

AD-A042 639

MASSACHUSETTS INST OF TECH CAMBRIDGE DEPT OF CHEMISTRY
HAZARDS IN FACTOR ANALYSIS.(U)

F/G 12/1

JUL 77 C G SWAIN, H E BRYNDZA, M S SWAIN

N00014-67-A-0204-0075
NL

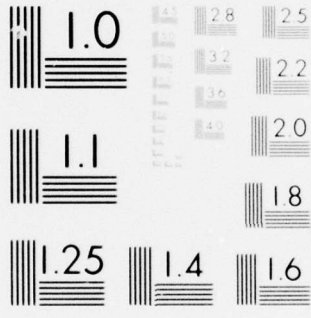
UNCLASSIFIED

12

| OF |
AD
A042639



END
DATE
FILMED
8-77
DDC



MICROCOPY RESOLUTION TEST CHART
NATIONAL BUREAU OF STANDARDS-1963-A

SECURITY CLASSIFICATION OF THIS PAGE (When Data Entered)

11 B.S.

REPORT DOCUMENTATION PAGE

READ INSTRUCTIONS BEFORE COMPLETING FORM

1. REPORT NUMBER 12, TR-2	2. GOVT ACCESSION NO.	3. RECIPIENT'S CATALOG NUMBER
4. TITLE (and Subtitle) HAZARDS IN FACTOR ANALYSIS.		5. TYPE OF REPORT & PERIOD COVERED Technical May 1, 1973-April 30, 1976
7. AUTHOR(s) C. Gardner/Swain, Henry E./Bryndza Marguerite S./Swain		6. PERFORMING ORG. REPORT NUMBER TR-2
9. PERFORMING ORGANIZATION NAME AND ADDRESS Department of Chemistry Massachusetts Institute of Technology Cambridge, Massachusetts 02139		8. CONTRACT OR GRANT NUMBER(s) N00014-67-A-0204-0075
11. CONTROLLING OFFICE NAME AND ADDRESS Office of Naval Research, Chemistry Program 800 North Quincy Street Arlington, Virginia 22217		10. PROGRAM ELEMENT, PROJECT, TASK AREA & WORK UNIT NUMBERS NR051-566
14. MONITORING AGENCY NAME & ADDRESS (if different from Controlling Office)		12. REPORT DATE 15 July 1977
		13. NUMBER OF PAGES 16
		15. SECURITY CLASS. (of this report) Unclassified
		15a. DECLASSIFICATION/DOWNGRADING SCHEDULE

16. DISTRIBUTION STATEMENT (of this Report)
Distribution of this document is unlimited
Technical rept. 1 May 73-30 Apr 76

DDDC
REF ID: A66116
AUG 10 1977
A

17. DISTRIBUTION STATEMENT (of the abstract entered in Block 20, if different from Report)

18. SUPPLEMENTARY NOTES
Technical Report No. 1 is submitted with this report. Both will be submitted to Science (A.A.A.S.).

19. KEY WORDS (Continue on reverse side if necessary and identify by block number)
Factor analysis
principal components
missing data
statistical analysis

20. ABSTRACT (Continue on reverse side if necessary and identify by block number)
Principal components and other standard factor analyses can yield misleading results if an assumed subsidiary condition is untrue or if data are missing. Principal components factor analysis is tested for its reliability, using a problem with known answers. Even when test data are complete (e.g., 70 data on 7 properties, dependent on both radii and heights of 10 cylinders), such analyses followed by varimax or other standard rotations give incorrect rank orders for factors (factor scores for radius and height for each of the

(OVER)

next page mt

DD FORM 1 JAN 73 1473

EDITION OF 1 NOV 65 IS OBSOLETE
S/N 0102-014-6601

220 007

SECURITY CLASSIFICATION OF THIS PAGE (When Data Entered)

ADA 042639

DDC FILE COPY

10 cylinders) and sensitivities (factor loadings for each of the 7 properties). When no data are missing, a transformation incorporating valid subsidiary conditions can be used instead of such rotations to obtain correct factors and sensitivities. However, when a moderate number (e.g., 20, or 29%) of the possible data are missing (randomly deleted), factors and sensitivities can have wrong rank orders and therefore be misleading even with this transformation. When data are missing, standard factor analysis is evidently unreliable and should be replaced by another method, such as that in the preceding article.

