

AVAILABILITY OF PERIODICALLY INSPECTED SYSTEMS WITH MARKOVIAN WEAR AND SHOCKS

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Abstract

We analyze a periodically inspected system with hidden failures in which the rate of wear is modulated by a continuous-time Markov chain and additional damage is induced by a Poisson shock process. We derive explicitly the system's lifetime distribution and mean time to failure, as well as the limiting average availability. The main results are illustrated through two numerical examples.

Keywords: Shock model; wear process; reliability; average availability

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1. Introduction

In this paper, we consider reliability and availability measures for a single-unit system that suffers degradation due to its operating environment and the impact of shocks of random magnitude occurring at random time intervals. The system is said to experience a *soft failure* when its cumulative level of degradation exceeds a fixed threshold value that may be considered, for example, as the component's nominal

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14. ABSTRACT We analyze a periodically inspected system with hidden failures in which the rate of wear is modulated by a continuous-time Markov chain and additional damage is induced by a Poisson shock process. We derive explicitly the system's lifetime distribution and mean time to failure, as well as the limiting average availability. The main results are illustrated through two numerical examples.					
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service life. However, it is possible that failures remain hidden until a routine (planned) inspection of the unit reveals the failure. Using a periodic inspection policy, if the unit is found to have suffered an intolerable level of degradation, it is instantaneously replaced with a new and identical unit. On the other hand, if the level of degradation observed upon inspection is acceptable (i.e., below the specified threshold value), no action is taken until the subsequent scheduled inspection. One example of a system maintained using this type of inspection policy is that of tool wear in manufacturing environments. Though the equipment may be operational, its cutting tool may have suffered sufficient degradation to render it ineffective in achieving engineering specifications. However, its condition may not be known until an actual measurement of the accumulated degradation is obtained. Hidden failures are also prevalent in the antennas of a satellite tracking station. In such systems the stations, situated in remote locations around the globe, may be difficult to access and inspect. Hence, periodic inspection policies are needed to ensure proper communication links with on-orbit satellites for such critical functions as surveillance, communications, navigation, and warning.

The impact of a random environment on reliability and availability measures has been examined extensively in the applied probability literature. Models including degradation due only to wear or only to random shocks are especially prevalent. The seminal work of Esary, et al. [3] has been extended in numerous directions by several authors. For instance, Çinlar [2] generalized most of the models of [3] by demonstrating that the joint process of the unit's wear level and the state of its ambient environment may be considered as a Markov-additive process and gave several such examples. The first example considered the case when the random environment is a general Markov process with a finite state space and the wear was assumed to increase as a Lévy process. Additionally, random shocks were assumed to occur only at environment transition epochs. The second example (see [2], pp. 201-202) is similar to the problem discussed here where the cumulative wear is a continuous additive functional of the operating environment, and the first time to failure is a type of first passage time for the degradation process. Though Råde [10], Shanthikumar and Sumita [11], and Nakagawa [9] provide extensions of [3] they did not incorporate the effect of the unit's operating environment on reliability measures. Igaki et al. [4], Skoilkakis [12], and Kharoufeh [5] present degradation models that explicitly incorporate the influence of

the unit's operating environment but do not consider the maintenance of such systems.

More recently, other researchers have placed an emphasis on studying maintained systems. For example, Klutke et al. [7] examined the availability of an inspected system whose inter-inspection times and wear rates are random. Subsequently, Klutke and Yang [8] derived an availability result for a system subject to constant degradation, shocks, and a deterministic inspection policy. In a model similar to the one presented in [5] and in this paper, Kiessler et al. [6] studied the limiting average availability of a system for which the wear rate depends on a continuous-time Markov chain (CTMC). Their model considers an i.i.d. sequence of nominal lifetimes that determine the failure criterion as opposed to the deterministic threshold value we consider. However, unlike our model, [6] does not explicitly provide reliability measures and does not include the degradation due to random shocks.

This paper extends the results of [5] and [6] by providing both reliability and availability measures for a system subject to Markovian wear *and* degradation due to random shocks. Additionally, we demonstrate that the results may be implemented numerically in a straightforward manner by employing standard Laplace-transform inversion algorithms. We specifically provide, in closed-form, the Laplace-Stieltjes transform of the unconditional and conditional lifetime distribution functions (d.f.) as well as the unconditional and conditional mean system lifetime. Using our main results, and exploiting the Markov-regenerative nature of the system, we compute the limiting average availability of the system under a periodic inspection policy using the results of [6]. Moreover, we illustrate the computational ease with which the results may be obtained numerically through two example problems.

The remainder of the paper is organized as follows. In section 2, we provide the formal mathematical model description and notation used throughout the paper. In section 3, we analyze the model and explicitly derive the lifetime d.f.s and the mean lifetimes as Laplace-Stieltjes transforms. In section 4, we provide all the components needed to compute the limiting average availability of the system while section 5 provides two illustrative numerical examples.

2. Mathematical Model

In this section, we describe the mathematical model and notation for the maintained system with hidden failures. A single-unit system is placed into service at time zero in perfect working order. The system accumulates degradation until a deterministic degradation threshold value x is exceeded, at which time the system is said to be failed. We denote the random time to achieve the threshold by T_x . The degradation incurred by the system is attributed to environment-induced wear, as well as shocks that occur at random time intervals. The wear- and shock-inducing mechanisms are assumed to behave independently of one another. We further describe the stochastic evolution of the system in what follows. All random variables are assumed to be defined on a complete probability space $(\Omega, \mathcal{A}, \mathbb{P})$.

