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20. Abstract (continued)

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THE USE OF HARMONIC ANALYSIS IN
SUBOPTIMAL ESTIMATOR DESIGN*

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Abstract

The state estimation problem for bilinear stochastic systems evolving on compact Lie groups and homogeneous spaces is considered. The problem is motivated by some applications involving rotational processes in three dimensions. The theory of harmonic analysis on compact Lie groups is used to define assumed density approximations which result in implementable suboptimal estimators for the state of the bilinear system. The results of Monte Carlo simulations are reported; these indicate that simple filters designed by these techniques perform well as compared to other filters.

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I. INTRODUCTION

Fourier series analysis has been applied in several recent studies [1] - [4] to estimation problems for stochastic processes evolving on the circle S^1 . Willsky [4] used Fourier series methods to define "assumed density" approximations for certain phase tracking and demodulation problems. In fact, a system designed using these techniques performed better than other estimators, including an optimal phase-lock loop.

In this paper we study bilinear systems evolving on compact Lie groups or homogeneous spaces [29]. The optimal estimator is in general infinite dimensional [7], and our approach to the design of suboptimal estimators is a generalization of that of Willsky [4], whose work is reviewed briefly in Section III. The basic approach involves the definition of an "assumed density" form for the conditional density of the system state at time t given observations up to time t . These densities are defined via the techniques of harmonic analysis on compact Lie groups [5],[6] (which generalize the Fourier series on the Lie group S^1). Our method differs from most previous assumed density approximations in that our approximation is defined on the appropriate compact manifold (as opposed to the usual Gaussian approximations, for example, which are defined on $\mathbb{R}^n$ [7]). This method also avoids the problem of merely truncating higher order terms in a harmonic expansion; as pointed out by Lo [18] and Willsky [4], such higher order terms will not be negligible, especially if the filter is performing well. For an alternative approach to discrete-time estimation problems on Lie groups and homogeneous spaces, see the work of Lo and Eshleman [18-20], who use exponential Fourier densities to avoid the truncation problem.

In Section II we review some general properties of stochastic bilinear systems and discuss the estimation problem for systems evolving

on compact Lie groups and homogeneous spaces. Section III contains the application of the technique to systems evolving on S^n , while Section IV contains the application to systems on $SO(n)$. Results of Monte Carlo simulations of the S^2 estimator are presented in Section V. Finally, the Appendix reviews some necessary concepts from the theory of harmonic analysis on compact Lie groups. Some of the concepts of this paper were introduced in [21] and [22], but no simulation results were presented.

II. ESTIMATION FOR STOCHASTIC BILINEAR SYSTEMS

The basic stochastic bilinear system (or linear system with state-dependent noise) considered here is described by the Ito stochastic differential equation [4], [9-17], [21], [25]

$$dx(t) = \left\{ [A_0 + \frac{1}{2} \sum_{i,j=1}^N Q_{ij}(t) A_i A_j] dt + \sum_{i=1}^N A_i dw_i(t) \right\} x(t) \quad (1)$$

where x is an n -vector or an $n \times n$ matrix, the A_i are $n \times n$ matrices, Q_{ij} is the (i,j) th element of Q , and w is a Brownian motion (Wiener) process with strength $Q(t)$ such that $E[w(t)w'(s)] = \int_0^{\min(t,s)} Q(\tau) d\tau$. Following the notation of [8-11], we define $\mathcal{L} = \{A_0, A_1, \dots, A_N\}_{LA}$ to be the smallest Lie algebra containing these matrices. The corresponding connected matrix Lie group is denoted by $G = \{\exp \mathcal{L}\}_G$. Then, if x is an $n \times n$ matrix and $x(t_0) \in G$, the solution $x(t)$ of (1) evolves on G (i.e., $x(t) \in G$ for all $t \geq 0$) in the mean-square sense and almost surely [15-17]. If x is an n -vector, then the solution of (1) evolves on the homogeneous space $G \cdot x(t_0)$.

Associated with the Ito equation (1) is a sequence of equations for the powers of the state $x(t)$ (see Brockett [9], [10]). If $N(n,p)$ denotes the binomial coefficient $\binom{n+p-1}{p}$, then given an n -vector x , we define $x^{[p]}$ to be the $N(n,p)$ -vector with components equal to the monomials (homogeneous polynomials) of degree p in $x_1, \dots, x_n$, the components of x , scaled so that

$\|x\|^{\tilde{p}} = \|x^{[p]}\|$. Given an $m \times n$ matrix A , we denote by $A^{[p]}$ the unique matrix which verifies

$$y = Ax \Rightarrow y^{[p]} = A^{[p]} x^{[p]}, \quad (2)$$

$A^{[p]}$ can be interpreted as a linear operator on symmetric tensors of degree p [9], and is known as the symmetrized Kronecker p^{th} power of A [26]. It is clear that if x satisfies the linear differential equation

$$\dot{x}(t) = Ax(t) \quad (3)$$

then $x^{[p]}$ also satisfies a linear differential equation

$$\dot{x}^{[p]}(t) = A_{[p]} x^{[p]}(t) \quad (4)$$

We regard this as the definition of $A_{[p]}$, which is the infinitesimal version of $A^{[p]}$. In fact, $A_{[p]}$ can be easily computed from A [25].

