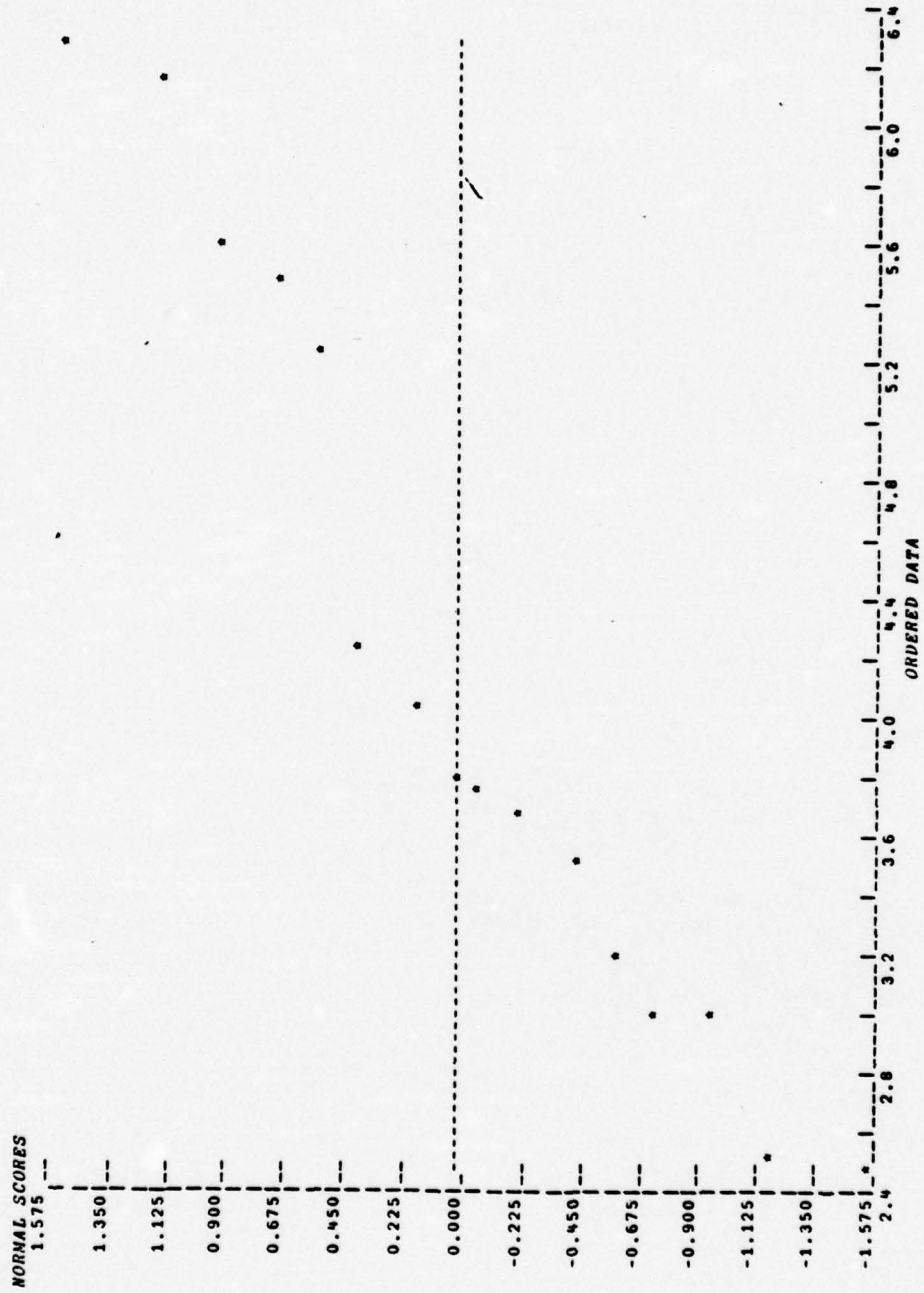


D. USAGE WITH COEFFICIENT OF VARIATION DATA OBTAINED FROM USING SECTIONING ROUTINE

In order to check for normality in the sectioned estimates obtained from using HISTS (sectioning routine) the following was done. The 16 coefficient of variation values obtained from using HISTS with telephone data 1 (column 5, figure 7) were entered as a vector into NORMP. Figure 14 shows that the plot is marginally linear. This demonstrates the need for formal tests to verify normality in the absence of a strictly linear plot (Wilk & Gnanadsikan, 1968).

FIGURE 14

PLOT OF COEF VAR VALUES USING 16 SECTIONS FROM FIGURE 7



VIII. THE INDEPENDENCE AND MARKOV CHAIN HYPOTHESES FOR THE TELEPHONE DATA

The telephone data used in the thesis (Lewis & Cox, 1966) actually consists of binary bits transmitted over telephone lines and the information that the bit transmitted at time i , $i = 0, 1, 2, \dots$ is in error or not. This information is characterized by a sequence of binary-valued random variables $x(i)$, $i = 0, 1, \dots$ where $x(i)=1$ means that the bit transmitted at time i is in error, while $x(i)=0$ means that the bit transmitted at time zero is correctly transmitted.

In telephone data 1 there are 672 ones and 1,105,476 zeros, and a much more compact and equivalent representation of the data is obtained via the sequence of random variables $y(j)$, $j=1, 2, \dots$ where $y(j)$ is one plus the number of correctly transmitted bits between the j^{th} and $(j-1)^{\text{st}}$ bit error, with the convention that $y(j)=1$ if the errors occur on adjacent transmitted bits, and $y(1)$ is the time from $i=0$ to the first incorrectly transmitted bit. The $y(j)$ are called the times-between-errors.

A null hypothesis for the error structure which could be examined is that errors occur independently at each bit with a fixed probability, i.e.

$$P\{x(i)=1\} = \pi(1) \quad i=0, 1, \dots$$

$$P\{x(i)=0\} = \pi(0) = 1-\pi(1) \quad i=0, 1, \dots$$

The $y(j)$'s then are independent and geometrically distributed, since

$$\begin{aligned}
 P\{y(j)=1\} &= P\{\text{if } (j-1)^{\text{st}} \text{ error at time } i; j^{\text{th}} \text{ at time } i+1\} \\
 &= \pi(1) \\
 P\{y(j)=2\} &= P\{\text{if } (j-1)^{\text{st}} \text{ error at time } i; j^{\text{th}} \text{ at time } i+2\} \\
 &= \pi(1)[1-\pi(1)] = \pi(1)\pi(0) \\
 P\{y(j)=k+1\} &= P\{\text{if } (j-1)^{\text{st}} \text{ error at time } i; j^{\text{th}} \text{ at time } i+1+k\} \\
 &= \pi(1)[1-\pi(1)]^k = \pi(1)[\pi(0)]^k
 \end{aligned}$$

Note that, using the geometric series summation formula,

$$\sum_{k=1}^{\infty} P\{y(j)=k\} = \frac{\pi(1)}{1 - (1-\pi(1))} = 1$$

$$E[y(j)] = \sum_{k=1}^{\infty} kP\{y(j)=k\} = \frac{1}{1-\pi(0)} = \frac{1}{\pi(1)}$$

Now assume that the Markov structure of the zero's and ones is described by the transition matrix

$$\underline{P} = \begin{pmatrix} P(0,0) & P(0,1) \\ P(1,0) & P(1,1) \end{pmatrix} = \begin{pmatrix} \rho+(1-\rho)\pi(1) & (1-\rho)\pi(0) \\ (1-\rho)\pi(1) & \rho+(1-\rho)\pi(0) \end{pmatrix}$$

Here $P(m,n) = P\{x(i+1)=n \mid x(i)=m\}$, and we have parameterized the chain in terms of the stationary probability of a one or zero, and a correlation parameter $0 \leq \rho < 1$. Note that there are only two degrees of freedom in the stochastic

matrix, since rows must sum to 1, and there is only one degree of freedom if the stationary probability $\pi(0)=1-\pi(1)$ is fixed. Note that the stationary probabilities in the 2-state case are given by

$$\pi(0) = \frac{P(1,0)}{2-P(0,0)-P(1,1)} \quad \pi(1) = \frac{P(0,1)}{2-P(0,0)-P(1,1)}$$

We now define the runs of ones or zeros i.e. for $\ell=0$ or $\ell=1$, let

$$T_{\ell} = \inf\{n \geq 1: x(i+n) \neq \ell\} - 1,$$

the length of a run of ℓ 's, starting after time i , where the length can be $0, 1, 2, \dots$.

For example if $x(i+1)=1$, then the length of runs of zeros starting after time i is zero, the length of runs of ones is at least one long. Note that it is possible to talk of a conditional runs structure, i.e. the length of a run of ones which is given to start after time i . The run length is then at least one long.

Now the probability of a run T_{ℓ} having length greater than k is, using the Markov property,

$$P\{T_{\ell} \geq k\} = P\{x(i+1)=x(i+2)=\dots x(i+k)=\ell\} = \pi(\ell) [P(\ell, \ell)]^{k-1}$$

$$\text{and} \quad P\{T_{\ell} = 0\} = 1 - \pi(\ell) \quad k=1, \dots$$

Thus, the run lengths are geometrically distributed and

$$E[T(\ell)] = \sum_{k=1}^{\infty} P\{T_{\ell} \geq k\} = \frac{\pi(\ell)}{1-P(\ell, \ell)} = \frac{\pi(\ell)}{(1-p)[1-\pi(\ell)]}$$

Note that $\rho=0$ gives the independence case, and while the runs of ones or zeros are geometrically distributed for both the independence or Markov dependent model, the mean run length is always longer for the Markov dependence, since

$$\frac{\pi(\ell)}{(1-\rho)[1-\pi(\ell)]} \geq \frac{\pi(\ell)}{[1-\pi(\ell)]} \quad 0 \leq \rho < 1$$

Thus, we could use the distributional properties of the runs to (1) check that either hypothesis is tenable or (2) if so, compare the estimated run lengths with the mean length $\hat{\pi}(\ell)/[1-\hat{\pi}(\ell)]$ predicted by the independence assumption. If the run lengths are not geometric, than another model must be postulated.