The random environment is characterized by an irreducible continuous-time Markov chain $\mathcal{Z} \equiv \{Z_t : t \geq 0\}$ on a finite state space $S \equiv \{1, 2, \dots, l\}$ with infinitesimal generator matrix $\mathbf{Q}$, transition probability functions $\pi_{i,j}(t)$, $i, j \in S$ and initial distribution (row) vector $\boldsymbol{\alpha}$. Whenever $Z_t = i$, the system accumulates wear at a rate $r(i)$, $r(i) > 0$ for $i = 1, 2, \dots, l$. Define $\mathbf{R}_D \equiv \text{diag}(r(1), r(2), \dots, r(l))$, the $l \times l$ diagonal matrix containing the state-dependent wear rates. The total accumulated wear up to time t , denoted W_t , is

$$W_t = \int_0^t r(Z_u) du. \quad (1)$$

The cumulative stochastic process $\mathcal{W} \equiv \{W_t : t \geq 0\}$, commonly known as a state-dependent wear process, has been studied extensively in the literature (cf. [3], [2], [5]) independent of a shock-inducing mechanism.

In addition to environment-induced wear, the system is damaged by shocks occurring at random time intervals. The damage caused by an individual shock is assumed to be relatively small; however, the cumulative effect of small shocks may be significant as in the case of fatigue deterioration resulting from mechanical vibrations. Denote by N_t the number of shocks occurring up to time t . The corresponding counting process $\{N_t : t \geq 0\}$ is assumed to be a temporally homogeneous Poisson process with rate parameter λ . The damage caused by the n th shock is a (nonnegative) random variable Y_n and $\{Y_n\}_{n=1}^\infty$ is an i.i.d. sequence with non-defective distribution function $F_Y(y) \equiv \mathbb{P}\{Y \leq y\}$. The total cumulative damage attributed to shocks up to time t is

the random variable,

$$\beta_t = \sum_{n=0}^{N_t} Y_n. \quad (2)$$

Therefore, the total degradation accrued by the system up to time t is the sum of degradation due to wear and that due to shocks given by

$$X_t = W_t + \beta_t, \quad t \geq 0. \quad (3)$$

The monotonicity of $\mathcal{X} \equiv \{X_t : t \geq 0\}$ and positivity of the degradation rates, $r(1), r(2), \dots, r(l)$, ensure the equivalence of events $\{X_t \leq x\}$ and $\{T_x > t\}$. Thus, the system's random lifetime is given by

$$T_x = \inf\{t > 0 : X_t \geq x\}, \quad (4)$$

or the first time the degradation process $\mathcal{X}$ crosses level x . Let $G(x, t) \equiv \mathbb{P}\{T_x \leq t\}$ denote the unconditional distribution function (d.f.) of the unit's lifetime and let its first moment be denoted by $\mathbb{E}[T_x]$. Note that (4) implies

$$G(x, t) = 1 - \mathbb{P}\{X_t \leq x\}.$$

In the remainder of this paper, we shall adopt the following notation. We let $\mathbb{E}_i[\cdot] = \mathbb{E}[\cdot | Z_0 = i]$ and $\mathbb{P}_i[\cdot] = \mathbb{P}[\cdot | Z_0 = i]$. Thus, the conditional distribution of the (first) lifetime is given by

$$G_i(x, t) = 1 - \mathbb{P}_i\{X_t \leq x\}.$$

The unconditional (conditional) mean (first) lifetime shall be denoted by $\mathbb{E}[T_x]$ ($\mathbb{E}_i[T_x]$). For the purpose of computing the limiting average availability of the maintained system, we obtain each of these quantities in the form of Laplace-Stieltjes transforms in section 3.

3. Lifetime Distribution and Mean Lifetime

In this section, we provide analytical expressions for the system's lifetime distribution function and the mean system lifetime in the form of Laplace-Stieltjes transforms (LSTs). To that end, we introduce the following definitions and notation. The complement of the lifetime distribution is

$$R(x, t) \equiv \overline{G}(x, t) = \mathbb{P}\{X_t \leq x\}. \quad (5)$$

where (5) holds due to the dual relationship $T_x = \inf\{t > 0 : X_t \geq x\}$. Define the joint probability distribution

$$R_{i,j}(x, t) = \mathbb{P}\{X_t \leq x, Z_t = j | Z_0 = i\}, \quad i, j \in S, \quad (6)$$

Our first main result characterizes the joint distribution of the process $(\mathcal{X}, \mathcal{Z})$.

Theorem 1. *The distribution function $R_{i,j}(x, t)$, $i, j \in S$ satisfies the partial differential equation*

$$\frac{\partial R_{i,j}(x, t)}{\partial t} + \frac{\partial R_{i,j}(x, t)}{\partial x} r(j) = \lambda \{ [R_{i,j}(\cdot, t) * F_Y](x) - R_{i,j}(x, t) \} + \sum_{k=1}^l q_{k,j} R_{i,k}(x, t) \quad (7)$$

for $x > 0$, $t \geq 0$ where $(*)$ denotes the convolution operator.

Proof. We begin by considering only the environment-induced wear of the system and subsequently incorporate shocks. To that end, set $V_{i,j}(x, t) \equiv \mathbb{P}\{W_t \leq x, Z_t = j | Z_0 = i\}$ and let $\epsilon > 0$ denote a small time increment. Because $\mathcal{Z}$ possesses the Markov property, is temporally homogeneous, and is independent of the degradation process $\mathcal{X}$, we may write (see [5])

$$\begin{aligned} V_{i,j}(x, t + \epsilon) &= \sum_{k=1}^l \mathbb{P}\{W_{t+\epsilon} \leq x, Z_{t+\epsilon} = j | Z_t = k, Z_0 = i\} \mathbb{P}\{Z_t = k\} \\ &= \sum_{k=1}^l \mathbb{P}\{Z_{t+\epsilon} = j | Z_t = k\} \mathbb{P}\{W_{t+\epsilon} \leq x, Z_t = k | Z_0 = i\} \\ &= \sum_{k=1}^l \pi_{k,j}(\epsilon) V_{i,k}(x - r(k)\epsilon, t). \end{aligned} \quad (8)$$