It can easily be shown that if x satisfies (1), then $x^{[p]}$ satisfies the Ito equation

$$dx^{[p]}(t) = \{A_{0[p]} + \frac{1}{2} \sum_{i,j=1}^N Q_{ij}(t) A_{i[p]} A_{j[p]}\} x^{[p]}(t) dt + \sum_{i=1}^N A_{i[p]} x^{[p]}(t) dw_i(t) \quad (5)$$

In addition, if the $n \times n$ matrix X satisfies (1), it is easy to show that $X^{[p]}$ also satisfies (5). As we shall see later in the section, this sequence of equations is a valuable tool in the study of state estimation.

The observation model considered in this paper consists of linear observations of the state corrupted by additive white noise, or

$$dz(t) = L(x(t)) dt + dv(t) \quad (6)$$

where L is a linear operator and v is a Wiener process. This bilinear system-linear observation model is useful in the study of certain practical problems, such as the S^2 satellite tracking and $SO(3)$ rigid body orientation estimation problems discussed in [13, Ch. 4] and [21, Sec. IV].

The remainder of the paper is devoted to the study of the estimation problem for two classes of systems of the form (1), (6), which are motivated by the aforementioned examples. The first system consists of the bilinear state equation

$$dX(t) = [A_0 + \frac{1}{2} \sum_{i,j=1}^N Q_{ij}(t) A_i A_j] X(t) dt + \sum_{i=1}^N A_i X(t) dw_i(t) \quad (7)$$

with linear measurements

$$dz_1(t) = X(t)h(t)dt + R^{\frac{1}{2}}(t)dv(t) \quad (8)$$

where $X(t)$ and $\{A_i\}$ are $n \times n$ matrices, $z_1(t)$ is a p -vector, w is a Wiener process with strength $Q(t) \geq 0$, v is a standard Wiener process independent of w , and $R > 0$. More general linear measurements can obviously be considered, but for simplicity of notation we restrict our attention to (8), which arises in the star tracking example of [13, Ch. 4]. We will make the crucial assumption that the Lie group $G = \{\exp \mathcal{L}\}_G$ is compact.

The second system consists of the bilinear state equation (1) with linear measurements

$$dz_2(t) = H(t)x(t)dt + R^{\frac{1}{2}}(t)dv(t) \quad (9)$$

where $x(t)$ is an n -vector, A_i are $n \times n$ matrices, and z_2 , v and w are as above. It will be assumed that x evolves on a compact homogeneous space [8], [13], [29]—i.e., the solution of (1) is

$$x(t) = X(t)x(0) \quad (10)$$

where X satisfies (7) with $X(0) = I$ and evolves on the compact Lie group $G = \{\exp \mathcal{L}\}_G$.

It is shown in [13] that, by a linear change of basis on the state space, (7) and (1) can be transformed into equations which evolve on the compact special orthogonal group $SO(n) = \{X \in \mathbb{R}^{n \times n} | X'X = I, \det X = +1\}$ and the compact homogeneous space $S^{n-1} = \{x \in \mathbb{R}^n | x'x = 1\}$ (the $(n-1)$ -sphere), respectively. Hence, we need only consider systems evolving on $SO(n)$ and S^{n-1} (this is equivalent to the assumption that $A_0, A_1, \dots, A_N$ are skew-symmetric [13]).

The estimation criterion which will be used for these two problems is the constrained least-squares estimator, which is analogous to the criterion used in [1], [4], and [21] for the phase estimation problem. That is, for (7)-(8) we wish to find $\tilde{X}(t|t)$ which minimizes the conditional error covariance

$$J_1 = E[\text{tr}\{(X(t) - \tilde{X}(t|t))'(X(t) - \tilde{X}(t|t))\} | z_1^t] \quad (11)$$

subject to the $SO(n)$ constraint $\tilde{X}(t|t)' \tilde{X}(t|t) = I$, where the notation (11) denotes the conditional expectation given the σ -field $\sigma\{z_1^t\}$ generated by the observed process $z_1^t \triangleq \{z_1(s), 0 \leq s \leq t\}$ up to time t . For (1), (9) we seek $\tilde{x}(t|t)$ which minimizes

$$J_2 = E[(x(t) - \tilde{x}(t|t))'(x(t) - \tilde{x}(t|t)) | z_2^t] \quad (12)$$

subject to the S^{n-1} constraint $\|\tilde{x}(t|t)\|^2 = \tilde{x}(t|t)' \tilde{x}(t|t) = 1$.

It is easily shown [13] that the optimal estimates are, respectively

$$\tilde{X}(t|t) = \pm \hat{X}(t|t) [\hat{X}(t|t)' \hat{X}(t|t)]^{-\frac{1}{2}} \quad (13)$$

$$\tilde{x}(t|t) = \frac{\hat{x}(t|t)}{\|\hat{x}(t|t)\|} \quad (14)$$

where the conditional expectation is denoted by the equivalent notations

$$\hat{x}(t|t) \triangleq E[x(t) | z_2^t] \triangleq E^t[x(t)] \quad (15)$$

The sign in (13) is chosen to insure that $\det \tilde{X}(t|t) = +1$ [23]. Thus in both cases we must compute the conditional expectation of the state ($X(t)$ or $x(t)$) given the past observations ($z_1^t = \{z_1(s), 0 \leq s \leq t\}$ or $\{z_2^t(s), 0 \leq s \leq t\}$).