→ such as the one presented herein. ↑

Hazards in Factor Analysis

Principal components and other standard factor analyses can yield misleading results if an assumed subsidiary condition is untrue or if data are missing.

C. Gardner Swain, Henry E. Bryndza, Marguerite S. Swain

TR-2

Dr. C.G. Swain is a professor of chemistry at the Massachusetts Institute of Technology, Cambridge, Mass. 02139. Mr. Bryndza is a senior student at M.I.T. Dr. M.S. Swain is a postdoctoral fellow at M.I.T.

ACCESSION FOR	
NTIS	White Section <input checked="" type="checkbox"/>
DDC	Buff Section <input type="checkbox"/>
UNANNOUNCED	<input type="checkbox"/>
JUSTIFICATION	
.....	
SI	
DISTRIBUTION/AVAILABILITY CODES	
Dist.	AVAIL. and/or SPECIAL
A	

Factor analysis is now widely used to identify and evaluate significant underlying factors from masses of numerical data (1). The methods were developed mainly by psychologists. By 1966 more than 2000 publications using factor analysis had appeared (3). Since then, applications have mushroomed into such fields as psychiatry (4), medicine (5), anthropology and biology (6), chemistry (7), geology (8) and national and international relations including military strategy (9). Textbooks on factor analysis abound (10,11). Applications are facilitated by the ready availability of computer programs for factor analysis in SPSS (12), BMD and BMDP (13) DATA-TEXT (14), P-STAT (15), SSP (16), IMSL (17), TSP (18), and other software packages for statistical analysis.

Factor analysis was designed to find the underlying determiners (factors) that can account for the correlations between the different sorts (series) of data measured. Recently we became concerned about whether or not the standard factor analysis procedures do in fact succeed in this goal, and if not, what modifications are needed to make them succeed. It is difficult or impossible to answer these questions from examination of past applications because the true underlying factors are generally not known in real problems. Thurstone's tests (19,20) were not adequate, as we shall show below. Therefore we devised the following synthetic problem where a correct set of underlying factors can be distinguished from an incorrect set.

We attempted to analyze data on seven properties of right circular cylinders (areas, masses, moments of inertia, etc.), all of which are determined precisely by only two underlying types of factors (radii r and heights h). To be successful, a factor analysis must separate these factors. By analysis of these data, it must calculate a number (factor) for each cylinder that is a pure measure of cylinder radius. This may be r , $2\pi r$, $4\pi r^2$, $\log r$, $2 \log r$ or $k \log r$, but must not be rh , r/h or any other mixture or hybrid of r and h . It must calculate a second factor for each cylinder that is a pure measure of cylinder height. This may be h , $\log h$ or any other fixed function of h that is unaffected by r . It must not calculate or indicate any additional factors. Success or failure in this simple synthetic example is therefore clearly defined, and can be used to test the standard procedures commonly employed and various modifications of them. We shall show that the standard procedures fail to meet this challenge, but that success can be attained by adding a novel kind of transformation (rotation), provided that there are no missing data.

It has been recognized that the principal problem with the use of factor analysis lies in discovering an appropriate transformation (rotation) capable of

separating the factors into pure (unhybridized) types. Rotation of axes and variation of the angle between axes have been tried in various combinations, but they have not been logically derived from the specific problem and do not give correct parameters. We shall show that the transformation equations must be derived from valid (true) subsidiary information related to the problem at hand. Unfortunately, no one has derived transformation equations from such problem-related subsidiary conditions previously. Instead, standard transformations such as varimax rotations have generally been used. Since we shall show that these fail to separate the factors in our simple example, we believe that they may have failed to separate them in all previous applications.

It has been recognized that missing data are undesirable. Gaps in the input data matrix introduce serious errors into the correlation coefficient matrix, well in advance of the factor extraction and transformation procedures. Unfortunately, many "factor analyses" have been carried out in spite of missing data. SPSS provides three different options for coping with missing data. We shall show that factors are not separated correctly when data are missing, using any of the standard methods for handling missing data, even with our novel transformation which works correctly with a full data set.