Note that when this mean time-between-errors is large as it is for telephone data 1 (figure 1; $E[y(j)] = 1,548$) the discreteness of the time scale can be ignored and the geometric distribution is indistinguishable from its continuous time analog, the exponential distribution.

That is approximation of the geometric distribution by an exponential distribution is valid can be seen from the fact that there are 672 errors ($x(i)$'s equal to one) in 1,106,148 transmitted bits, so that an estimate of $\pi(1)$, which is the maximum likelihood estimate under the independence hypothesis, is

$$\hat{\pi}(1) = \frac{\# x(i)'s = 1}{\text{total \# bits transmitted}} = \frac{\# x(i)'s = 1}{\# x(i)'s=1 + \# x(i)'s=0}$$

In the present data

$$\hat{\pi}(1) = \frac{672}{1,106,148} = .0006075$$

Now this geometric hypothesis will be examined, but it is clear from figure 1 that the hypothesis is not true. The distribution is in fact highly skewed and has been examined by Lewis & Cox, 1966.

An alternative model to independent bit errors is that the dependence structure is Markovian. One could examine this hypothesis with time-series methods but a method which is adaptable for use with the histogram routine and which examines both the independence and Markov assumptions is to look at runs of ones and zeros in the $x(i)$. Under both hypothesis these runs have geometrically distributed lengths.

The alternating conditional runs of ones for telephone data 1 are shown in figure 15 and for runs of zeros are shown in figure 16. Also, HISTLIST was used on the conditional runs and figure 17 shows the runs of ones and figure 18 shows the runs of zero.

To test the hypothesis that the runs of ones in telephone data 1 is geometrically distributed the following was done.

Using figures 15 and 17 the following data was obtained:

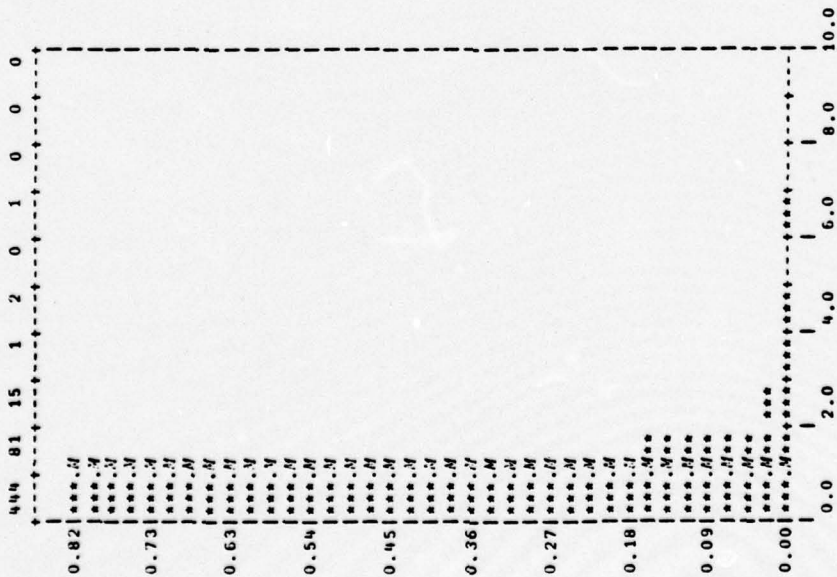
MEAN	= 1.235294	# of runs = 1	= 444
VARIANCE	= .346008	# of runs = 2	= 81
		# of runs = 3	= 15
		# of runs = 4	= 1
		# of runs = 5	= 2
		# of runs \geq 6	= 1

FIGURE 15

RUNS OF ONE FOR TELEPHONE DATA 1

SAMPLE SIZE = 544

FREQUENCIES



CELL WIDTH = 1.000000E00

CENTRAL TENDENCY		SPREAD		HIGHER CENTRAL MOMENTS		DISTRIBUTION	
MEAN	1.235294E00	VARIANCE	3.46080E-01	M3	7.990531E-01	MINIMUM	1.000000E00
MEDIAN	1.000000E00	STD DEV	5.882245E-01	M4	3.219455E00	.10 QUANTILE (HINGE)	1.000000E00
TRIMEAN	1.000000E00	COFF VAR	4.761817E-01	SKEWNESS	3.925965E00	.25 QUANTILE (HINGE)	1.000000E00
MIDMEAN	1.000000E00	MEAN DEV	2.352941E-01	KURTOSIS	2.388968E01	.50 QUANTILE (MEDIAN)	1.000000E00
MIDRANGE	4.000000E00	RANGE	6.000000E00	BETA1	7.946519E-01	.75 QUANTILE (HINGE)	1.000000E00
GEOM MEAN	1.156667E00	MTDSPREAD	0.000000E00	BETA2	3.196802E00	.90 QUANTILE	2.000000E00
HARM MEAN	1.109541E00					MAXIMUM	7.000000E00

FIGURE 16

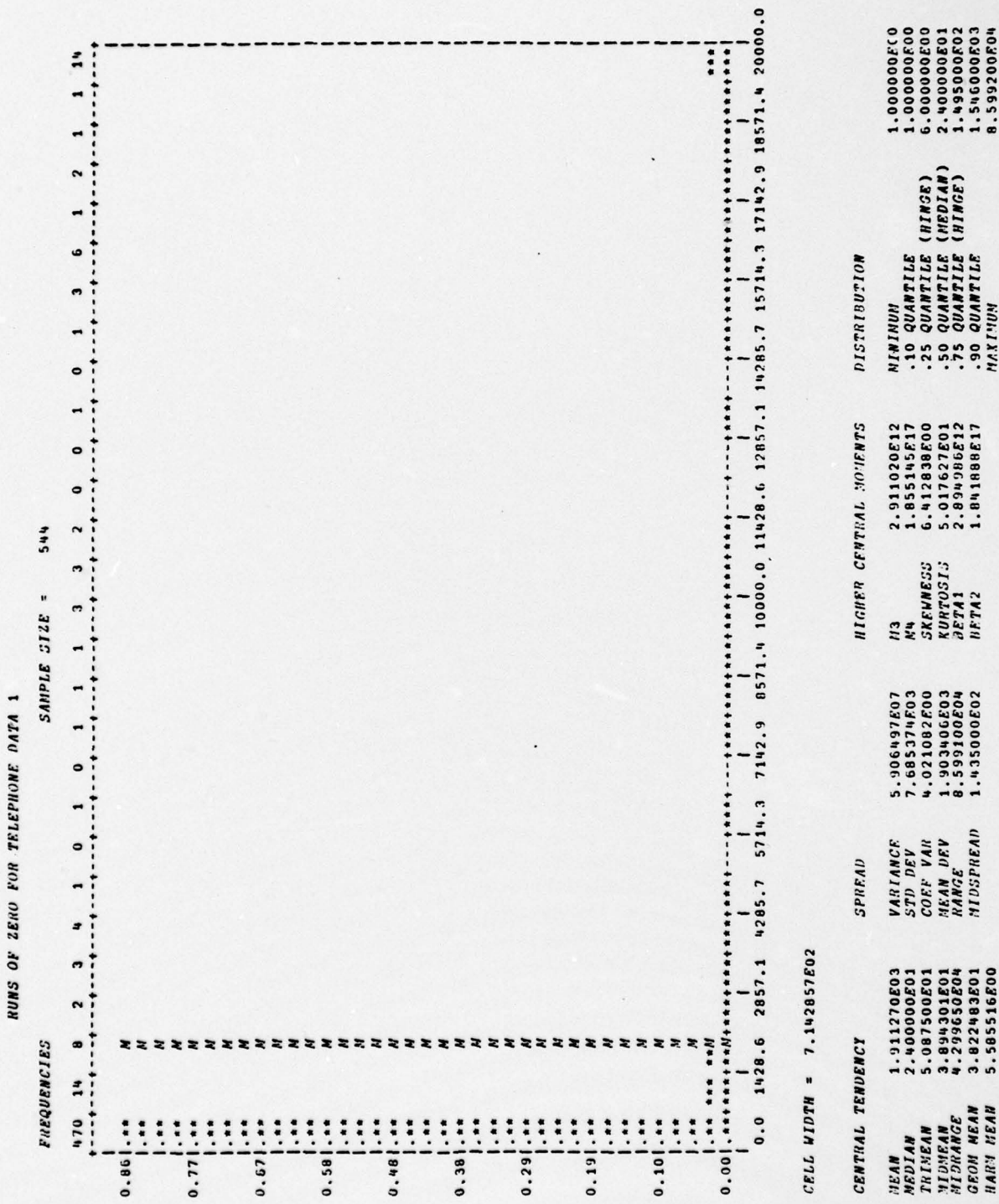


FIGURE 17

HISTLIST
 HISTLIST PRINTS THE SERIAL NUMBER OF THE COMPRESSED
 DATA, THE ORDERED DATA COMPRESSED, AND THE NUMBER OF
 LIKE OCCURENCES. ENTER YOUR DATA IN VECTOR FORM.
 U:

ONE

IF YOU WANT TO TITLE YOUR DATA TYPE YOUR TITLE.
 IF YOU DO NOT WANT A TITLE JUST HIT THE CARRIAGE
 RETURN.

RUNS OF ONE

IF YOU WANT YOUR OUTPUT TO APPEAR ON THE OFFLINE
 PRINTER TYPE 1 . IF YOU WANT YOUR OUTPUT TO APPEAR
 ON YOUR TERMINAL TYPE 0 .