The equations for computing the conditional expectation can be derived from the general nonlinear filtering equation [7] and the moment equation (5). The resultant equations for the SO(n) system (7)-(8) are

$$\begin{aligned} dE[X_v^{[p]}(t)] = & [A_{0[p]} + \frac{1}{2} \sum_{i,j=1}^N Q_{ij}(t) A_{i[p]} A_{j[p]}] \otimes I E[X_v^{[p]}(t)] dt \\ & + \{E[X_v^{[p]}(t)h'(t)X(t)] - E[X_v^{[p]}(t)]h'(t)E[X(t)]\} R^{-1}(t) dv_1(t) \end{aligned} \quad (16)$$

$$dv_1(t) = dz_1(t) - \hat{X}(t|t)h(t)dt \quad (17)$$

where $\otimes$ denotes Kronecker product and $X_v^{[p]}$ is the vector containing the elements of the matrix $X^{[p]}$ in lexicographic order [26], [32, p.64]. For the S^{n-1} system (1), (9), we have

$$\begin{aligned} dE[X^{[p]}(t)] = & [A_{0[p]} + \frac{1}{2} \sum_{i,j=1}^N Q_{ij}(t) A_{i[p]} A_{j[p]}] E[X^{[p]}(t)] dt \\ & + \{E[X^{[p]}(t)x'(t)] - E[X^{[p]}(t)]E[x'(t)]\} H'(t)R^{-1}(t) dv_2(t) \end{aligned} \quad (18)$$

$$dv_2(t) = dz_2(t) - H(t)\hat{x}(t|t)dt. \quad (19)$$

The structure of these equations is quite similar to that of [4]—i.e., each estimator consists of an infinite band of filters, and the filter for the p^{th} moment is coupled only to those for the first and $(p+1)^{st}$ moments. Therefore, we are led to the design of suboptimal estimators. The technique proposed here is motivated by the highly successful use of folded normal assumed density approximations in the phase tracking problem [4]; filters designed using this technique performed very well as compared with other suboptimal estimators. We will describe similar techniques for the design of suboptimal estimators on S^n and SO(n).

We first review the notions of Brownian motion and Gaussian densities on Lie groups and homogeneous spaces. Yosida [28] proved that the fundamental solution of

$$\frac{\partial p(x,t)}{\partial t} - \gamma \Delta p(x,t) = 0 \quad (20)$$

where $\gamma > 0$ and Δ is the Laplace-Beltrami operator (Laplacian) on a Riemannian homogeneous space M [5,13,29], is the density (with respect to the Riemannian measure) of a Brownian motion on M^1 .

According to [5], the fundamental solution of (20) is given by

$$p(x,t;x_0,t_0) = \sum_i \phi_i(x) \phi_i(x_0) e^{-\lambda_i(t-t_0)\gamma} \quad (21)$$

where λ_i and ϕ_i are the eigenvalues and the corresponding eigenfunctions of the Laplacian (see the Appendix). The function $p(x,t;x_0,t_0)$ is the solution to (20) with initial condition equal to the singular distribution concentrated at $x = x_0$. Also, Grenander [27] defines a Gaussian (normal) density to be the solution of (20) for some t .

The folded normal density $F(\theta;\eta,\gamma)$ used by Willsky as an assumed density approximation for the phase tracking problem is indeed a normal density on S^1 in the sense of Grenander [4]; in fact, the trigonometric polynomials $e^{-in\theta}$ are eigenfunctions of the Laplacian on S^1 . Motivated by the success of Willsky's suboptimal filter, we will design suboptimal estimators for the $SO(n)$ and S^n bilinear systems by employing normal assumed

¹ Yosida defines a Brownian motion process to be a temporally and spatially homogeneous Markov process on M which satisfies a continuity condition of Lindeberg's type.

conditional densities of the form

$$p(x,t) = \sum_i \phi_i(x) \phi_i(\eta(t)) e^{-\lambda_i \gamma(t)} \quad (22)$$

where $\eta(t)$ and $\gamma(t)$ are parameters of the density which are to be estimated.²

² In order to assure the existence of a conditional density, it is sufficient to assume that the system is "controllable from the noise" [10,13,17].

III. ESTIMATION ON S^n

In this section the suboptimal estimation technique discussed in the previous section will be used in order to design filters for the S^n estimation problem (1), (9). The optimal constrained least-squares estimator is described by (14) and (18)-(19). First, the suboptimal estimator for S^2 will be described in detail; then the generalization to S^n will be discussed.

In our discussion of estimation on S^2 , we will refer to a point on S^2 in terms of the Cartesian coordinates $x \triangleq (x_1, x_2, x_3)$ or the polar coordinates (θ, ϕ) (see the Appendix, in which harmonic analysis on S^n is summarized). According to the Appendix, the spherical harmonics $\{Y_{\ell m}\}$ (defined in (A.11)-(A.12)) are the eigenfunctions of the Laplacian Δ_{S^2} (defined in (A.10)), and all spherical harmonics of degree ℓ have the same eigenvalue $-\ell(\ell+1)$. Thus the assumed density approximation is a normal density on S^2 of the form (22), as discussed in the previous section:

$$p(\theta, \phi, t) = \sum_{\ell=0}^{\infty} \sum_{m=-\ell}^{\ell} Y_{\ell m}(\theta, \phi) Y_{\ell m}^*(\eta(t), \lambda(t)) e^{-\ell(\ell+1)\gamma(t)}. \quad (23)$$

where $*$ denotes complex conjugate. In other words,

$$c_{\ell m}(t) \triangleq E^t[Y_{\ell m}^*(\theta(t), \phi(t))] = Y_{\ell m}^*(\eta(t), \lambda(t)) e^{-\ell(\ell+1)\gamma(t)}. \quad (24)$$

