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**TECHNIQUES FOR SECONDARY
INTEGRATED FORECASTING
ITEM CLASSES
PART I - ACTIVE ITEMS**



**U.S. ARMY
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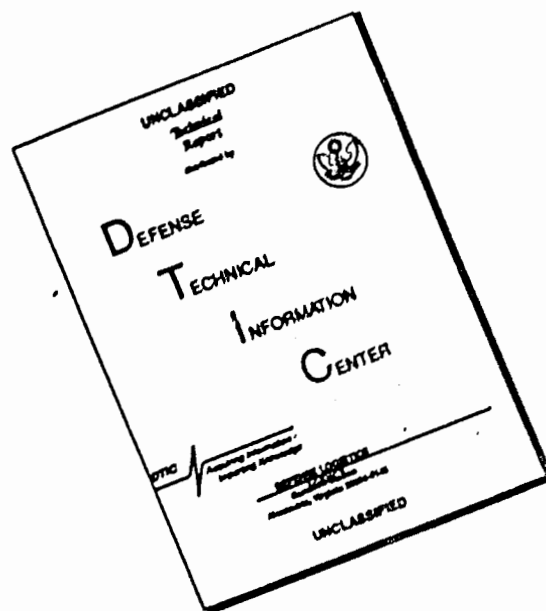
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classes and for interfacing across classes could be developed. However, the evaluation results as described in this report present no clear cut improvement to the current method for active items. In Part II, a new forecast procedure is recommended for inactive items.

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SUMMARY

This report is one of the final in a series on forecasting methods and forecast performance for Army demands in the wholesale supply system. An extended data base of 48 quarters is used and a summary of results from both old and new forecast methods is presented. The analysis is more intense than in [8]; forecast algorithms are used on various item activity classes with the intent of detecting patterns which could indicate where certain algorithms work best. It was hoped that ultimately a synthesis of procedures for forecasting by item classes would be developed. However, the evaluation results as found in the report present no clear cut improvement to the current method for active items. In [4], a new forecast procedure is recommended for inactive items.

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CHAPTER I

INTRODUCTION

1.1 Overview

This report is one of the final in a series on forecasting methods and forecast performance for Army item demands in the wholesale supply system. An extended data base of 48 quarters is used and both old and new algorithms are evaluated. A more intensive analysis is made herein than in [8]; forecast algorithms are used on various item activity classes with the intent of detecting patterns which could indicate where certain algorithms work best. It was hoped that ultimately a synthesis of procedures for forecasting by item classes and for interfacing across classes could be developed. However, the evaluation results as described in this report present no clear cut improvement to the current method for active items. In [4], a new forecast procedure is recommended for inactive items.

1.2 Scope

The specific reasons and scope for this report are listed:

- a. IRO continues to maintain and expand a data base of aviation parts' demand. The data base of report [3], quarterly data for 9700 items from 1967-1973, has been extended to a file of 13900 items for years 1967-1977. The analysis herein concentrated upon the years 1971-1977, the Vietnam war having less impact than in the earlier file.
- b. Several of the algorithms in [8] utilized forecast parameters (e.g. Kalman k-factors) which were obtained from statistical properties of aggregate time series of demand and of flying hours over that earlier 7-year time period. In this report we update these parameters (as well as compute parameters for new types of algorithms) based on a later 7-year period 1971-1977.
- c. Reports, [7], of superior performance of moving average forecasts incorporating a Trigg tracking signal by the TARCOM Systems Analysis Group led IRO to modify such a Trigg algorithm and to test it with the expanded data base.
- d. Preliminary success of the Trigg tracking signal led IRO to develop other refined algorithms which utilized other switching signals for fixing the length of past history used in a current forecast.

e. Changes have been made in the supply simulator since it was used as one of the performance measuring tools in [8]. Additional improvements developed during the course of the work have been included in the final simulation analyses.

1.3 Findings

For the cursory reader, the table on page 6 is probably the best culmination of the many results and tables in the body of the report. At a glance one can see the basic breakout of the analysis per stratification class. (See bottom of page). A 95% confidence interval was plotted for the mean difference in performance (as measured by the simulator) per item between the alternative algorithms and standard (1794) for the indicated stratification classes. In each case, none of the alternative algorithms performed significantly better than the standard (each confidence interval contains zero).

The final candidate forecast algorithms considered for this analysis were:

1794: the current 8 quarter moving average

KAL: a modified Kalman filter (see [9]) which is similar to exponential smoothing.

IROTRIGG: a switching scheme between an 8 quarter and 4 quarter moving average, with some extra weight given to the current quarter's demand.

MED4: a 4 quarter moving median.

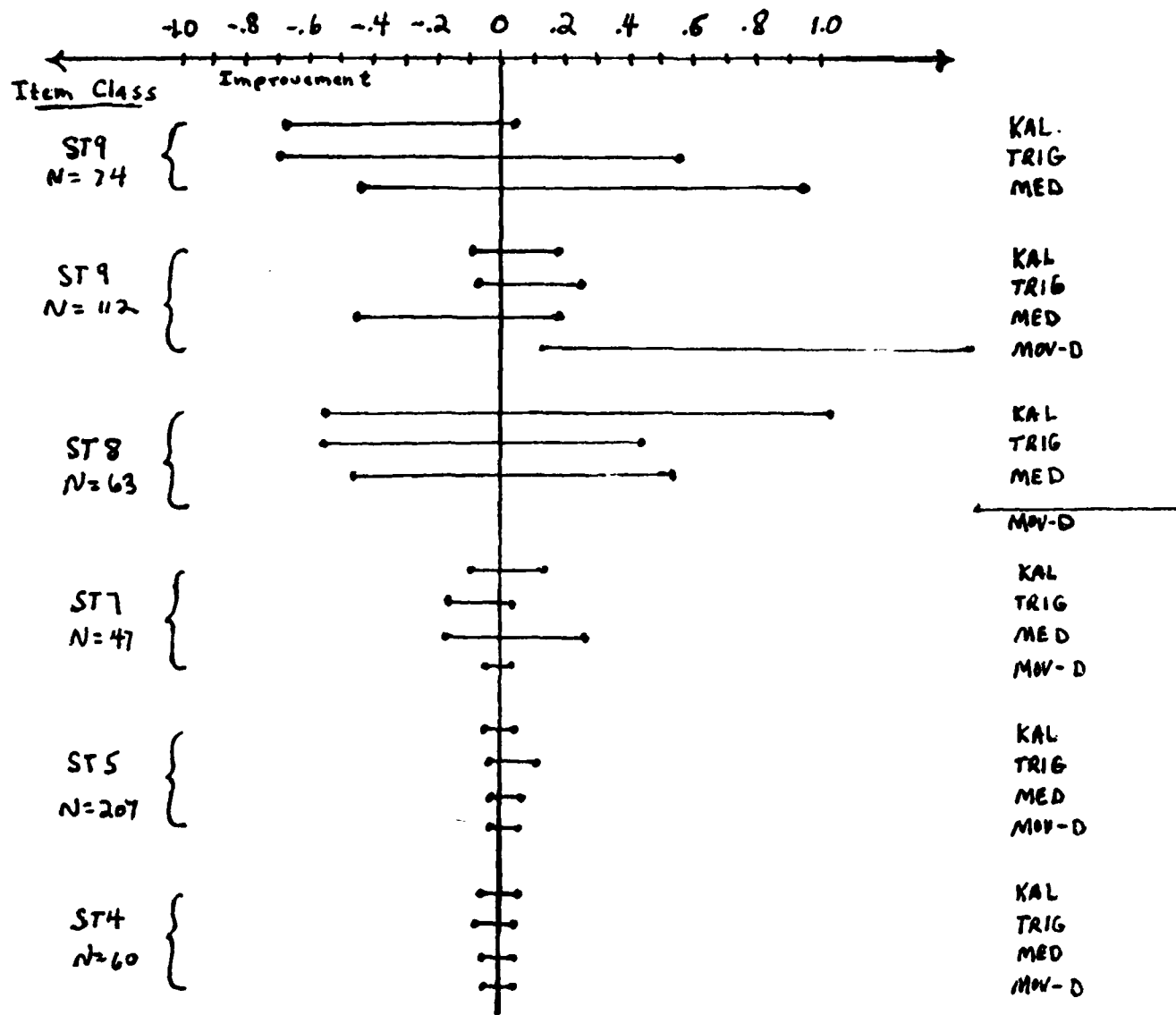
MOVD: a 8 quarter moving average on demands only.

It should be noted that these findings are limited to the more active classes of items. Report [4] has been written which deals specifically with items exhibiting low demands.

		YEARLY DOLLAR DEMAND		
STRATIFICATION CLASSES	YEARLY REQUISITIONS	0 - \$5000	\$5000 - \$50000	> \$50000
	0-3	ST 1 N=335	ST 2 N=100	ST 3 N=4
	3-12	ST 4 N=124	ST 5 N=230	ST 6 N=17
	>12	ST 7 N=98	ST 8 N=64	ST 9 N=115

95% Confidence Intervals For The Mean of Individual Differences Between Time Weighted Backorders of (test policy - std Policy)

1794 = Std
Policy



CHAPTER II

DATA

2.1 Description

The IRO demand history file includes 11 years of requisitions and demand by quarter accumulated from the AVSCOM Demand Return & Disposal (DRD) files from 1967 thru 1977. Flying hours covering the same period as the demand data were obtained from DCSLOG. The file contains a sample of 20,865 items from all those in the system in 1966 and subsequently entered.

The data base is limited to recurring demands for which program data was available. SSA and Grant Aid demands were eliminated as were items not purchased thru central procurement, based on the last recorded IMPC code. Every attempt was made to drop items subject to logistical transfer.

Previous IRO forecasting projects used an older 7-year data base [3], compiled in much the same way from the DRD files. The only significant difference in the new data base is the inclusion of items with trivial demand, essential for forecasting demand for inactive items.

2.2 Classification

Each item was classified as low dollar value (LDV) or high dollar value (HDV) according to whether the demand rate averaged over the 11 years was less than or greater than or equal to \$50,000; and the requisition rate was less than or greater than or equal to 100 per year. Items with over a million dollars of demand per year were dropped.

The items were further divided into dynamic (DYN) and non-dynamic (NON) based on the Federal Stock Class (FSC). The dynamic components were considered to be those that experience high rotation rates; i.e. rotor blades, transmissions, and turbine engines. For more detail see Cohen [3].

The data breaks out into the following four groups:

HDVDYN	86
HDVNON	262
LDVDYN	1169
LDVNON	<u>19348</u>
	20865

Some of the simulation and other forecasting work was done using the last 7 years of the 11 year data base, thus eliminating the 1967-1970 period

subject to the Vietnam war. LDV and HDV divisions were recreated based on the demand in the last 7 years. The new breakdown of items for this period 1971-1977 is:

HDVDYN	54
HDVNON	224
LDVDYN	1199
LDVNON	<u>19384</u>
	20861

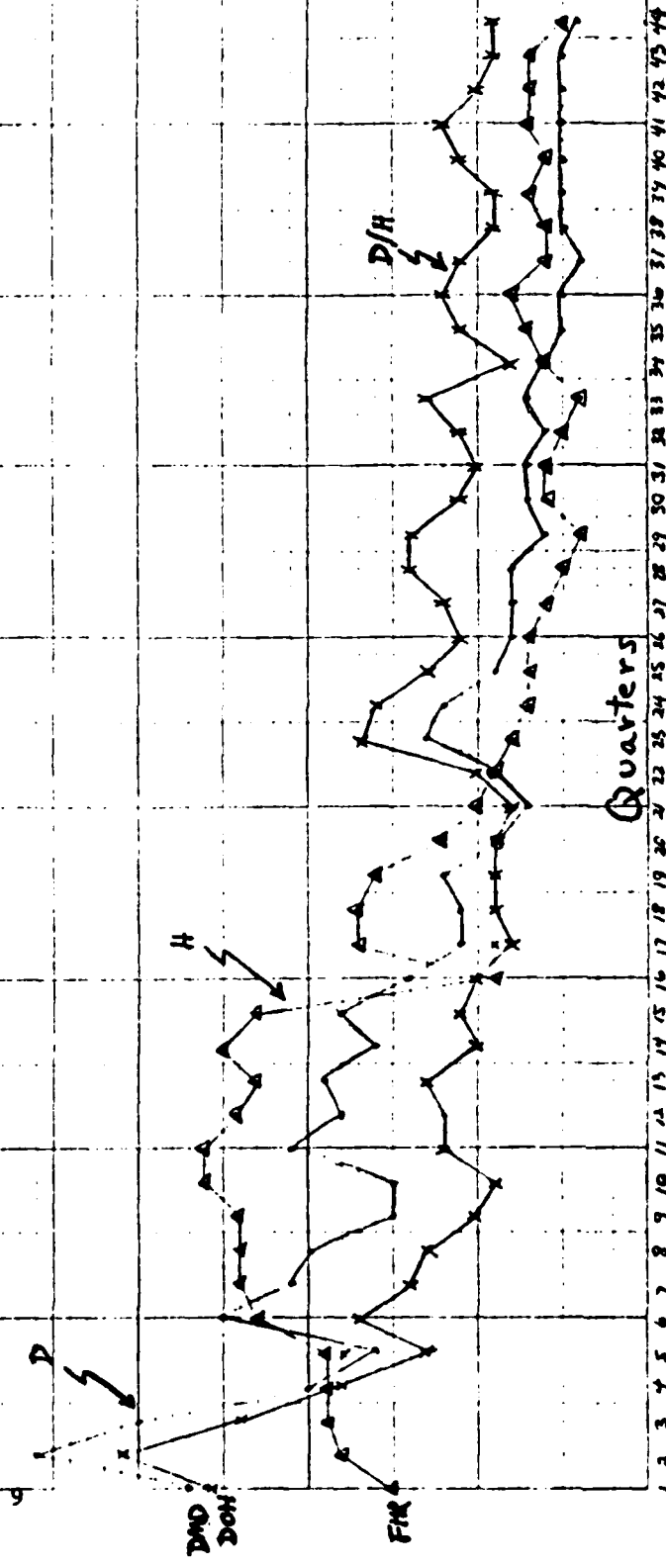
Four items were eliminated which had over one million dollar demand, based on the last 7 years.

2.3 Aggregate Series

The graph on page 9 illustrates the aggregate series of all the items for demand (D), flying hours (H), and demand divided by flying hours (D/H).

Normalized, Aggregate Time Series 1967-1977
 Demand, Flying Hours, Demand per Flying Hour

•	Δ	x
D	H	D/H



CHAPTER III

KALMAN FILTERS

3.1 Basic Forecast Methods

As noted in references [8], [9], "Kalman Filters" are, for a general class of statistical processes, optimal forecast procedures, in the sense of minimizing mean square forecast error. These algorithms have a general exponential smoothing structure where the smoothing weights themselves are variable and updated.

In this section the most important Kalman algorithms that were tested are briefly summarized and designated with code names. Complete mathematical descriptions are found in the Appendix to this chapter. All algorithms operate on the demand per flying hour time series D/H in order to predict a rate value. Forecasts for demand in a future period are then made by multiplying the rate estimate D/H by the program (flying hours) for that future period.