Our findings are presented under four headings: (1) selection of a problem suitable for testing factor analysis, one where success or failure can be recognized but otherwise similar to a real application; (2) the transformation needed after the initial extraction process; (3) a demonstration that extraction followed by this transformation yields correct parameters when no data are missing; (4) a demonstration that seriously erroneous and misleading parameters can result if a moderate fraction of the pertinent data are missing (e.g., owing to their never having been measured, or to lost records, or incomplete returns, or deliberate omission due to concern about accuracy) even if all of the used data are perfectly accurate.

Selection of a Test Problem

The hazards in factor analysis first became apparent to us when we attempted to apply principal components (1) to a chemical problem, to identify and evaluate the two solvent characteristics most responsible for changes in thermodynamic, kinetic and spectral data when the solvent is varied. Missing data (never measured) were extensive in this problem, comprising 75% of the possible combinations of the 25 reactions and 60 solvents considered, but the data that were available covered

ranges of more than a power of ten for each reaction and were believed to be individually accurate and precise to about $\pm 15\%$. Principal components gave solvent parameter values in absurd rank orders using any of the programmed procedures for handling missing data. Therefore we undertook to examine the method further, using a similar but simpler problem. Since logarithms of the measurements constituting the data in our chemical problem appeared to be linear functions of two solvent factors (anion-stabilizing ability and cation-stabilizing ability), we chose a test problem where logarithms of the data were known to be accurately linear functions of two factors, i.e., the cylinder problem that was solved correctly, without or with missing data, by our "DOVE" (dual obligate vector evaluation) procedure outlined in the previous article (2).

Thurstone attempted to justify factor analysis by means of a cylinder problem more than 30 years ago (19), in one of the few recorded tests (20) that involved applying it to a problem with answers known in advance, but neither his cylinder problem nor any of the other test problems had any missing data. Thurstone's problem was based on diameters and heights carefully preselected to have zero correlation in spite of small sample size. Real data samples seldom have this property. Nevertheless, his selection of an application to cylinder properties was an inspired choice for testing factor analyses because it embodies the most essential features of a real two-mode problem in an especially simple and easily understood form.

The Transformation

Slopes ($\underline{a}_1, \underline{b}_1$) and factors ($\underline{x}_j, \underline{y}_j$) obtained initially from principal component analysis with two modes ($\underline{n}=2$) correspond to formulation of the data as $\underline{p}_{1j} = \underline{a}_1 \underline{x}_j + \underline{b}_1 \underline{y}_j$ under the conditions that the factors are orthogonal, i.e. $\sum_j \underline{x}_j \underline{y}_j = 0$ for all j for which data exist (e.g., for our 10 cylinders), and that the slopes are normalized, i.e. $\underline{a}_1^2 + \underline{b}_1^2 = 1$ for each \underline{i} (property). However, this orthogonality condition is undesirable, and a condition of zero covariance or zero correlation coefficient between the factors would also be undesirable, because none of these is likely to be even approximately true for a small sample (only 10 cases). Furthermore, in many other problems none of these would be a good approximation even for thousands or even an infinite number of cases. For example, in our solvent effect problem a weak negative correlation between the two factors is expected and perfectly reasonable. The second condition, $\underline{a}_1^2 + \underline{b}_1^2 = 1$, is also generally untrue. Therefore we should replace these two untrue conditions by two valid conditions.

Our transformation is carried out as follows. The old \underline{a}_1 and \underline{b}_1 values from factor analysis are unstandardized to make them comparable with the observed data, \underline{z}_{1j} , by multiplication of each by the standard deviation of \underline{z}_{1j} for its reaction (i) from the mean of \underline{z}_{1j} for that i . Corresponding \underline{x}_j and \underline{y}_j values are calculated by least squares from the original data \underline{z}_{1j} and the newly unstandardized \underline{a}_1 and \underline{b}_1 . Values of \underline{c}_1 are calculated by simple least squares to fit the observed data using $\underline{p}_{1j} = \underline{a}_1 \underline{x}_j + \underline{b}_1 \underline{y}_j + \underline{c}_1$. Finally, new parameters meeting the new conditions are calculated by the transformation equations (10-15) using \underline{t} values derived previously (2, example using eq. 5). This transformation of the factors and slopes to meet the new six subsidiary conditions is necessarily considerably more complicated than the usual "rotations" of factor analysis which only rotate the axes in the plane on which the slopes are displayed.