In order to truncate the optimal estimator (18)-(19) after the $\hat{x}^{[N]}(t|t)$ equation using the assumed density (23), we must compute $E^t[x^{[N]}(t)x'(t)]$, or equivalently, $\hat{x}^{[N+1]}(t|t)$, in terms of $\hat{x}^{[p]}(t|t)$, $p = 1, 2, \dots, N$. However, if $\hat{x}(t|t)$ is known, so are $c_{10}(t)$ and $c_{11}(t)$. A

simple computation [13], [22] then shows that $\{c_{N+1,m}^{(N+1)}, \dots, c_{N+1,N+1}^{(N+1)}\}$ can be computed from

$$\begin{aligned} c_{N+1,m}^{(N+1)}(t) &= Y_{N+1,m}^*(\eta(t), \lambda(t)) e^{-(N+1)(N+2)\gamma(t)} \\ &= (-1)^m \left[\frac{(N+1-m)!}{(N+1+m)!} \frac{2N+3}{4\pi} \right]^{\frac{1}{2}} P_{N+1,m} \left(\frac{c_{10}(t)}{(c_{10}^2(t) + 2|c_{11}(t)|^2)^{\frac{1}{2}}} \right) \\ &\quad \cdot \left(\frac{c_{11}(t)}{c_{11}^*(t)} \right)^{m/2} \left[\frac{4\pi}{3} (c_{10}^2(t) + 2|c_{11}(t)|^2) \right]^{\frac{1}{2}(N+1)(N+2)} \end{aligned} \quad (25)$$

where $P_{\ell m}$ are the associated Legendre functions (see (A.11)-(A.12)). Finally, it is shown in [13] that there exists a nonsingular matrix T such that

$$T x^{[N+1]} = \begin{bmatrix} Y_{N+1}(x) \\ x^{[N-1]} \end{bmatrix} \quad (26)$$

Thus $\hat{x}^{[N+1]}(t|t)$ can be computed from $\{c_{N+1,m}^{(N+1)}, -(N+1) \leq m \leq N+1\}$ and $\hat{x}^{[N-1]}(t|t)$. The optimal estimator (18) is truncated by substituting this approximation for $\hat{x}^{[N+1]}(t|t)$ into the equation for $\hat{x}^{[N]}(t|t)$. Notice that the entire procedure for truncating the optimal estimator can equivalently be performed on the infinite set of coupled equations for the generalized Fourier coefficients $c_{\ell m}(t)$, using the approximation (24).

We note that one can show that

$$\alpha(t) \triangleq \sqrt{\|\hat{x}(t|t)\|} \leq 1$$

and this quantity can be used as a measure of our confidence in our estimate. If $\hat{x}(t|t)$ satisfies the assumed density (23),

$$\alpha(t) = \|\hat{x}(t|t)\| = e^{-\gamma(t)} \quad (27)$$

so $\gamma = 0$ (zero "variance") implies $\alpha = 1$, and $\gamma = \infty$ (infinite "variance") implies $\alpha = 0$ (see [4] for the S^1 analog).

Example 1: Suppose that we truncate the optimal S^2 estimator (18) after $N = 1$ —i.e., we approximate $\hat{x}^{[2]}(t|t)$ using the above approximation. Assume that $Q(t) = I$ and $\{A_i, i=1,2,3\}$ are given. Then the resulting suboptimal estimator is (for $Q(t) = I$)

$$\begin{aligned} d\hat{x}(t|t) = & [A_0 + \frac{1}{2} \sum_{i=1}^N A_i^2] \hat{x}(t|t) dt \\ & + P(t)H'(t)R^{-1}(t)[dz_2(t) - H(t)\hat{x}(t|t)dt] \end{aligned} \quad (28)$$

where the "covariance" matrix $P(t)$ is given by

$$P_{ii}(t) = \hat{x}_i^2(t|t) \left(\frac{2}{3} \|\hat{x}(t|t)\| - 1 \right) - \frac{1}{3} (\hat{x}_j^2(t|t) + \hat{x}_k^2(t|t)) \|\hat{x}(t|t)\| + \frac{1}{3} \quad (29)$$

for $i \neq j, i \neq k, j \neq k$, and

$$P_{ij}(t) = \hat{x}_i(t|t)\hat{x}_j(t|t) (\|\hat{x}(t|t)\| - 1) \quad (30)$$

for $i \neq j$. It is shown in [37] that the matrix $P(t)$ of (29)–(30) is positive semidefinite, and thus can be viewed as a covariance matrix. The results of Monte Carlo simulations to evaluate the performance of this estimator are presented in Section VII.

The extension to S^n of this technique for constructing suboptimal estimators is straightforward. The procedure uses the spherical harmonics on S^n . In polar coordinates, a point on S^n can be described by $(\theta_1, \theta_2, \dots, \theta_{n-1}, \phi) \triangleq (\theta, \phi)$, where $0 \leq \theta_j \leq \pi$ and $0 \leq \phi \leq 2\pi$. Also, the spherical harmonics are denoted by

$$\begin{aligned} Y_{\ell, (m)}(\theta, \phi) & \triangleq Y_{\ell, m_1, \dots, m_{n-1}}(\theta_1, \dots, \theta_{n-1}, \phi) \\ & = e^{\pm i m_{n-1} \phi} \prod_{k=0}^{n-2} (\sin \theta_{k+1})^{m_{k+1}} C_{m_k - m_{k+1}}^{m_{k+1} + \frac{1}{2}(n-k-1)} (\cos \theta_{k+1}) \end{aligned} \quad (31)$$

where $\ell \geq m_1 \geq \dots \geq m_{n-1} \geq 0$ and C_j^i are the Gegenbauer polynomials [33] (that is, the functions $Y_{\ell, (m)}$ are eigenfunctions of the Laplace-Beltrami operator with eigenvalue $-\ell(\ell+n-1)$). Hence the assumed density approximation on S^n is

$$p(\theta, \varphi, t) = \sum_{\ell, (m)} Y_{\ell, (m)}(\theta, \varphi) Y_{\ell, (m)}^*(\eta(t), \lambda(t)) e^{-\ell(\ell+n-1)\gamma(t)} \quad (32)$$