Next we incorporate the impact of Poisson shocks occurring at homogeneous rate λ . The probability mass function (p.m.f.) of the number of shocks occurring up to time ϵ is $p_n(\epsilon) \equiv \mathbb{P}\{N_\epsilon = n\}$, $n \geq 0$. The magnitude of cumulative damage caused by n -independent shocks is given by

$$\beta_n = \sum_{i=1}^n Y_i. \quad (9)$$

Since shock magnitudes form an i.i.d. sequence with distribution function F_Y , we note that

$$\mathbb{P}\{\beta_n \leq y\} \equiv F_Y^{(n)}(y),$$

where $F_Y^{(n)}$ denotes the n -fold convolution of F_Y . Conditioning on the number of shocks in the interval $(t, t + \epsilon)$ and the magnitude of damage due to those shocks, equation (8) may be replaced by

$$R_{i,j}(x, t + \epsilon) = \sum_{k=1}^l \sum_{n=0}^{\infty} \int_0^{\infty} \pi_{k,j}(\epsilon) R_{i,k}(x - r(k)\epsilon - y, t) p_n(\epsilon) F_Y^{(n)}(dy). \quad (10)$$

Substituting the appropriate expressions for $p_0(\epsilon)$ and $p_1(\epsilon)$ in (10) and noting that, for ϵ small, $\mathbb{P}\{N_\epsilon = n\} = o(\epsilon)$, $n \geq 2$, we may write

$$R_{i,j}(x, t + \epsilon) = \sum_{k=1}^l \pi_{k,j}(\epsilon) \times \left(R_{i,k}(x - r(k)\epsilon, t)(1 - \lambda\epsilon) + \lambda\epsilon \int_0^{\infty} R_{i,k}(x - r(k)\epsilon - y, t) F_Y(dy) \right) + o(\epsilon). \quad (11)$$

The transition probability functions for the $\mathcal{Z}$ process, $\pi_{i,j}(\epsilon)$, $i, j \in S$, can be written as

$$\pi_{i,j}(\epsilon) = \delta_{i,j} + \epsilon q_{i,j} + o(\epsilon) \quad (12)$$

where $\delta_{i,j}$ assumes the value 1 when $i = j$ and the value 0 when $i \neq j$. Substituting (12) into (11) and simplifying gives

$$\begin{aligned} R_{i,j}(x, t + \epsilon) = & (1 - \lambda\epsilon) R_{i,j}(x - r(j)\epsilon, t) + \epsilon(1 - \lambda\epsilon) \sum_{k=1}^l q_{k,j} R_{i,k}(x - r(k)\epsilon, t) \\ & + \lambda\epsilon \int_0^{\infty} R_{i,j}(x - r(j)\epsilon - y, t) F_Y(dy) \\ & + \lambda\epsilon^2 \sum_k q_{k,j} \int_0^{\infty} R_{i,k}(x - r(k)\epsilon - y, t) F_Y(dy) + o(\epsilon). \end{aligned} \quad (13)$$

Rearranging and simplifying the terms of (13), dividing through by the time increment ϵ , and allowing $\epsilon \rightarrow 0$, shows that

$$\begin{aligned} \frac{\partial R_{i,j}(x, t)}{\partial t} + \frac{\partial R_{i,j}(x, t)}{\partial x} r(j) = & \\ - \lambda R_{i,j}(x, t) + \sum_{k=1}^l q_{k,j} R_{i,k}(x, t) + \lambda \int_0^{\infty} R_{i,j}(x - y, t) F_Y(dy) \end{aligned} \quad (14)$$

where the right-most term of (14) is the convolution of distributions $R_{i,j}$ and F_Y .

We next set out to solve the partial differential equations of (7) by the method of Laplace transforms. To that end, define the $l \times l$ matrix $\mathbf{R}(x, t) = [R_{ij}(x, t)]$. Using

matrix notation, (7) may be written as

$$\frac{\partial \mathbf{R}(x, t)}{\partial t} + \frac{\partial \mathbf{R}(x, t)}{\partial x} \mathbf{R}_D = \lambda \{ [\mathbf{R}(\cdot, t) * F_Y](x) - \mathbf{R}(x, t) \} + \mathbf{R}(x, t) \mathbf{Q} \quad (15)$$

where $[\mathbf{R}(\cdot, t) * F_Y](x)$ denotes the convolution of each element of $\mathbf{R}(x, t)$ with F_Y . We next provide a Laplace-Stieltjes transform (LST) solution with respect to the spatial variable x for the partial differential equation (15). Define the matrix LST of $\mathbf{R}(x, t)$ with respect to x as

$$\tilde{\mathbf{R}}(u, t) = \int_0^\infty e^{-ux} \mathbf{R}(dx, t) \quad (16)$$

and the $l \times l$ diagonal matrix $\tilde{\mathbf{F}}_D(u)$ with each diagonal element identically equal to

$$\tilde{F}_Y(u) = \int_0^\infty e^{-uy} F_Y(dy), \quad (17)$$

the LST of F_Y with respect to y .

Theorem 2. *The Laplace-Stieltjes transform of the matrix $\mathbf{R}(x, t)$ with respect to the spatial variable x is*

$$\tilde{\mathbf{R}}(u, t) = \exp \left[\left(\mathbf{Q} + \lambda(\tilde{\mathbf{F}}_D(u) - \mathbf{I}) - u\mathbf{R}_D \right) t \right]. \quad (18)$$

Proof. Taking the Laplace-Stieltjes transform of both sides of (15) with respect to x yields the first order, ordinary differential equation (ODE) in t ,

$$\frac{d\tilde{\mathbf{R}}(u, t)}{dt} + \tilde{\mathbf{R}}(u, t) \left(u\mathbf{R}_D + \lambda\mathbf{I} - \mathbf{Q} - \lambda\tilde{\mathbf{F}}_D(u) \right) = \mathbf{0}. \quad (19)$$

The general solution of the ordinary differential equation is obtained via the use of an integrating factor

$$\tilde{\mathbf{R}}(u, t) \exp \left[\left(u\mathbf{R}_D + \lambda\mathbf{I} - \mathbf{Q} - \lambda\tilde{\mathbf{F}}_D(u) \right) t \right] = \boldsymbol{\psi},$$

where $\boldsymbol{\psi}$ is a matrix of constants of integration. The final result is obtained by applying the initial condition $\tilde{\mathbf{R}}(u, 0) = \mathbf{I}$ and rearranging terms.

