

**Aleatory Uncertainty and Scale Effects in Computational  
Damage Models for Failure and Fragmentation**

by OE Strack, RB Leavy, and RM Brannon

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## **Aleatory Uncertainty and Scale Effects in Computational Damage Models for Failure and Fragmentation**

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# Aleatory uncertainty and scale effects in computational damage models for failure and fragmentation

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## SUMMARY

Stress concentrations near grain boundaries, precipitates, and similar micro-heterogeneities nucleate instabilities leading to macroscale fracture. As it is not practical to model each flaw explicitly, their ensemble effect is modeled statistically. Accounting for this aleatory uncertainty requires smaller specimens (e.g., small finite elements) to have generally higher and more variable strengths, which is necessary for the initial failure probability of a finite domain to be unaffected by its discretization into elements. Localization itself, which might be attributed to constitutive instability, requires realistic numerical perturbations to predict bifurcations such as radial cracking in axisymmetric problems. These perturbations, stemming from microscale heterogeneity, are incorporated in simulations by imposing statistical spatial variability in the parameters of an otherwise conventional (deterministic and scale-independent) damage model. This approach is attractive for its algorithmic simplicity and straightforward calibration from standard strength tests. In addition, it results in virtually no loss of efficiency or robustness relative to deterministic models and accommodates general three-dimensional loading. Despite these advantages, some significant challenges remain and are discussed. However, it is demonstrated that including aleatory uncertainty with associated scale effects significantly improves predictiveness on large-scale computational domains, where it is impractical to resolve each crack or localization zone. Copyright © 2014 John Wiley & Sons, Ltd.

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**KEY WORDS:** constitutive equations; damage; fracture and fragmentation; fuzzy probabilistic methods; impact; rate dependence; scale/size effects; mesh dependence; verification; validation

## 1. INTRODUCTION

This paper has been submitted as part of a special issue in honor of Professor Ted Belytschko, whose contributions are pervasive in the finite-element community. As an example, the production finite-element code that we use in this paper [1] is based on his seminal publications on hexahedral elements [2, 3]. Among his numerous contributions in the area of fracture mechanics, the extended FEM [4] has enabled accurate modeling of discontinuities in finite-element meshes (often introduced by cracks). As computational resources continue to evolve, methods such as extended FEM might ultimately be used ubiquitously to model failure. Meanwhile, damage models, such as the one summarized in this paper, remain useful for modeling ballistic impact and similar high-energy events. We have selected this topic for this special issue because of Professor Belytschko's continual support of our efforts to include strength statistics and scale effects in engineering damage simulations.

This paper describes a practical method for extending a computational framework to accommodate aleatory uncertainty and associated size effects (e.g., spatial variability in strength and other material properties) without prohibitively increasing computational overhead or model development

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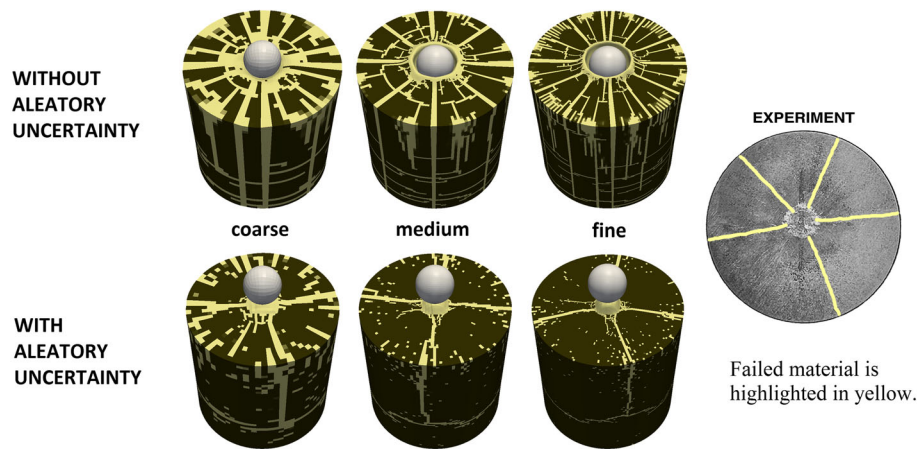