- KALMAN - Original Kalman filter algorithm investigated in reference [8]. Updated k factors for D/H for the period 1971-1977 are used.
- KALNEW1 - Kalman filter with a modification to the formula for updating the weights. New $k_{D/H}$ factors are used. (See Section 3.2)
- KALMANS - A switching signal is used to choose either the current observation of D/H or the KALMAN estimate.
- KALMANSMA2 - As in KALMANS but replace current observation by current 2 quarter average.
- KALNEW1S - As in KALMANS but uses the KALNEW1 estimate.
- KALNEW1SMA2 - As in KALNEW1S but replace current observation by current 2 quarter average.
- KALNEW1SMA3 - As in KALNEW1S but replace current observation by current 3 quarter average.
- KALREL - Current weighting value in original KALMAN procedure is adjusted by a "relevance" function which tends to put more weight on current observations.
- KAL2SPK - An a priori two spike distribution on zero value and a non-zero value for $k_{D/H}$ is assumed. Bayesian updating is applied to the spike probabilities and the estimates from 2 Kalman filters (for $k = 0$ and $k \neq 0$) are appropriately weighted.

3.2 K-factors for Kalman Algorithms

A parameter k_y is needed to obtain the changing values of weights applied to past and current observations of a time series variable y in a Kalman forecasting procedure. The details on the mathematical theory of Kalman k factors and on the statistic formed from empirical forecast errors can be found in [9]. Suffice it here to say that the mean square error on y over L periods when using a moving average of B periods is a function of k and of a variance q^2 of the process mean. The functional relations can be solved to find estimates of k .

The tables present the average k values for D and D/H processes for items falling in various requisition classes. The 1967-1973 table is a refinement of the table for that same period in [8], a result of correcting a not completely innocuous bias in data processing; k_D did not change much, but $k_{D/H}$ now does not increase continuously as requisition activity goes up. The period 1971-1977 indicates generally higher values of k_D , $k_{D/H}$; hence relatively more stability in the processes D , D/H than in the Vietnam era. There has also been a shift in the column trends in the later era, an indication of a change in demand patterns by requisition class. The 1971-1977 k values are probably more representative of "normality" but should be updated about every three years or with changes in war or economic environment.

3.3 Appendix A - Mathematical Addendum

Basic Notation

y_n = observed value of process in period n .

For all algorithms listed, Demand/Flying Hours, D/H was the utilized observation variable y in the empirical analysis.

$\bar{x}_n$ = mean of process in period n .

$\hat{x}_n$ = estimate of mean of process at end of period n .

$\hat{y}_n(l)$ = forecast at end of period n of the process value l periods later.

Except for a few cases with assumed deterministic trend components in the process model y (which did not perform well), our processes assume stochastic fluctuations in the future around the current mean; therefore a best MSE forecast $\hat{y}_n(l) = \hat{x}_n$ is used in the cases below.

Parameters for D, D/H Processes - Years 1967-1973
(by requisition class)

Upper Bound	# Items	Avg Req.	DEMAND		smoothed q^2	DEMAND per Flying Hour $K_{D/H}$
			K_D	q_D^2	q^2	
1	1637	.61	0	-		1.561
2	1554	1.52	3.368	9.65	10 3.2	5.446
3	1314	2.50	4.654	28.47	15 3.9	8.250
4	973	3.49	4.728	76.05	20 4.5	9.527
5	703	4.50	5.749	75.9.	40 6.3	11.987
6	549	5.51	4.724	33.27	80 8.9	10.016
8	866	6.96	4.730	87.90	110 10.5	10.246
12	1053	9.74	4.635	190.19	190 13.8	10.523
18	844	14.69	4.058	290.33	290 17.0	11.029
∞	1429	41.39	2.642	1869.97	1200 34.6	10.145

Parameters for D, D/H Processes - Years 1971-1977
(by requisition class)

Upper Bound	# Items	Avg Req.	DEMAND		smoothed q^2	DEMAND per Flying Hour K _{D/H}
			K _D	q^2		
1	7176	.59	0	-		0
2	1723	1.45	11.503	2.79	5	18.358
3	916	2.46	12.409	17.91	15	17.381
4	595	3.50	11.321	15.51	20	16.076
5	430	4.49	10.550	42.52	35	17.232
6	359	5.48	10.164	32.51	50	13.287
8	499	6.94	9.691	114.07	93	13.265
12	596	9.83	9.550	115.79	140	13.760
18	525	14.68	8.168	214.39	205	11.425
∞	1089	52.03	7.128	576.77	575	10.356

H_{n+L} = flying hours in future period $n+L$.

$H(n,L)$ = total flying hours, period $n+1$ to $n+L$.

$\hat{D}(n,L)$ = forecast of demand over L periods based on information thru end of period n

$$= \sum_{i=1}^{n+L} \hat{y}_n(i) H_{n+i} = \hat{x}_n H(n,L)$$

KALMAN

$$\hat{x}_n = \hat{x}_{n-1} + G_n (y_n - \hat{x}_{n-1}) \quad (A1)$$

$$G_{n+1} = \frac{1 + k G_n}{1 + k G_n + k (H_n^2 / H_{n+1}^2)} \quad (A2)$$

$$\hat{y}_n(i) = \hat{x}_n$$

k is updated every 4 quarters based on the current 8 quarter moving average estimate of yearly requisitions. (Table lookup)

KALNEW1

As above with

$$G_{n+1} = \frac{1 + k G_n (H_{n+1}^2 / H_n^2)}{1 + k G_n (H_{n+1}^2 / H_n^2) + k} \quad (A3)$$

KALMANS

In reference [] the formula for the MSE of a one period (quarter) forecast for moving average of M periods applied to the Dynamic Mean process is

$$MSE_M = q^2 \left[k + k/M + \frac{(M+1)(2M+1)}{6M} \right] \quad (A4)$$

So in terms of two moving average mean square errors, solving (A4) for k ,

$$K = \frac{(\frac{MSE_N}{N} \cdot C_M) - (\frac{MSE_M}{M} \cdot C_N)}{(\frac{N+1}{N} \cdot C_M) - (\frac{M+1}{M} \cdot C_N)} \quad (A5)$$

$$\text{where } C_I = 1 + \frac{(I-1)(2I-1)}{6I}$$

For $N = 4$, $M = 8$ and $R = MSE/MSE_4$

Then (A5) leads to

$$k \leq .5 \text{ if and only if } R \geq 1.5 \quad (A6)$$

We approximate the effect of k inferred to be quite small by putting all weight ($G_n=1$) in a Kalman algorithm on the current observation y_n .

The above relations lead to the following heuristic, hybrid algorithm employing a switching signal.

Let

$\hat{z}_n(j) = j$ quarter moving average after period n .

$$MSE_n(j) = (1-\alpha) MSE_{n-1}(j) + \alpha (\hat{z}_{n-1}(j) - y_n)^2 \quad (A7)$$

Then

if $[MSE_n(8)/MSE_n(4) > 1.5], \hat{y}_n(l) = y_n$

otherwise $\hat{y}_n(l) = \hat{x}_n = \hat{x}_{n-1} + G_n(y_n - \hat{x}_{n-1})$

where G_n is given by (A2).

α was chosen to be .8 based on empirical testing.

KALMAN SMA2

The reasoning applied here to obtain a heuristical switching signal algorithm is similar to that of KALMANS. We wish to know when it is appropriate to use an MA2(2 quarter moving average) on y (i.e. D/H).

Note in (A2), that when flying hours are stable ($H_n \approx H_{n+1}$), G approaches .67 for $k = .7$. An exponential smoothing weight of .67 is equivalent (in the sense of average weighting of all past history) to a moving average base $B = 2$.

For $N = 2$, $M = 8$ and $R \equiv MSE_8/MSE_2$, then (A5) leads to:

$$k \leq .7 \text{ if and only if } R \geq 1.728 \quad (A8)$$

Then using the notation of (A7), define KALMAN SMA2 as

$$\text{If } [MSE_n(8)/MSE_n(2) > 1.728], \hat{y}_n(l) = \frac{y_n + y_{n-1}}{2}$$

$$\text{Otherwise } \hat{y}_n(l) = \hat{x}_n = \hat{x}_{n-1} + G_n(y_n - \hat{x}_{n-1})$$

where G_n is given by (A2)

KALNEW1S

Use G_n computed via (A3) in KALMANS procedure.

KALNEW1SMA2

Use G_n computed via (A3) in KALMANSMA2 procedure.

KALNEW1SMA3

Since $k = 2$ implies steady state $C = .5$, which in turn indicates a moving average base $B = 3$ (see reasoning in KALMANSMA2), equation (A5) is used with $N = 3$, $M = 8$ to find

$$k \leq 2 \text{ if and only if } \text{MSE}_8 / \text{MSE}_3 \geq 1.288$$

Hence the following heuristic for KALNEW1SMA3:

$$\text{if } [\text{MSE}_n(8) / \text{MSE}_n(3) > 1.288], y_n(l) = \frac{y_n + y_{n-1} + y_{n-2}}{3}$$

$$\text{Otherwise } \hat{y}_n(l) = \hat{x}_n = \hat{x}_{n-1} + G_n(y_n - \hat{x}_{n-1})$$

where G_n is given by (A3)

KALREL

$$G_{n+1} = 1 + \frac{r}{1 + A_{n+1}}, \quad r = 1 \quad (\text{A9})$$

$$\text{where } A_{n+1} = \frac{1 + k G_n}{1 + k H_n^2 / H_{n+1}^2} \quad (\text{A10})$$

If $0 \leq r < 1$, then more weight is applied to the observation y_{n+1} . The factor r can be defined as the value of a relevance function, which measures in some manner the relevance of the current observation. Extending the theory of [3a] to our models, a relevant relevance function may be formulated, viz.,

$$r = \min \left\{ 1, \left(\frac{2.5 H_{n+1} \hat{q}_D}{k H_n^2} \right) \cdot \frac{(1 + k G_n + k H_n^2 / H_{n+1}^2)}{|y_{n+1} - \hat{x}_n|} \right\} \quad (\text{A11})$$

where $\hat{q}_D$ = standard deviation of demand mean based on items' requisition class. See Table

2.5 is an adjustable value which here bounds the difference between the new estimate $\hat{x}$ and the observation to be no more than 2.5 standard deviations of the process.

This procedure hedges towards y_{n+1} as a new estimate if the error magnitude $|y_{n+1} - \hat{x}_n|$ is large.

KAL2SPK

A two-spike distribution for the item's k-factor is assumed - around values $k = 0$ and $k = k_{D/H}$, where the latter is tabulated by the item's requisition class. The probabilities z_n of $k = k_{D/H}$ are updated as below and used in combining the Kalman estimate $\hat{x}_n$ from (A1) for $k \neq 0$ with y_n , the Kalman estimate for $k = 0$, i.e.

$$\hat{y}_n(k) = z_n \hat{x}_n + (1-z_n) y_n \quad (A12)$$

If w_n , v_n are proportional to the probabilities p (error in period n/k) for $k = k_{D/H}$ and $k = 0$ respectively and assuming normality of errors, then it can be shown for the Dynamic Mean model,

$$w_n = \frac{1}{\sqrt{B}} \exp \left[- (y_n - x_{n-1})^2 / 2 BV_0 \right] \quad (A13)$$

$$v_n = \frac{1}{\sqrt{B}} \exp \left[- (y_n - y_{n-1})^2 / 2 V_0 \right] \quad (A14)$$

where

$$V_0 = \hat{q}_D^2 / H_{n-1}^2, \hat{q}_D^2 = \text{tabulated variance of demand mean.}$$

$$B = \left(1 + k \frac{H_{n-1}^2}{H_n^2} + k G_{n-1} \right)$$

BV_0 = theoretical variance of error.

Straightforward application of Bayesian updating yields new z_n

$$z_n = \frac{z_{n-1} w_n}{z_{n-1} w_n + (1-z_{n-1}) v_n} \quad (A15)$$

REGKBNEWSMA2

This algorithm, see [9], is the closest weighted average analogue to a Kalman filter's weighting of past history; in addition a switch is incorporated ala' KALMANSMA2. Hence referring to the section of KALMANSMA2

$$\text{if } [\text{MSE}_n(8)/\text{MSE}_n(2) > 1.728], \hat{y}_n(l) = \frac{y_n + y_{n-1}}{2}$$

otherwise

$$\hat{y}_n(l) = \sum_{j=1}^B w_{n-j+1} y_{n-1+1} \quad (\text{A16})$$

$$\text{where } w_{n-j+1} = H_{n-j+1}^2 / \sum_{i=1}^B H_{n-i+1}^2$$

Base period B is determined from the nearest integer $[\cdot]$

$$B = \left\lceil \sqrt{\frac{1+6k}{2}} \right\rceil \quad (\text{A17})$$

and k is updated every 4 quarters based on the current 8 quarter moving average of yearly requisitions (Table lookup on $k_{D/H}$).

3.4 Appendix B - Structural Changes in KALMAN to obtain KALNEW1

With a slight change in the assumed process model, the Dynamic Mean model (see [9] Chapter II), the updating formula for the weight in period $n+1$ applied to the current observed value y_{n+1} of the process becomes

$$G_{n+1} = \frac{q_{n+1}^2 + r_n^2 G_n}{q_{n+1}^2 + r_n^2 G_n + r_{n+1}^2} \quad (\text{B1})$$

instead of equation 2.5 in [9]. If as in [9] Chapter IV, it is assumed that r_n^2 , the variance of $y_n \approx D_n/H_n$, varies with $1/H_n^2$, and that $r_n^2/q_n^2 = k_{D/H}$, then

$$G_{n+1} = \frac{1 + (H_{n+1}^2/H_n^2) k_{D/H} G_n}{1 + (H_{n+1}^2/H_n^2) k_{D/H} G_n + k_{D/H}} \quad (\text{B2})$$

as opposed to

$$G_{n+1} = \frac{1 + k G_n}{1 + k G_n + k H_n^2/H_{n+1}^2} \quad (\text{B3})$$

as given in [].

Equation (B2) is the basis for the KALNEW algorithms. Note as k becomes large (B2) and (B3) behave similarly.

CHAPTER IV

IRO TRIGG TRACKING SIGNAL

4.1 Background

The IRO-TRIGG is a modified version of the TACOM New Parameter method described in [7]. It utilizes a TRIGG tracking signal to determine whether the series is stationary (constant mean) or not and uses an adjusted eight or four quarter moving average respectively as the forecast.

The tracking signal is the ratio between a weighted average of the algebraic (signed forecast) errors and absolute (non signed) error, while using a standard eight quarter moving average (MA8) as the forecast. If the series is stationary, the MA8 will do well in forecasting the constant mean of the system, and the over and under forecast errors will be on average algebraically sum to zero. On the other hand, if the series is following a trend, then the MA8 will always underforecast (trend positive) or over forecast (trend negative) and thus the algebraic sum of the errors will differ by the absolute sum only in sign.

Thus if the signal is close to ± 1 , then the series is following a trend whereas if the signal is close to 0, then the series is stationary with a constant mean.