That is, $c_{\ell, (m)}(t) \triangleq E^t[Y_{\ell, (m)}^*(\theta(t), \varphi(t))]$ is assumed to be

$$c_{\ell, (m)}(t) = Y_{\ell, (m)}^*(\eta(t), \lambda(t)) e^{-\ell(\ell+n-1)\gamma(t)}. \quad (33)$$

The procedure for truncating the filter (18) is identical to the S^2 case. If $\hat{x}(t|t)$ is known, so are $c_{1, (m)}(t)$, and these can be used to compute $\gamma(t)$, $\eta(t)$, and $\lambda(t)$. Then $\{c_{N+1, (m)}(t)\}$ can be computed from (33), and $\hat{x}^{[N+1]}(t|t)$ can be computed from $\{c_{N+1, (m)}(t)\}$ and $\hat{x}^{[N-1]}(t|t)$. The estimator is truncated by substituting this approximate expression for $\hat{x}^{[N+1]}(t|t)$ into the equation (18) for $\hat{x}^{[N]}(t|t)$.

IV. ESTIMATION ON SO(n)

In this section we discuss the construction of suboptimal estimators for the SO(n) estimation problem (7)-(8). We will only consider the SO(3) problem; the results are extended to SO(n) in [13]. The concepts of harmonic analysis on SO(3) presented in the Appendix will be used extensively.

Consider the sequence $\{D^\ell, \ell = 0, 1, \dots\}$ of irreducible unitary representations of SO(3), as defined in (A.5)-(A.6). Theorem A.1 implies that, for fixed ℓ , the matrix elements $\{D_{mn}^\ell; -\ell \leq m, n \leq \ell\}$ are eigenfunctions of the Laplacian $\Delta_{SO(3)}$ (defined in (A.4)) with the same eigenvalue λ_ℓ ; also, all eigenfunctions of the Laplacian can be written as linear combination of the $\{D_{mn}^\ell\}$. Hence, the assumed density which will be used to truncate the optimal estimator (16)-(17) is a normal density on SO(3) of the form (22):

$$p(R, t) = \sum_{\ell=0}^{\infty} \sum_{m,n=-\ell}^{\ell} D_{m,n}^\ell(R) D_{m,n}^\ell(\eta(t))^* e^{-\lambda_\ell \gamma(t)} \quad (34)$$

where $R, \eta(t) \in SO(3)$ and $\gamma(t)$ is a scalar. That is,

$$c_{mn}^\ell(t) \triangleq E^t[D_{mn}^\ell(\eta(t))^*] \quad (35)$$

is assumed to be

$$c_{mn}^\ell(t) = D_{mn}^\ell(\eta(t))^* e^{-\lambda_\ell \gamma(t)} \quad (36)$$

The procedure for truncating the filter (16) is similar to the S^n case, although we make use of some additional concepts from representation theory. If $\hat{X}(t|t)$ is known, so are $\{c_{mn}^1(t); -1 \leq m, n \leq 1\}$, since D^1 is equivalent to the self-representation of SO(3). Define the matrix $C^\ell(t)$ with elements $c_{mn}^\ell(t)$, $-\ell \leq m, n \leq \ell$; then

$$A(t) \triangleq \bar{C}^1(t) C^1(t) = [D^1(\eta(t))] [D^1(\eta(t))]^* e^{-2\lambda_1 \gamma(t)}$$

$$= I \cdot e^{-2\lambda_1 \gamma(t)} \quad (37)$$

since D^1 is unitary (here $\bar{C}$ is the hermitian transpose of C). Thus $\gamma(t)$ can be computed from

$$\gamma(t) = -\frac{1}{2\lambda_1} \log \left[\frac{1}{3} \text{tr } A(t) \right]. \quad (38)$$

Then the elements of $\eta(t)$ can be computed from (36) and (38), since $D^1(\eta(t))$ is similar to $\eta(t)$. Once $\gamma(t)$ and $\eta(t)$ have been computed, $\{c_{mn}^{N+1}; -(N+1) \leq m, n \leq N+1\}$ are computed from the formula (36).

In order to truncate (16) after the N^{th} moment equation, we must approximate $E^t[X_v^{[N]}(t)h'(t)X(t)]$; however, this matrix consists of time-varying deterministic functions multiplying elements of $\hat{X}^{[N+1]}(t|t)$, so we will show how to approximate this matrix. The symmetrized Kronecker p^{th} power $X^{[p]}$ operating on the symmetric tensors $x^{[p]}$ such that $\|x^{[p]}\| = \|x\|^p = 1$ furnishes a representation of $SO(3)$ which is reducible [26]. It is shown in [13] that there is a nonsingular matrix T such that

$$TX^{[p]}T^{-1} = \begin{bmatrix} D^p(X) & 0 \\ 0 & X^{[p-2]} \end{bmatrix}. \quad (39)$$

The matrix T is related to the Clebsch-Gordan coefficients [6], but T can also be computed by the method of Gantmacher [34, p. 160]. It is clear from the decomposition (39) that $\hat{X}^{[N+1]}(t|t)$ can be computed from $C^{N+1}(t)$ and $\hat{X}^{[N-1]}(t|t)$. The optimal estimator (16) is truncated by substituting this approximation into the equation for $\hat{X}^{[N]}(t|t)$.