The initial probability vector of the environment process $\mathcal{Z}$ is $\boldsymbol{\alpha} = [\alpha_i]$ where $\alpha_i \equiv \mathbb{P}\{Z_0 = i\}$, $i \in S$. Let $\mathbf{e}$ denote a column vector of ones and denote by $\mathbf{e}_i$ a column vector whose i th element is unity and all other elements are zero. An explicit expression for the unconditional d.f. of the system's lifetime is

$$G(x, t) \equiv \mathbb{P}\{T_x \leq t\} = 1 - \boldsymbol{\alpha} \mathbf{R}(x, t) \mathbf{e} \quad (20)$$

and the conditional distribution is

$$G_i(x, t) \equiv \mathbb{P}_i\{T_x \leq t\} = 1 - \mathbf{e}_i^T \mathbf{R}(x, t) \mathbf{e} \quad (21)$$

where $\mathbf{e}_i^T$ denotes the transpose of $\mathbf{e}_i$.

Define the Laplace-Stieltjes transforms of $G(x, t)$ and $G_i(x, t)$ with respect to x as

$$\tilde{G}(u, t) = \int_0^\infty e^{-ux} G(dx, t) \quad (22)$$

and

$$\tilde{G}_i(u, t) = \int_0^\infty e^{-ux} G_i(dx, t), \quad (23)$$

respectively.

Theorem 3. *The Laplace-Stieltjes transform of the unconditional and conditional first lifetime distributions, $G(x, t)$ and $G_i(x, t)$, with respect to x are, respectively,*

$$\tilde{G}(u, t) = 1 - \boldsymbol{\alpha} \exp\left((\mathbf{Q} + \lambda(\tilde{\mathbf{F}}_D(u) - \mathbf{I}) - u\mathbf{R}_D)t\right) \mathbf{e} \quad (24)$$

and

$$\tilde{G}_i(u, t) = 1 - \mathbf{e}_i^T \exp\left((\mathbf{Q} + \lambda(\tilde{\mathbf{F}}_D(u) - \mathbf{I}) - u\mathbf{R}_D)t\right) \mathbf{e}, \quad (25)$$

with $\text{Re}(u) > 0$.

Next we examine the unconditional (conditional) mean system lifetime, $\mathbb{E}[T_x]$ ($\mathbb{E}_i[T_x]$) for a single-unit system subject to Markovian wear and a Poisson shock process. The LSTs of $\mathbb{E}[T_x]$ and $\mathbb{E}_i[T_x]$, with respect to x , are respectively,

$$\tilde{\mathbb{E}}[T_u] = \int_0^\infty e^{-ux} d_x \mathbb{E}[T_x]$$

and

$$\tilde{\mathbb{E}}_i[T_u] = \int_0^\infty e^{-ux} d_x \mathbb{E}_i[T_x],$$

which are obtained via Theorem 4.

Theorem 4. *The Laplace-Stieltjes transform of the unconditional mean lifetime with respect to x is*

$$\tilde{\mathbb{E}}[T_u] = \boldsymbol{\alpha} \left(u\mathbf{R}_D - \mathbf{Q} - \lambda(\tilde{\mathbf{F}}_D(u) - \mathbf{I})\right)^{-1} \mathbf{e}, \quad (26)$$

and the conditional mean lifetime is

$$\tilde{\mathbb{E}}_i[T_u] = \mathbf{e}_i^T \left(u\mathbf{R}_D - \mathbf{Q} - \lambda(\tilde{\mathbf{F}}_D(u) - \mathbf{I}) \right)^{-1} \mathbf{e}, \quad (27)$$

with $\text{Re}(u) > 0$.

Proof. By (24), the d.f. of the system lifetime is

$$\tilde{G}(u, t) = 1 - \boldsymbol{\alpha} \exp \left[\left(\mathbf{Q} + \lambda(\tilde{\mathbf{F}}_D(u) - \mathbf{I}) - u\mathbf{R}_D \right) t \right] \mathbf{e}.$$

For brevity, substitute $\mathbf{A} = \mathbf{Q} + \lambda(\tilde{\mathbf{F}}_D(u) - \mathbf{I}) - u\mathbf{R}_D$ to obtain the LST of $\tilde{G}(u, t)$ with respect to t as

$$\begin{aligned} \tilde{G}(u, s) &= \int_0^\infty e^{-st} \tilde{G}(u, dt) \\ &= 1 - \boldsymbol{\alpha} \mathbf{A} (s\mathbf{I} - \mathbf{A})^{-1} \mathbf{e}. \end{aligned} \quad (28)$$

In the usual way, the mean system lifetime is obtained by evaluating at $s = 0$ the first partial derivative of $\tilde{G}(u, s)$ with respect to s . That is

$$\begin{aligned} \tilde{\mathbb{E}}[T_u] &= - \left. \frac{\partial \tilde{G}(u, s)}{\partial s} \right|_{s=0} \\ &= -\boldsymbol{\alpha} (-\mathbf{A})^{-1} \mathbf{e} \\ &= \boldsymbol{\alpha} (u\mathbf{R}_D - \mathbf{Q} - \lambda(\tilde{\mathbf{F}}_D(u) - \mathbf{I}))^{-1} \mathbf{e}. \end{aligned}$$

The conditional mean lifetime is obtained analogously by simply replacing $\tilde{G}(u, t)$ by $\tilde{G}_i(u, t)$.