4.2 Computation

Let:

$$MA8(n) = 1/8 \sum_{i=n-7}^n Y_i, \text{ where } Y_i = \frac{\text{Demand}}{\text{Flying Hours}} \text{ in quarter } i$$

$$MA4(n) = 1/4 \sum_{i=n-3}^n Y_i$$

$$AMA8(n) = \alpha_1 Y_n + (1-\alpha_1)MA8(n) \quad (0 \leq \alpha_1 \leq 1)$$

$$AMA4(n) = \alpha_2 Y_n + (1-\alpha_2)MA4(n) \quad (0 \leq \alpha_2 \leq 1)$$

$$ERROR(n) = Y_n - MA8(n-1)$$

$$MALE(n) = \alpha ERROR(n) + (1-\alpha)MALE(n-1) \quad (0 \leq \alpha < 1)$$

(Mean Algebraic Error)

$$MABE(n) = \alpha / ERROR(n) + (1-\alpha)MABE(n-1)$$

(Mean Absolute Error)

$TS(n) = MALE(n)/MABE(n)$
(Tracking Signal)

Forecast

$$\hat{Y}_{n+1} = \begin{cases} AMA8(n) & \text{if } |TS(n)| < \beta \\ AMA4(n) & \text{if } |TS(n)| \geq \beta \end{cases} \quad (0 < \beta < 1)$$

Initialization for Tracking Signal

Ref [].

Let:

$$MA7(7) = 1/7 \sum_{i=1}^7 Y_i$$

$$ERROR(8) = Y_8 - MA7(7)$$

$$MALE(7) = 0$$

$$MABE(7) = (1.25/7) \sum_{i=1}^7 Y_i - MA7(7) /$$

Then

$$TS(8) = \frac{\alpha ERROR(8)}{\alpha |ERROR(8)| + (1-\alpha)MABE(7)}$$

4.3 IRO-TRIGG Parameters

The IRO-TRIGG forecast procedure depends on certain parameters which have to be determined prior to the forecast. These parameters may be selected subjectively or empirically by experimentation. Remarks and observations concerning these parameters are as follows:

Remarks

a. The α_1 and α_2 parameters define the weights given to the past observation when computing $AMA8(n)$ and $AMA4(n)$, i.e.

$$AMA8(n) = W_1 \sum_{i=n-7}^{n-1} Y_i + W_2 Y_n$$

$$\text{where } W_1 = 1/8 (1-\alpha_1)$$

$$\text{and } W_2 = 1/8 (1+\alpha_1)$$

$$\text{Hence as } \alpha_1 \rightarrow 0, AMA8(n) \rightarrow MA8$$

$$\text{and as } \alpha_1 \rightarrow 1, AMA8(n) \rightarrow Y_n$$

Similarly

$$\text{AMA4}(n) = W_1^* \sum_{i=n-3}^{n-1} Y_i + W_2^* Y_n$$

$$\text{where } W_1^* = 1/4 (1-\alpha_2)$$

$$\text{and } W_2^* = 1/4 (1-3\alpha_2)$$

$$\text{also as } \alpha_2 \rightarrow 0, \text{AMA4}(n) \rightarrow \text{MA4}$$

$$\alpha_2 \rightarrow 1, \text{AMA4}(n) \rightarrow Y_n$$

So if α_2 is small and α_1 is large

then the AMA8(n) will do better than AMA4(n) in forecasting a non-stationary (trend) series,

also AMA4(n) would do better than AMA8(n) in forecasting a stationary series.

Since this is counter to the logic of the switching process, careful judgement should be made when selecting the combination of values for α_1 and α_2 .

b. The α parameter determines the amount of weight given to the latest error when using MA8 for computing the tracking signal.

$$\text{TS}(n) = \frac{\alpha \text{ ERROR} + (1-\alpha) \text{ MALE}(n-1)}{\alpha |\text{ERROR}| + (1-\alpha) \text{ MABE}(n-1)}$$

(Tracking Signal)

$$\text{Now as } \alpha \rightarrow 1, \text{TS}(n) \rightarrow \frac{\text{ERROR}}{|\text{ERROR}|} \rightarrow \pm 1$$

independent to the type of series being forecasted. Therefore for α large,

the TRIGG signal may incorrectly identify a trend condition.

c. The β or threshold parameter identifies the region in which the tracking signal indicates a stationary or non-stationary series. The larger the β the more confidence is given to correctly identifying a non-stationary series, but less confidence is given to correctly identifying a stationary series. From the previous paragraph, it is obvious that α and β are related and that as α gets large β should likewise get large in order to maintain the same level of confidence.

Observations: (Results from experimentation with data base)

- a. We found that $\alpha_1 = \alpha_2 = .15$ worked as well or better than any other values selected for experimentation and these values were somewhat robust.
- b. The best α level was found to be .66, contrary to the small α implied in the remarks. The switching process was very sensitive to changes in this parameter. For $\alpha = .66$, the tracking signal indicated a non-stationary process 74% of the time.
- c. It was surprising that changes in the β level made little difference in the statistical results. β of .5 was chosen.
- d. The IRO-TRIGG with parameters $\alpha_1 = \alpha_2 = .15$, $\alpha = .66$, and $\beta = .5$ was the final model chosen to represent the tracking signal technique. This method performed better than the other alternatives not listed in the report.

4.4 Structural Change to the TACOM New Parameter Algorithm

As mentioned earlier the IRO-TRIGG forecast algorithm is a modified version of the TACOM New Parameter algorithm cited in [7]. The changes as described below were made in an effort to make the forecast technique more consistent with the underlying theory of the TRIGG tracking signal.

a. Computing the Tracking Signal. Instead of using the errors from the actual forecast as done by TACOM, the IRO version uses the errors from an eight quarter moving average forecast in computing the tracking signal. By so doing the tracking signal becomes a monitor for trends. The eight quarter moving average will lag behind any trend in the data which will result in a tracking signal close to $\pm$ one. Since the tracking signal is computed independently of the actual forecast, it is not effected by any switch in the forecast technique and will continue to indicate a trend as long as there is one.

b. α_1 and α_2 Constraints. The TACOM New Parameter algorithm utilizes two empirically found parameters α_1 and α_2 which adjust the amount of weight given to the current observation in both the eight and four quarter moving average computations ($0 \leq \alpha_1 \leq 1$), ($0 \leq \alpha_2 \leq 1$). The IRO-TRIGG version also uses these parameters but constrains their values so that the adjusted four quarter average still responds faster to changes (trends) in the data than the adjusted eight quarter average. (This will not be the case if α_1 is sufficiently large and α_2 is sufficiently small as noted under remarks, section a.).

c. An Additional Empirical Parameter α . When computing the TRIGG tracking signal, an exponentially smoothed algebraic error is compared with an exponentially smoothed absolute error. The TACOM New Parameter algorithm uses a smoothing constant of .66 for these calculations. The IRO-TRIGG version considers this α smoothing constant as an empirical parameter which needs to be estimated.

CHAPTER V

CURRENT AND OTHER FORECAST METHODS

5.1 Current - 1794

The current Army method of forecasting estimates the demand per program (flying hours) at end of period* n using

$$\hat{x}_n = \frac{\sum_{j=1}^8 D_{n-j+1}}{\sum_{j=1}^8 H_{n-j+1}}$$

This may be written in terms of a weighted moving average on $y = D/H$:

$$x_n = \sum_{j=1}^8 W_{n-j+1} Y_{n-j+1}$$

$$W_{n-j+1} = \frac{H_{n-j+1}}{\sum_{i=1}^8 H_{n-i+1}}$$

5.2 Moving Median (MED4)

In an effort to eliminate the effects of spikes in the data, a simple four quarter moving median on D/H as described below was used.

Let $s(n) = \{y_{n-3}, y_{n-2}, y_{n-1}, y_n\}$ be the last four observations

and let

$(y_{1,n}, y_{2,n}, y_{3,n}, y_{4,n})$ be the ordered array of this set

where $y_{i,n} \leq y_{j,n}$ for $i < j$

Then

$$MED4(n) = \frac{y_{2,n} + y_{3,n}}{2}$$

and

$$\hat{y}_{n+1} = MED4(n)$$

Several data experiments were performed on the length of the base period and the four quarter base appeared to perform best.

* A period = one quarter of a year, so 8 quarters or 2 years of history is used.

5.3 Moving Average on Demand

To support Cohen's [3] findings that the use of program data (flying hours) improves forecasts a simple eight quarter moving average was applied to the demand series D.

Let $\{x_{n-7}, x_{n-6}, \dots, x_n\}$ be the latest eight observation

$$\text{Then } \hat{x}_{n+1} = \sum_{i=0}^7 x_{n-i} / 8$$

CHAPTER VI

EXPERIMENTAL DESIGN

6.1 Overview

From our previous experiences we have found that there is no clear cut way to evaluate forecast algorithms in an inventory management system. With this in mind, an ad hoc sequential step wise experiment was designed where both forecast methods and evaluation procedures were eliminated and/or refined after each step. As a result both a best forecast algorithm and a best evaluation method may be determined. The details of each step are as follows:

6.2 Step 1: (13 forecast algorithms, 9 statistical error measures)

For the first step, the four major data sets described on page were used. After appropriate initialization, each forecast algorithm was used to make a one quarter and a four quarter forecast for every quarter of each series. These forecasts were compared to the actual demand of the series and the errors were rolled up within classes of series. Various error measures were computed for 13 forecast algorithms, 6 series classes, and 2 forecast horizons. The experimental layout is as follows:

Data Sets:

Low Dollar Value Non-Dynamic (LDVNON) consisting of 54 items

Low Dollar Value Dynamic (LDVDYN) consisting of 224 items

High Dollar Value Dynamic (HDVDYN) consisting of 1199 items

High Dollar Value Non-Dynamic (HDVNON) consisting of 19384 items

Forecast Horizons

1 quarter

4 quarters

ERROR Measures:

Let x_{ij} be the demand (for the i th item) in the j th quarter

F_i = index set of forecasts for i th item

N_i = cardinal size of F_i (number of times a forecast was made)

E_{ij} = error ($x_{ij} - \hat{x}_{ij}$), $j \in F_i$

AVG_i = average demand for item i

(double 12 month moving average starting after the
first non-zero demand)

Simple Averages

The first error measures considered were simple averages of traditional measures.

$$\text{MAD} = \frac{1}{N} \sum_{i=1}^N \frac{1}{n_i} \sum_{j \in F_i} |E_{ij}|$$

(Mean Absolute Error)

$$\text{MSE} = \frac{1}{N} \sum_{i=1}^N \frac{1}{n_i} \sum_{j \in F_i} (E_{ij})^2$$

(Mean Square Error)

$$\text{BIAS} = \frac{1}{N} \sum_{i=1}^N \frac{1}{n_i} \sum_{j \in F_i} E_{ij}$$

Percent Error Measures

The simple averages give more weight to items with high demand frequency. Since the items were stratified into homogeneous classes it was desirable to give equal weights to each item in the class, hence the following percent error measures were considered.

$$\text{AVG \% of Forecast} = \frac{1}{N} \sum_{i=1}^N \frac{1}{n_i} \sum_{j \in F_i} \frac{|E_{ij}|}{\hat{x}_{ij}}$$

$$\text{AVG \% of Actual} = \frac{1}{N} \sum_{i=1}^N \frac{1}{n_i} \sum_{j \in F_i} \frac{|E_{ij}|}{x_{ij}}$$

$$\text{AVG \% of Both} = \frac{1}{N} \sum_{i=1}^N \frac{1}{n_i} \sum_{j \in F_i} \frac{|E_{ij}|}{1/2 (x_{ij} + \hat{x}_{ij})}$$

Relative Error Measures

Now since many of the series were quite variable, the denominator of the percent error measures did not reflect the steady state demand of the item hence the following relative measures were considered.

$$\text{RELATIVE BIAS} = \frac{1}{N} \sum_{i=1}^N \frac{1}{n_i} \sum_{j \in F_i} \frac{E_{ij}}{\text{AVG}_i}$$

$$\text{RELATIVE MAD} = \frac{1}{N} \sum_{i=1}^N \frac{1}{n_i} \sum_{j \in F_i} \frac{|E_{ij}|}{\text{AVG}_i}$$

$$\text{RELATIVE MSE} = \frac{1}{N} \sum_{i=1}^N \frac{1}{n_i} \sum_{j \in F_i} \frac{E_{ij}^2}{\text{AVG}_i^2}$$

Forecast Algorithms: (refer to Chapters III, IV, V for definition)

KALMAN	IROTRIGG15-75
KALNEW1	IROTRIGG25-25
KALMANS	REGKBNEWSMA2
KALNEWIS	CURRENT 1794
KALNEW1SMAZ	KALREL
KALNEW1SMA3	KAL2SPK
KALMANSMA2	

Series Classifications:

- Class 1 Average Annual Dollar demand between \$0 and \$5000
- Class 2 Average Annual Dollar demand between \$5000 and \$50000
- Class 3 Average Annual Dollar demand greater than \$50000
- Class 4 Average Annual Number of Requisitions between 0-3
- Class 5 Average Annual Number of Requisitions between 3-12
- Class 6 Average Annual Number of Requisitions greater than 12

6.3 Step 2: (5 forecast algorithms, 6 error measures)

This step consisted of evaluating five forecast algorithms over a four quarter forecast horizon. The data sets and series classifications were the same as Step 1. Four of the five algorithms were the best ranked ones from the previous step and the error measures were those which appeared most consistent and/or easiest to understand. The fifth algorithm, MED4, was suggested after Step 1 was completed. The algorithm and error measures are as follows:

ERROR MEASURES: (refer to pages 27, 28, 29 for definition)

REL MAD

REL MSE

% of Actual

% of Forecast

% of Both

Bias

Algorithms:

1794

KALNEW1

KALNEW13MA3

IROTRIGG

MED4

6.4 Step 3: (Statistical vs Simulation Evaluation)

This step is the most complex of the three. Samples from five of nine stratification classes were taken and both statistical and simulation analysis were used to evaluate the remaining four algorithms. The following table gives a count of items sampled in the 3 x 3 dollar demand versus requisition stratification.

		Yearly Dollar Demand		
		0 - \$5000	\$5000 - \$50000	> \$50000
Yearly Requisitions	0 - 3	Strat Class	2	3
		1		
	N = 335	N = 100	N = 4	
	3 - 12	Analyzed	Analyzed	
		4	5	6
	N = 124	N = 230	N = 17	
> 12	Analyzed	Analyzed	Analyzed	
	7	8	9	
N = 98	N = 64	N = 115		

Forecast Algorithms:

The five remaining forecast algorithms considered for the analysis were:

KALMAN

1794

IROTRIGG

MED4

MOVD

Moving D was added to the list of those previous tested to determine if the results from Cohen's report [3] still appear valid. That is, do forecast algorithms utilizing program data perform better than those forecasting on demand only.

Statistics:

In an effort to better relate error measures to inventory performance, the following procedural changes were made to the way the statistics were collected.

(1) Forecast only after a demand; this is the only time a reorder point may be triggered and where the forecast is actually used (alternative would be periodic review which wasn't considered).

(2) Use only the item's PLT as a forecast horizon; again this is what would be used in an inventory system.

(3) Use a simple eight quarter average for the forecast if the item had been inactive for a year prior to the demand triggering the forecast; this would handle the migration of an item from an active strat class to an inactive one without unduly penalizing the algorithm which would normally work well in an active class and does poorly in the less active class.

Along with these procedural changes, additional inventory measures were considered. An overforecast error in predicting demand impacts the inventory control system differently than does an underforecast. Overforecasts result in carrying too much stock and increase the possibility of being stuck with obsolete items, whereas underforecasts increase the possibility of not satisfying a customer's orders and in the case of the Army may reduce the readiness of a weapon system. Since there is not a natural tradeoff between these two types of errors, the following separate measures were developed.