We note here that, due to the decomposition (39), the estimation equations and the truncation procedure could have been expressed solely in terms of the irreducible representations $D^p(X(t))$. However, we have chosen to work with the $X^{[p]}$ equations primarily for ease of notation. For large N , the D^p equations would provide significant computational savings over the $X^{[p]}$ equations, as these are redundant; however, the practical implementation of this technique will probably be limited to small values of N .

V. SIMULATION RESULTS

As an illustration of the techniques presented in the previous sections, the first order filter (FOF) of Example 1 (Section III) was evaluated by means of digital Monte Carlo simulations. It was compared to both the extended Kalman filter (EKF) [7] and the Gustafson-Speyer linear, minimum-variance quadrature filter (LQF) [24]. Identical noise sequences were used to allow direct comparisons.

The system considered was the S^2 system, i.e.,

$$dx(t) = Fx(t) dt + \sum_{i=1}^3 A_i x(t) dw_i(t) \quad (40)$$

$$dz(t) = x(t) dt + r^{1/2} dv(t) \quad (41)$$

where $F = \sum_{i=1}^3 f_i A_i + \frac{1}{2} q \sum_{i=1}^3 A_i^2$, and $\{A_i, i=1,2,3\}$ are the skew-symmetric matrices

$$A_1 = \begin{bmatrix} 0 & 0 & 0 \\ 0 & 0 & -1 \\ 0 & 1 & 0 \end{bmatrix} \quad A_2 = \begin{bmatrix} 0 & 0 & 1 \\ 0 & 0 & 0 \\ -1 & 0 & 0 \end{bmatrix} \quad A_3 = \begin{bmatrix} 0 & -1 & 0 \\ 1 & 0 & 0 \\ 0 & 0 & 0 \end{bmatrix} \quad (42)$$

Also, $w(t)$ has strength qI ; and v has strength I . In this experiment, the nominal angular velocities $\{f_i, i=1,2,3\}$ were chosen to be 100.0, and q and r were varied.

For all three filters, the normalized estimate $\tilde{x}(t) = \hat{x}(t|t)/\|\hat{x}(t|t)\|$ was used. The filters have an identical structure for the approximate $\hat{x}$ equation:

$$d\hat{x}(t|t) = F\hat{x}(t|t) + \frac{1}{r} P(t) [dz(t) - \hat{x}(t|t)dt] \quad (42)$$

However, for the FOF, $P(t)$ is given by the highly nonlinear memoryless equations (29)–(30). In the EKF, $P(t)$ satisfies the Riccati equation

$$\frac{d}{dt} P(t) = FP(t) + P(t)F' + qG(\hat{x}(t|t))G'(\hat{x}(t|t)) - \frac{1}{r} P(t)P'(t) \quad (43)$$

where

$$G(\hat{x}) = [A_1 \hat{x}, A_2 \hat{x}, A_3 \hat{x}] \quad (44)$$

Since the Riccati equation (43) is a function of $\hat{x}(t|t)$, the $P(t)$ calculation in the EKF requires extensive on-line computation, which represents a considerable burden.

In the LQF, $P(t)$ is given as the solution of the coupled dynamic equations

$$\frac{d}{dt} P(t) = FP(t) + P(t)F' + \Delta(X(t), t) - \frac{1}{r} P(t)P'(t) \quad (45)$$

$$\frac{d}{dt} X(t) = FX(t) + X(t)F' + \Delta(X(t), t) \quad (46)$$

where $\Delta(X(t), t)$ is a diagonal matrix with i th component

$$\Delta(X(t), t)_i = q \sum_{k, \ell, m=1}^3 (A_m)_{ik} (A_m)_{i\ell} X_{k\ell}(t) \quad (47)$$

Notice that these equations for $P(t)$ and $X(t)$ can be calculated off-line, but the LQF thus has a considerable storage requirement. Because $P(t)$ in the FOF is only given by a memoryless nonlinearity, this filter requires considerably less storage than the LQF and less on-line computation than the EKF.

Our approach to the statistical analysis of the Monte Carlo simulations closely parallels that of Bucy and his associates [1], [3], [35]. The steady-state mean-squared error

$$\mu_2 = E[\|x(t) - \hat{x}(t)\|^2] = \sum_{i=1}^3 E[(x_i(t) - \tilde{x}_i(t))^2] \quad (48)$$

where $\tilde{x}_i(t)$ denotes the estimate of the i th component of the state $x_i(t)$, was used as the performance criterion. If $\{x^n\}$ and $\{\tilde{x}^n\}$, $i = 1, \dots, N$, are sequences of independent realizations of $x(t)$ and $\hat{x}(t)$, respectively, then the statistic

$$\mu_2 = \frac{1}{N} \sum_{n=1}^N \|x^n - \tilde{x}^n\|^2 \quad (49)$$

is an approximation to μ_2 for sufficiently large N . In fact, by the Central Limit Theorem [36, p. 278], $\hat{\mu}_2$ is asymptotically normal with

$$E[\hat{\mu}_2] = \mu_2 \quad (50)$$

$$\text{var} [\hat{\mu}_2] = \frac{1}{N} \left\{ \sum_{i=1}^3 (\mu_4)_i + 2(\mu_4)_{12} + 2(\mu_4)_{13} + 2(\mu_4)_{23} - (\mu_2)^2 \right\} \quad (51)$$