To fit an equation of the form $\underline{p}_{1j} = \underline{a}_1 \underline{x}_j + \underline{b}_1 \underline{y}_j + \underline{c}_1$ to experimental data uniquely, six independent subsidiary conditions must be specified (2, phase 2). We choose $\underline{x}_5=0$, $\underline{y}_5=0$, $\underline{a}_1=1$ and $\underline{b}_3=1$ as the four relatively unimportant ones (which define reference points and unit sizes). For the two critical statements, we again choose $\underline{b}_1=0$ and $\underline{a}_3=\underline{a}_1/2$ for logical reasons explained previously (2, example).

Factor Analysis with No Missing Data

Table 1 and Figure 1 show selected parameters after various procedures. We have plotted slopes here rather than factors because slopes (factor loadings) are more commonly reported by users of factor analysis than the factors (factor scores) themselves. From a complete data set, correct slopes are obtained after our transformation (column 4), although the values before transformation (column 2) or after standard rotations seem quite different. Values in parentheses are those assigned by the subsidiary conditions. Use of $\underline{a}_7=-\underline{a}_1$ instead of $\underline{a}_3=\underline{a}_1/2$ as the sixth condition yields identical values for all the parameters. Introduction of small ($\sim 2\%$) random errors into the 70 data causes only small ($\sim 2\%$) errors in the factors and slopes.

Since most of the thousands of publications applying factor analysis have certainly not used our complicated transformation or its equivalent, but only varimax or other similar kinds of rotations contained in statistical analysis computer packages (12-18), it is instructive to note what kind of results such standard factor analyses give.

Assumptions built into the principal components direct solution cause the distribution of data in space to determine the order in which modes are extracted and force the factors to be orthogonal and the sum of squared slopes to be unity, although these conditions may not correspond at all with reality. Other extraction or direct

solution methods assume other conditions, similarly arbitrary or untrue. The very popular varimax rotation technique then "simplifies" the distribution of slopes so as to concentrate each slope on a few i 's and in one mode, enlarging slopes for those i 's and reducing them for others in that mode, again regardless of whether this makes physical sense or not. Other rotation techniques enhance this compartmentalized maldistribution of slopes even more.

Columns 2-4 show that 100% of the variance is explained (86.2% by the first mode, 13.8% by the second). Thus, principal components calculates correctly that there are two modes. However, this has no bearing on the correctness of the unrotated or rotated slopes.

Column 3 shows the result of varimax rotation, because this is the rotation method used in most published applications (21). The second-mode slopes b_1 are positive for $i=2, 4$ and 6 , as they must be if they are to make any physical sense, because the values of these properties do increase with increasing cylinder height. However, they certainly appear wrong for $i=1$ and 7 . The sensitivity of the first property, logarithm of the total area of the flat faces, to the second factor, logarithm of cylinder height, is deduced to be also positive and relatively large (0.56), although we happen to know that it is in fact exactly zero. The sensitivity of the last property, the log of electrical resistance between the ends, to $\log h$ is deduced to be even larger, near-unity, but negative (-0.82). Accordingly, electrical resistance of wires might be expected to decrease with increasing length, with a nearly reciprocal dependence. This would be a real boon for long-range power transmission and could lead to some interesting science fiction, but electrical engineers should not be misled by these results. Neither should those who would save energy be misled, by a thermal resistance analogy, into using close-packed lateral arrays of wide but very short cylinders (discs) for space-saving thermal roof or wall insulation in building construction.

Experimental verification of hypotheses is sufficiently practicable in physics and engineering that there has been little desire for and use of factor analysis in those fields. The greatest dangers are therefore in psychology, psychiatry, medicine, and political science, where experimental design is more difficult and where so many conclusions have accordingly been drawn from just such standard factor analyses with conventional rotations. It seems distinctly possible that progress in those fields has been hampered rather than helped in the past by such numerology, because such misleading results are obtained even when no data are missing. Such answers can be worse than no answers at all. Our test might be criticized because

of the small number of types of factors in our example ($n=2$, only two modes of interaction), whereas the problems to which factor analysis are applied usually have n equal to 3 or more. However, the untrue conditions corresponding to these artificial manipulations of the slopes, which are so obviously invalid in our example, are still assumed when there are more factors, although the absurdity of the results is then obscured by their greater complexity.