4. Limiting Average Availability

In this section, we study the asymptotic behavior of the average availability of the maintained system with hidden failures. The main result for the limiting average availability (Theorem 5) is contained in Kiessler et al. [6]; however, we review this result within the framework of our model which differs from that of [6] in a few important ways. In our model, a failure occurs when the cumulative degradation (due to wear and shocks) reaches or exceeds a deterministic threshold. The model in [6] does not consider shocks. Moreover, the numerical approach we use for computing the limiting average availability is significantly simpler. Specifically, we apply the Laplace-Stieltjes

transforms of section 3 and numerical inversion to simplify the computation of the required components.

The system is inspected periodically at times in the set $I = \{k\tau : k = 1, 2, \dots\}$ for some $\tau > 0$. If upon inspection the system is found to be failed (i.e., the cumulative degradation of the system is at least as large as the tolerable level x), then it is instantaneously replaced with a new and identical unit; however, if the system is found to be non-failed, then no action is taken. Failures occurring between inspection epochs cannot be detected nor repaired; thus, failures remain hidden until the next inspection epoch, and the process $\mathcal{X}$ remains at the failed level until the subsequent inspection epoch. We further assume that inspections correctly diagnose the level of degradation. Define $\Psi \equiv \{\psi(t) : t \geq 0\}$, the right-continuous stochastic process describing the system state by

$$\psi(t) = 1(X_t < x),$$

where $1(E)$ denotes the indicator function for event E . As defined in [6], the system's limiting average availability is given by

$$\bar{A} = \lim_{t \rightarrow \infty} t^{-1} \int_0^t \mathbb{E}[\psi(w)] dw. \quad (29)$$

Let L_n denote the lifetime of the n th system and let R_n denote the n th replacement epoch (with $R_0 \equiv 0$). We further define $\xi_n = Z_{R_n}$, the state of the environment at the time of the n th replacement, and F_n , the n th failure epoch. The process $\{\xi_n : n \geq 0\}$ is an irreducible discrete-time Markov chain with one-step transition probability matrix $\mathbf{P}$ and stationary distribution $\mathbf{p} = [p_i]$, $i = 1, 2, \dots, l$. It is important to note that, by its nature, $\Psi \equiv \{\psi(t) : t \geq 0\}$ is *not* regenerative with respect to the sequence of failure or replacement epochs. This is due to the fact that successive unit lifetimes do not form an i.i.d. sequence of random variables unless the successive units are placed into service in identical environment states. Kiessler, et al. [6] showed that the bivariate process, $\{(\xi_n, R_n) : n \geq 0\}$, is a Markov renewal process and that $\{\psi(t) : t \geq 0\}$ is Markov-regenerative with respect to $\{(\xi_n, R_n) : n \geq 0\}$. Consequently, the limiting average availability is obtained by Theorem 5 (see [6], p. 704).

Theorem 5. *The limiting average availability is given by*

$$\bar{A} = \lim_{t \rightarrow \infty} t^{-1} \int_0^t \mathbb{E}[\psi(w)] dw = \frac{\sum_{i=1}^l p_i \mathbb{E}_i[F_1]}{\sum_{i=1}^l p_i \mathbb{E}_i[R_1]}. \quad (30)$$

In order to compute the limiting average availability, we will obtain the transition probability matrix $\mathbf{P}$ as the limit of the semi-Markov kernel of the process $\{(\xi_n, R_n) : n \geq 0\}$. Subsequently, we obtain the stationary distribution $\mathbf{p}$. Additionally, we require the conditional expectations $\mathbb{E}_i[R_1]$ and $\mathbb{E}_i[F_1]$. Fortunately, the Laplace-Stieltjes transform of $\mathbb{E}_i[F_1]$ may be obtained by (27), and the conditional distribution of F_1 is given by (25). Thus, our next objective is to provide an expression for $\mathbb{E}_i[R_1]$. However, we first need to establish that F_1 is a bounded random variable. Let us re-order the wear rates so that

$$r_1 < r_2 < \cdots < r_l.$$

That is, r_1 corresponds to the minimum attainable wear rate. The following lemma holds for both the conditional and unconditional first unit lifetimes.

Lemma 1. *The random variable F_1 is bounded above by xr_1^{-1} .*

Proof. Let $D \equiv \{r(i) : i = 1, 2, \dots, l\}$ and let $\mathbf{r} \equiv (r_1, r_2, \dots, r_l)^T$ denote the column vector of ordered wear rates. Further define the set

$$A \equiv \{\mathbf{a} \in \mathbb{R}_+^{1 \times l} : \mathbf{a}\mathbf{e} = xr_1^{-1}\}.$$

We first note that

$$\inf\{\mathbf{a}\mathbf{r} + b : \mathbf{a} \in A, b \geq 0\} \leq x \tag{31}$$

since for $\mathbf{a} = xr_1^{-1}\mathbf{e}_1$ and $b = 0$, $\mathbf{a}\mathbf{r} + b = x$. Moreover,

$$\begin{aligned} \inf\{\mathbf{a}\mathbf{r} + b : \mathbf{a} \in A, b \geq 0\} &= \inf\{\mathbf{a}\mathbf{r} : \mathbf{a} \in A\} \\ &\geq \inf\{\mathbf{a}(r_1\mathbf{e}) : \mathbf{a} \in A\} \\ &= x. \end{aligned} \tag{32}$$

By (31) and (32) we see that

$$x = \min\{\mathbf{a}\mathbf{r} + b : \mathbf{a} \in A, b \geq 0\}. \tag{33}$$

Let $\omega \in \Omega$, and define the quantity

$$w_n(\omega) \equiv \int_0^{x/r_1} 1(r(Z_t(\omega)) = r_n) dt.$$

Summing over all possible environment states, we see that

$$\begin{aligned}
\sum_{n=1}^l w_n(\omega) &= \sum_{n=1}^l \int_0^{x/r_1} 1(r(Z_t(\omega)) = r_n) dt \\
&= \int_0^{x/r_1} \sum_{n=1}^l 1(r(Z_t(\omega)) = r_n) dt \\
&= \int_0^{x/r_1} 1(r(Z_t(\omega)) \in D) dt \\
&= \int_0^{x/r_1} dt \\
&= xr_1^{-1},
\end{aligned}$$

since $r(Z_t(\omega)) \in D$ for almost every $t \geq 0$. It follows that the row vector $\mathbf{w}(\omega) \equiv (w_1(\omega), \dots, w_l(\omega)) \in A$.