U:

0

RUNS OF ONE

SERIAL NUMBER	ORDERED DATA	NUMBER OF OCCURENCES	PER CENT
1	1.000000	444 *****	0.816
445	2.000000	81 *****	0.149
526	3.000000	15 **	0.028
541	4.000000	1	0.002
542	5.000000	2	0.004
544	7.000000	1	0.002

FIGURE 18A

RUNS OF ZERO

SERIAL NUMBER	ORDERED DATA	NUMBER OF OCCURENCES	PER CENT
1	.CCCCC0	54	0.099
2	.CCCCC0	28	0.051
3	.CCCCC0	26	0.040
4	.CCCCC0	27	0.031
5	.CCCCC0	11	0.020
6	.CCCCC0	10	0.019
7	.CCCCC0	11	0.022
8	.CCCCC0	14	0.026
9	.CCCCC0	14	0.017
10	.CCCCC0	10	0.019
11	.CCCCC0	11	0.020
12	.CCCCC0	6	0.011
13	.CCCCC0	6	0.011
14	.CCCCC0	6	0.011
15	.CCCCC0	6	0.015
16	.CCCCC0	6	0.015
17	.CCCCC0	6	0.009
18	.CCCCC0	12	0.022
19	.CCCCC0	15	0.002
20	.CCCCC0	15	0.009
21	.CCCCC0	15	0.003
22	.CCCCC0	15	0.006
23	.CCCCC0	15	0.006
24	.CCCCC0	15	0.004
25	.CCCCC0	15	0.003
26	.CCCCC0	15	0.003
27	.CCCCC0	15	0.003
28	.CCCCC0	15	0.003
29	.CCCCC0	15	0.003
30	.CCCCC0	15	0.003
31	.CCCCC0	15	0.003
32	.CCCCC0	15	0.003
33	.CCCCC0	15	0.003
34	.CCCCC0	15	0.003
35	.CCCCC0	15	0.003
36	.CCCCC0	15	0.003
37	.CCCCC0	15	0.003
38	.CCCCC0	15	0.003
39	.CCCCC0	15	0.003
40	.CCCCC0	15	0.003
41	.CCCCC0	15	0.003
42	.CCCCC0	15	0.003
43	.CCCCC0	15	0.003
44	.CCCCC0	15	0.003
45	.CCCCC0	15	0.003
46	.CCCCC0	15	0.003
47	.CCCCC0	15	0.003
48	.CCCCC0	15	0.003
49	.CCCCC0	15	0.003
50	.CCCCC0	15	0.003
51	.CCCCC0	15	0.003
52	.CCCCC0	15	0.003
53	.CCCCC0	15	0.003
54	.CCCCC0	15	0.003
55	.CCCCC0	15	0.003
56	.CCCCC0	15	0.003
57	.CCCCC0	15	0.003
58	.CCCCC0	15	0.003
59	.CCCCC0	15	0.003
60	.CCCCC0	15	0.003
61	.CCCCC0	15	0.003
62	.CCCCC0	15	0.003
63	.CCCCC0	15	0.003
64	.CCCCC0	15	0.003
65	.CCCCC0	15	0.003
66	.CCCCC0	15	0.003
67	.CCCCC0	15	0.003
68	.CCCCC0	15	0.003
69	.CCCCC0	15	0.003
70	.CCCCC0	15	0.003
71	.CCCCC0	15	0.003
72	.CCCCC0	15	0.003
73	.CCCCC0	15	0.003
74	.CCCCC0	15	0.003
75	.CCCCC0	15	0.003
76	.CCCCC0	15	0.003
77	.CCCCC0	15	0.003
78	.CCCCC0	15	0.003
79	.CCCCC0	15	0.003
80	.CCCCC0	15	0.003
81	.CCCCC0	15	0.003
82	.CCCCC0	15	0.003
83	.CCCCC0	15	0.003
84	.CCCCC0	15	0.003
85	.CCCCC0	15	0.003
86	.CCCCC0	15	0.003
87	.CCCCC0	15	0.003
88	.CCCCC0	15	0.003
89	.CCCCC0	15	0.003
90	.CCCCC0	15	0.003
91	.CCCCC0	15	0.003
92	.CCCCC0	15	0.003
93	.CCCCC0	15	0.003
94	.CCCCC0	15	0.003
95	.CCCCC0	15	0.003
96	.CCCCC0	15	0.003
97	.CCCCC0	15	0.003
98	.CCCCC0	15	0.003
99	.CCCCC0	15	0.003
100	.CCCCC0	15	0.003

FIGURE 18C

457	33000000	1	0.002
456	33000000	1	0.002
455	33000000	1	0.002
454	33000000	1	0.002
453	33000000	1	0.002
452	33000000	1	0.002
451	33000000	1	0.002
450	33000000	1	0.002
449	33000000	1	0.002
448	33000000	1	0.002
447	33000000	1	0.002
446	33000000	1	0.002
445	33000000	1	0.002
444	33000000	1	0.002
443	33000000	1	0.002
442	33000000	1	0.002
441	33000000	1	0.002
440	33000000	1	0.002
439	33000000	1	0.002
438	33000000	1	0.002
437	33000000	1	0.002
436	33000000	1	0.002
435	33000000	1	0.002
434	33000000	1	0.002
433	33000000	1	0.002
432	33000000	1	0.002
431	33000000	1	0.002
430	33000000	1	0.002
429	33000000	1	0.002
428	33000000	1	0.002
427	33000000	1	0.002
426	33000000	1	0.002
425	33000000	1	0.002
424	33000000	1	0.002
423	33000000	1	0.002
422	33000000	1	0.002
421	33000000	1	0.002
420	33000000	1	0.002
419	33000000	1	0.002
418	33000000	1	0.002
417	33000000	1	0.002
416	33000000	1	0.002
415	33000000	1	0.002
414	33000000	1	0.002
413	33000000	1	0.002
412	33000000	1	0.002
411	33000000	1	0.002
410	33000000	1	0.002
409	33000000	1	0.002
408	33000000	1	0.002
407	33000000	1	0.002
406	33000000	1	0.002
405	33000000	1	0.002
404	33000000	1	0.002
403	33000000	1	0.002
402	33000000	1	0.002
401	33000000	1	0.002
400	33000000	1	0.002
399	33000000	1	0.002
398	33000000	1	0.002
397	33000000	1	0.002
396	33000000	1	0.002
395	33000000	1	0.002
394	33000000	1	0.002
393	33000000	1	0.002
392	33000000	1	0.002
391	33000000	1	0.002
390	33000000	1	0.002
389	33000000	1	0.002
388	33000000	1	0.002
387	33000000	1	0.002
386	33000000	1	0.002
385	33000000	1	0.002

If the runs of ones are geometric then $\text{prob}\{x(i)=k\} = (1-p)p^{k-1}$ $k=1,2,\dots$. Thus, this is the "geometric plus one" distribution.

$$\mu = E[X] = \frac{1}{(1-p)}$$

$$\sigma^2 = \text{VAR}[X] = \frac{1}{(1-p)^2}$$

$$C(X) = \frac{\text{VAR}[X]^{\frac{1}{2}}}{E[X]} = p^{\frac{1}{2}}$$

To find p set $E[X] = 1.235294 = 1/(1-p)$

$$p = \underline{.1904761}$$

Therefore, if the data is "geometric plus one" then

$$\begin{aligned} \text{EXPECTED VAR}[X] &= .1904761/ (.8095329)^2 \\ &= \underline{.2906572} \end{aligned}$$

Thus, the expected variance is .2906572 and the observed variance from HIST is .3460080. Also, the expected coefficient of variance is

$$\text{EXPECTED } C(X) = (.1904761)^{\frac{1}{2}} = \underline{.4364356}$$

And, the observed coefficient of variation is .4761817.

Therefore, at this point there seems to be a fairly close agreement between the runs of one and a "geometric plus one" distribution with $p = .1904761$.

As further proof a Chi-square test for goodness of fit was run on the runs. By using the formula

$$\text{prob}\{X = x\} = (1-p)p^{x-1} \text{ for } x=1,2,3,4,5,\dots$$

<u>PROBABILITY</u>	<u>EXPECTED</u>	<u>OBSERVED</u>
P(X=1) = .8095239	440.38	444
P(X=2) = .1541949	83.88	81
P(X=3) = .0293704	15.98	15
P(X=4) = .0055943	3.04	1
P(X=5) = .0010655	.58	2
P(X>6) = .0002510	.14	1
	19.74	19

Note, to use Chi-square not more than 20% of the cells should have expected frequencies less than 5 and no cell should have an expected frequency less than one. Therefore, the above frequencies must be combined into 3 cells.

$$\chi^2 = \sum_{i=1}^3 \frac{(\text{obs}_i - \text{ex}_i)^2}{\text{ex}_i} = \underline{.1562799}$$

And, $\chi^2_{.05,2} = 5.99$. Thus, the null hypothesis that the runs of one are "geometric plus one" with $p = .1904761$ can not be rejected.

A similar procedure was done with the runs of greater than one. By using figure 15 the following information can be obtained:

MEAN = 1911.27
 VARIANCE = 59,064,970
 COEF.VAR.= 4.021082

And, by using the same method as previously done and solving for p one gets $p = \underline{.9994767}$.

$$\text{EXPECTED VAR}[X] = .9994767 / (.0005233)^2 = \underline{3,651,213}$$

This expected variance differs greatly from the observed variance. Also, the expected coefficient of variation is

computed to be

$$\text{EXPECTED } C(X) = (.9994767)^{\frac{1}{2}} = \underline{.9997383}$$

This compares with the observed coefficient of variation of 4.021082 . Because of the gross departures of the variance and the coefficient of variation in the geometric hypothesis, one can conclude that the runs of length greater than 1 are not geometrically distributed.