where $(\mu_4)_i \triangleq E[(x_i(t) - \tilde{x}_i(t))^4]$ and

$$(\mu_4)_{ij} = E[(x_i(t) - \tilde{x}_i(t))^2 (x_j(t) - \tilde{x}_j(t))^2].$$

Thus, for large N , with probability 0.9974, the 3σ confidence interval is given by

$$|\mu_2 - \hat{\mu}_2| \leq 3\sqrt{\text{var}(\hat{\mu}_2)}, \quad (52)$$

or equivalently,

$$\Pr \left\{ \frac{\hat{\mu}_2}{1 + 3\sqrt{\alpha}} \leq \mu_2 \leq \frac{\hat{\mu}_2}{1 - 3\sqrt{\alpha}} \right\} = 0.9974 \quad (53)$$

where

$$\alpha = \frac{\text{var}(\hat{\mu}_2)}{(\hat{\mu}_2)^2}. \quad (54)$$

In the Monte Carlo simulations, α was estimated from the samples (using sample means as in (49)), and approximate confidence intervals were thus computed.

In the experiment, 15 sample paths, each of which contained 1000 steps of length .001 seconds, were run in each simulation. The first 200 samples in each sample path were discarded to allow the transients to decay, so the remaining 800 samples represented steady-state. If all the steady-state errors were averaged as in (49), this would lead to 12000 samples of the steady-state error. However, as noted in [4], [24] and [35], adjacent errors in each path are correlated, so the effective Monte Carlo length is somewhere in the range between $N=1200$ and $N=12000$. The three standard deviation confidence intervals were calculated for both values of N .

The results of the simulations are presented in Table I. The 3σ confidence intervals I_1 (for $N = 12000$) and I_2 (for $N = 1200$) are shown. The results of this approximate statistical analysis of the Monte Carlo simulations indicate that, for this simple example, the FOF performs comparably to the LOF, and better than the EKF. The FOF seems to perform better in comparison to

the other filters as q increases, due to the increasing dominance of the bilinear noise term in the system equation (40). These results are significant, due to the fact that the FOF designed here requires considerably less storage and computation than the other filters (no additional differential equations or storage for $P(t)$ are required).

		M.S. Error $\hat{u}_2$	I_1 (N=12000)	I_2 (N=1200)
$q = 0.01$ $r = 0.01$	FOF	9.65	[9.41, 9.91]	[8.92, 10.52]
	EKF	10.21	[9.96, 10.48]	[9.45, 11.11]
	LQF	9.49	[9.26, 9.75]	[8.78, 10.34]
$q = 0.01$ $r = 1.00$	FOF	11.69	[11.39, 12.01]	[10.79, 12.77]
	EKF	11.75	[11.44, 12.07]	[10.84, 12.82]
	LQF	11.69	[11.39, 12.01]	[10.79, 12.77]
$q = 1.00$ $r = 0.01$	FOF	170.38	[165.26, 175.82]	[155.19, 188.87]
	EKF	193.75	[188.74, 199.04]	[178.74, 211.51]
	LQF	170.65	[165.58, 176.03]	[155.60, 188.92]

Table I. Monte Carlo M.S. Estimation Error(s)
($\times 10^{-3}$)

VIII. CONCLUSIONS

The state estimation problem for bilinear stochastic systems evolving on compact Lie groups and homogeneous space has been considered. The techniques of harmonic analysis on compact Lie groups have been applied to the design of suboptimal estimators for such systems. Monte Carlo simulations of a simple example indicate that a computationally simple filter designed by these methods performs favorably as compared to two other filters.

APPENDIX - HARMONIC ANALYSIS ON COMPACT LIE GROUPS

In this appendix we summarize some facts from the theory of harmonic analysis on Lie groups. For details of group representations see references [5,6,13,29,30].

Definition A.1: A finite-dimensional matrix representation of a compact Lie group G is a continuous homomorphism D which maps G into the group of nonsingular linear transformations on a finite dimensional vector space V . The representative ring of G is the ring generated over the field of complex numbers by the set of all continuous functions D_{ij} on G which are matrix elements of some irreducible representation D .

The Peter-Weyl theorem is the major result in harmonic analysis on compact Lie groups [5,6,29,30]. It gives the direct sum decomposition of $L_2(G)$, the space of square integrable functions with respect to the Haar measure dg :

$$L_2(G) = \bigoplus_{\alpha \in \Lambda} V_{\alpha} \quad (\text{A.1})$$

where Λ is the set of equivalence classes of finite dimensional irreducible representations of G , and V_{α} denotes the vector space spanned by the $(n_{\alpha})^2$ functions $\{D_{ij}^{\alpha}; i, j=1, \dots, n_{\alpha}\}$.

The Laplace-Beltrami operator (Laplacian) on a Lie group or Riemannian homogeneous space is discussed in [5], [13] and [29]. Some examples will be presented here, but one very important result for our purposes is the following.

Theorem A.1 [5, p. 40], [29, p. 257]: Let G be a compact Lie group, and let V_α be defined as in equation (A.1). Then each function $\varphi \in V_\alpha$ is an eigenfunction of the Laplacian Δ , and all $\varphi \in V_\alpha$ have the same eigenvalue λ_α . Conversely, each eigenfunction φ of the Laplacian is an element of the representative ring.