Notation

For the given i^{th} demand series $\{x_{ij}\}$ and its corresponding $\{r_{ij}\}$ requisition series (the number of requisitions at time j)

let $D_{i,t}(\ell) = \sum_{j=1}^{\ell} x_{i,t+j}$ be the demands over ℓ periods from time t

$R_{i,t}(\ell) = \sum_{j=1}^{\ell} r_{i,t+j}$ be the number of requisitions over ℓ periods from time t

$EL_{i,t}(\ell) = (\hat{D}_{i,t}(\ell) - D_{i,t}(\ell)) = \sum_{j=1}^{\ell} (\hat{x}_{i,t+j} - x_{i,t+j})$ be the errors over lead time ℓ

UP_i = unit price of the i^{th} item

Overforecast Measure

$$OF(i) = \sum_{j \in F_i} \max \left[\frac{EL_{i,j}(8)}{D_{i,j}(8)}, 0 \right] D_{i,j}(8) UP_i$$

$$= \sum_{j \in F_i} \max [EL_{i,j}(8), 0] UP_i$$

is the cost of the extra stock purchased for periods of eight quarters for the i^{th} item.

$$OF = \frac{\sum_{i=1}^N OF(i)}{\sum_{i=1}^N UP_i \sum_{j \in F_i} D_{ij}(8)}$$

is a percent of the total dollar demand spent on extra stock.

The base period of 8 is used to represent the long term effect of an overforecast.

Underforecast Measure

$$UF(i) = \sum_{j \in F_i} \max \left[-\frac{EL_{ij}(PLT)}{D_{ij}(PLT)}, 0 \right] R_{ij}(PLT)$$

is an estimate of the number of requisitions not satisfied for the i^{th} item, i.e. if demand is underforecasted, say by 20%, then it is implied that 20% of the requisitions will not be satisfied.

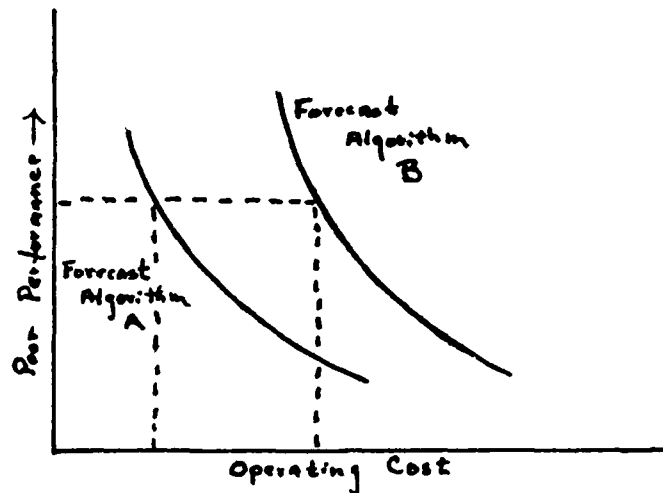
$$UF = \frac{\sum_{i=1}^N UF(i)}{\sum_{i=1}^N \sum_{j \in F_i} (R_{ij}(PLT))}$$

is an estimate of the percent of the total requisition not satisfied over all the items.

The base period is the procurement lead time of the item which is the quickest time stock could be replenished after a new order is placed. In an underforecast situation the reorder point will probably be crossed within a PLT.

Simulator

The final analysis was done with the IRO Simulator of the Army wholesale supply system, a description of which is found in Cohen [2]. Algorithms are compared in the form of cost-performance curves; the curves are traced thru several " λ " points for each forecast procedure, the lambda (λ) values reflecting an operating policy which relates to the cost of a backorder. Actual demand and flying hour time series for items in any of the various data groupings are used in particular simulation runs. All algorithms have the same starting conditions prior to accumulating performance statistics. To do this during the warmup period (2 year) all algorithms utilize an 8 quarter moving average on demand. Of course, also during warmup, the algorithms obtain their various starting values where needed.



To determine if the curves are statistically different, a fixed cost analysis is done on the difference between the performance of the current policy and the test policies for each item. Histograms of these differences are displayed and statistical tests are used for comparison. Details of the test procedure are found in Section 7.3.

There have been many changes made to the simulator since used in reference [8]. The most noteworthy are:

- (1) Excess costs (projected from assets above RO at end of simulator run) are accumulated and averaged into operating costs only on items coded terminal or obsolete or on items with trivial (nearly zero AYD) demand. However, end of simulation stock is also stratified into 1 to 15 years of supply over all items and presented as a simulator output for each forecast policy.

- (2) Previously, forecasts could be updated in between quarters (at time of buys) using moving, interpolated, quarters formed from actual quarterly data. Now forecasts are updated only on the actual quarters.

- (3) The current PCER tables utilized in the VSL module were previously adjusted for lead time by a theoretical factor of $1/SL$ from base values. Now, the PCER base values do not change, in order to reflect conservatively the empirical observation that percentage error increased with lead time L .

- (4) The constraint that safety level be no more than the expected lead time demand quantity is lifted.

- (5) The simulator now incorporates the effects of phased deliveries.

- (6) The point estimate of an item's order size in a current interpolated time interval (obtained previously from the interpolated requisition and demand history) is now smoothed (averaged) with previous order size computations.

CHAPTER VII

ANALYSIS AND RESULTS

The results from the data experiments described in the previous chapters are analyzed in the next few sections.

In each step algorithms were eliminated based on their performance when compared to the alternative methods while using several error measures. The final analysis was performed via the IRO simulator on the resulting five techniques.

7.1 Step 1 (Rankings based on 9 statistical error measures)

The tables in this section summarize the comparative performance (rankings) of the algorithms from which initial screening decisions were made. For each of the four data base groups, and within stratification classes (dollar demand and average yearly requisitions) for each group, the algorithms are ranked ("1" being best) based on values across the many error measures described in Chapter VI. Rankings are done for both 1 quarter and 4 quarter forecast error measures.

After a study of the ranking patterns in these tables, KALNEW1, CURRENT, KALNEW1SMA3, IROTRIGG 25-25* were chosen for further statistical and simulation investigation. KALNEW1SMA3 was chosen over KALNEW1SMA2 because of our conservative tendency to use the last 3 quarters (MA3) of data rather than only the last 2 quarters. Current 1794, of course, is chosen as the base for improvement. MED4 was developed later in our study.

*The "25-25" refers to the α_1 , α_2 values in hundredths.

Item Group: LDVDYN28

Rankings, 1 Qtr Error
4 Qtr Error

Stratification Class

	SD 0 - 5000	SD - 50000	SD - ∞	Reg 0 - 3	Reg - 12	Reg - ∞
# in Class	547	77	2	341	131	154
Algorithm						
KALNEWI	8 3	7 7		9 9	4 1	1 1
KALMANS	5 7	8 8		5 5	7 7	6 7
KALNEWIS	6 6	9 9		4 4	6 8	7 8
KALNEWISMA2	3 4	3 2		2 2	2 4	4 4
KALNEWISMA3	4 5	1 1		6 6	3 2	2 2
KALMANSMA2	2 2	2 4		3 3	1 3	5 3
IROTRIGG15-75	7 8	10 10		7 7	10 10	10 10
IROTRIGG25-25	1 1	4 3		1 1	5 6	3 5
REGKBNEWSMA2	9 9	6 6		8 8	8 5	8 6
CURRENT 1794	10 10	5 5		10 10	9 9	9 9

Item Group: HDVDYN28

Rankings, 1 Qtr Error
4 Qtr Error

Stratification Class						
	SD 0 - 5000	SD - 50000	SD - ∞	Req 0 - 3	Req - 12	Req - ∞
# in Class	2	13	34	3	12	34
Algorithm						
KALMAN		11 10	11 7		8 2	11 9
KALNEW1		9 9	3 1		1 3	6 1
KALMANS		6 1	5 3		2 4	4 3
KALNEWIS		5 4	9 10		10 10	8 10
KALNEWISMA2		3 3	1 6		3 7	2 5
KALNEWISMA3		7 5	4 4		5 6	3 2
KALMAN SMA2		4 2	2 5		4 8	1 4
IROTRIGG15-75		1 7	10 11		11 11	9 11
IROTRIGG25-25		2 6	6 8		6 5	5 6
REGKBNEWSMA2		8 8	7 7		9 9	7 8
CURRENT 1794		10 11	8 2		7 1	10 7
KALREL *						
KAL2SPK *						

* Ranking poor in general

Item Group: HDVNON28

Rankings, 1 Qtr Error
4 Qtr Error

Stratification Class						
	SD 0 - 5000	SD - 50000	SD - ∞	Req 0 - 3	Req - 12	Req - ∞
# in Class	11	73	128	4	27	181
Algorithm						
KALMAN	5 3	8 8	1 4		9 9	2 2
KALNEW1	2 1	7 7	3 1		8 8	1 1
KALNEWIS	7 7	6 6	9 9		6 6	9 9
KALNEWISMA2	3 5	2 2	5 6		5 3	5 5
KALNEWISMA3	1 4	1 1	4 5		1 2	4 3
KALMANSMA2	4 8	3 3	6 7		4 5	6 7
IROTRIGG15-75	10 9	9 9	10 10		7 7	10 10
IROTRIGG25-25	6 2	4 5	2 2		3 4	3 4
REGKBNEWSMA2	8 10	5 4	8 8		2 1	7 8
CURRENT 1794	9 6	10 10	7 3		10 10	8 6
KALREL *						
KALZSPK *						

* Ranking poor in general

7.2 Step 2 (Ranks from four quarter error measures)

Four tables are presented in this section, one for each of the four item groupings - LDV dynamic and non-dynamic items, HDV dynamic and non-dynamic items. Each table presents rankings of the five candidate items for two stratifications, by annual dollar demand and by annual demand frequency (requisitions). The relative performance rankings in terms of 6 error measures are tabulated; it should be noted that all but the last are relative or "percent" error measures.

The algorithm MED4 was statistically evaluated and ranked only for the measures REL MAD and REL MSE. These two measures, incidently, are the most consistent, in the sense that their rank orderings most frequently agree with a consensus rank ordering across all the measures in a stratification class.

A pattern of some note in the tables: 1794 and KALNEW1 are often ranked closely ("paired") compared to the KALNEW1SMA3 - IROTRIGG pair. The latter pair tend to perform well in less active classes (I, II, V, VI), their tracking signals reacting to fluctuations, while the 1794-KALNEW1 algorithms weight more past history and hence perform well on items with more stable D/H values, i.e. the active classes (II, III, V, VI).

There is no dominant algorithm across all strat classes and tables. In Step 3 we focus upon the statistical and simulator performances of the algorithms in a three by three stratification.

Item Group: LDVNON28

Rankings within class

4 Quarter Error Measures

Algorithm	Strat Class, upper bound	REL MAD	REL MSE	% Actual	% Forecast	% Both	Bias
CURRENT 1794	S 5000	4	4	2	2	4	4
KALNEW1		3	3	1	1	2	2
KALNEWISMA3		2	1	3	4	1	3
IROTRIGG		1	2	4	3	3	1
MED4		1-	1-				
CURRENT 1794	A 50000	2	2	3	2	3	4
KALNEW1		1	3	2	1	1	2
KALNEWISMA3		3	4	1	4	2	3
IROTRIGG		4	1	4	3	4	1
MED4		1-	1-				
CURRENT 1794	H 80						
KALNEW1							
KALNEWISMA3							
IROTRIGG							
MED4							
CURRENT 1794	Reg 3	4	4	3	4	4	2
KALNEW1		3	3	4	2	2	1
KALNEWISMA3		2	1	2	1	1	3
IROTRIGG		1	2	1	3	3	4
MED4		1-	1-				
CURRENT 1794	Reg 12	2	3	4	1	2	2
KALNEW1		1	1	3	2	1	4
KALNEWISMA3		3	2	1	4	3	1
IROTRIGG		4	4	2	3	4	3
MED4		3.5	3.5				
CURRENT 1794	Reg 8	3	3	4	4	4	3
KALNEW1		1	1	3	1	1	4
KALNEWISMA3		2	2	1	2	2	2
IROTRIGG		4	4	2	3	3	1
MED4		2.5	2.5				

Item Group: HDVDYN28

Rankings within class

4 Quarter Error Measures

Algorithm	Strat Class, upper bound	REL MAD	REL MSE	% Actual	% Forecast	% Both	BIAS
CURRENT 1794	5						
KALNEWI	5000						
KALNEWISMA3							
IROTRIGG							
MED4							
CURRENT 1794	5000	4	4	4	3	3	3
KALNEWI		3	3	3	3+	4	1
KALNEWISMA3		1	1+	1	1	1	2
IROTRIGG		1+	1	1+	1+	2	4
MED4		1-	1-				
CURRENT 1794	8	1+	2	4	1	1	4
KALNEWI		1	1	3	2	2	3
KALNEWISMA3		3	3	1	4	3	2
IROTRIGG		4	4	2	3	4	1
MED4		5	5				
CURRENT 1794	Key 3						
KALNEWI							
KALNEWISMA3							
IROTRIGG							
MED4							
CURRENT 1794	Key 12	1	1	1	1	1	2
KALNEWI		2	2	2	2	2	1
KALNEWISMA3		3	3	4	4	4	3
IROTRIGG		4	4	3	3	3	3+
MED4		5	5				
CURRENT 1794	Key 8	3+	4	4	1+	4	3
KALNEWI		1	1	3	1	2	1
KALNEWISMA3		1+	1+	1	4	1	2
IROTRIGG		3	3	2	2	3	4
MED4		2	2				

Item Group: LDVDYN28

Rankings within class

4 Quarter Error Measures

Algorithm	Strat Class, upper bound	REL MAD	REL MSE	% Actual	% Forecast	% Both	/Bias/
CURRENT 1794	I 5000	4	2 ⁺	4	2	4	4
KALNEWI		3	3	3	1	1	2
KALNEWISMA3		1 ⁺	2	1	3	1 ⁺	1
IROTRIGG		1	1	1 ⁺	3 ⁺	3	3
MED4		1 ⁻	1 ⁻				
CURRENT 1794	I 50000	3	2	4	2	4	4
KALNEWI		4	4	3	1	1	3
KALNEWISMA3		1	2 ⁺	1	2 ⁺	2	1
IROTRIGG		2	1	2	3	3	2
MED4		5	5				
CURRENT 1794	II 8						
KALNEWI							
KALNEWISMA3							
IROTRIGG							
MED4							
CURRENT 1794	III Key 3	4	3	4	3	4	4
KALNEWI		3	3 ⁺	3	1	2	3
KALNEWISMA3		2	2	2	2	1	2
IROTRIGG		1	1	1	3 ⁺	3	1
MED4		1 ⁻	1 ⁻				
CURRENT 1794	IV Key 12	4	4	4	2	4	4
KALNEWI		1	1	2	1	2	2
KALNEWISMA3		2	2	1	4	1	1
IROTRIGG		3	3	2 ⁺	3	3	3
MED4		2.5	2.5				
CURRENT 1794	V Key 8	4	4	4	4	4	1
KALNEWI		1	1	3	1	1	2
KALNEWISMA3		2	2	1	3	2	3
IROTRIGG		3	3	2	2	3	4
MED4		2.5	2.5				

Item Group: HDVNON28

Rankings within class

4 Quarter Error Measures

Algorithm	Stat Class, Upper bound	REL MAD	REL MSE	% Actual	% Forecast	% Both	BIAS
CURRENT 1794	1	2	4	1 ⁺	4	4	4
KALNEW1	ADD 5000	1	1	1	2	1 ⁺	3
KALNEWISMA3		4	2	4	3	3	2
IROTRIGG		1 ⁺	3	3	1	1	1
MED4		5	2.5				
CURRENT 1794	4	4	4	4	4	4	4
KALNEW1	ADD 50000	3	3	3	1	1 ⁺	3
KALNEWISMA3		1	1	1	3	3	2
IROTRIGG		1 ⁺	2	2	2	1	1
MED4		1 ⁻	1 ⁻				
CURRENT 1794	1	2 ⁺	4	4	4	1	4
KALNEW1	8	1	2	3	1	2	3
KALNEWISMA3		2 ⁺	4	2	3	3 ⁺	1
IROTRIGG		2	1	1	2	3	2
MED4		1.5	1 ⁻				
CURRENT 1794	1						
KALNEW1	Key 3						
KALNEWISMA3							
IROTRIGG							
MED4							
CURRENT 1794	4	4	4	4	1	4	4
KALNEW1	Key 12	3	3	3	3	3	3
KALNEWISMA3		1	1	2	4	1 ⁺	2
IROTRIGG		2	2	1	2	1	1
MED4		1 ⁻	1 ⁻				
CURRENT 1794	2	3	4	4	4	3	4
KALNEW1	Key 8	1	1	3	1	1	3
KALNEWISMA3		2 ⁺	2	1	3	2 ⁺	2
IROTRIGG		3	2 ⁺	2	2	2	1
MED4		5	5				

7.3 Step 3 (Simulation and final statistical results)

In this section the performance results via simulated cost effectiveness and statistical error measures are analyzed. The final four algorithm candidates, MED4, 1794, KAL1 (formerly coded KALNEW1), Trigg (IROTRIGG), along with MOVD (moving average on demands) were used to compute forecasts in various simulation runs over five classes (active items) of the 3 x 3 data stratification described on page 29. The simulated cost - performance curves are captured in pages 44 - 50 where performance is measured in terms of the average of the time weighted backorders as opposed to average days wait which was reported in the previous studies. (This transformation of performance measures has no impact on the simulation results but does make the comparative analysis easier, ref [6]).