IX. DOCUMENTATION ON ROUTINES

A. LOCATION IN APL LIBRARY

The descriptions and routines that have been presented are all available in the APL workspace library 2 DATALFNS . Providing the user is properly logged on the terminal and in the APL mode, all that is necessary is to type)LOAD 2 DATALFNS . If the user then types DESCRIBE, a short description of the six routines presented and instructions on how to obtain the detailed information that is available in each of the "HOW" variables would be printed.

B. WORKSPACE LOADING PROCEDURES

Each of the routines was designed to stand alone. That is, if the user desires just to use HIST , all that is necessary is to type)COPY 2 DATALFNS HISTGRP into a clear workspace. HISTGRP contains the principal routine HIST and only the additional routines necessary for HIST to operate. Thus, the user does not clutter his workspace with any unneeded functions. It is this group structure that maintains the orderliness of the workspace. And, the ability to copy a particular group into a clear workspace provides more space for data and executions of the functions.

The following is the group structure in library 2 DATALFNS .

<u>GROUP</u>	<u>PRINCIPAL ROUTINE</u>	<u>OTHER NECESSARY ROUTINES</u>	<u>VARIABLES</u>
HISTGRP	HIST	APLNAME, APLOT, AUTOS, CMS, DFT, ECDF, ECODE, EFT, OF, OUT, WRITE	
HISTLISTGRP	HISTLIST	APLNAME, CMS, ECODE, DFT, OF, OUT, WRITE	
HISTSGRP	HISTS	DFT, EFT	
HISTJACKGRP	HISTJACK	DFT, EFT, TOT	
EXPONPGRP	EXPONP	AND, AUTOSCALE, INITIAL, MLOT, MSGS, VS, MULTILOT, SET Δ AP, TICMARK	<u>BS</u>
NORMPGRP	NORMP	AND, AUTOSCALE, INITIAL, MLOT, MSGS, VS, MULTILOT, SET Δ AP, TICMARK	<u>BS</u>
DESCGRP (Descriptive group)			DESCRIBE, HISTHOW, HISTHOW, HISTLISTHOW, HISTJACKHOW, EXPONPHOW, NORMPHOW
VARIGRP (Variable group)			TELDAT1, TELDAT2, YROVR

C. ROUTINE LISTING

The above mentioned routines were either created by the author, adapted from existing fortran routine HISTG/F, or borrowed from the current APL library to supplement the author created routines.

1. Author Created Routines

HISTLIST, HISTS, HISTJACK, EXPONP, NORMP, APLLOT,
AUTOS, OUT, TOT

2. Adapted from Fortran Library Routine HISTG/F

HIST, ECDF

3. Borrowed Routines to Supplement Author Created Routines

AND, APLNAME, AUTOSCALE, CMS, DFT, ECODE, EFT,
INITIAL, MPLOT, MSGS, MULTILOT, NDTRI, OF, SETΔAP, TICMARK,
VS, WRITE


```

[44] TAB14: (A[11]<10)/LAST
[45] +(A[11]>28)/LAST
[46] DELTA+(HSCALE+A[3]-A[2])+A[1]
[47] XLABEL+A[2], (A[2]+DELTA*A[1])
[48] F+/(A[1])*. = [(X-A[2])+DELTA
[49] F[1]+/(X*XLABEL[1])+F[1]
[50] F[A[1]]+/(X>XLABEL[A[1]+1])+F[A[1]]
[51] C[9]+(C[8]+/(X-C[1])+(X/N)*2)+(N+pX)-1)*0.5
[52] C[2]+0.5*X/(X-C[1])+(X/N)*2, 1+{N+2}
[53] C[11]+/(X-C[2])*N
[54] C[22]+X[(N+10)]
[55] C[29]+N-C[28]+((N+4)
[56] M2+1+M1+(1-((4|M)+2))
[57] C[23]+(X[C[28]]+1)+(X[C[28]]*M1)+M2
[58] C[24]+C[2]
[59] C[25]+(X[C[29]]+(X[C[29]]+1)*M1)+M2
[60] C[26]+X[(0.9*N)]
[61] C[13]+C[25]-C[23]
[62] C[5]+(C[27]+C[21])*0.5
[63] C[3]+0.25*(C[23]+C[24]+C[24]+C[25])
[64] C[15]+((X-C[1])*3)*N+((N-1)*(N-2))
[65] C[19]+/(X-C[1])*3)+N
[66] C[16]+((X-C[1])*4)*((N-1)*(N-2)*(N-3))
[67] C[20]+/(X-C[1])*4)+N
[68] C[16]+C[16]-(C[8]*C[8]*3*(N-1)*(N+3))+((N*(N-2)*(N-3))
[69] C[18]+ 3+C[16]+(C[8]*C[8])
[70] C[29]+N+1-C[28]+2+{N+4}
[71] SUN+C[23]+C[25]
[72] SUNA+/(X[(C[28]-1)])
[73] SUNB+/(X[(C[29]])]
[74] SUN+SUM+(SUNB-SUNA)
[75] C[4]+SUM+(3+C[29]-C[28])
[76] C[17]+C[15]+C[9]*3
[77] C[6]+C[7]+0
[78] SUMA+5
[79] SUNB+7
[80] +(X[1]≤0)/TAB
[81] C[6]+((X)/(X))+N
[82] C[7]+N+/(X)
[83] SUNA+SUNB+7
[84] TAB: +(N>300)/TAB4
[85] +(N>C[14]+pA1+((pV)p(1M)≤1)/V+X[(V=M+{V+X}/X*. =X)/TAB2
[86] TAB4: C[14]+0
[87] .SUNB+6
[88] TAB2: C[10]+0

```

```

[89] +(CL1) < 1E 30) / TAB3
[90] C[10] + C[9] * IC[1]
[91] TAB3: E1 ← CENTRAL TENDENCY
[92] A1 ← MEAN MEDIAN SPREAD HIGHER CENTRAL MOMENTS DISTRIBUTION
[93] A2 ← VARIANCE STD DEV TRIMEAN MIDMEAN MIDRANGE GEOM MEAN HARM MEAN
[94] A3 ← M3 SKENNESS KURTOSIS BETA1 RANGE MIDSPEED MODE
[95] A4 ← MINIMUM .10 QUANTILE BETA2 .25 QUANTILE (HINGE) .50 QUANTILE (MEDIAN)
[96] A42 ← .75 QUANTILE (HINGE) .90 QUANTILE MAXIMUM
[97] A4 ← A41, A42
[98] E1 ← 13 7 EFT E2 ← ((SUMA), 1) P E1 ← C[11], C[12], C[13], C[14], C[15], C[16], C[17]
[99] E2 ← 13 7 EFT E3 ← ((SUMB), 1) P E2 ← C[18], C[19], C[10], C[11], C[12], C[13], C[14]
[100] E3 ← E3 + 6 1 P E3 ← C[15], C[16], C[17], C[18], C[19], C[20]
[101] E4 ← 13 7 EFT E4 + 7 1 P E4 ← C[21], C[22], C[23], C[24], C[25], C[26], C[27]
[102] A1 ← ((SUMA), 11) P A1
[103] A2 ← ((SUMB), 11) P A2
[104] A3 ← 7 10 P A3
[105] A4 ← 7 23 P A4
[106] C ← 7 4 P C + , ,
[107] A PLOT
[108] TAB5: OL OUT T ←
[109] OL OUT 'CELL WIDTH = ', 13 7 EFT DELTA
[110] E3 ← 7 13 P E3 + E3, T
[111] OL OUT 2 7 P T
[112] OL OUT T1
[113] OL OUT T
[114] + (SUMA = 5) / TAB6
[115] + (SUMB = 6) / TAB7
[116] TAB8: OL OUT A1, E1, C, A2, E2, C, A3, E3, C, A4, E4
[117] + (OL = 0) / 0
[118] CMS 'FINIS HIST APLPF'
[119] CMS 'O PRINTCC HIST APLPF'
[120] + (O = ECODE) / EX1
[121] 'HISTOGRAM SENT TO PRINTER'
[122] CMS 'ERASE HIST APLPF'
[123] → 0
[124] EX1: 0, P [ ] ← 'PRINTING FAILED. TRY AGAIN OR SEE APL PROGRAMMER.'
[125] TAB6: A1 ← 7 11 P A1
[126] E1 ← 7 13 P E1
[127] A1[6:] ← A1[7:] + E1[6:] + E1[7:] + , ,
[128] + (SUMB = 7) / TAB8
[129] TAB7: A2 ← 7 11 P A2
[130] E2 ← 7 13 P E2
[131] A2[7:] ← E2[7:] + , ,
[132] → TAB8
[133] LAST: NUMBER OF CELLS GIVEN IS NOT PERMISSABLE'

```

```

VHISTLIST[U]V
[1] HISTLIST;A;B;C;D;E;F;G;I;J;K;N;O;S;TT;APLN;OL
[2] 'HISTLIST PRINTS THE SERIAL NUMBER OF THE COMPRESSED
[3] 'DATA, THE ORDERED DATA COMPRESSED, AND THE NUMBER OF
[4] 'LINE OCCURENCES. ENTER YOUR DATA IN VECTOR FORM.'
[5] X+[]
[6] .
[7] 'IF YOU WANT TO TITLE YOUR DATA TYPE YOUR TITLE.'
[8] 'IF YOU DO NOT WANT A TITLE JUST HIT THE CARRIAGE'
[9] 'RETURN.'
[10] TT+[]
[11] .
[12] 'IF YOU WANT YOUR OUTPUT TO APPEAR ON THE OFFLINE'
[13] 'PRINTER TYPE 1 . IF YOU WANT YOUR OUTPUT TO APPEAR'
[14] 'ON YOUR TERMINAL TYPE 0 .'
[15] OL+[]
[16] +(OL=0)/TABA
[17] APLN+APLNAME 'HISTLIST APLPF P1 V'
[18] CMS 'ERASE HISTLIST APLPF'
[19] (20+1) WRITE APLN
[20] TABA:+(TT=0)/TAB10
[21] OL OUT 2 7 p'
[22] OL OUT TT
[23] OL OUT 2 7 p'
[24] TAB10:X+X[AX]
[25] C+(pX)p1
[26] K+1
[27] E+1p1
[28] J+0
[29] F+10
[30] TAB2:J+J+1
[31] +(X[J]=X[J+1])/TAB1
[32] K+K+1
[33] E+E,(J+1)
[34] F+F,X[J]
[35] +((J+1)=pX)/TAB3
[36] +TAB2
[37] TAB1:C[K]+C[K]+1
[38] +((J+1)=pX)/TAB3
[39] +TAB2