Figure 1. Tungsten-carbide (WC) sphere impacting a silicon carbide (SiC-N) ceramic at 375 m/s. All simulations use the same constitutive model on the same three mesh resolutions, but those in the bottom row are initialized using experimental data for aleatory uncertainty and associated size effects, giving relatively mesh-insensitive radial cracking patterns similar to those in the (truly representative) experiment [7].

time relative to “conventional” (deterministic and scale-independent) damage models. As illustrated in Figure 1, a conventional model’s predictions can be made more realistic and less mesh sensitive when it is run using properties that include statistical variability and scale effects that are evident in standard laboratory strength testing. Heterogeneity is applied to an existing conventional smeared-damage model at the host-code level so that few (if any) alterations to the underlying material model’s source code are needed. Merits and drawbacks of the conventional damage model used for this study [5, 6] receive only slight attention in this paper, whereas the physical foundations, algorithms, and calibration procedures for uncertainty and size effect are covered in detail.

Aleatory uncertainty refers to natural spatial variation in flaw or inclusion morphology that causes local regions (at a scale far smaller than a finite element<sup>‡</sup>) to be weaker or stronger than surrounding material. Because it is impractical to precisely measure and model each flaw size and orientation, certain material properties such as strength are distributed statistically in a manner consistent with macroscale observations for a variety of specimen sizes. Slight spatial variability, which is relatively inconsequential in stable deformations [9, 10], activates localized zones of damage (e.g., shear bands or fractures) as the intensity of loading approaches an unstable state.

Conventional plasticity and damage models have long been known to admit bifurcation to a localized state (when, e.g., the acoustic tensor has a zero eigenvalue; cf. [12, 13]). In the laboratory, however, damage begins at stress levels well below this mathematical stability limit [14, 15]. The situation is analogous to buckling: a column will buckle at a load well below the theoretical limit load if its heterogeneities (holes and pits) are large enough to induce non-infinitesimal perturbations in the stress field. If mesh morphology, material properties, and loading in a simulation are perfectly homogeneous, then conventional deterministic damage models (including those with bifurcation and decohesion criteria) are mathematically incapable of predicting heterogeneous fragmentation of the type depicted in Figure 2. Without perturbations, an idealized simulation of homogeneous loading (i.e., run at “infinite precision”) would reach a stability limit at the same moment for all elements, producing unrealistic failure in every element in a computational domain, which is an undesired material response that has thwarted numerous attempts to simulate dynamic failure of armor ceramics [16, 17]. Because constitutive instabilities appear to be extremely sensitive to the nature of perturbations, realistic perturbation fields must be used rather than relying on spurious mesh-dependent and platform-dependent disturbances such as numerical round-off (cf. [18, 19]).

<sup>‡</sup>Simulations presented in this paper were run using a production-quality finite-element code [1], but the methods apply equally well to other PDE solvers. The same techniques have, for example, been confirmed to give similar results when finite elements are reinterpreted as particle domains in a material-point-method code [8].

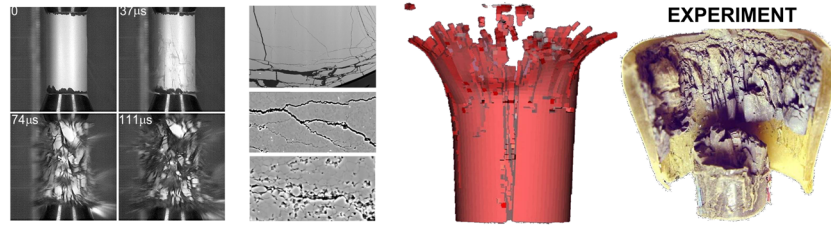


Figure 2. Heterogeneous dynamic response to nominally homogeneous (and initially quasistatic) unconfined compression of a ceramic (SiC-N). The simulation uses aleatory uncertainty in strength (i.e., macroscale statistical variability attributed to micro-heterogeneity) to capture the tendency of these experiments to fail at only one end, typically leaving an intact fragment at the other end [11].