Recall that the total damage due to shocks on the interval $[0, x/r_1)$ is given by $\beta_{x/r_1}(\omega)$. Using equation (3), we may write

$$\begin{aligned}
X_{x/r_1}(\omega) &= W_{x/r_1}(\omega) + \beta_{x/r_1}(\omega) \\
&= \int_0^{x/r_1} r(Z_t(\omega)) dt + \beta_{x/r_1}(\omega) \\
&= \int_0^{x/r_1} \sum_{n=1}^l r_n 1(r(Z_t(\omega)) = r_n) dt + \beta_{x/r_1}(\omega) \\
&= \sum_{n=1}^l \int_0^{x/r_1} r_n 1(r(Z_t(\omega)) = r_n) dt + \beta_{x/r_1}(\omega) \\
&= \sum_{n=1}^l w_n(\omega) r_n + \beta_{x/r_1}(\omega) \\
&= \mathbf{w}(\omega) \mathbf{r} + \beta_{x/r_1}(\omega).
\end{aligned}$$

By (33) we may conclude that

$$x = \min\{X_{x/r_1}(\omega) : \omega \in \Omega\}.$$

Since $X_{x/r_1}(\omega) \geq x$, it is clear that

$$|F_1(\omega)| = F_1(\omega) = \inf\{t > 0 : X_t(\omega) \geq x\} \leq xr_1^{-1},$$

i.e., the non-negative random variable F_1 is bounded above by xr_1^{-1} .

Lemma 1 implies that $G_i(x, t) = 1$ for all $t \geq xr_1^{-1}$. Equipped with this result, we are now prepared to establish the mean of the first replacement epoch and, subsequently, the limiting average availability.

Lemma 2. *The conditional expectation of the first replacement epoch is given by*

$$\mathbb{E}_i[R_1] = \tau \left(\gamma - \sum_{n=0}^{\gamma-1} G_i(x, n\tau) \right) \quad (34)$$

where $\gamma \equiv \min\{n \geq 1 : n\tau \geq xr_1^{-1}\}$ and $r_1 \equiv \min\{r(i) : i = 1, 2, \dots, l\}$.

Proof. The first replacement time R_1 depends explicitly on F_1 by

$$R_1 = \inf\{w \in I : w > F_1\}.$$

By construction, F_1 corresponds to the lifetime of the first unit, and as such, has the (conditional) distribution function $G_i(x, t)$, $i \in S$. By conditioning on F_1 , we obtain

$$\mathbb{E}_i[R_1] = \int_0^\infty \mathbb{E}_i[R_1 | F_1 = t] G_i(x, dt). \quad (35)$$

Because a failure remains hidden until the first inspection following the failure, the integrand of (35) is

$$\mathbb{E}_i[R_1 | F_1 = t] = \begin{cases} \tau, & 0 \leq t < \tau \\ 2\tau, & \tau \leq t < 2\tau \\ \vdots & \\ n\tau, & (n-1)\tau \leq t < n\tau \\ \vdots & \end{cases}. \quad (36)$$

Substituting (36) into (35), we may write

$$\begin{aligned} \mathbb{E}_i[R_1] &= \tau \int_0^\tau G_i(x, dt) + 2\tau \int_\tau^{2\tau} G_i(x, dt) + 3\tau \int_{2\tau}^{3\tau} G_i(x, dt) + \dots, \\ &= -\tau \sum_{n=0}^{\infty} G_i(x, n\tau). \end{aligned} \quad (37)$$

By applying Lemma 1 it can be shown that the infinite series, $-\sum_n G_i(x, n\tau)$, is convergent with sum

$$\sum_{n=0}^{\infty} -G_i(x, n\tau) = \gamma - \sum_{n=0}^{\gamma-1} G_i(x, n\tau), \quad (38)$$

where $\gamma = \min\{n \geq 1 : n\tau \geq xr_1^{-1}\}$. Therefore,

$$\mathbb{E}_i[R_1] = \tau \left(\gamma - \sum_{n=0}^{\gamma-1} G_i(x, n\tau) \right). \quad (39)$$

Finally, we turn our attention to the transition probability matrix, $\mathbf{P}$, for the discrete-time Markov chain $\{\xi_n : n \geq 0\}$. Before deriving an expression for the elements of $\mathbf{P}$, we need to establish the conditional distribution of R_1 given the initial environment state.

Lemma 3. *Assuming the initial environment state is $i \in S$, the probability mass function of R_1 is given by*

$$\mathbb{P}_i\{R_1 = n\tau\} = \Delta_i(x, n\tau), \quad n \geq 1 \quad (40)$$

where $\Delta_i(x, n\tau) \equiv G_i(x, n\tau) - G_i(x, (n-1)\tau)$ for $i \in S$ and $n \geq 1$.

Proof. Recall that $R_1 = \inf\{w \in I : w > F_1\}$. Thus, we obtain the probability distribution of R_1 by conditioning on the first failure epoch given the initial state of Z . For $n \geq 1$,

$$\begin{aligned} \mathbb{P}_i\{R_1 = n\tau\} &= \int_0^\infty \mathbb{P}_i\{R_1 = n\tau | F_1 = t\} G_i(x, dt) \\ &= \int_0^\infty 1(t \in [(n-1)\tau, n\tau]) G_i(x, dt) \\ &= \int_{(n-1)\tau}^{n\tau} G_i(x, dt) \\ &= G_i(x, n\tau) - G_i(x, (n-1)\tau). \end{aligned}$$

Finally, we obtain the one-step transition probability matrix for $\{\xi_n : n \geq 0\}$ by computing the limit of the semi-Markov kernel matrix of the Markov renewal process $\{(\xi_n, R_n) : n \geq 0\}$.