```

```

[40] TAB3:F+F,X[J+1]
[41] A'SERIAL NUMBER ORDERED DATA NUMBER OF OCCURRENCES
[42] B+L(0.5+60*((I/C)+(PX)))
[43] D+8[(B+2)
[44] B+(2[(B-4))P'
[45] OL OUT A.B,PER CENT
[46] OL OUT
[47] J+0
[48] TAB4:J+J+1
[49] G+L(0.5+60*I+(C[J]+(PX)))
[50] B+G+Dp'
[51] G+Gp'*
[52] I+3 DFT I
[53] S+ 10 0 DFT E[J]
[54] O+ 16 6 DFT F[J]
[55] N+ 10 0 DFT C[J]
[56] OL OUT S, 'O,' 'N,' 'G.B,I
[57] +(J=K)/TAB5
[58] +TAB4
[59] TAB5:+(OL=0)/O
[60] CMS 'FINIS HISTLIST APLPF'
[61] CMS 'O PRINTCC HISTLIST APLPF'
[62] +(O*ECODE)/EX1
[63] 'HISTLIST SENT TO PRINTER'
[64] CMS 'ERASE HISTLIST APLPF'
[65] +0
[66] EX1:0,P[]+'PRINTING FAILED. TRY AGAIN OR SEE APL PROGRAMMER.'

```

```

VHISTS[0]V
V HSTS: X: P: SE: ARRAY: J: FN: A: B: C: D: E: F: G: H: I: K: SD: VAR: MED: MIN: MAX: SDS: STS: KURT: SKEW: CVAR: MEAN: VRS: MNS: S2: M3: M4: N
'TYPE THE NUMBER OF SECTIONS YOU DESIRE ( INTEGER
'BETWEEN 2 AND 28 ) BE SURE TO PICK YOUR NUMBER OF
'SECTIONS SO AS TO MINIMIZE THE NUMBER OF DATA
'POINTS THAT WILL HAVE TO BE DISCARDED. (HISTS
'PLACES THE DATA INTO THE EQUAL NUMBER OF SECTIONS'
'YOU INDICATE DISCARDING ANY DATA LEFT OVER )'
SE=I
.
'ENTER YOUR DATA TO BE SECTIONED IN VECTOR FORM'
K=I
.
P=0
SDS+VRS+MNS+STS+7p0
TAB10: S2=I(PX)ISE
MEAN+VAR+SD+CVAR+SKEW+MED+MIN+KURT+MAX+(SE)P0
ARRAY+(ISE).(S2)PX
J=0
TAB3: J+J+1
+(J>SE)/TAB2
MAX[J]+I/(ARRAY[J: ])
MIN[J]+I/(ARRAY[J: ])
SD[J]+(VAR[J]+(+(/ARRAY[J: ]-MEAN[J])+(+/ARRAY[J: ])+N)*2)+(N-SZ)-1)*0.5
FN+PN[AFN]
MED[J]+0.5*(+/FN[(FN+2).1+(LN+2)])
M3+M4+(SE)P0
M3[J]+(+(/((ARRAY[J: ]-MEAN[J])*3))*N)+(N-1)*(N-2)
M4[J]+(+(/((ARRAY[J: ]-MEAN[J])*4))*N)*(3+N*(N-2))+(N-1)*(N-2)*(N-3)
M4[J]+M4[J]- (VAR[J]*VAR[J]*3*(N-1)*(N+N-3))+(N*(N-2)*(N-3))
SKEW[J]+M3[J]+SD[J]*3
KURT[J]+3+M4[J]+(VAR[J]*VAR[J])
CVAR[J]+SD[J]+MEAN[J]
+TAB3
TAB2: +(P=1)/TAB12
ARRAY+MEAN.MED.VAR.SD.CVAR.SKEW.KURT
ARRAY+(7.(SE))PARRAY
J=0
TAB4: J+J+1
SDS[J]+(VRS[J]+(+(/ARRAY[J: ]-MNS[J])+(+/ARRAY[J: ])+N)*2)+(N-SE)-1)*0.5
STS[J]+SDS[J]+I((N)*0.5)
+(J=7)/TABS
+TAB4
TABS: A+SECTION MEAN MEDIAN VARIANCE STD DEV COEF VAR
B+SKEWNESS KURTOSIS MINIMUM MAXIMUM
A.B

```

```

[46] ' '
[47] TAB12:J+0
[48] TAB6:J+J+1
[49] K+ 2 0 DFT J
[50] A+ 11 5 EFT MEAN[J]
[51] B+ 11 5 EFT MED[J]
[52] C+ 11 5 EFT VAR[J]
[53] D+ 11 5 EFT SD[J]
[54] E+ 11 5 EFT CVAR[J]
[55] F+ 11 5 EFT SKEW[J]
[56] G+ 11 5 EFT KURT[J]
[57] H+ 11 5 EFT MIN[J]
[58] I+ 11 5 EFT MAX[J]
[59] +(P=1)/TAB7
[60] ' ' 'K' '
[61] 'A.' 'B.' 'C.' 'D.' 'E.' 'F.' 'G.' 'H.' 'I'
[62] +(J=SE)/TAB11
[63] 'TAB6
[64] TAB11:P+SE+1
[65] 'TAB10
[66] TAB7: 2 7 P '
[67] 'UNSECTIONED' 'A.' 'B.' 'C.' 'D.' 'E.' 'F.' 'G.' 'H.' 'I'
[68] 2 1 P
[69] 'SUMMARY FOR SECTIONED DATA'
[70] ' '
[71] ' '
[72] ' '
[73] J+0
[74] TAB8:J+J+1
[75] A+ 11 5 EFT MNS[J]
[76] B+ 11 5 EFT VRS[J]
[77] C+ 11 5 EFT SDS[J]
[78] D+ 11 5 EFT STS[J]
[79] E+'MEAN
[80] E+ 7 12 P E
[81] E[J].A.' 'B.' 'C.' 'D
[82] +(J=7)/0
[83] 'TAB8

```

KURTOSIS

SKEWNESS

COEF VAR

STD DEV

VARIANCE

MEDIAN

MEAN

STD DEV

STD:(SECS)*.5'

```

V HISTJACK(U)
V HISTJACK;X:SEC1;PSV;SZ;A;B;C;J;G;ARRAY;BRRAY;K;FN;S;CVAR;KURT;MAX;MEAN;MEANS;MED;MIN;M3;M4;N;SD;SKEW;VAR;VARSA
'TYPE THE NUMBER OF GROUPS YOU DESIRE (INTEGER
BETWEEN 2 AND 50) BE SURE TO PICK YOUR NUMBER
OF GROUPS SO AS TO MINIMIZE THE NUMBER OF DATA
POINTS THAT WILL HAVE TO BE DISCARDED. (HISTJACK
PLACES THE DATA INTO THE EQUAL NUMBER OF GROUPS
YOU INDICATE DISCARDING ANY DATA LEFT OVER)'
SEC1=I
. .
. ENTER YOUR DATA TO BE JACKKNIFED IN VECTOR FORM'
X=I
. .
MEANS+VARSA+S+ 7 1 p0
PSV+((7). (SEC1))p0
SZ+(pX)-((I(pX),SEC1))
MEAN+VAR+SD+CVAR+SKEW+MED+MIN+KURT+MAX+(SEC1+1)p0
ARRAY+(1. (pX))pX
J+G+1
B+pX
TAB3:+(J>(SEC1+1))/TAB2
MAX(J)+I/(ARRAY[G;])
MIN(J)+I/(ARRAY[G;])
SD(J)+((VAR(J))-((I+I)/ARRAY[G;])-(MEAN(J))^2)/(N+1)*0.5
FN+ARRAY[G;]
FM+FM[AFN]
MED(J)+0.5*(+/(FN((N+2).1+I(N+2)))
M3+M4+(SEC1+1)p0
M3(J)+((+/(ARRAY[G;])-(MEAN(J))^2)*N)/((N-1)*(N-2))
M4(J)+((+/(ARRAY[G;])-(MEAN(J))^4)*N)/((3+N*(N-2))*((N-1)*(N-2))*(N-3))
M4(J)+M4(J)-(VAR(J))*VAR(J)*3*(N-1)*(N+3)/((N-1)*(N-2))*(N-3)
SKEW(J)+M3(J)+SD(J)*3
KURT(J)+ 3+M4(J)+((VAR(J))*VAR(J))
CVAR(J)+SD(J)+I*MEAN(J)
G+(J+J+1)-1
B+SZ
+(G>2)/TAB3
C+I(pX)+A+SEC1
BRRAY+(A). (C)pX
ARRAY+(A). (SZ)p0