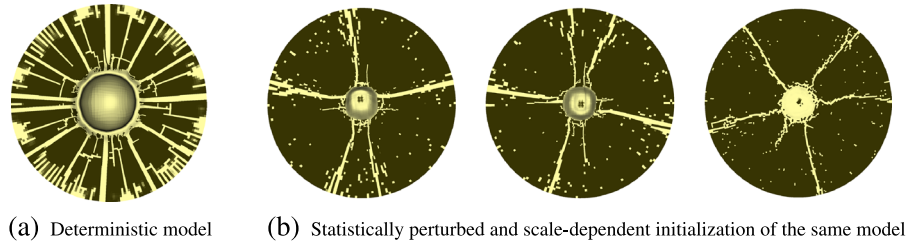


Figure 3. (a) Spurious mesh-texture-biased radial cracking in a conventional damage model compared with (b) realistic randomly oriented cracking (with statistically reproducible spacing) with aleatory uncertainty. For a SiC-N cylinder impacted by a WC sphere at 375 m/s, the first two simulations in sub-figure (b) use the same mesh as sub-figure (a)—each with a different random seed, whereas the third uses a “paved” mesh.

A damage model is successful if it is able to reproduce (without mesh sensitivity) observed scale-dependent and rate-dependent statistical bifurcation into damaged and intact zones with realistic expended fragmentation energy.

In a systematic investigation of correlations between standard material properties (such as hardness, grain size, porosity, etc.) and ballistic performance of a variety of SiC-N ceramics, Ray *et al.* [20] found that “a wide variety of SiC-based materials can give good ballistic results, contradicting some of the theories about what is important to improve ceramics for armor.” Their investigation revealed a possible correlation between the Weibull modulus and ballistic performance, but they were quick to note that this result might have been coincidental because there was also no correlation to the weak tail of the strength distribution. If these conclusions are independently validated through further testing, then variability itself (not the value of the lowest strength) might be the key factor in ballistic performance, suggesting that the salient numerical requirement is to realistically seed an instability, perhaps of the nature discussed by Grinfeld [21].

For axisymmetric loading, such as indentation of a macroscopically homogeneous cylinder, symmetry-breaking radial cracks originate from micromorphological instability [21]. Just as a mathematical perturbation-free analysis of column buckling produces unrealistic simultaneous failure at every point in the column, a perturbation-free mathematical analysis of axisymmetric loading produces simultaneous failure of all points equidistant from the symmetry axis or excessive cracking in a deterministic simulation such as Figure 3(a). Round-off noise represents a numerically infinitesimal perturbation *many* orders of magnitude smaller than the perturbations evident in experimental data. As illustrated in Figure 3, measured statistics in strength should be used to establish heterogeneous material property fields of statistically homogeneous media. When simulations rely exclusively on round-off error to stimulate instabilities, the results might break symmetry without explicit control of perturbations, but such predictions have no physical basis, as they are typically caused by the following: (1) mesh-dependent or platform-dependent differences in the fields or (2) order of operation errors in the code (where, for example the stress field at certain locations is updated at the wrong time). In Figure 3(b), realistic random crack orientations exhibit reduced

texture bias on the radial-trisection meshes. Timmel [22] notes that other mesh types might promote realistic failure patterns, but algorithms that minimize this mesh-texture bias are preferable. The “paved” mesh in the rightmost image in Figure 3(b), for example, encourages slight curvature in the crack paths without any other pronounced differences in comparison with the radial trisection results.