Hence, harmonic analysis on a compact Lie group can be performed either in terms of the representative ring or the eigenfunctions of the Laplacian, since these two sets of functions are the same. In this paper we are primarily concerned with the application of these results to the special orthogonal group $SO(n)$ and the n -sphere S^n .

The Lie group $SO(n)$ is defined by

$$SO(n) = \{X \in \mathbb{R}^{n \times n} \mid X'X = I, \det X = +1\} \quad (A.2)$$

The theory of representations of $SO(n)$ is discussed in [13], [26]; for this paper we need only consider $SO(3)$. Any matrix R in $SO(3)$ can be described in local coordinates in terms of the Euler angles ϕ, θ, ψ , domain $0 \leq \phi < 2\pi, 0 \leq \theta \leq 2\pi, 0 \leq \psi < 2\pi$ [6]. An element of $SO(3)$ will thus be denoted by $R(\phi, \theta, \psi)$ or just (ϕ, θ, ψ) .

In the Euler angle coordinates, the (unnormalized) Haar measure is

$$d\mu(\phi, \theta, \psi) = \sin\theta d\phi d\theta d\psi \quad (A.3)$$

and the corresponding Laplace-Beltrami operator is given by [31]

$$\Delta_{SO(3)} = \frac{1}{\sin\theta} \frac{\partial}{\partial\theta} \left(\sin\theta \frac{\partial}{\partial\theta} \right) + \frac{1}{\sin^2\theta} \left(\frac{\partial^2}{\partial\phi^2} - 2\cos\theta \frac{\partial^2}{\partial\phi\partial\psi} + \frac{\partial^2}{\partial\psi^2} \right) \quad (A.4)$$

Talman [6] computes a sequence $D^l(\phi, \theta, \psi)$, $l = 0, 1, \dots$, of unitary irreducible representations of $SO(3)$; its matrix elements are given by

$$D_{mn}^l(\phi, \theta, \psi) = i^{m-n} e^{-im\phi} d_{mn}^l(\theta) e^{-in\psi} \quad (A.5)$$

where

$$d_{mn}^{\ell}(\theta) = \sum_t (-1)^t \frac{[(\ell+m)! (\ell-m)! (\ell+n)! (\ell-n)!]^{\frac{1}{2}}}{(\ell+m-t)! (t+n-m)! t! (\ell-n-t)!} \cdot \cos^{2\ell+m-n-2t} \left(\frac{\theta}{2} \right) \sin^{2t+n-m} \left(\frac{\theta}{2} \right) \quad (\text{A.6})$$

for $-\ell \leq m, n \leq \ell$. Here t is summed over all nonnegative integers such that the arguments of the factorial functions in (A.6) are nonnegative; i.e.,

$$m-n \leq t \leq \ell+m, \quad 0 \leq t \leq \ell-n.$$

In fact, these are (up to equivalence) all of the irreducible representations of $SO(3)$. The Peter-Weyl Theorem yields the decomposition

$$L_2 SO(3) = \bigoplus_{\ell} V_{\ell} \quad (\text{A.7})$$

where H_{ℓ} is the vector space spanned by the $(2\ell+1)^2$ functions $\{D_{mn}^{\ell}; m, n = -\ell, \dots, \ell\}$.

The n -sphere $S^n = \{x \in \mathbb{R}^{n+1} | x'x = 1\}$ is diffeomorphic to the homogeneous space $SO(n)/SO(n-1)$. Harmonic analysis on S^n is studied in terms of the spherical harmonics [6, 30, 33]. The space H_{ℓ} of spherical harmonics of degree ℓ can be characterized as the eigenspace of the $SO(n+1)$ -invariant Δ_{S^n} on S^n with eigenvalue $-\ell(n-1+\ell)$; other equivalent characterizations are given in [13].

In particular, we consider the 2-sphere S^2 . Any point (x_1, x_2, x_3) on S^2 can be expressed in the polar coordinates (θ, ϕ) , where $0 \leq \theta \leq \pi$, $0 \leq \phi < 2\pi$, by defining

$$x_1 = \cos \theta; x_2 = \sin \theta \cos \phi; x_3 = \sin \theta \sin \phi. \quad (\text{A.8})$$

The Riemannian measure invariant under the action of $SO(3)$ is

$$d\mu(\theta, \phi) = \sin \theta \, d\theta \, d\phi. \quad (\text{A.9})$$

The corresponding invariant Laplace-Beltrami operator is

$$\Delta_{S^2} = \frac{1}{\sin \theta} \left[\frac{\partial}{\partial \theta} \left(\sin \theta \frac{\partial}{\partial \theta} \right) + \frac{1}{\sin \theta} \frac{\partial^2}{\partial \varphi^2} \right]. \quad (\text{A.10})$$

The normalized spherical harmonics of degree ℓ on S^2 are defined by [6]

$$Y_{\ell m}(\theta, \varphi) = (-1)^m \left[\frac{(\ell-m)!}{(\ell+m)!} \frac{(2\ell+1)}{4\pi} \right]^{\frac{1}{2}} P_{\ell m}(\cos \theta) e^{im\varphi} \quad (\text{A.11})$$

$$Y_{\ell, -m}(\theta, \varphi) = (-1)^m Y_{\ell m}^*(\theta, \varphi) \quad (\text{A.12})$$

for $\ell=0,1,\dots$ and $m=0,1,\dots,\ell$, where $P_{\ell m}(\cos \theta)$ are the associated Legendre functions and $*$ denotes complex conjugate. Notice that $Y_{\ell m}$ is an eigenfunction of Δ_{S^2} with eigenvalue $-\ell(\ell+1)$.

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