```

```

[34J 101
[35] +TAB3
[36] TAB2:PSV[1;]+(A*MEAN[1])-(A-1)*MEAN[1+1SEC1])
[37] PSV[2;]+(A*MED[1])-(A-1)*MED[1+1SEC1])
[38] PSV[3;]+(A*VAR[1])-(A-1)*VAR[1+1SEC1])
[39] PSV[4;]+(A*SD[1])-(A-1)*SD[1+1SEC1])
[40] PSV[5;]+(A*CVAR[1])-(A-1)*CVAR[1+1SEC1])
[41] PSV[6;]+(A*SKEW[1])-(A-1)*SKEW[1+1SEC1])
[42] PSV[7;]+(A*KURT[1])-(A-1)*KURT[1+1SEC1])
[43] MEANS*((+/PSV)+A)
[44] VARSA*((+/PSV*2)-(((+/PSV)*2)+1SEC1))+1SEC1-1)
[45] S*((VARSA+SEC1)*0.5
[46] A+GROUP MEAN MEDIAN VARIANCE STD DEV COEF VAR
[47] C+SKEWNESS KURTOSIS MINIMUM MAXIMUM
[48] . .
[49] A.C
[50] . .
[51] J+1
[52] A+ 9 11 p' .
[53] TAB4:J+J+1
[54] A[1;]+ 11 5 EFT MEAN[J]
[55] A[2;]+ 11 5 EFT MED[J]
[56] A[3;]+ 11 5 EFT VAR[J]
[57] A[4;]+ 11 5 EFT SD[J]
[58] A[5;]+ 11 5 EFT CVAR[J]
[59] A[6;]+ 11 5 EFT SKEW[J]
[60] A[7;]+ 11 5 EFT KURT[J]
[61] A[8;]+ 11 5 EFT MIN[J]
[62] A[9;]+ 11 5 EFT MAX[J]
[63] K+ 2 0 DFT(J-1)
[64] + (J=1)/TAB6
[65] . .K. . 'A[1;]. 'A[2;]. 'A[3;]. 'A[4;]. 'A[5;]. 'A[6;]. 'A[7;]. 'A[8;]. 'A[9;]
[66] +(J=(SEC1+1))/TAB5
[67] +TAB4
[68] TAB5:J+0
[69] +TAB4

```

```

[70] TAB6: 2 1 p '
[71] 'UNGROUPED 'A[1:],' 'A[2:],' 'A[3:],' 'A[4:],' 'A[5:],' 'A[6:],' 'A[7:],' 'A[8:],' 'A[9:]
[72] 2 1 p '
[73] 'SUMMARY FOR JACKKNIFE DATA'
[74] '
[75] '          JACKKNIFE ESTIMATE      VARIANCE      (VAR*GROUPS)*.5'
[76] A+48P '
[77] A+ 'JACKKNIFE ESTIMATE OF STD DEV'
[78] A+ 'OF MEAN OF PSEUDO-VALUES'
[79] '
[80] A+ 'MEAN      MEDIAN      VARIANCE      STD DEV      COEF VAR      SKEWNESS      KURTOSIS '
[81] A+ 7 9 pA
[82] J+0
[83] C+ 3 11 p '
[84] TAB7: J+J+1
[85] C[1:]+ 11 5 EFT MEANS[J]
[86] C[2:]+ 11 5 EFT VARSA[J]
[87] C[3:]+ 11 5 EFT S[J]
[88] A[J:],' 'C[1:],' 'C[2:],' 'C[3:]
[89] +(J=7)/TAB8
[90] +TAB7
[91] TAB8:+0

```

```

VEXPONP[[]]V
EXPONP: X: J: TT: SM: EC: EL: TL: EC2: R90: US: GL: RT: PL: ST: Q
'EXPONP ORDERS THE DATA YOU GIVE AND COMPUTES THE
'EMPIRICAL LOG SURVIVER FUNCTION FOR THE DATA.'
'A PLOT OF THE LOG SURVIVER FUNCTION FOR THE DATA'
'IS THEN PRINTED TO SEE IF THERE IS A LINEAR FIT.'
'
'IF YOU WANT TO TITLE YOUR PLOT TYPE YOUR TITLE.'
'IF YOU DO NOT WANT A TITLE JUST HIT THE CARRIAGE.'
'RETURN.'
'
TT: B
'
'ENTER YOUR DATA IN VECTOR FORM.'
X: I
SM: 3 10
'
PC: ' *+0'
EL: 'ORDERED DATA'
TL: 'EXPONENTIAL SCORES'
PC2: ' *'
R90: US: GL: 0
ST: 1 1.25 1.5 2 2.5 3 4 5 7.5 10
RT: 10
RT: 10
X: I[AX]
Y: 0(((PAX)+1)-1((PAX)))+((PAX)+1))
(10 10 .(PAX)) MPELOT Y VS X
'
VNDTRI[[]]V
R: NDTRI P: W
R: 0 0
+((0.5<|P-0.5),0=|P-0.5)/L1,L2
W: 0((P1-P)*2)*0.5
L6: R[1]+(W-((+/(2.515517 0.802853 0.010328)*M*-1+13))+/(1 1.432788 0.189269 0.001308)*W*-1+14))*P-
0.5
R[2]+0.3989423**0.5*R[1]*R[1]
+0
L1: R[1]+(-P-0.5)*10*74
+0
L2: R[2]+0.3989423
'
VECDF[[]]V
ECDP: XN: BN: FN: LIN: I: FMA: J: U: LOW: XINC: MAXM1: TRY
XN: (BN+((W)*0.5)+C[12]): N
LOW: I+1
FMA: 10
XINC: HSCALE: (A[1]*4)
FN: (MAXM1-((A[1]*4)-1))P0
TAB60: +(I>MAXM1)/TAB50
U: A[2]+I*XINC
J: LOW
TAB61: +(J>N)/TAB53
TRY: BN*(U-X[J])
+(TRY>1)/TAB52
+(TRY<1)/TAB53
FNL: FNL[I]+1-|TRY
J: J+1
TAB61
TAB52: LOW-J
J: J+1
TAB61
TAB53: FNL[I]+FNL[I]*XN
FMA: FMA|FNL[I]
I: I+1
TAB60
TAB50: I+1
TAB62: +(I>MAXM1)/TAB54
LIN: I((D[1]+0.5)-(D[1]-1))*(FNL[I]+FMA))
ARRAY[LIN;I]+P'
I: I+1
TAB62
TAB54: +0
'

```

```

VEXPONP[[]]V
EXPONP: X: J: TT: SM: EC: EL: TL: EC2: R90: US: GL: RT: PL: ST: Q
'EXPONP ORDERS THE DATA YOU GIVE AND COMPUTES THE
'EMPIRICAL LOG SURVIVER FUNCTION FOR THE DATA.'
'A PLOT OF THE LOG SURVIVER FUNCTION FOR THE DATA'
'IS THEN PRINTED TO SEE IF THERE IS A LINEAR FIT.'
'
'IF YOU WANT TO TITLE YOUR PLOT TYPE YOUR TITLE.'
'IF YOU DO NOT WANT A TITLE JUST HIT THE CARRIAGE.'
'RETURN.'
'
TT: B
'
'ENTER YOUR DATA IN VECTOR FORM.'
X: I
SM: 3 10
'
PC: ' *+0'
EL: 'ORDERED DATA'
TL: 'EXPONENTIAL SCORES'
PC2: ' *'
R90: US: GL: 0
ST: 1 1.25 1.5 2 2.5 3 4 5 7.5 10
RT: 10
RT: 10
X: I[AX]
Y: 0(((PAX)+1)-1((PAX)))+((PAX)+1))
(10 10 .(PAX)) MPELOT Y VS X
'
VNDTRI[[]]V
R: NDTRI P: W
R: 0 0
+((0.5<|P-0.5),0=|P-0.5)/L1,L2
W: 0((P1-P)*2)*0.5
L6: R[1]+(W-((+/(2.515517 0.802853 0.010328)*M*-1+13))+/(1 1.432788 0.189269 0.001308)*W*-1+14))*P-
0.5
R[2]+0.3989423**0.5*R[1]*R[1]
+0
L1: R[1]+(-P-0.5)*10*74
+0
L2: R[2]+0.3989423
'