In addition to needing realistic statistical strength perturbations in simulations, an equally important feature is the well-known concomitant dependence of strength on specimen size [23]. As material failure *tends* to originate at the weakest point in a material, large samples will be weaker (on average) because they are more likely to contain a critically large or critically oriented flaw. Because the location of the weakest points in a material may not coincide with the points of highest stress, localized material failure does not generally originate at the highest stress concentrations, nor does it necessarily precipitate cascading formation of macrocracks. With aleatory uncertainty in simulations, non-cascading local damage (i.e., non-weakest-link behavior) is evident in the “speckled” failure points in Figure 3(b) (which are very similar to those observed in computed tomography data; cf. [24], which likewise do not all coalesce into macrocracks).

Using statistics in constitutive modeling dates back to the 19th century, where work was driven by geomaterials because of their heterogeneity even at the macroscale. Statistics have been also used extensively to predict concrete fracture [25] and munitions fragmentation or “hot-spot” formation in energetic media [26–28]. Convergence of the statistical distribution of fragments generated by a uniformly expanding ring was demonstrated by including Weibull variability of the yield strength (without a size effect) [29]. Weibull statistics and size effects applied to cohesive zone models improve their numerical convergence and predictive properties [30], without necessarily implying that the overall strength is itself Weibull distributed.

In Section 2, we summarize a recent study [31] showing that strength distributions strongly depend on the tensorial character of the loading mode, which leads to the mathematically rigorous concept in Section 3 of “quantile” (isoprobability) surfaces that generalize the limit functions of deterministic damage theories. Stress dependence of failure is characterized in Section 4 and Section 5 from examination of experimental data, which suggest that both scale effects and aleatory uncertainty decrease with increasing pressure. These observations motivate using a Weibull strength distribution *only for spherical (isotropic) tension*, whereas non-Weibull distributions (as well as decreased statistical variability with pressure) are naturally predicted for states having a nonzero stress deviator. Data delocalization methods accounting for non-uniform stress fields in laboratory testing are discussed in Section 6, along with recent work on so-called data relocalization needed in simulations to compensate for the inability of low-order shape functions to describe stress variation on an element [32]. The constitutive damage model used in our case studies is briefly discussed in Section 7 (though our approach applies to any damage model). Finally, although failure initiation receives the most attention in this paper, Section 8 emphasizes the need for including scale-dependent damage rate effects during the failure progression phase of the simulation. Failure initiation (e.g., the first growth of a microcrack leading to initiation of stiffness degradation) is herein treated as a weakest link phenomenon, but there is no assumption that initiation of failure will necessarily progress to a complete loss of stiffness or strength. Thus, overall structural failure is not generally a weakest link event.

## 2. EXTENSION OF WEIBULL THEORY FOR GENERALIZED STRESS STATES

Weakest link theory (which we do not assume applies to all loading modes) is governed by Weibull strength distributions, leading to Weibull scale effects in which large specimens (and hence large finite elements) are weaker, on average, than small specimens (or small finite elements). For general loading trajectories, our phenomenological extension of Weibull theory will be shown to exhibit realistic non-Weibull statistical strength distributions and non-Weibull scale effects.

### 2.1. Failure theories and implications for specimen size dependence

If it were possible to create a sample containing exactly one small flaw, then onset of failure under a given tensor stress state  $\mathbf{T}$  (depicted in Figure 4 as a 3D Mohr’s circle) relates to the likelihood that