The table on page 51 contains the statistics from the various error measures for each of the five data classes and for both a one quarter and PLT forecast horizon. The second part of the table compares the within class ranks of the statistical measures and the simulation results which were ranked at a fixed cost as described in the next section on Final Analysis.

ST4

Start Class
(ST4)

1794
KAL
SALG
MEL
MA

Ave
Back 2.55m

44

5

7

1

300

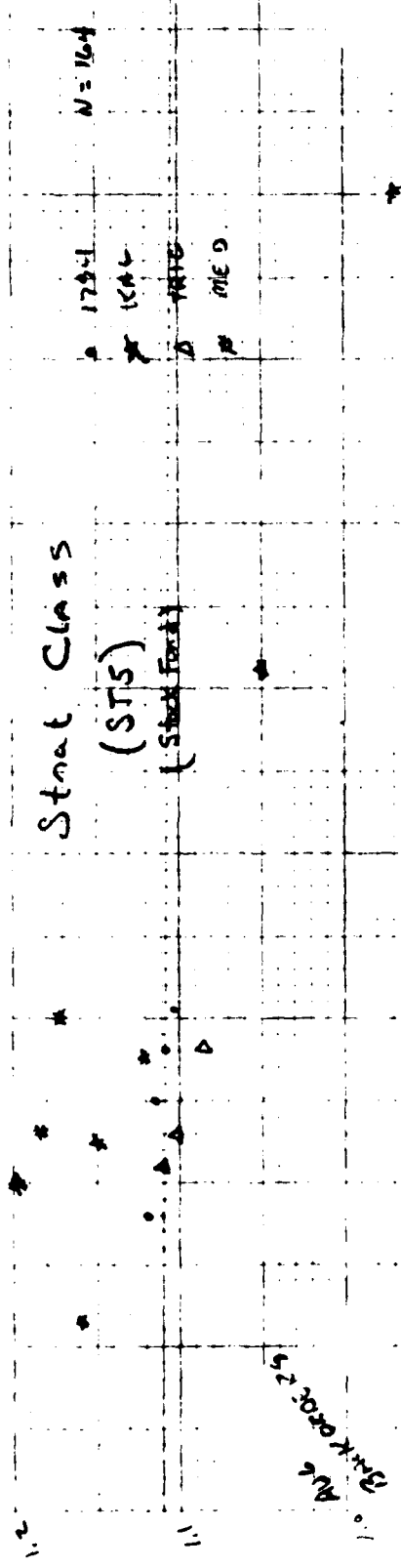
400

500

600

700

Ave Cost Per Item



AUG 1955

Stat Class (STS)

N = 207

AVG BACKCUT DEPTH

Strat Class
(ST5)

1794

VAL

TDIC

MEB

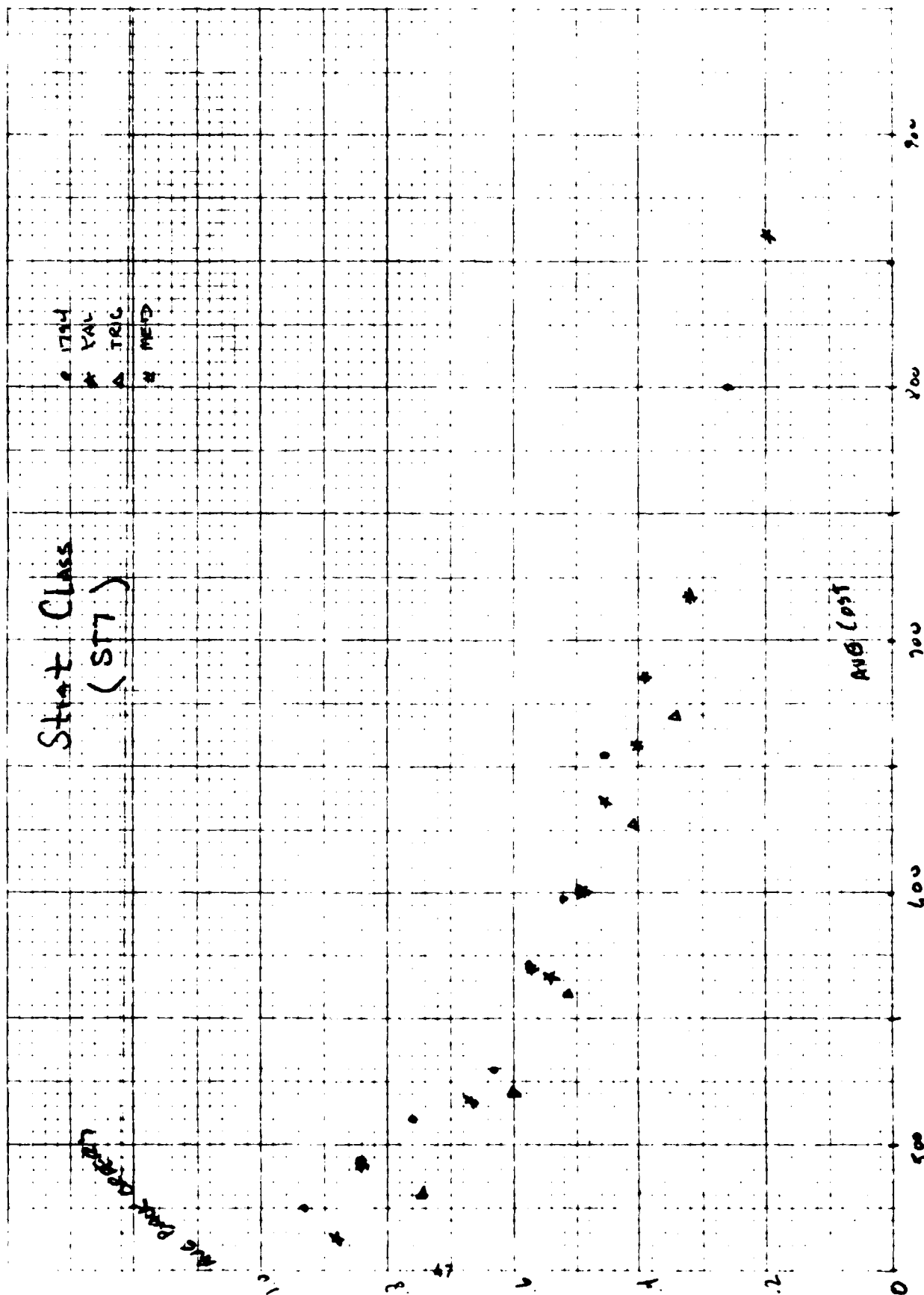
AVG COST

3100

3100

3100

3100



Strat Class (ST8)

MEMS

MEM
KAL
A TRG
MED
MA

Aug 1964

10

2

48

6

8

2

40

50

60

70

Aug 1964

0.1

0.2

0.3

0.4

0.5

0.6

0.7

0.8

0.9

1.0

1.1

1.2

1.3

1.4

1.5

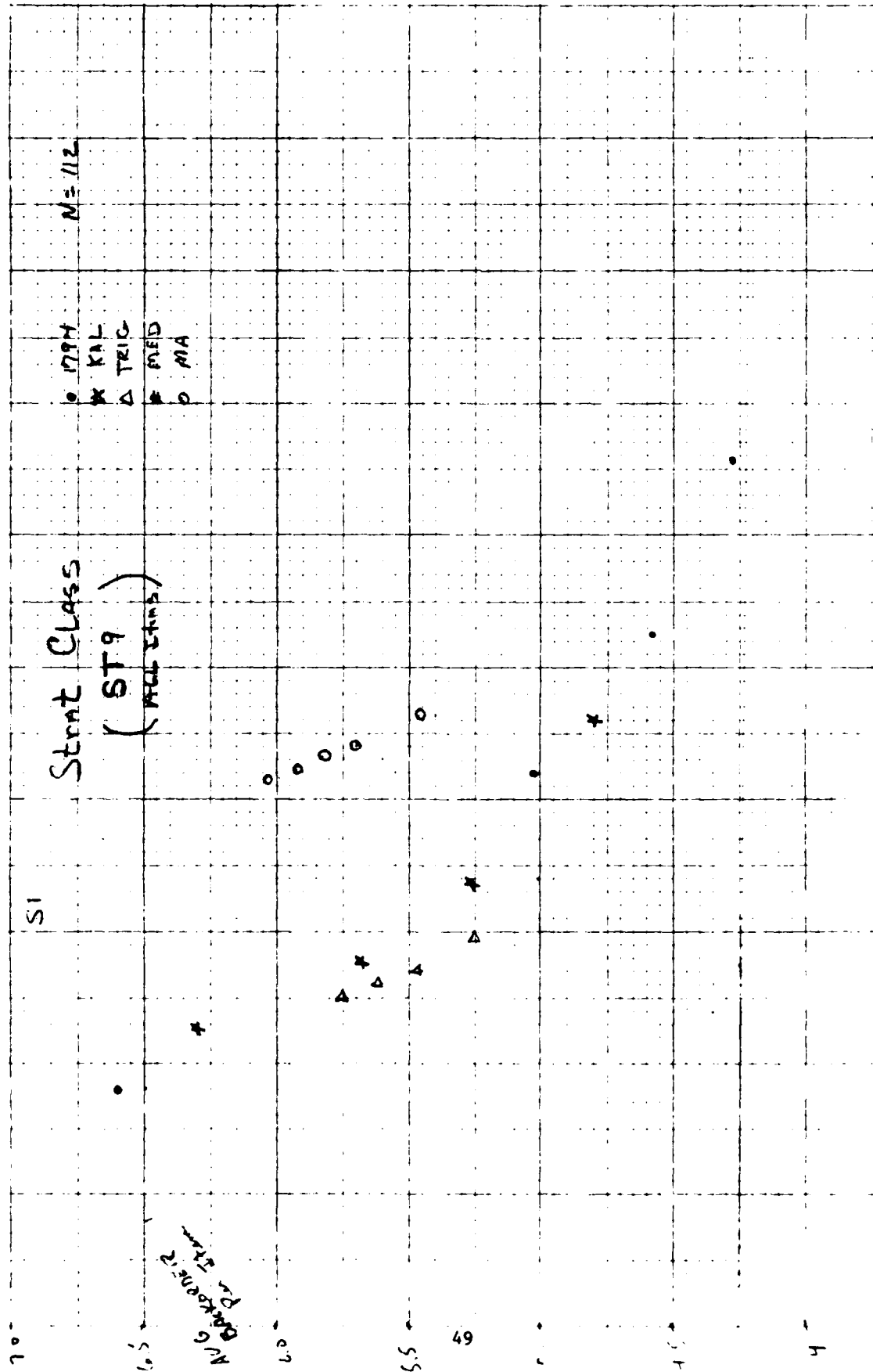
1.6

1.7

1.8

1.9

2.0



AUS C.E. Per. Thru

N/6
Baker
per. Thru

Strat Class
(ST9
(Stock Fund)

N=24

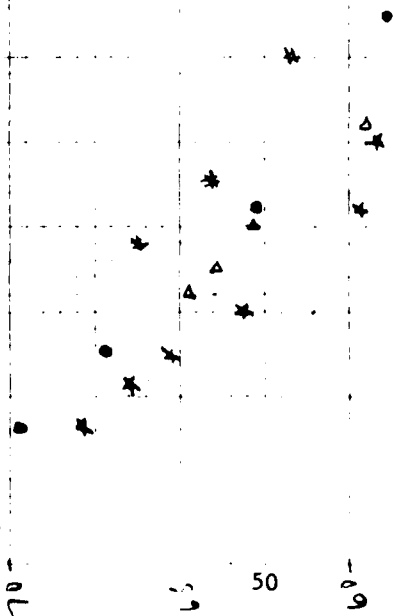
1794

KAL

TRIO

WED

AUC
Brickwork



AUC Cost

$\times 10^2$

118

Statistical Error Measures

STATION	ERROR		RELATIVE ERROR		MAD		REL MAD		REFL		MSE
	1st gL	Pt	1st gL	Pt	1st gL	Pt	1st gL	Pt	1st gL	Pt	
1741	-0.117	.717	-.047	.028	4.625	9.910	1.032	.823	2.301	1.384	
S KAL	.021	1.178		.077	4.504	9.414	1.00	.790	2.086	1.247	
T TAC	.342	1.918	.096	.171	4.722	9.984	1.050	.858	2.202	1.424	
MED	-.555	-.0369	-.011	-.031	4.514	9.471	1.024	.804	2.281	1.813	
MAD	.453	2.435	.045	.137	4.98	10.95	1.08	.87	2.54	1.642	
1741	.615	5.323	.053	.137	2.444	20.494	1.060	.810	2.624	1.278	
S KAL	1.249	8.765	.093	.155	7.882	22.950	1.056	.792	2.607	1.150	
T TAC	.857	6.54	.187	.242	2.486	21.754	1.102	.865	2.725	1.344	
MED	-1.128	-.019	-.032	.025	7.246	20.022	1.022	.778	2.500	1.104	
MAD	.743	6.204	.103	.200	7.75	21.031	1.084	.845	2.473	1.317	
1741	-2.880	-10.314	-.072	-.069	15.593	38.124	.731	.622	1.025	.630	
S KAL	-2.667	-10.335	-.057	-.056	14.483	34.651	.692	.564	.912	.531	
T TAC	-1.778	-7.477	-.011	-.009	14.291	34.223	.690	.583	.926	.601	
MED	-3.818	-13.045	-.114	-.116	14.639	34.971	.703	.599	1.017	.629	
MAD	-.824	-4.090	-.008	.006	15.460	37.243	.735	.616	1.038	.630	
1741	-6.319	-21.329	-.020	-.032	44.61	107.121	.443	.377	3.73	2.02	
S KAL	-4.657	-17.778	-.010	-.021	41.57	100.45	.409	.344	3.12	2.13	
T TAC	1.594	3.302	.013	.005	42.61	103.96	.409	.349	3.10	2.14	
MED	-1.626	-6.227	-.037	-.051	42.649	104.571	.410	.348	3.09	2.10	
MAD	-.201	.033	.049	.057	43.98	110.12	.474	.424	4.71	4.00	
1741	1.108	-4.619	-.009	-.055	23.49	58.25	.481	.370	4.07	3.74	
S KAL	.558	-5.098	-.019	-.053	22.019	54.05	.473	.371	4.10	3.41	
T TAC	2.014	2.361	.027	.004	21.57	53.53	.484	.389	4.53	3.79	
MED	-.778	-4.94	-.028	-.059	22.198	57.02	.490	.407	4.82	3.10	
MAD	2.521	1.902	.623	-.007	23.478	61.698	.489	.402	4.76	3.10	

Since there are obvious discrepancies between the simulation ranks and statistical ranks a decision had to be made as to what results should be used for the final analysis. Much experimentation was done with the simulator but inconsistencies with the statistics continued to plague the results. A discussion with Prof Mueller of the University of Ghent (Belgium) revealed that work being done on an unpublished PhD thesis using Monte Carlo methods indicate that minimum mean squared error forecast techniques are not necessarily optimal methods when applied to inventory management models. This fact along with our belief that the IRO simulator best represents the Army management system compelled us to use the simulated results for our final decisions.