To obtain meaningful slopes (column 4) from factor analysis we have had to resort to our "transformation", instead of any conventional kind of rotation, in order to replace the untrue conditions by true ones, even when there are no missing data. Unfortunately, in other studies where factor analyses have been used, untrue conditions have generally not been replaced by a subsequent transformation incorporating true subsidiary conditions. Results of such studies are therefore unreliable except under special circumstances (22).

Factor Analysis with 20 Missing Data

Parameters deduced from 50 (accurate) data instead of 70, also with our transformation using $a_3 = a_1/2$ as the sixth subsidiary condition, are shown in columns 6-10 of Table I. The 20 data that are missing (randomly deleted) are indicated in Table 2 (2). The available factor analysis programs such as those in SPSS (12) provide options for the handling of missing data: listwise deletion, replacement by zeros, or pairwise deletion. Since listwise deletion eliminated all i except $i=3$, and replacement by zeros gave worse results than pairwise deletion, columns 5-6 were obtained via the third option. The correlation coefficient between the 50 observed and calculated data is now only 0.833. The rank order for b_i is now $5 < 4 < 1 < 6 < 2 < 3 < 7$ instead of the true order, $5 < 1 < 2 = 3 = 4 = 5 = 6 = 7$, putting $i=4$ between 5 and 1 instead of above them. The diamonds (\diamond) shown for several cylinders in Fig. 1 are not acceptably close to the circles. The calculated order of increasing $\log h (=y_j)$ is $7 < 3 < 10 < 4 < 9 < 1 < 2 < 6 < 8 < 5$ instead of the true order, $7 < 9 < 3 < 10 < 1 < 2 < 8 < 4 < 5 < 6$. With $a_7 = -a_1$ as the sixth subsidiary condition, the calculated order of increasing $\log h$ is still different, namely $7 < 4 < 10 < 2 < 3 < 9 < 6 < 1 < 8 < 5$. Such results based on accurate but incomplete data could be dangerously misleading in real problems, and therefore worse than useless.

Although SPSS does not provide convenient facilities for substituting the mean of data for an i in place of missing data for that i , use of means is claimed by many authors to be preferable to pairwise deletion. Column 7 shows the result (for $a_3 = a_1/2$). The correlation coefficient between observed and calculated data for the

50 is now 0.813, not much different with this example than pairwise deletion. The a_i parameters differ by more than 50% for $i=2, 4$ and 6 , although these should all be identical. A partial rank order for $\log h$ is $1<3<10<4<2<6<8<5$ vs. the true order, $3<10<1<2<8<4<5<6$. The percent of the variance explained by increasing numbers of modes is 68.5 (1), 82.9 (2), 94.2 (3), 97.8 (4), 99.4 (5), 99.9 (6) and 100.0 (7), which would mislead us into believing that at least three modes are involved and should be considered, if we did not know that there are only two.

In a relatively expensive but still futile attempt to find a practical way to obtain correct parameters from these 50 accurate data by factor analysis, we resorted to iteration of the combination of principal components extraction plus the complete transformation described above. Missing data for any i were replaced by the mean of data for that i before extraction in the first cycle, but by data predicted from the latest parameters (after transformation) in each subsequent cycle. This iterative procedure has serious disadvantages compared with DOVE. First, it expands each multiple regression to involve the latest estimates of all the missing data as well as the measured data, increasing the number of calculations involved in the time-consuming summations by 40% in the present example, and fourfold in our solvent effect problem where 75% of the data are missing. Second, although most of the parameters appear to be converging, it still gives a serious number of bad parameter values after five iterations (after 10 job steps and computer charges 10 times those required to obtain correct data and parameters to more than 6 decimal places by DOVE). At this point, the % variance explained has leveled off, but the correlation coefficient has begun to drop, as shown in columns 7-10 of Table 1. Lest anyone believe that these parameters are now nevertheless good enough because 98.7% of the variance is explained, we should note that the correlation coefficient is only 0.917 and that the fit to the observed data is becoming worse, that cylinder 1 is still deduced to be shorter than cylinder 10 ($\log h_1 < \log h_{10}$) whereas it is actually taller, the sensitivities to $\log h$ appear to vary more than 50% for properties such as 2, 4, 6 and 7, for which they are truly identical, and those for 4, 5 and 6 are getting worse. Doubtless there would be other inversions in rank orders if the distribution of the missing data were more biased (more associated with particular i 's and j 's), as they often are in real problems, rather than being random and relatively uniform as in this example. The squares (\square) in Fig. 1 show that the b_2 parameters for $i=4$ and 6 are bad and currently becoming worse. It would be interesting to follow this progression to its limit, where all values might possibly end up correct, but they are changing too slowly to make this practicable.