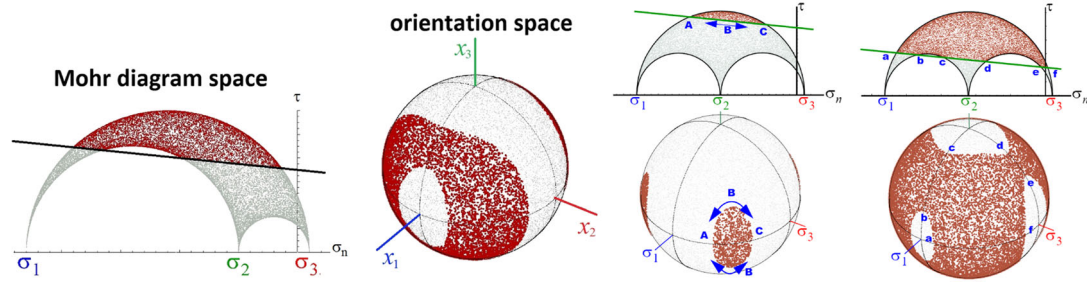


Figure 4. Uncertainty in a single crack's orientation (quantified by its unit normal, which is a point on the unit sphere) allows the stress state (Mohr's circle) to fall outside a failure line, depicted here as a straight line corresponding to a criterion that the crack will fail if the resolved shear stress (minus friction) reaches a critical value. Statistically variable crack sizes and failure modes perturb the failure line itself (cf. [31]).

the flaw is critically oriented or critically large. Figure 4, for example, shows that even a simplistic Mohr–Coulomb failure criterion has uncertainty of failure if the crack orientation is random [31].

Similar observations have been made using more sophisticated single-crack failure theories [33], and the same concepts have been applied to account for variable slip-plane orientations in plasticity [34]. Some flaw orientations are safe from failure, whereas others (dark regions on the spheres in Figure 4) are not. Brannon and Gowen [31] limit their scope to failure initiation in a statistically isotropic medium (for which flaw orientation is initially uniformly random so that the initial failure probability is the ratio of solid angle of the zone of critical orientations to the area of the sphere). Using analytical methods, they derive failure statistics (probabilities that samples will fail under specific loading conditions) from uncertainties in flaw orientations, and they illustrate the corresponding concept of strength quantile surfaces that generalize the notion of a limit failure surface used in classical damage theories.

Dienes *et al.* [28] and Nemeth [35] accommodate intrinsic and deformation-induced anisotropy by allowing flaw size distributions to be correlated to flaw orientation either from the outset or through oriented crack growth. Macon *et al.* [36] further point out that induced anisotropy (leading to what Nemeth [35] refers to as “equatorial” flaw distributions) may represent a way to model dilatation in triaxial compression in a way that avoids anomalous predictions of hydrostatic strengthening under dilatation seen in some isotropic hardening models (e.g., [37]). Rather than further expanding on the physical motivations for these so-called unit-sphere representations of failure statistics, we focus on actual implementations of these theories in engineering simulations.

## 2.2. Weibull theory (and its extension) for generalized stress states

Consider a stress state  $\mathbf{T} = \sigma \mathbf{N}$ , where  $\sigma$  is the magnitude of stress and  $\mathbf{N}$  is a unit tensor called the stress trajectory or loading mode. For example, uniaxial stress has  $\mathbf{N} = \text{DIAG}[1, 0, 0]$ , whereas spherical tension has  $\mathbf{N} = \text{DIAG}[1, 1, 1]/\sqrt{3}$  and biaxial tension has  $\mathbf{N} = \text{DIAG}[1, 1, 0]/\sqrt{2}$ . For simplicity, our discussions might presume a straight path through stress space (where  $\mathbf{N}$  is constant), but the concept of strength quantile surfaces in six-dimensional (6D) stress space [31] admits more tortuous tensorial paths through stress space in which the scalar  $\sigma$  (called strength) is merely the value of the path parameter at failure. Just as deterministic plasticity yield surfaces generally evolve with plastic flow, strength quantile surfaces are also expected to change in a history-dependent manner so that probabilistic strength,  $\sigma$ , generally depends implicitly on evolving internal state variables such as crack density in addition to the loading mode. Let  $\bar{\sigma}$  be the *median* strength observed when testing a sample of volume  $\bar{V}$  under a given loading mode  $\mathbf{N}$ . To account for microscale heterogeneities (e.g., stress concentrations at flaws) and to accommodate associated scale effects, a theory is required for predicting the probability,  $P^{\text{fail}}(\mathbf{T}, V)$ , that a sample of a volume  $V$  (generally distinct from  $\bar{V}$ ) will begin failing under an applied stress  $\mathbf{T}$ . The  $P^{\text{fail}}(\mathbf{T}, V)$  function is essentially a cumulative distribution function (CDF) in 6D stress space, so it is a Heaviside step function in a deterministic theory (i.e., its gradient is singular at the limit surface). Otherwise, for non-deterministic theories, the directed derivative along a stress trajectory gives the probability