Theorem 6. *The (i, k) th element of the transition probability matrix $\mathbf{P} = [p_{i,k}]$, $i, k \in S$, is given by*

$$p_{i,k} = \sum_{n=1}^{\gamma} \pi_{i,k}(n\tau) \Delta_i(x, n\tau), \quad (41)$$

where $\Delta_i(x, n\tau)$ is obtained by

$$\Delta_i(x, n\tau) = \mathcal{L}^{-1}\{u^{-1}[\tilde{G}_i(u, n\tau) - \tilde{G}_i(u, (n-1)\tau)]\}, \quad (42)$$

$\mathcal{L}^{-1}$ denotes the inverse Laplace transform operator, and the integer γ is as defined in Lemma 2.

Proof. Let $K_{i,k}(t) \equiv \mathbb{P}\{\xi_1 = k, R_1 \leq t | \xi_0 = i\}$ denote the (i, k) th element of the semi-Markov kernel, $\mathbf{K}(t)$, of the Markov renewal process $\{(\xi_n, R_n) : n \geq 0\}$. Let us further define

$$\tilde{K}_{i,k}(s) = \mathbb{E}_i[e^{-sR_1} \cdot 1(Z_{R_1} = k)]$$

with $\text{Re}(s) > 0$. By conditioning on R_1 , we may write,

$$\begin{aligned} \mathbb{E}_i[e^{-sR_1} \cdot 1(Z_{R_1} = k)] &= \sum_{n=1}^{\infty} \mathbb{E}_i[e^{-sR_1} \cdot 1(Z_{R_1} = k) | R_1 = n\tau] \mathbb{P}_i\{R_1 = n\tau\} \\ &= \sum_{n=1}^{\infty} e^{-sn\tau} \mathbb{E}_i[1(Z_{n\tau} = k)] \Delta_i(x, n\tau) \\ &= \sum_{n=1}^{\infty} e^{-sn\tau} \mathbb{P}_i\{Z_{n\tau} = k\} \Delta_i(x, n\tau) \\ &= \sum_{n=1}^{\infty} e^{-sn\tau} \pi_{i,k}(n\tau) \Delta_i(x, n\tau). \end{aligned} \quad (43)$$

By the final value theorem of Laplace transforms, we may obtain $p_{i,k}$ as

$$\begin{aligned} p_{i,k} &= \lim_{t \rightarrow \infty} K_{i,k}(t) \\ &= \lim_{s \rightarrow 0} \tilde{K}_{i,k}(s) \\ &= \lim_{s \rightarrow 0} \sum_{n=1}^{\infty} e^{-sn\tau} \pi_{i,k}(n\tau) \Delta_i(x, n\tau). \end{aligned}$$

Recall from Lemma 1 that $\Delta_i(x, n\tau) = 0$ for all $n \geq \gamma + 1$, and so we have

$$\begin{aligned} p_{i,k} &= \lim_{s \rightarrow 0} \sum_{n=1}^{\infty} e^{-sn\tau} \pi_{i,k}(n\tau) \Delta_i(x, n\tau) \\ &= \lim_{s \rightarrow 0} \sum_{n=1}^{\gamma} e^{-sn\tau} \pi_{i,k}(n\tau) \Delta_i(x, n\tau) \\ &= \sum_{n=1}^{\gamma} \lim_{s \rightarrow 0} e^{-sn\tau} \pi_{i,k}(n\tau) \Delta_i(x, n\tau) \\ &= \sum_{n=1}^{\gamma} \pi_{i,k}(n\tau) \Delta_i(x, n\tau). \end{aligned}$$

Comparing equation (41) with Corollary 1 of [6], p. 708, we note that our expression requires the evaluation of only a finite sum and (numerical) Laplace transform inversion

to obtain $\Delta_i(x, n\tau)$, $\mathbb{E}_i[F_1]$, and $\mathbb{E}_i[R_1]$. Fortunately, numerical inversion algorithms abound for this task, and the truncation point γ is well defined by the parameters τ and r_1 . By contrast, the results of [6] require evaluation of an infinite Fourier series to compute $p_{i,k}$ and the conditional expectations.

5. Numerical Examples

In this section, we illustrate the computation of the limiting average availability by means of two examples. For both examples, we compute the conditional expectations, $\mathbb{E}_i[F_1]$ and $\mathbb{E}_i[R_1]$, the matrix $\mathbf{P}$, the stationary distribution $\mathbf{p}$, and the limiting average availability, $\bar{A}$. The results were obtained via numerical Laplace-transform inversion using the robust technique of Abate and Whitt [1]. The algorithm was coded and executed in the Matlab[®] computing environment. In each of the two cases, $\mathcal{Z}$ is an irreducible continuous-time Markov chain on finite state space S with generator matrix $\mathbf{Q}$.