From these and other experiences with iterative factor analyses, we have found that it is critically important not to rely on convergence or constancy of any missing datum, parameter, % variance, or correlation coefficient as a criterion of correctness, but to test any proposed program by a problem like the present example that has its answers known in advance.

We have not found any published literature documenting a claim that correct results have ever been obtained by an application of principal components or any other correlation coefficient-based factor analysis to data sets with missing data. We conclude that although DOVE is able to solve such problems by simply omitting missing data, standard factor analyses can neither omit them nor obtain satisfactory estimates of the missing data or parameters by any efficient successive approximation process.

How to Solve Problems with Missing Data

Correct parameters cannot be obtained by principal components or any other kind of factor analysis based on correlation coefficients when data are missing. Usually the best way to obtain correct parameters is to measure the missing data to obtain a full data set, or narrow the scope of the study to a full subset having no missing data and then use principal components analysis followed by a valid transformation. If neither of these is practicable, we recommend use of the iterative DOVE procedure illustrated in the previous article, which works correctly in spite of missing data.

Reproduction of the data is not a sufficient test that a procedure is correct, because an infinite number of sets of answers (parameters) are consistent with the data. It is indispensable to test any new or untested procedure on a problem having the same number of modes and with answers known in advance, because errors in assumptions or logic are almost certain to escape detection if only real problems with unknown answers are analyzed.

Summary

Principal components factor analysis is tested for its reliability, using a problem with known answers. Even when test data are complete (e.g., 70 data on 7 properties, dependent on both radii and heights of 10 cylinders), such analyses followed by varimax or other standard rotations give incorrect rank orders for factors (factor scores for radius and height for each of the 10 cylinders) and sensitivities

(factor loadings for each of the 7 properties). When no data are missing, a transformation incorporating valid subsidiary conditions can be used instead of such rotations to obtain correct factors and sensitivities. However, when a moderate number (e.g., 20, or 29%) of the possible data are missing (randomly deleted), factors and sensitivities can have wrong rank orders and therefore be misleading even with this transformation. When data are missing, standard factor analysis is evidently unreliable and should be replaced by another method, such as that in the preceding article.

1. This article focuses on 1973-1976 studies using standard or conventional methods of factor analysis, especially the relatively simple and popular method usually called principal components. Cf. (2) which illustrates a practical alternative approach capable of yielding correct parameters even when data are missing.
2. P.F. Strong, C.G. Swain, M.S. Swain, preceding article.
3. Beginning with K. Pearson, Phil. Mag., ser. 6, 2, 559 (1901), they include 530 before 1940 [D. Wolfle, Psychometric Monogr., 3, 44-66 (1940)], 786 for 1940-52 [B. Fruchter, Introduction to Factor Analysis (van Nostrand, New York, N.Y., 1954, pp. 221-266)], and 640 even excluding psychology [in 10, pp. 522-600] for 1952-66.
4. R.N. Kavanagh, T.M. Darcey, D.H. Fender, Electroencephalogr. Clin. Neurophysiol., 40, 633 (1976). C.S. Newmark, T.R. Faschingbauer, A.J. Finch, P.C. Kendall, J. Clin. Psychol., 31, 449 (1975); D.J. Tosi and D.M. Eshbaugh, ibid., 32, 322 (1976).
5. A. Cammarata and G.K. Menon, J. Med. Chem., 19, 739 (1976); M.L. Weiner and P.H. Weiner, ibid., 16, 655 (1973).
6. A.M. Henderson and D.L. Greene, J. Dent. Res., 54, 344 (1975); A.V. Lombardi, Am. J. Phys. Anthropology, 42, 99 (1975); D.G. Barker, ibid., 44, 27 (1976); E.H. Bryant and W.R. Atchley, Eds., Benchmark Papers in Systematic and Evolutionary Biology, No. 2 (Wiley, New York, N.Y., 1975, pp. 1-345); G. Darland, Applied Microbiol., 30, 282 (1975); C.S. Moore, Ann. Bot. (London), 39, 113 (1975); M.V. Angel and M.J.R. Fasham, J. Mar. Biol. Assoc. U.K., 55, 709 (1975); H.J.B. Birks and M. Saarnisto, Boreas (Oslo), 4, 77, (1975); J.M. Perkins, Heredity, 32, 189 (1974).
7. D.L. Duewer, B.R. Kowalski, J.L. Fasching, Anal. Chem., 48, 2002 (1976); G.L. Ritter, S.R. Lowry, T.L. Isenhour, C.L. Wilkins, ibid., 48, 591 (1976); D.G. Howery, American Laboratory (Internat. Scientific Communications, Inc., Fairfield, Ct.), 8 (2), 14 (1976), 32 references; R.W. Rozett and E.M. Petersen, Anal. Chem., 48, 817 (1976), 29 references; J.B. Justice, Jr. and T.L. Isenhour, ibid., 47, 2286 (1975); R.N. Carey, S. Wold, J.O. Westgard, ibid., 47, 1824 (1975), 12 references; P.H. Weiner, H.L. Liao, B.L. Karger, ibid., 46, 2182 (1974); J.E. Davis, A. Shepard, N. Stanford, L.B. Rogers, ibid., 46, 821 (1974); D.G. Howery, P.H. Weiner, J.S. Blinder, J. Chromatogr. Sci., 12, 366 (1974),