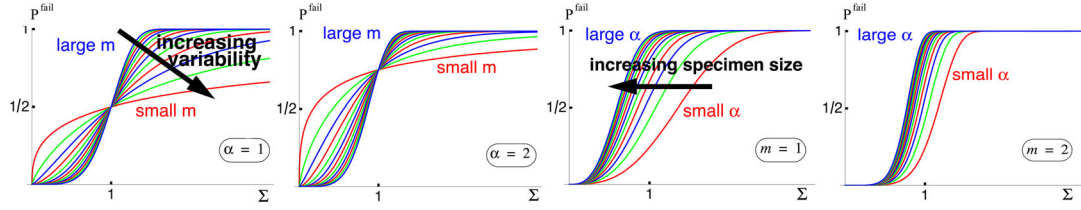


Figure 5. CDFs for normalized failure stress  $\Sigma$  for various normalized sample volumes,  $\alpha$ , and various values of the Weibull modulus  $m$ .

density function (PDF) for trajectory-dependent strength as  $p(\sigma, V) \equiv (\partial P^{\text{fail}} / \partial \mathbf{T}) : \mathbf{N}$ , which is the key observation needed to construct the physically meaningful tensor-based function,  $P^{\text{fail}}(\mathbf{T}, V)$ , whose level sets are dubbed as “aleatory quantile surfaces”.

For a fixed stress trajectory  $\mathbf{N}$ , define non-dimensionalized stress and size ratios:

$$\Sigma = \frac{\sigma}{\bar{\sigma}} \quad \text{and} \quad \alpha = \frac{V}{\bar{V}}. \quad (1)$$

Weibull theory (which we soon argue does not apply to general loading modes but is nevertheless useful for illustrating concepts in a familiar context) predicts that a sample of non-dimensionalized volume  $\alpha$  will begin failing under non-dimensionalized stress  $\Sigma$  with probability

$$P^{\text{fail}} = 1 - P^{\text{safe}} = 1 - 2^{-\alpha \Sigma^m}, \quad (2)$$

which is the initial CDF for strength, in which  $m$  is a constant called the Weibull modulus (or shape parameter). The *complementary* CDF  $P^{\text{safe}}(\mathbf{T}, V)$  represents the probability that a sample of volume  $V$  will *not* have begun failing under an applied stress  $\mathbf{T}$ . Plots of Equation (2) are shown in Figure 5 for various relative sample volumes and Weibull moduli, demonstrating that small values of  $m$  correspond to large variability in strength. A purely deterministic model corresponds to a unit step function in the limit  $m \rightarrow \infty$ . Because the plots for small specimens (small  $\alpha$ ) fall to the right of those for larger specimens, small specimens have, on average, higher strengths. Equivalently, because curves for small specimens fall below those of larger specimens, a small specimen is less likely to fail at any given applied stress. Furthermore, the plots show that small specimens have greater spread in their strengths.

### 2.3. Mode, median, and standard deviation of the Weibull distribution

For the Weibull distribution introduced in Equation (2), the median  $\Sigma^{\text{median}}$  is the 50% quantile, representing the stress level at which half of the specimens in a large population will have initiated failure. Hence,  $\Sigma^{\text{median}}$  is the value of  $\Sigma$  at which  $P^{\text{fail}} = 1/2$ , and referring to Equation (2), it follows that the median strength depends on specimen size according to

$$\Sigma^{\text{median}} = \alpha^{-1/m}. \quad (3)$$

The probability density function (PDF),  $p(\Sigma, \alpha)$  is the derivative of the CDF with respect to  $\Sigma$ :

$$p(\Sigma, \alpha) \equiv \frac{\partial P^{\text{fail}}}{\partial \Sigma} = (\ln 2) \left( 2^{-\alpha \Sigma^m} \right) \alpha m \Sigma^{m-1} \quad (4)$$