```

```

V NORMP[0]V
NORMP: X: TT: BL: SM: TL: PC2: I: J: R: R90: RS: QL: BT: PI: ST: D: EC
'NORMP ORDERS THE DATA YOU GIVE AND COMPUTES THE'
'INVERSE OF THE UNIT NORMAL CUMULATIVE DISTRIBU-'
'TION FOR THE DATA. A PLOT OF THE INVERSE OF THE'
'UNIT NORMAL CUMULATIVE DISTRIBUTION VS THE ORDER-'
'ED DATA IS THEN PRINTED TO SEE IF THERE IS A'
'LINEAR FIT.'
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V
'IF YOU WANT TO TITLE YOUR PLOT TYPE YOUR TITLE.'
'IF YOU DO NOT WANT A TITLE JUST HIT THE CARRIAGE'
'RETURN.'
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V
'ENTER YOUR DATA IN VECTOR FORM'
X+0
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V
'ORDERED DATA'
SM+ 3 10
R90+RS+QL+0
D+130
ST+ 1 1.25 1.5 2 2.5 3 4 5 7.5 10
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V
'NORMAL SCORES'
PC+*-1'
X+I[AX]
R+((PX),2)P0
I+((PX))+((PX)+1)
J+0
TAB3: J+J+1
R[J:]+NDTRI I[J]
+(J-(PX))/TAB2
+TAB3
TAB2: J+((R[1:1]),101P0
I+X[X[1]],X[1])+((X[(PX)]]-X[1])+100)*,100)
+(X[1]<0AX[(PX)]>0)/TAB4
(10 10 ,(PX),(101)) MPELOT J VS I
+TAB5
TAB4: J+J,R[1:1],(R[1:1])+((R[(PX):1]]-R[1:1]))*(48)*,148)
I+I+.49P0
(10 10 ,(PX),(101),(49)) MPELOT J VS I
TAB5: +0
V
V NORMP[0]V
NORMP: X: TT: BL: SM: TL: PC2: I: J: R: R90: RS: QL: BT: PI: ST: D: EC
'NORMP ORDERS THE DATA YOU GIVE AND COMPUTES THE'
'INVERSE OF THE UNIT NORMAL CUMULATIVE DISTRIBU-'
'TION FOR THE DATA. A PLOT OF THE INVERSE OF THE'
'UNIT NORMAL CUMULATIVE DISTRIBUTION VS THE ORDER-'
'ED DATA IS THEN PRINTED TO SEE IF THERE IS A'
'LINEAR FIT.'
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V
'IF YOU WANT TO TITLE YOUR PLOT TYPE YOUR TITLE.'
'IF YOU DO NOT WANT A TITLE JUST HIT THE CARRIAGE'
'RETURN.'
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V
'ENTER YOUR DATA IN VECTOR FORM'
X+0
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V
'NORMAL SCORES'
PC+*-1'
X+I[AX]
R+((PX),2)P0
I+((PX))+((PX)+1)
J+0
TAB3: J+J+1
R[J:]+NDTRI I[J]
+(J-(PX))/TAB2
+TAB3
TAB2: J+((R[1:1]),101P0
I+X[X[1]],X[1])+((X[(PX)]]-X[1])+100)*,100)
+(X[1]<0AX[(PX)]>0)/TAB4
(10 10 ,(PX),(101)) MPELOT J VS I
+TAB5
TAB4: J+J,R[1:1],(R[1:1])+((R[(PX):1]]-R[1:1]))*(48)*,148)
I+I+.49P0
(10 10 ,(PX),(101),(49)) MPELOT J VS I
TAB5: +0
V
V AUTOS[0]V
AUTOS
+((PX)>80)/TAB70
A[1]+16
+TAB71
TAB70: A[1]+L(28((PX)+5))
TAB71: A[2]+C[21]
A[3]+C[27]
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V EFT[0]V
EFT: X: D: E: H: J: K: L: Q: S: T: U: Y
D+.0123456789.E
+(V/W)*L*W+.W+(H+0)*L+1<PPX)/EFTERR+0*K+2
+(3 2 1 <PPX)/EFTERR+K+0), 2 3 +I26
X+((V/ 1 2 =PH)/ 1 2)*((1.0.X)PX
X+(PH*PX)PX
+((A/(PH)* 1 2 .2*E+1PH*PX),1=PH)/(EFTERR*K+1),2+I26
W+(W+6*(V/X<0)+V/.1>X),W
+(V/6>-[1] W-q(E.2)PW)/EFTERR+0*K+2
EFTLP: +(E<H+H+1)/EFTEND
[1] S+1+[10*(Y+0=Y+X:H)]
[2] U+1+[10*(Y+0=Y+10.5+(10*Q-15)+Y+10*(Q+W[2:H]))-S
[3] J+((T-4)P1),4P0)\1+10|(I|Y:10>U>Q).+10* 1+PH, 4+T+W[1:H]
[4] J[:T- 2 1]+1+10|(I|S-U>Q).+ 10 1
[5] J[: (10+T-4+Q).T]+13
[6] J[: (1.U.T.T-3)+Q(4,K)P(K011),(13*0>Y.S-1),K012
[7] J[: T-3]+J[: (1PH*U+1),(U+1+Q)]
[8] J[: T- 2 1 0]+(-S50)PHJ[: T- 2 1 0]
[9] +EFTLP.PZ[: (+/W[1:H-1])+T]+D[J]
[10] EFTEND:+L/0
[11] +0*PZ+.2
[12] EFTERR: EFT ,(3 6 P' RANK LENGTHDOMAIN')(K+1;), ' PROBLEM.'
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VAPLOT(UJV
V APLOT;I,J;LINE;CROB;PROB;VERT;H1;PLABEL;DIB;FSCALE;DID;DIT;DIS;IQ1T;IQ2T;IQ3T;NMAX;MNT;RT;PRBMX;INCR;ARRAY;FMAX
-(POTT)=(P,0)/TAB5A
OL OUT(18P',),TT
OL OUT 1P'
TAB5A:ARRAY+(D[1]),(4*A[1])P',
FSCALE=(D[1]-1)*FMAX*(F[A])P[F]
I←0
J←3
TAB12:→(I=A[1])/TAB15
I←I+1
J←J+4
LINE←1((D[1]-0.5)-(FSCALE*F[I]))
→(LINE*(D[1]-1))/TAB13
ARRAY[LINE+(D[1]-LINE);J]+*
ARRAY[LINE+(D[1]-LINE);J+1]+*
ARRAY[LINE+(D[1]-LINE);J+2]+*
ARRAY[LINE+(D[1]-LINE);J+3]+*
ARRAY[(D[1]);J+3]+*
→TAB12
TAB13:→(F[I]≠0)/TAB14
ARRAY[(D[1]);J]+*
ARRAY[(D[1]);J+1]+*
ARRAY[(D[1]);J+2]+*
ARRAY[(D[1]);J+3]+*
→TAB12
TAB14:ARRAY[(D[1]);J]+*
ARRAY[(D[1]);J+1]+*
ARRAY[(D[1]);J+2]+*
ARRAY[(D[1]);J+3]+*
→TAB12
TAB15:PROB+(D[1],4)P',
INCR+(PRBMX+FMAX*M)19
CROB+PRBMX.(PRBMX-INCR*18).0
CROB+ 4 2 DFT CROB+ 10 1 P CROB
PROB[D[2]+D[3]*10;]←CROB[110;]
VERT+(D[1],1)P',
RT+(NMAX+A[1])*4+(A[3]-A[2])
IQ1T+(0.5+(C[23]-A[2])*RT)
IQ2T+(0.5+(C[24]-A[2])*RT)
IQ3T+(0.5+(C[25]-A[2])*RT)
MNT+L(0.5+(C[1]-A[2])*RT)
→(MNT*1)/TAB21
MNT+1
TAB21:→(MNT*NMAX)/TAB22
MNT+NMAX
TAB22:→(IQ1T*1)/TAB23

```

```

[46] IQ1T+1
[47] TAB23:-(IQ2T>1)/TAB24
[48] IQ2T+1
[49] TAB24:-(IQ3T>1)/TAB25
[50] IQ3T+1
[51] TAB25:-(IQ1T<NMAX)/TAB26
[52] IQ1T+NMAX
[53] TAB26:-(IQ2T<NMAX)/TAB27
[54] IQ2T+NMAX
[55] TAB27:-(IQ3T<NMAX)/TAB28
[56] IQ3T+NMAX
[57] TAB28:ARRAY[(D[1]);IQ1T]+'. '
[58] ARRAY[(D[1]);IQ2T]+'. '
[59] ARRAY[(D[1]);IQ3T]+'. '
[60] ARRAY[(D[1]);MNT]+'.M'
[61] ARRAY[(D[1]);NMAX]+'.|'
[62] +(B=0)/TAB34
[63] ECDF
[64] TAB3A:H1+(5p' ').H1+'FREQUENCIES',H1+(32p' '),H1+'SAMPLE SIZE = '
[65] OL OUT H1, 6 0 DFT(pX)
[66] OL OUT 1p' '
[67] OL OUT H1+(4p' ').H1+ 4 0 DFT F
[68] OL OUT ' ',+(4*A[1])p'---+'
[69] OL OUT ' ',H1+((4*(A[1]-1))p' '),H1+' |'
[70] OL OUT PROB,VERT,ARRAY
[71] DIS+1((pXLABEL)+2)
[72] DIT+(pXLABEL)+2
[73] +((DIT-DIS)=0)/TAB40
[74] TAB41:DIS+(8*DIS)p' |
[75] OL OUT ' ',DID,|'
[76] +TAB42
[77] TAB40:DIS+DIS-1
[78] +TAB41
[79] TAB42:DIS+DIS+1
[80] XLABEL+XLABEL_1+2*1DIB]
[81] +([XLABEL>99999)/TAB31
[82] +([XLABEL< 99999)/TAB31
[83] +(DELTA<0.1)/TAB31
[84] LABEL+((pXLABEL).1)pXLABEL
[85] OL OUT ' ',PLABEL+PLABEL+(PLABEL+ 7 1 DFT PLABEL),((pXLABEL).1)p' '
[86] +0
[87] TAB31:XLABEL+XLABEL_1+2*1([((pXLABEL)+2)))
[88] LABEL+((pXLABEL).1)pXLABEL
[89] OL OUT ' ',PLABEL+PLABEL+(10 4 EFT PLABEL),((pXLABEL).6)p' '
[90] +0

```

```

V MULTIPLOT [ ] V
MULTIPLY I; J; L; T; FT; U; K; M; N; L; L1; L2; L3; L4; L5; E; TM; HM; Q; Q1; Q8; R
D=2L1+C.613,C+6I 3 120
MSG 'OFF'
+(-R90)/PL2.ST+6pK+R+0
(SM[2]-p.K+H[1]) TICMARKOP[2]+1,
PL2:L+(1.QLA.AHM+0=(SM[2]+1 2)0.1.H[2])\2
8 TICMARKpL3+PL3+1+HS.L2+P+1Q+~HS.pC+H[1]
L5+P-1+HS.L4+Q+1-P2.pM+(pP)[1]-1
L1+((HS.HS+A+0>Q8+D-8)^(A)/PL4+ 2 1
TM+TM.[1.5] TM+1+TM
PL3:E+-I+0A+L+K+R
+(L1+1Q<pL[D+1+D/X[;2]]+J+2+(D+X[;1]=N+|K-C)/A).L2
+P.L[TT]+(E/L)[T+(RAL51)/1oL]
D+(E/1oE)[T.(U+L[TT]>2)/T~D]
L[(-U)/T]+(-U,U/1)/J+J,U/M+1
PE4:+(AV120D)/P.E+1vXL+L
+(L5+1~v/U+(P2VT+J+1QJ+1,J[TT])^D+0.D[T+U[D(U+AJ)]])L4
+(A/U+1+/(D0.=D+U/D)^J0.=J)/Q.J[T/1oJ+U/J]+(T+U/~T)/M-1
U+1+(T=1QJ+1,J+J[TT])*U=1QJ+0.D+D[T+U[D(U+AJ)]]
Q:+P+1Q8<pT+~(1+T+pJ)εD+(D-1).D+D+2x1T+pD+U/D
I+T/I
I[D]+(U/M),U/J
L+I[(E+~xI)\L
E:+(xN)/PL5+v/T+(2 1 xN)ε.TM
+(PL5+1).L+L[E\1+1.v/HM
PL5:L+L[0.T.((1 4+pL)0QLA+1T),0
PT[TM[;1]M;].P[1+L]
+(05R+xC+C-1)/L3
(SM[2]-1) TICMARK-R90
+U+(ST[3 4],1)/ 1 3 4 +I26
'SCALE FACTOR FOR ORDINATE: '10*ST[5]
+U+1+U
'SCALE FACTOR FOR ABSCISSA: '10*ST[6]
MSG 'ON'