- 14 references; D.H. McCloskey and S.J. Hawkes, ibid., 13, 1 (1975); J.T. Bulmer and H.F. Shurvell, J. Phys. Chem., 77, 256 (1973).
8. L.G. Closs and I. Nichol, Can. J. Earth Sci., 12 1316 (1975); H. Hoetzl, Dtsch. Geol. Ges. Z., 126, 121 (1975).
 9. J.E. Vincent, Factor Analysis in International Relations (Univ. of Florida Press, Gainesville, Fla., 1971); D.R. Hall and R.J. Rummel, The Dimensionality of Nations Project, research report no. 16 (Univ. of Hawaii, 1969); C. Wall and R.J. Rummel, ibid., report no. 20 (1969); B.M. Russett, Ed., Peace, War, and Numbers (Sage, Beverly Hills, Calif., 1972); P. Warwick, Soc. Sci. Research, 14 241 (1975).
 10. R.J. Rummel, Applied Factor Analysis (Northwestern Univ. Press, Evanston, Ill., (1970).
 11. D. Child, Essentials of Factor Analysis (Holt, Rinehart and Winston, London 1970); R.L. Gorsuch, Factor Analysis (Saunders, Philadelphia, Pa., 1974); H.H. Harman, Modern Factor Analysis (Univ. of Chicago Press, Chicago, Ill., ed. 3, 1976); A.L. Comrey, A First Course in Factor Analysis (Academic Press New York, N.Y., 1973); S.A. Mulaik, Foundations of Factor Analysis (McGraw-Hill, New York, N.Y., 1972); W.H. Guertin and J.P. Bailey, Jr., Introduction to Modern Factor Analysis (Edwards, Ann Arbor, Mich., 1970); D.N. Lawley and A.E. Maxwell, Factor Analysis as a Statistical Method (Am. Elsevier, New York, N.Y., 1971); W.W. Cooley and P.R. Lohnes, Multivariate Data Analysis (Wiley New York, N.Y., 1971, pp. 96-167); D.F. Morrison, Multivariate Statistical Methods (McGraw-Hill, New York, N.Y., ed. 2, 1976, pp. 266-343); A.S.C. Ehrenberg, Data Reduction (Wiley, New York, N.Y., 1975 pp. 263-282); K. Enselin, A. Ralston and R.S. Wilf, Eds., Statistical Methods for Digital Computers (Wiley, New York, N.Y., Vol. 13, 1977, pp. 9-10, 123-202).
 12. N.H. Nie, C.H. Hull, J.G. Jenkins, K. Steinbrenner and D.H. Bent, Statistical Package for the Social Sciences, (McGraw-Hill, New York, N.Y., 1975, ed. 2, pp. 468-514.
 13. W.J. Dixon, BMD Biomedical Computer Programs (Univ. of California Press, 1973); W.J. Dixon, BMDP Biomedical Computer Programs (Univ. of California Press, 1975).
 14. D.G. Armor and A.S. Couch, Data-Text Primer (Macmillan, New York, N.Y., 1972, pp. 82-91).
 15. R. Buhler, P-STAT (Princeton Univ., 1975).
 16. System/360 Scientific Subroutine Package Programmer's Manual, Version III (Internat. Bus. Machines Corp., manual GH20-0205-4, 1968).