7.4 Final Results (Evaluation of Simulation Results)

In an effort to determine if there is a statistical difference between the cost-performance curves generated by the simulator the following fixed cost analysis was employed, details of which are in [6].

For each class of items, a fixed (current) cost is computed by running the simulator using the standard forecast algorithm (1794) and the λ value the Commodity Command presently uses. The resulting cost represents the inventory cost presently incurred to manage the class of items. Also during this run, the performance of each item is arrayed in a data file for future analysis.

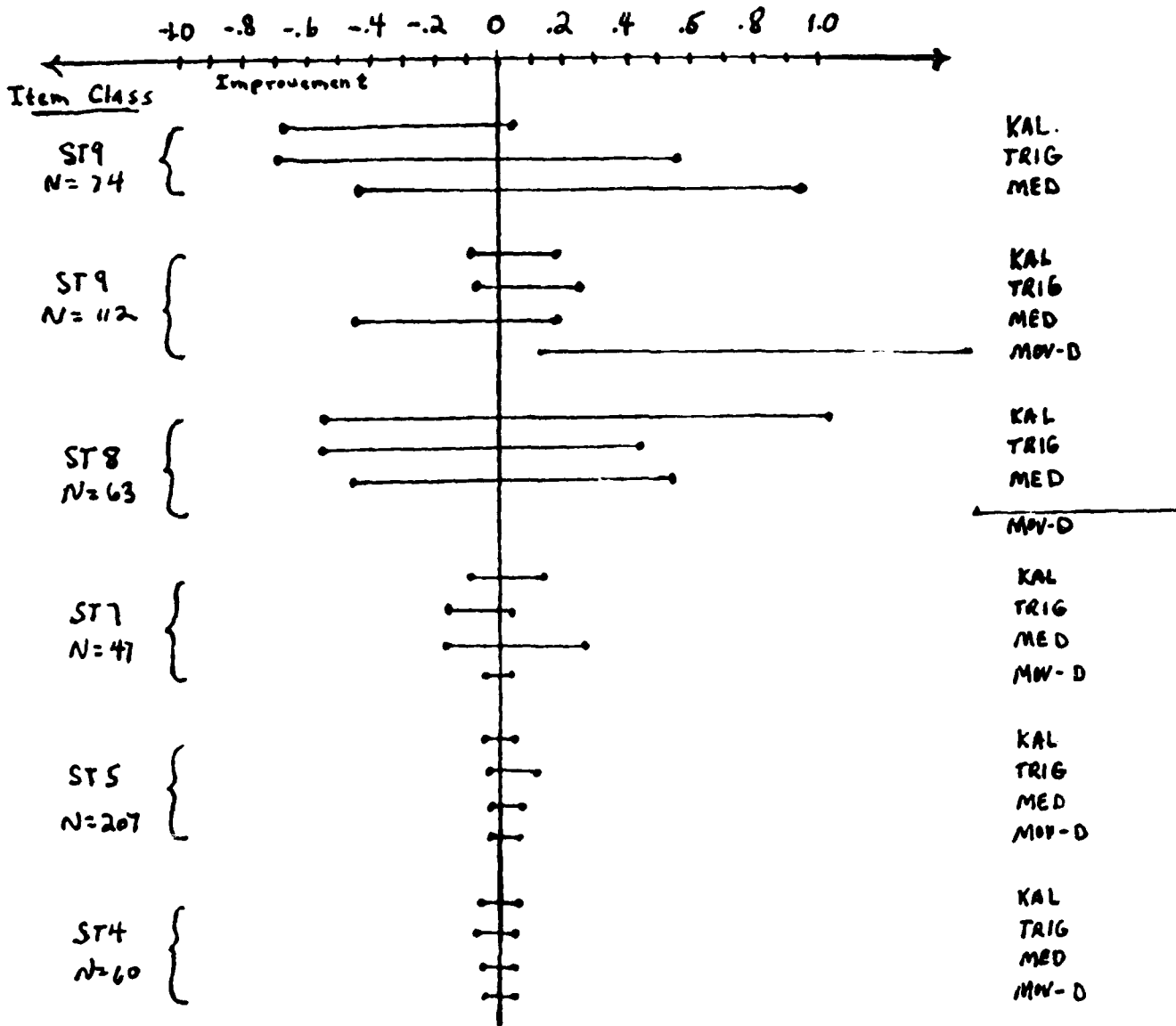
For the alternative policies (forecast algorithms) several λ s are used to generate the cost-performance curve for each policy. Using the shape of this curve, a spline technique is used to determine the performance of each item using the alternative policy at the fixed cost. Distributions (histograms) of item performance for each policy at the current cost are found in Appendix A.

To better measure the difference between the alternative methods and the standard, the difference between the individual item performance for the alternative and the standard were computed. (It is shown in the basic statistical literature that these differences will be less variable due to the elimination of extraneous effects and will measure only the difference in methods.) The distributions of these differences are found in Appendix B.

Statistically testing for zero means for each of these difference distributions is equivalent to testing for a difference between the standard (1794) curve and the alternatives as plotted on pages 44 to 50. The results of these tests are captured on the next page where 95% confidence intervals are

95% Confidence Intervals For The Mean of Individual Differences Between Time Weighted Backorders of (test policy - std policy)

1794 = Std
Policy



displayed for each distribution of differences for each class of items tested. (Note if these intervals do not contain zero, then the null hypothesis of zero mean is rejected at an α level of 5%) The rankings found on page 52 taken from this table by looking at the mid-point of each interval and ranking them from left to right.

Findings

(1) Using the IRO simulator as described in this report, there is no statistical difference in the simulated performance for each of the five stratification classes for the following comparison.

KALMAN vs 1794

IROTRIGG vs 1794

MED4 vs 1794

(2) MOVD performed worse than 1794 and the other alternative algorithms for the more active items - Class 8 and 9.

Note: MOVD was the only algorithm not using program data.

(3) There was no difference between MOVD and 1794 for less active classes 7, 6, 5, 4.

CHAPTER VIII

CONCLUSION

8.1 Findings

In an earlier report by Orr [8], the candidate of final choice for forecasting was a Kalman filter algorithm. This choice was based upon its dominance in statistical forecast accuracy and its savings over the current method as projected from cost-performance curves produced from simulator runs. There was little difference amongst the algorithms' curves for LDV items, and so most of the savings was driven by HDV items.

From the current viewpoint it appears that the savings might not have been statistically significant. At the least, the algorithm's simulated performance has not been robust against the changes that have occurred (see below) since that report (although its relative forecast accuracy has held up). Presently we find "Kalman" working well on some of our groupings of items and the current (1794) method (among other algorithms) working well on other items. The improved statistical tools [6] we use now show, in any case, that no algorithm's cost-performance curve is significantly better than another's - for active items. In several cases a few items can influence the performance rankings.

It is not possible to isolate the impact of individual changes made since the earlier report when comparing the differing results. However, such changes were:

a. The data base was extended from 7 to 11 years and more items were captured. The last 7 years of the 11 were used in the current analysis, so much of the Vietnam era from the earlier time series was not influential.

b. It is possible that a different small group of items might now be driving the HDV savings.

c. The statistical analysis program was overhauled to be more flexible and capture various error measures.

d. Changes in the simulator were made. Several of these could narrow the potential difference in performance amongst the algorithms, e.g., excess cost savings were accumulated only for some items, a standard moving average forecast superseded all algorithms in periods of very little activity, forecast updates were made only after quarters with demand.

e. Some changes to the algorithms themselves were made. Theoretical adjustments were made to the "Kalman" to produce several versions; also, the "k" parameters were updated to reflect the later 7 years of history that were used.

8.2 Postscript

This table consolidates statistical results of one of the more meaningful error measures, MAD/AYD, for the four algorithms - by data group and by requisition class. This relative error when multiplied by an algorithm's current forecast of average yearly demand, AYD, yields an estimate of the mean absolute deviation (error) in a year's demand. The theory and formulas in [9] can be used to convert this estimate to a variance of lead time demands, the latter a necessary variable in computing safety levels in the VSL EOQ module of CCSS.

It is apparent that the choice of a final algorithm dictates the percent errors to be used in VSL EOQ. The present PCER tables in that module should be superseded by an expanded, refined version of these tables by requisition class (and perhaps by dollar demand). The MAD/AYD for the inactive class (0-3 requisitions) would be based upon current work on algorithms for inactive items, and not upon those values in this current table, which for some entries are suspect.

8.3 Recommendations

For active items (greater than 3 requisitions a year) the current (1794) forecast procedure should not be replaced.

The percent error (PCER) tables in the VSL EOQ module should be overhauled using the statistical byproducts of this current research and of the inactive item research.

A plan for consolidating the current algorithm (for active items), the pending algorithm (for inactive items, including NSO & insurance) and the consequent lead time demand variance procedure should be instituted.

Relative Error, 1 Year Forecasts
MAD/AYD by Requisition Class
(0-3) - (3-12) - (>12)

Data Base	# Items	1794	KALI	TRIG	MED4
LDVNON28	9467	1.637	1.396	1.174	.781
	2301	.797	.779	.837	.805
	1242	.512	.476	.514	.510
LDVDYN28	341	1.525	1.337	1.194	.750
	131	.802	.718	.772	.745
	154	.530	.465	.502	.495
HDVNON28	4	1.042	1.331	1.019	1.096
	27	1.043	.839	.779	.687
	181	.369	.358	.372	.381
HDVDYN28	3	10.173	6.273	1.115	.568
	12	.482	.522	.576	.643
	34	.388	.357	.380	.376

APPENDIX A

HISTOGRAMS OF INDIVIDUAL ITEM PERFORMANCE AS DETERMINED
BY SIMULATOR

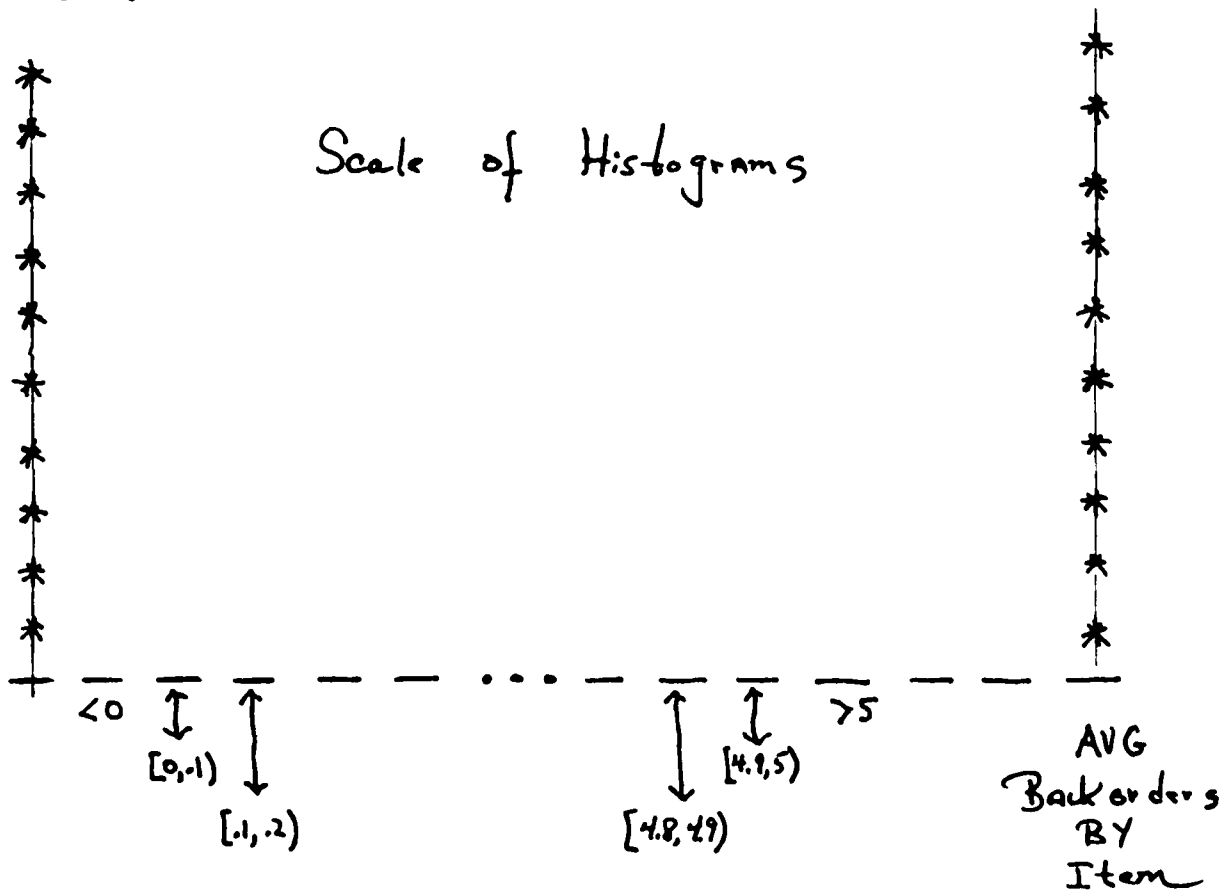
FORECAST METHOD

1794
KAL
IROTRIGG
MED4

DATA STRATS

ST9 (stock fund)
ST8
ST7
ST5

Frequency



[illegible]