```

```

VDFTE[0]V
V Z+W DFT X;D;E;F;G;H;I;J;K;L;Y
D+ 0123456789.-
+(V/H=L+H+(H+0)*L+1<0X)/DFTE+0*E+2
+(3 2 1 <0X)/(DFTE+P+0). 2 3 +I26
+(2+I26).0X+(V/ 1 2 =0W)0 1 2)0(1.0.X)0X
X+(0 1 1 /0X)0X
+((A/(0W)± 1 2 .2*E+100X).1=0W)/(DFTE+P+1).3+I26
I+1/0..[100|X+1>|X
W+(2+I+W+(W=0)+V/X<0).W
+(V/2>-/[1] W+0(E.2)0W)/DFTE+0*P+2
Z+((K-100X).+W[1;])0
X+(0.5+X*10*(0X)0W[2;])
DFTLP:+(E<H+H+1)/DFTE
J+1+10(|Y+X[;H])+.510*1+0+I+W[1;H]
J+(J)*G+.0(00J)0(.0(J=1)0.A(1I)).51I-F+1).(K*1+F+W[2;H])0 1
+(A/05Y)/2+I26
J[1+(0J)] 1+(I-+/(K.I)0G)+I* 1+K]+12*Y<0
J+(K.I)0J
+(0=F)/3+I26
J+J[(100G).(G+-/W[;H])±IF]
J[;G]+11
+DFTLP.0Z[:(+ /W[1;H-1])±I]+D[1+J]
DFTE+L/0
+0*02+.Z
DFTE:'DFT',(3 6 0' RANK LENGTHDOMAIN')(F+1;). ' PROBLEM.'
V
VAPLNAME[0]V
V FID+APLNAME A;K;REM
A REMOVE EXTRA BLANKS
A-1+(K+0X+- 'A)/A+- 'A.'
A FIND END OF FILENAME
K+(A.' ')-11
A IF ONE WORD - SYNTAX ERROR
+(K=0A)/ER1
A EXTRACT FILENAME
FID+8+K+A
A AND REMAINDER
REN+(K+1)+A
A FIND END OF FILETYPE
K+(REN.' ')-11
A ADD FILETYPE TO FILE
FID+FID.(8+K+REN)
A EXTRACT 2ND REMAINDER
REN+(K+1)+REN
A CHECK SPECIAL MODES
+((A/'SY'=2+REN)V(A/'* '=2+REN))/L1
A NODELETTER='P' UNLESS OTHERWISE
FID+FID.'ABCTP'['ABCT'1+REN]
A NODENO='1' UNLESS OTHERWISE
FID+FID.'0234561'['023456'11+1+REN]
-L2
L1:FID+FID.2+REN
A RECTYPE='P' UNLESS V SPECIFIED
L2:FID+FID.' '.'FV'[( 'V'=-1+REN)+11]
A CONVERT TO EBCDIC INTEGER
FID+2 OF FID
-L29
-0
ER1:'FILETYPE MISSING'
V

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VAUTOSCALE[0]V
V AUTOSCALE;C;D
C+C+(X[1;]+X[1;]=0)*0=C+(I/X)-D+I/X
F+F+G+10+1100P+|C+H+SM*L((6.PI)*20A)+SM+[16]SM 1 5
F+G*ST[+/F+.SST+10.ST[127]]
X+(C0 0.5)+(C0F)*X-(C+0X)0G-G*(0<G-C+C)0>G+C+|D+C+F+SM+2
V

```

```

VMPLOT[]V
A MPLOT X;C;D;F;G;H;P;P2;HS;A;ST
INITIAL
AUTOSCALE
SETAAP
MULTIPLY H+SM*(1/X-LX)+SM

```

```

VOUT[]V
OL OUT R;I;J;MAX
+(2=ppR)/L1
R+(1.pR)pr
L1:+(OL=1)/OFF
+0.pU+R
OFF:MAX+1+pr+ ,R
J+20[1+pr
I+1
L2:(J+R(I;)) WRITE APLN
+(MAX>I+1)/L2

```

```

VINITIAL[]V
INITIAL
+(0=x/(2pA).D+px), 2 1 <ppX)/0.PL2- 1 0
+PL2.D+px-(1.p.X).[1.5] X
X+(D-2+D)px
PL2:X+R90φ(φ 0 1 +X).[1.5](C+x/D+D- 0 1)px[;1]

```

```

VAND[]V
L+A AND B;C;D
+((2=ppA)V3=ppB).0=ppB)/ 17 3
B+.B
+((3=ppB)^1=1ppB).2=ppA)/ 17 7
A+.A
+Λ/((pA)=1.D).1=D+1p-2φpB)/16
A+((DxρA)D[ρA].1)ρA
+(1=ppB)/9
B+((pB)[(1=ppB)×1ppA].1)ρB
+((Λ/D×1.1ppA).1=D+1p-2φpB)/ 16 11
B+((3=ppB)ρ1).(1ppA).1φpB)ρB
+(3=ppB)/14
L+((C+1φpA)ρ0).(1φpB)ρ1)\B
+0xprL[;1C]+A
L+(1.((C+1φpA)ρ0).(1+1φpB)ρ1)\B
+0xprL[;1+1C]+A
+0=pr\ARGUMENTS OF AND ARE NOT CONFORMABLE.'
'AN ARGUMENT OF AND IS OF IMPROPER RANK.'

```

```

VTOT[]V
TOT
ARRAY[1;]+.BRRAY[1+(A-1);]
K+1
TAB1:K+K+1
ARRAY[K;]+((.BRRAY[K+(A-K);])).(.BRRAY[(K-1);])
+(K=(A-1))/TAB2
+TAB1
TAB2:ARRAY[A;]+.BRRAY[(A-1);]
+0

```

```

VTICMARK[]V
V U TICMARK ISV;C:I;J:L;T;NB;VT;O;N;E;K
  [1] + (PL3 - I - K + R90 = ISV), BP * 1 - ISV * 0 1
  [2] (X.P.TI)/BS[2], ((OF[0.5 * H[2]] - P.TI)P', ), IT.1 + RE
  [3] ((R90 * 0 [8] + 10.5 * H[2]) - P.TI)P', ), TI
  [4] C + pNB + G[J] + F[J] * M + 0.5 * M[J] * x [1] * H[J] + SM[J + 2 - ISV]
  [5] PL3: N = (C.VT + V/N)P (N + 0 > PT - [NB * 10 * U - I + [10 * E + [10 * (NB + NB = 0)] \ '- '
  [6] L * (U - U - N + -VT) + 1 - VT + (P' 0' A. = PT * N, ' 0123456789' [1 + Q(U * P10) * I * PT]) : 0
  [7] + ((U > T + VT + I / I, (L + L * I), (I * 50) * 2 + L - I), L * 2 * U - VT + L * I) / 3 2 + I * 26
  [8] + (I * 26) - I * P, ST[6 - K] + I - 1
  [9] + PL3.pNB + (10 * - L) * (L + 10.5 * NB * 10 * (L + 3 * U + U - VT) - ST[6 - K] + [ / (NB * 0) / E
  [10] PT + ( - (1 * U + J) * (I + J + VT), 1 1 + J + U - T) \ (C, U)P (, VT * 0) \ (, O - (I - O) * < N * U) / , PT
  [11] ST[4 - K] + ST[6 - K] * P * PT [ ; I + J + VT] + ' .
  [12] + ISV / 0 * P * PT * ((C + 1), ~VT)P', ), (1 1 + C, U) + PT
  [13] (SM[2] - 9)P', ), PT
  [14] + R90 / 0
  [15] BP: (( (~R90) * 0 [8] + 10.5 * H[2]) - P.BL)P', ), BL
  [16] (X.P.BI)/BS[2], ((OF[0.5 * H[2]] - P.BI)P', ), BI.1 + RE
V

```

```

VSETAAP[]V
V SETAAP
  [1] D + pA + (A > 0) / A + A, C + / A + 2 * A, D[2]P D[1]
  [2] + (D > 1) / 4, pP + ' _1', (D * PE), ((P2 + (X.P.PE2) ^ ~HS + RE ^ ~R90) + PE2), 1 + RE
  [3] + A - pA + ~P2
  [4] A + 1 + / (1 * C) * . > (D * . zD + 1 * D) * * A
V

```

```

VVSE[]V
V M * A VS B; C; D
  [1] + (( (pB + B) < pPB), 2 1 0 < pPA) / 8 8 4 3
  [2] A + ((pB), 1) * pA
  [3] A + ((x / pA), 1) * pA
  [4] + (A / (pB) * 1, 1 * pPA) / 9
  [5] M + (0, (1 * pPA) * 1) \ A
  [6] M [ ; 1 ] + B
  [7] + 0 * pPM + (1, pM) * pM
  [8] + 0 = 0 \ 'AN ARGUMENT OF VS IS OF IMPROPER RANK.'
  [9] 'ARGUMENTS OF VS ARE NOT CONFORMABLE.'
V

```

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