17. IMSL Library 1 (IBM Series), ed. 5, 1975 and IMSL Library 2, (Honeywell Series), ed. 4, 1974 (Internat. Mathematical and Statistical Libraries, Houston, Texas).
18. B.H. Hall, Time Series Processor, Version 2 (Harvard Institute of Economic Research, Harvard Univ., Cambridge, Ma., 1976).
19. L.L. Thurstone, Multiple-Factor Analysis (Univ. of Chicago Press, Chicago, Ill., 1947, pp. 117-124).
20. The eight synthetic problems or "artificial experiments" listed in 10, pp. 528-9.
21. Four SPSS extraction procedures (PA1, PA2, Rao, and alpha) with six rotations (varimax, equimax, quartimax, and oblique with $\delta=-1, 0, +1$) were carried out. Image extraction failed (singular correlation matrix) even with 0.1-1% data error (absolute magnitude changed by 1 in third significant figures, up for odd $i+j$, down for even $i+j$). None of the combinations gave realistic sets of factors. Quartimax also made b_2 positive and b_7 negative. Highly oblique ($\delta=+1$) structure factors were only slightly negative for b_2 but varied more than threefold from b_4 to b_7 . Most other sets had numerous incorrect signs.
22. One-mode problems, such as the chemical one of best fitting the Hammett equation with a single type of substituent constant [S. Wold and M. Sjöström, Chem. Scr., 2, 49 (1972); S. Wold, ibid., 5, 97 (1974); M. Sjöström and S. Wold, ibid., 6, 114 (1974)], do not lead to such errors and confusion, because the number of critical conditions is zero for them, hence there is no possibility of choosing the critical conditions incorrectly.
23. We thank Professor Albert J. Petkau, Department of Mathematics, University of British Columbia, for discussions and many helpful suggestions.
24. This work was supported by research grants from the Office of Naval Research and the Petroleum Research Fund of the American Chemical Society. Mr. Bryndza acknowledges a Hertz Foundation scholarship and computer charge credits by the M.I.T. student information processing board.

Table 1. Second-mode slopes, b_i , from principal components analyses.

i	70 true data		50 true data		50 true + 20 calcd data				
	not rotated	varimax rotated	after trans-formation	not rotated	after trans-formation	0	3	4	5
1	-0.04	0.56	(0.00)	0.03	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)
2	0.22	0.33	1.00	0.26	0.78	0.93	0.88	0.88	0.90
3	0.38	0.16	(1.00)	0.47	(1.00)	(1.00)	(1.00)	(1.00)	(1.00)
4	0.11	0.43	1.00	0.02	-0.11	0.49	1.01	1.21	1.39
5	-0.74	0.98	-1.00	-0.80	-0.67	-0.59	-0.79	-0.74	-0.67
6	0.22	0.33	1.00	0.13	0.36	1.25	1.28	1.39	1.46
7	0.40	-0.82	1.00	0.42	1.18	0.85	1.06	1.11	1.08
Σ^*	100.0	100.0	100.0	101.7	101.7	82.9	98.2	98.7	98.7
cct	1.000	1.000	1.000	0.833	0.833	0.813	0.926	0.925	0.917

* Percent of the variance explained by two modes. † Correlation coefficient based on only the true data used.

Legend for Figure 1

Figure 1. A plot of correct relative sensitivity to log height vs. relative sensitivity to log radius for 7 cylinder properties, calculated from 70 data by principal components followed by a valid transformation ● . Also shown are the most seriously incorrect values calculated similarly except from 50 true data (with 20 data or 29% of the possible data missing) via pairwise deletion ◆ , replacement by means ▲ , or five subsequent iterative replacements □ . Other calculated points falling closer to the correct values have been omitted.