Example 1: $l = 2$

The first case is a simple environment which toggles between two distinct states ($S = \{1, 2\}$) according to its generator matrix

$$\mathbf{Q} = \begin{bmatrix} -25/3 & 25/3 \\ 25/3 & -25/3 \end{bmatrix}.$$

The diagonal matrix of state-dependent wear rates is

$$\mathbf{R}_D = \text{diag}(13/12, 1/4),$$

while the Laplace-Stieltjes transform of the shock-magnitude distribution function is

$$\tilde{F}_Y(u) = \frac{4.0}{4.0 + u}.$$

Shocks occur according to a Poisson process with rate parameter $\lambda = 0.50$. The threshold degradation level for failure is $x = 1.0$ units of degradation. Inspections occur every $\tau = 0.10$ time units. Applying equations (27) and (34), we compute

$$\begin{aligned} \mathbb{E}_1[F_1] &= 1.2976 \\ \mathbb{E}_2[F_1] &= 1.3609 \end{aligned}$$

and

$$\begin{aligned}\mathbb{E}_1[R_1] &= 1.3475 \\ \mathbb{E}_2[R_1] &= 1.4109.\end{aligned}$$

Selecting the truncation point γ by Theorem 6, the transition probability matrix is

$$\mathbf{P} = \begin{bmatrix} 0.5002 & 0.4998 \\ 0.4998 & 0.5002 \end{bmatrix}$$

with the obvious stationary distribution

$$\mathbf{p} = \begin{pmatrix} 0.500 & 0.500 \end{pmatrix}.$$

Subsequently, the limiting average availability is computed as

$$\bar{A} = \frac{\sum_{i=1}^2 p_i \mathbb{E}_i[F_1]}{\sum_{i=1}^2 p_i \mathbb{E}_i[R_1]} = \mathbf{0.96378}.$$

Example 2: $l = 5$

In this case, the environment process is a five-state, continuous-time Markov chain with state space $S = \{1, 2, 3, 4, 5\}$. The generator matrix for the process is

$$\mathbf{Q} = \begin{bmatrix} -0.500 & 0.125 & 0.125 & 0.125 & 0.125 \\ 0.400 & -2.00 & 0.400 & 0.600 & 0.600 \\ 0.025 & 0.025 & -0.100 & 0.025 & 0.025 \\ 0.050 & 0.050 & 0.050 & -0.200 & 0.050 \\ 1.500 & 1.000 & 1.000 & 1.500 & -5.000 \end{bmatrix}.$$

The diagonal matrix of state-dependent wear rates is

$$\mathbf{R}_D = \text{diag}(1.0, 2.0, 3.0, 4.0, 10.0),$$

while the Laplace-Stieltjes transform of the shock-magnitude distribution function is

$$\tilde{F}_Y(u) = \left(\frac{0.20}{0.20 + u} \right)^8.$$

Shocks occur according to a Poisson process with parameter $\lambda = 0.25$. The threshold degradation level for failure of the system is $x = 100$ units of degradation. Inspections

occur every $\tau = 5.00$ time units. Applying (41) gives

$$\mathbf{P} = \begin{bmatrix} 0.1393 & 0.0300 & 0.4810 & 0.3368 & 0.0129 \\ 0.1262 & 0.0291 & 0.4868 & 0.3454 & 0.0126 \\ 0.1045 & 0.0258 & 0.6115 & 0.2471 & 0.0112 \\ 0.1255 & 0.0301 & 0.4181 & 0.4133 & 0.0130 \\ 0.1268 & 0.0291 & 0.4876 & 0.3439 & 0.0126 \end{bmatrix}.$$

Table 1 summarizes the stationary probabilities and the conditional expectations for this case.

TABLE 1: Stationary probabilities and conditional means (Ex. 2).

Measure	$i = 1$	$i = 2$	$i = 3$	$i = 4$	$i = 5$
p_i	0.1160	0.0277	0.5308	0.3135	0.0120
$\mathbb{E}_i[R_1]$	11.6861	11.4400	11.4752	11.2014	11.3457
$\mathbb{E}_i[F_1]$	9.1931	8.9485	8.9836	8.7116	8.8542

From these values, we compute the limiting average availability

$$\bar{A} = \frac{\sum_{i=1}^5 p_i \mathbb{E}_i[F_1]}{\sum_{i=1}^5 p_i \mathbb{E}_i[R_1]} = \mathbf{0.7817}.$$

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References

- [1] ABATE, J. AND W. WHITT. (1995). Numerical inversion of Laplace transforms of probability distributions. *ORSA Journal on Computing*, **7**, 36-43.
- [2] ÇINLAR, E. (1977). Shock and wear models and Markov additive processes, in: I.N. Shimi, C.P. Tsokos, *The Theory and Applications of Reliability*, Academic Press, New York, 1977, 193-214.
- [3] ESARY, J., W. MARSHALL, AND F. PROSCHAN. (1973). Shock models and wear processes. *The Annals of Probability*, **1**, 627-649.

- [4] IGAKI, N., U. SUMITA, AND M. KOWADA. (1995). Analysis of Markov renewal shock models. *Journal of Applied Probability*, **32**, 821-831.
- [5] KHAROUFEH, J. (2003). Explicit results for wear processes in a Markovian environment. *Operations Research Letters*, **31**, 237-244.
- [6] KIESSLER, P., KLUTKE, G. AND YANG, Y. (2002). Availability of periodically inspected systems subject to Markovian degradation. *Journal of Applied Probability*, **39**, 700-711.
- [7] KLUTKE, G., WORTMAN, M. AND AYHAN, H. (1996). The availability of inspected systems subject to random deterioration *Probability in the Engineering and Informational Sciences*, **10**, 109-118.
- [8] KLUTKE, G. AND YANG, Y. (2002). The availability of inspected systems subject to shocks and graceful degradation. *IEEE Transactions on Reliability*, **51**, 371-374.
- [9] NAKAGAWA, T. (1979). Replacement problem of a parallel system in a random environment. *Journal of Applied Probability*, **16**, 203-205.
- [10] RÅDE, J. (1976). Reliability in a random environment. *Journal of Applied Probability*, **13**, 407-410.
- [11] SHANTHIKUMAR, J. AND SUMITA, U. (1983). General shock models associated with correlated renewal sequences. *Advances in Applied Probability*, **17**, 600-614.
- [12] SKOULAKIS, G. (2000). A general model for a reliability system. *Journal of Applied Probability*, **37**, 925-935.
- [13] WORTMAN, M., AND KLUTKE, G. (1994). On maintained systems operating in a random environment. *Journal of Applied Probability*. **31**, 589-594.