CLASS	5	10	15	20	25	30	35	40	45	50
-------	---	----	----	----	----	----	----	----	----	----

FREQUENCY-----										
30	*									I *
29	*									I *
28	*									I *
27	*									I *
26	*									I *
25	*									I *
24	*									I *
23	*									I *
22	*									I *
21	*									I *
20	*									I *
19	*									I *
18	*									I *
17	*									I *
16	*									I *
15	*									I *
14	*									I *
13	*									I *
12	*	I								I *
11	*	I								I *
10	*	I								I *
9	*	I								I *
8	*	I								I *
7	*	I								I *
6	*	I								I *
5	*	I								I *
4	*	I								I *
3	*	I	I	I						I *
2	*	IIII	I							I *
1	*	IIIIIIII	I	I	I	I	III	I	III	I *

CLASS		5	10	15	20	25	30	35	40	45 50

ST9 - Shuk
Fora
KAL

MEAN= 5.966345876083

VARIANCE= 92.61153326723

```

25 * * * * * I *
24 * * * * * I *
23 * * * * * I *
22 * * * * * I *
21 * * * * * I *
20 * * * * * I *
19 * * * * * I *
18 * * * * * I *
17 * * * * * I *
16 * * * * * I *
15 * * * * * I *
14 * * * * * I *
13 * * * * * I *
12 * * * * * I *
11 * * I * * I *
10 * * I * * I *
9 * * I * * I *
8 * * I * * I *
7 * * I * * I *
6 * * I * * I *
5 * II * * * * I *
4 * II * * * * I *
3 * II * * * * I *
2 * II * I I * I * I * I * I * I *
1 * IIII VIIII I * III IIII III IIII II I III

```

CLASS	0	5	10	15	20	25	30	35	40	45	50
-------	---	---	----	----	----	----	----	----	----	----	----

519 Stock
FUN

6²: 13.6

STRAT CLASS 8

FREQUENCY-----										
18	*	I								*
17	*	I								*
16	*	I								*
15	*	I								*
14	*	I								*
13	*	I								*
12	*	I								*
11	*	I								*
10	*	I								*
9	*	I								*
8	*	I								*
7	*	I								*
6	*	I								*
5	*	I								*
4	*	I	I	I						*
3	*	I	I	I						*
2	*	I	I	I	II		I	I		*
1	*	I	IIIIIIII	II	II	II	I	II	I	*
CLASS	5	10	15	20	25	30	35	40	45	50
MEAN-	4.927899047419									

VARIANCE- 104.4875355832

FREQUENCY-----										
16	*									*
15	*									*
14	*									*
13	*									*
12	*									*
11	*									*
10	*	I								*
9	*	I								*
8	*	I								*
7	*	I								*
6	*	I								*
5	*	I								*
4	*	I								*
3	*	I	II		II		II			*
2	*	I	IIIIII		III		II			*
1	*	I	IIIIIIII	IIII	I	III	I		I	*
CLASS	5	10	15	20	25	30	35	40	45	50
MEAN-	5.154220381447									

VARIANCE- 81.70849224575

STRAT CLASS 2

N=62

FREQUENCY-----										
14	*									I *
13	*									I *
12	*									I *
11	*									I *
10	*	I								I *
9	*	I								I *
8	*	I								I *
7	*	I								I *
6	*	II	I							I *
5	*	II	I							I *
4	*	II	I							I *
3	*	II	I	I		I				I *
2	*	IIIIII				I				I *
1	*	IIIIII	III	II	III	I	I	I	I	I *

CLASS	5	10	15	20	25	30	35	40	45	50
MEAN=	4.85372414978									
VARIANCE=	84.96466955536									

TRIG

FREQUENCY-----										
14	*									I *
13	*									I *
12	*									I *
11	*									I *
10	*	I								I *
9	*	I								I *
8	*	I								I *
7	*	I								I *
6	*	I								I *
5	*	I								I *
4	*	I	I	I						I *
3	*	I	IIII	I						I *
2	*	I	IIIIII	I				I	I	I *
1	*	IIIIIIII	III	III		I	I	I	I	I *

CLASS	5	10	15	20	25	30	35	40	45	50
MEAN=	4.951909449766									
VARIANCE=	92.52434274389									

MED

FREQUENCY

25 * I
24 * I
23 * I
22 * I
21 * I
20 * I
19 * I
18 * I
17 * I
16 * I
15 * I
14 * I
13 * I
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ST7

1774

MEAN= .63373

VARIANCE= 1.115241971574

CLASS 5 10 15 20 25 30 35 40 45 50

FREQUENCY

25 * I
24 * I
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3 * II I I
2 * II I I I
1 * IIII IIII I I

ST7

KAL

CLASS 5 10 15 20 25 30 35 40 45 50

MEAN= .6705007185431

VARIANCE= 1.43568143392

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24 * I
23 * I
22 * I
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3 * IIII
2 * IIII I
1 * IIIII IIII I II I II I I I

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MEAN - 575723795.1587

FREQUENCY

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22 * I
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3 * I I
2 * I I I I II
1 * IIIII IIIII II II I I

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CLACS	5	10	15	20	25	30	35	40	45	50
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STRAT CLASS 5

FREQUENCY

41 * I
40 * I
39 * I
38 * I
37 * I
36 * I
35 * I
34 * I
33 * I
32 * I
31 * I
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29 * I
28 * I
27 * I
26 * I
25 * I
24 * I
23 * I
22 * I
21 * I
20 * I
19 * I
18 * I
17 * I I I
16 * I I I
15 * I I I
14 * III I
13 * III I
12 * III I II
11 * III I II
10 * IIII II
9 * IIII II
8 * IIII II
7 * IIIII
6 * IIIII I I
5 * IIIII I III
4 * IIIII I
3 * IIIII I I
2 * IIIII I II I
1 * IIIII III II I I I

ST5

1774

CLASS 5 10 15 20 25 30 35 40 45 50

MEAN= .8839957004831

VARIANCE= 1.15650812053

STRAT CLASS 5

FREQUENCY-----										
43	*	I							STS	*
42	*	I							KAL	*
41	*	I								*
40	*	I								*
39	*	I								*
38	*	I								*
37	*	I								*
36	*	I								*
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16	*	I								*
15	*	I	I							*
14	*	I	I	I						*
13	*	I	I	I	I					*
12	*	I	I	I	I	I				*
11	*	I	I	I	I	I	I			*
10	*	I	I	I	I	I	I	I		*
9	*	I	I	I	I	I	I	I	I	*
8	*	I	I	I	I	I	I	I	I	*
7	*	I	I	I	I	I	I	I	I	*
6	*	I	I	I	I	I	I	I	I	*
5	*	I	I	I	I	I	I	I	I	*
4	*	I	I	I	I	I	I	I	I	*
3	*	I	I	I	I	I	I	I	I	*
2	*	I	I	I	I	I	I	I	I	*
1	*	I	I	I	I	I	I	I	I	*

CLASS	5	10	15	20	25	30	35	40	45	50
MEAN=	.8997160002407									
VARIANCE=	1.241705621852									

STRAT CLASS 5

FREQUENCY-----										
44	*	I								
43	*	I								
42	*	I								
41	*	I								
40	*	I								
39	*	I								
38	*	I								
37	*	I								
36	*	I								
35	*	I								
34	*	I								
33	*	I								
32	*	I								
31	*	I								
30	*	I								
29	*	I								
28	*	I								
27	*	I								
26	*	I								
25	*	I								
24	*	I								
23	*	I								
22	*	I								
21	*	I								
20	*	I								
19	*	I								
18	*	I								
17	*	I								
16	*	I								
15	*	I	I							
14	*	I	II	I						
13	*	I	II	I						
12	*	I	III	I						
11	*	IIII	I							
10	*	IIII	II							
9	*	IIII	III							
8	*	IIIIIIIIII								
7	*	IIIIIIIIII								
6	*	IIIIIIIIII	I	I						
5	*	IIIIIIIIII	IIII							
4	*	IIIIIIIIII	IIIIII							
3	*	IIIIIIIIII	IIIIII							
2	*	IIIIIIIIII	IIII	I	III	I	I			I
1	*	IIIIIIIIII	IIII	I	II	II	I	I	I	I

CLASS	5	10	15	20	25	30	35	40	45	50
MEAN=	.8694979853693									
VARIANCE=	1.252062547115									

STS
TRIG

STRAT CLASS 5

[illegible]

```

CLASS      5      10      15      20      25      30      35      40      45      50
MEAN=      .9331464133597
VARIANCE=   1.327564530822

```

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APPENDIX B

HISTOGRAM OF DIFFERENCES BETWEEN INDIVIDUAL PERFORMANCE OF ALTERNATIVE ALGORITHMS AND STANDARD (1794)

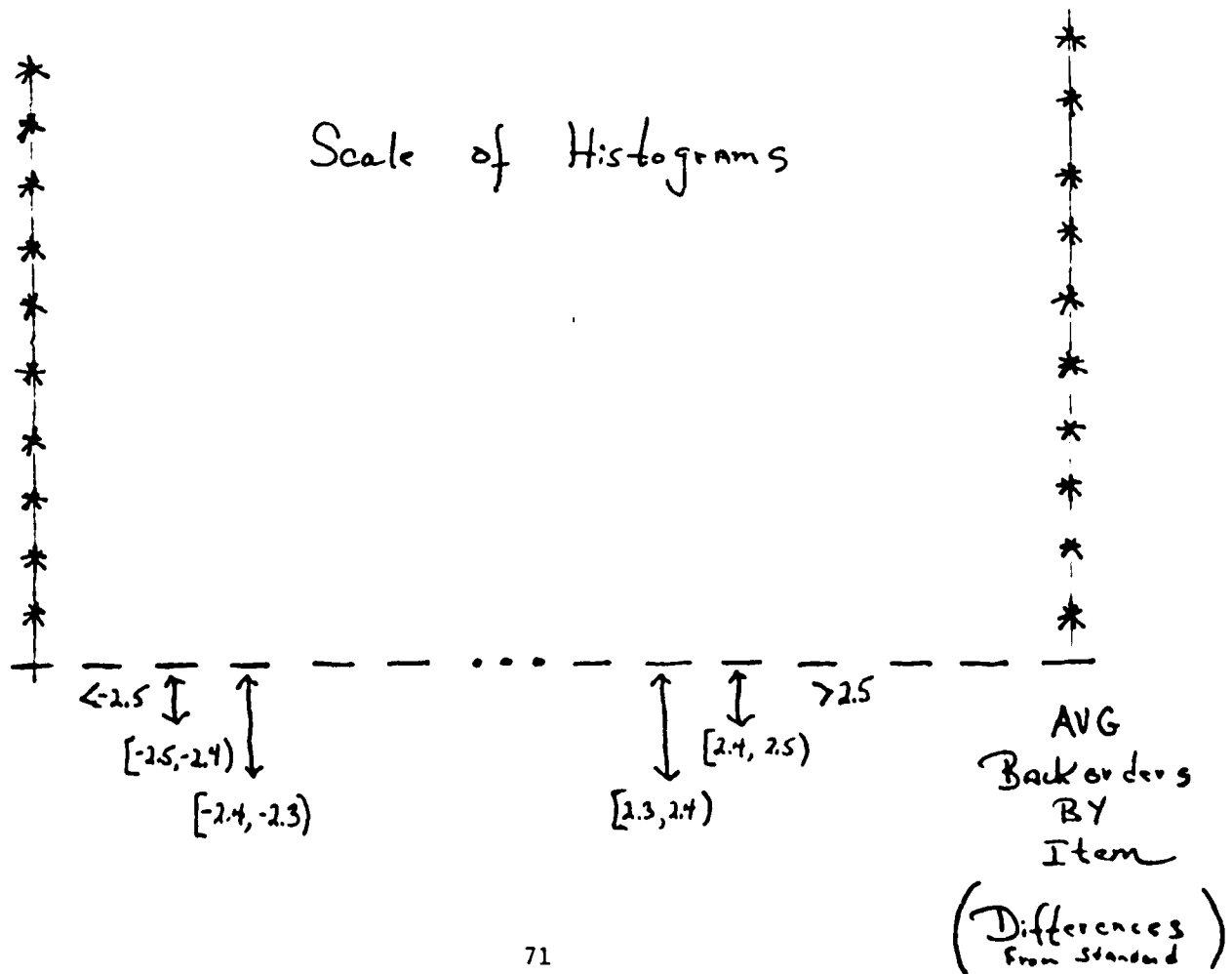
ALTERNATIVE METHOD

KAL
IROTRIGG
MED4
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DATA STRATS

ST9
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ST7
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Frequency



$N = 14$

[illegible]

CLASS	5	10	15	20	25	30	35	40	45	50
-------	---	----	----	----	----	----	----	----	----	----

MEAN = - .7085357455797 2% Conf. Int. interval for μ
VARIANCE = 2.07247292340 (-.641, .025)

VARIANCE = 2.072632302748

[illegible]

CLASS	5	10	15	20	25	30	35	40	45	50
-------	---	----	----	----	----	----	----	----	----	----

MEAN = -.04810898829984 Confidence Interval for μ
 (-.665, .569)
 VARIANCE = 7.084044187073

VARIANCE = 7.094064197033

1 FREQUENCY

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2 #1

1 #1

STG
MED4

MEDIAN

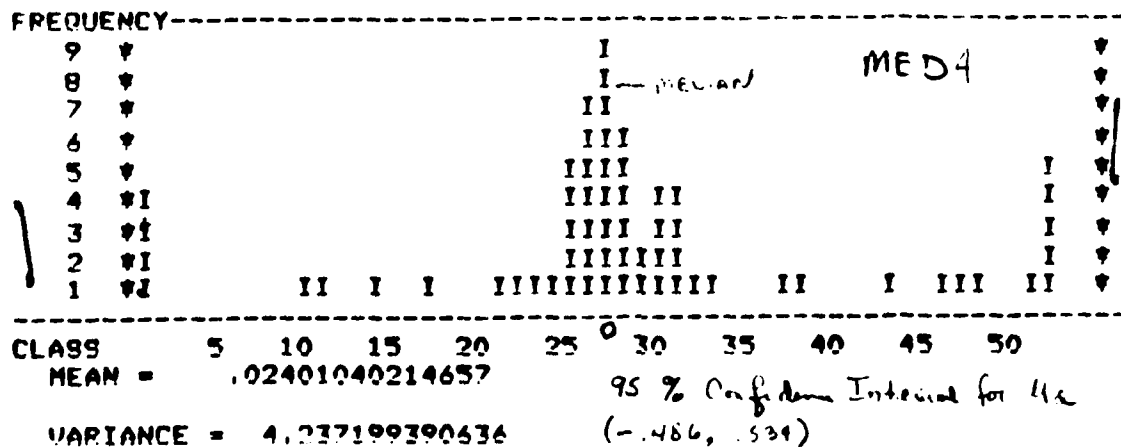
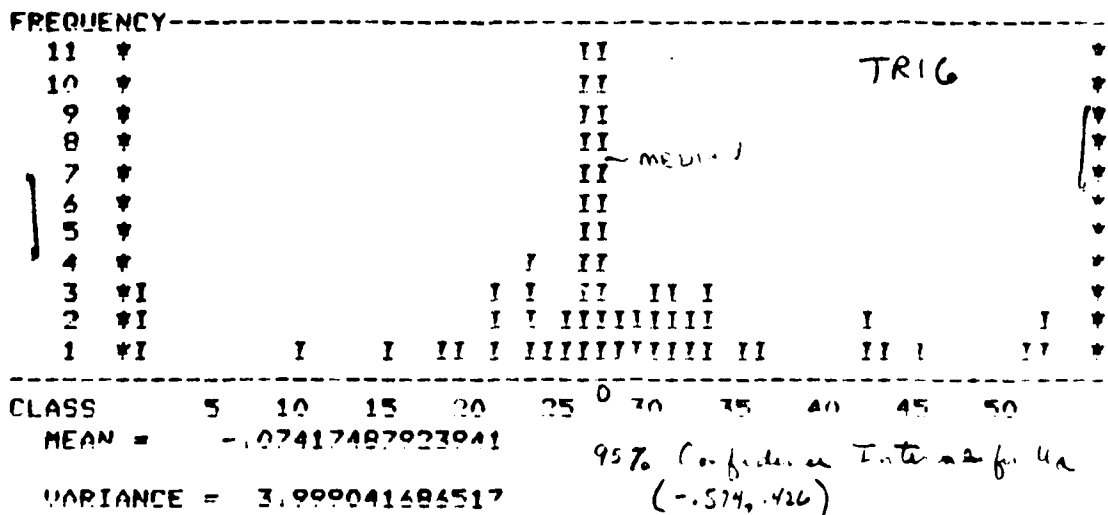
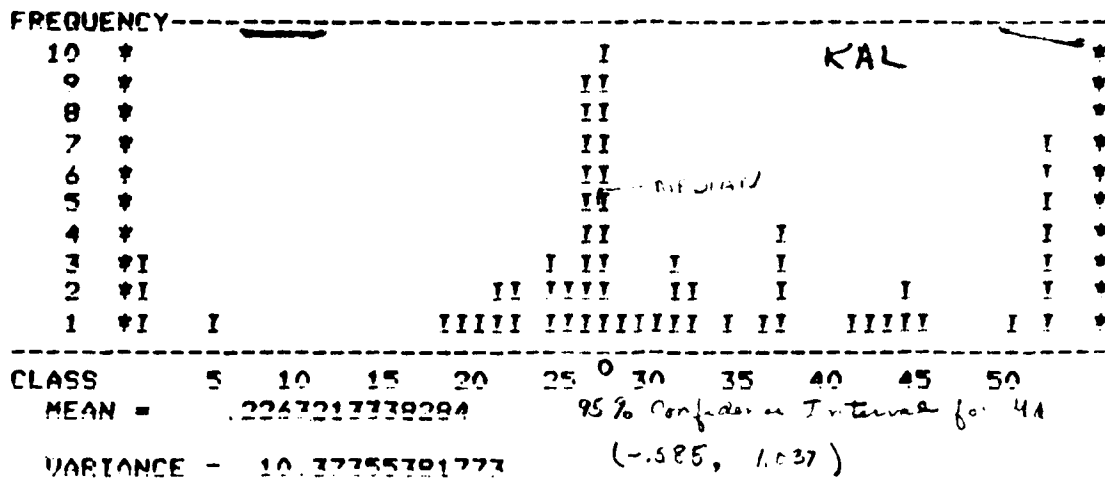
0 5 10 15 20 25 30 35 40 45 50

95% Confidence Interval for μ_d
(-428, 928)

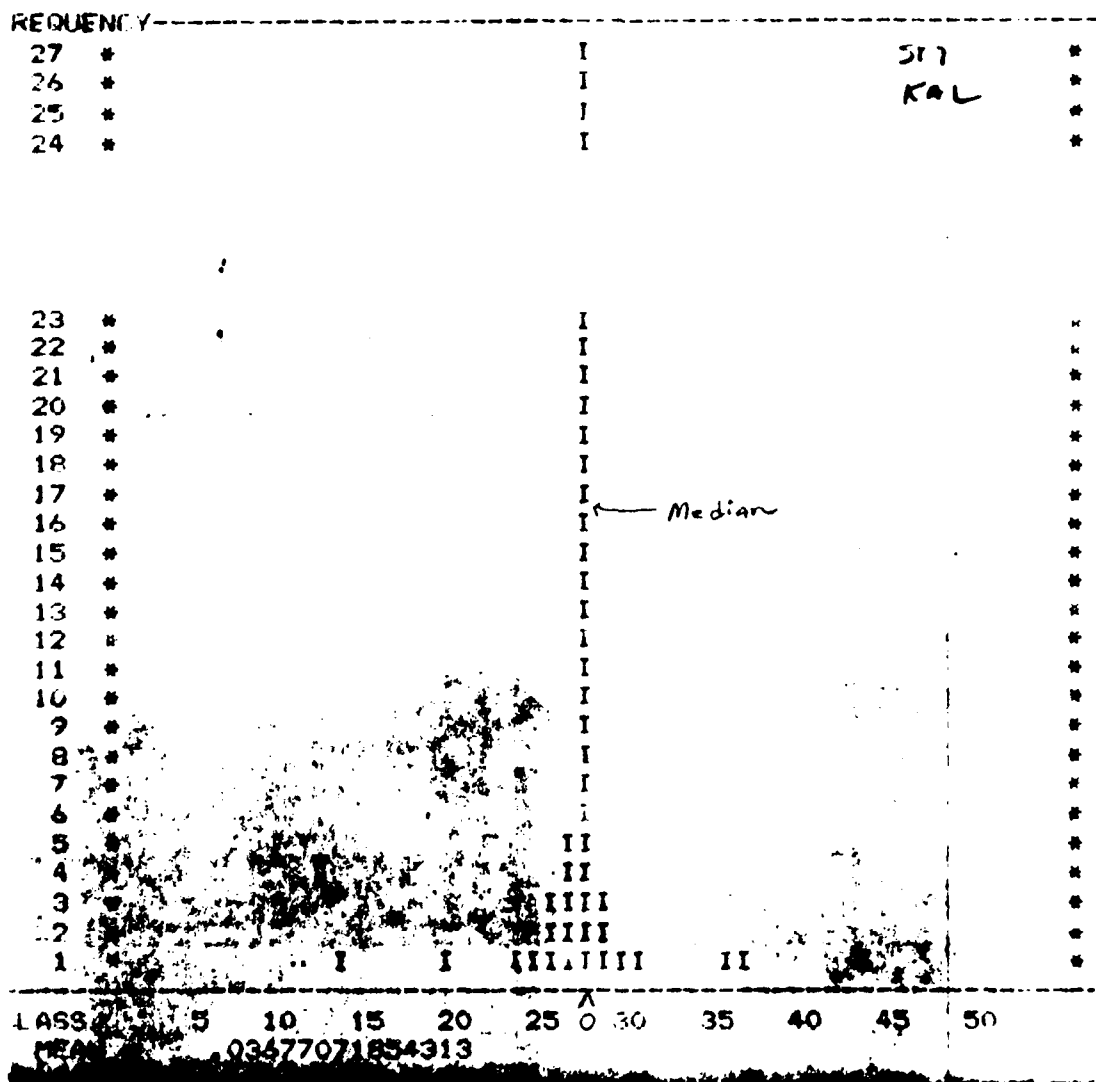
SECRET - 100-201070

STRAT CLASS 8

N:63



N=47



$G^2 = .1603$

95% confidence interval for μ_d

$(-.080, .153)$

Z FROM DEGREE 19

$\bar{X} = .0617$

$G^2 = .262$

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MEDIAN

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CLASS      5    10    15    20    25    30    35    40    45    50
MEAN =     -05550012046613
VARIANCE =  .05571826937446

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FREQUENCY-

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ST 7
MED 4

MEYIARD

I I I I I I I I I I I I .

CLASS 5 10 15 20 25 30 35 40 45 50
 MEAN = .07457221271823
 VARIANCE = .5582554976329

95% Confidence Interval for μ_d
 (-.1443, .2933)

NOT RECORDED TO BAC

Frequency

20	*	I
19	*	I
18	*	I
17	*	I
16	*	I
15	*	I
14	*	I
13	*	I
12	*	I
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8	*	I
7	*	I
6	*	I
5	*	I
4	*	I
3	*	I
2	*	I
1	*	I
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-1	*	I
-2	*	I
-3	*	I
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-5	*	I
-6	*	I
-7	*	I
-8	*	I
-9	*	I
-10	*	I
-11	*	I
-12	*	I
-13	*	I
-14	*	I
-15	*	I
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-17	*	I
-18	*	I
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-97	*	I
-98	*	I
-99	*	I
-100	*	I

ST 1
MOVED
X-450

CLASS 5 10 15 20 25 30 35 40 45 50
MEAN - .02508341344056
VARIANCE - .04456217941227

Frequency

1	*	I
2	*	I
3	*	I
4	*	I
5	*	I
6	*	I
7	*	I
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ST 2
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CLASS 5 10 15 20 25 30 35 40 45 50
MEAN - .02508341344056
VARIANCE - .04456217941227

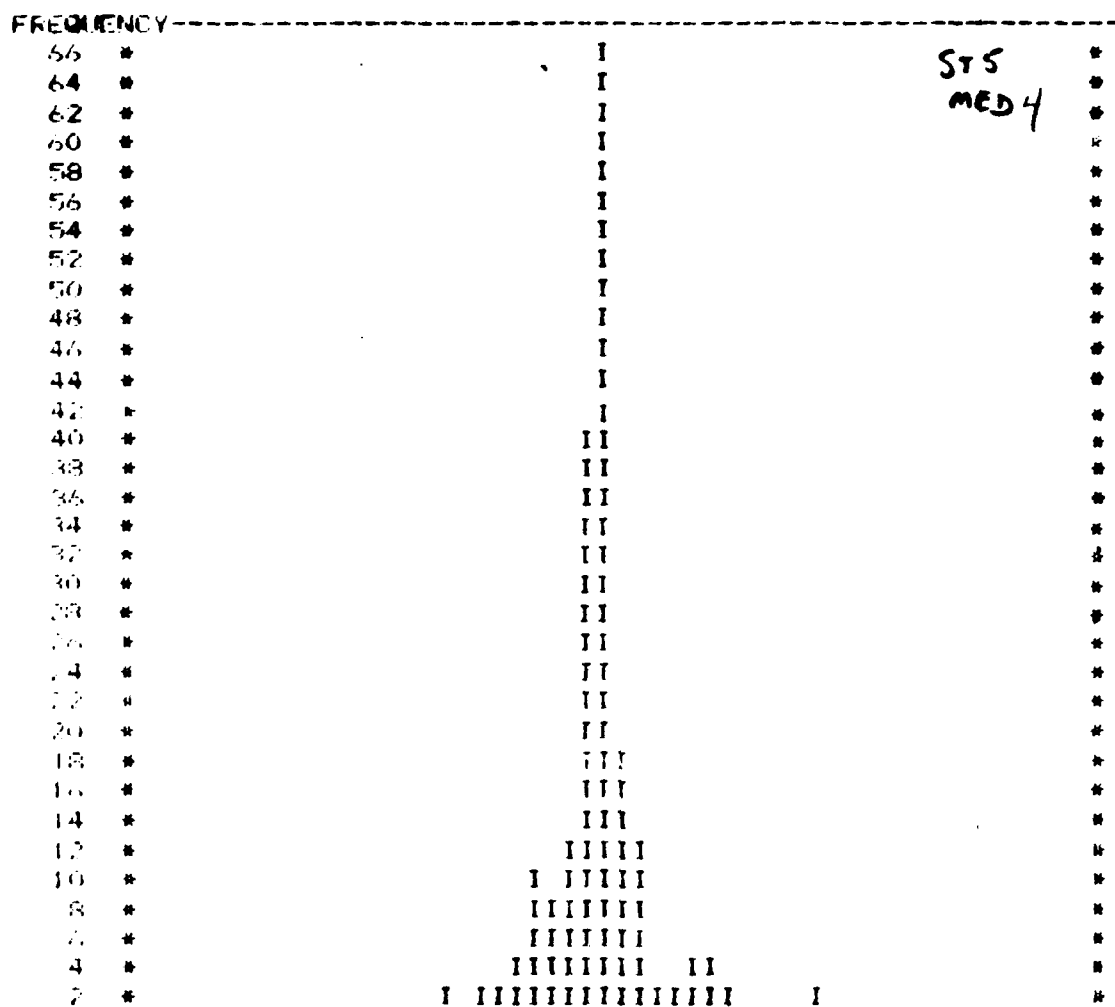
12:09

FREQUENCY-----				
67	*	I	STS	*
75	*	II	KAL	*
63	*	II		*
61	*	II		*
59	*	II		*
57	*	II		*
55	*	II		*
53	*	II		*
51	*	II		*
49	*	II		*
47	*	II		*
45	*	II		*
43	*	II		*
41	*	II		*
39	*	II		*
37	*	II		*
35	*	II		*
33	*	II		*
31	*	II		*
29	*	II		*
27	*	II		*
25	*	II		*
23	*	II		*
21	*	II		*
19	*	II		*
17	*	II		*
15	*	III		*
13	*	IIII		*
11	*	IIII		*
9	*	IIIII		*
7	*	IIIIII		*
5	*	IIIIII		*
3	*	IIIIII I I		*
1	*	IIIIIIIIIIII I		*

CLASS	5	10	15	20	25	0	30	35	40	45	50
MEAN =	.0157202997576					11.104	$\bar{X} = .0198$		} 43		
VARIANCE =	.05188707045175						$S^2 = .065$		} ZEROS DELETED		

95% Confidence Interval for μ
 (-.016, .046)

STRAT CLASS 5



CLASS	5	10	15	20	25	30	35	40	45	50
-------	---	----	----	----	----	----	----	----	----	----

MFAN = .04915071287663

VARIANCE = .176623729334

$$\bar{x} = -.0183$$

$G^2: .0774$

95% Confidence Interval for μ_2
(-1.0082, .1063)

N- 207

STRAT CLASS 5

FREQUENCY				
69	*	I	STS	*
67	*	I	TRIG	*
65	*	I		*
63	*	II		*
61	*	II		*
59	*	II		*
57	*	II		*
55	*	II		*
53	*	II		*
51	*	II		*
49	*	II		*
47	*	II		*
45	*	II		*
43	*	II		*
41	*	II		*
39	*	II		*
37	*	II		*
35	*	II		*
33	*	II		*
31	*	II		*
29	*	II		*
27	*	II		*
25	*	III		*
23	*	III		*
21	*	III		*
19	*	III		*
17	*	III		*
15	*	III		*
13	*	III		*
11	*	IIII		*
9	*	IIII		*
7	*	I IIII		*
5	*	IIIIII		*
3	*	I IIIII I II		*
1	*	II IIIIIIIIIII II		*

CLASS	5	10	15	20	25	30	35	40	45	50
MEAN =	- .0144977151132									
VARIANCE =	.04173208020638									

N=147

$\bar{Y} = .062$
 $S^2 = .221$

47
 ZEROS
 DENIED

95% Confidence Interval for μ_a
 (- .0482, .0194

FREQUENCY

72	*	I
70	*	I
68	*	I
66	*	I
64	*	I
62	*	I
60	*	I
58	*	II
56	*	II
54	*	II
52	*	II
50	*	II
48	*	II
46	*	II
44	*	II
42	*	II
40	*	II
38	*	II
36	*	II
34	*	II
32	*	II
30	*	II
28	*	II
26	*	II
24	*	II
22	*	III
20	*	III
18	*	IIII
16	*	IIII
14	*	IIII
12	*	IIII
10	*	IIII
8	*	IIII
6	*	IIIIII
4	*	IIIIIIII I
2	*	I IIIIIIIIII

MOVED
D
STS

CLASS	5	10	15	20	25	30	35	40	45	50
MEAN =	.007430825590905									
VARIANCE =	.1128929523846									

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