The so-called Weibull scale parameter  $\Sigma^{\text{scale}}$ , the mode  $\Sigma^{\text{mode}}$  (at the peak of the PDF), and the mean,  $\Sigma^{\text{mean}}(\alpha) = \int_0^\infty \Sigma p(\Sigma, \alpha) d\Sigma$  are found from

$$\frac{\Sigma^{\text{scale}}}{\Sigma^{\text{median}}} = (\ln 2)^{\frac{1}{m}}, \quad \frac{\Sigma^{\text{mode}}}{\Sigma^{\text{median}}} = \left( \frac{m-1}{m \ln 2} \right)^{\frac{1}{m}}, \quad \text{and} \quad \frac{\Sigma^{\text{mean}}}{\Sigma^{\text{median}}} = \frac{\Gamma \left[ 1 + \frac{1}{m} \right]}{[\ln 2]^{1/m}} \quad (5)$$

where  $\Gamma$  is the gamma function. Variability (spread) in the data may be quantified by the standard deviation,  $\Sigma^{\text{s.d.}}$ , illustrated in Figure 6(a) and defined by

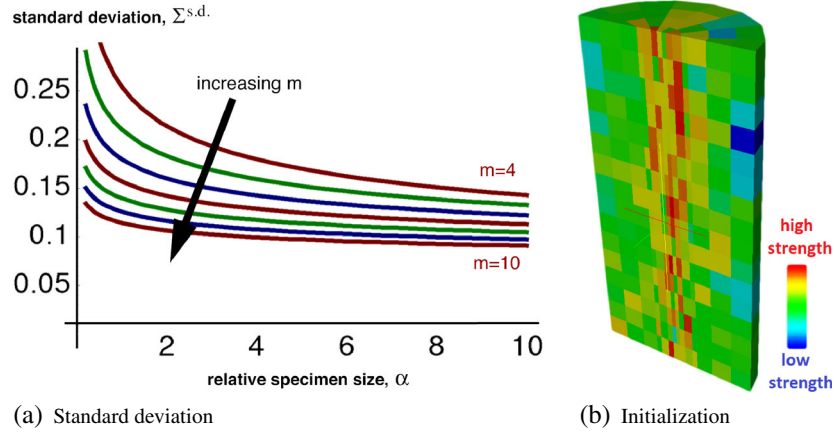


Figure 6. The graphs in (a) imply that, in addition to higher strength, greater variability must apply at smaller elements within a computational domain, such as those near the centerline in (b), giving an unavoidable impression of heterogeneity in what is nevertheless a *statistically homogeneous* strength distribution so that a physical domain's failure initiation probability under homogeneous loading is unaffected by the meshing of that domain.

$$\frac{\Sigma^{s.d.}}{\Sigma^{median}} = \sqrt{\left(\Gamma\left[\frac{2+m}{m}\right]\right) - \left(\Gamma\left[1 + \frac{1}{m}\right]\right)^2} \approx \frac{\pi}{\sqrt{6m}} \text{ for large } m \quad (6)$$

The plot of the standard deviation  $\Sigma^{s.d.}$  in Figure 6(a) illustrates our earlier assertion that one should not rely on round-off error to initiate mathematical instabilities. Even for the largest Weibull modulus in Figure 6(a), the standard deviation in strength can be on the order of 15%. In fact, for the left-hand side of Equation (6) to equal a typical round-off error of  $10^{-16}$ , the Weibull modulus would need to be more than 10 orders of magnitude larger than reported for any polycrystalline material to our knowledge. Hence, perturbations that precipitate instabilities or bifurcations associated with localization in a numerical code must be realistically seeded in order to obtain realistic results. A deterministic model unrealistically has no variability (zero standard deviation) in strength.

Hild *et al.* [38] identify a characteristic sample (element) size marking a statistical-deterministic transition such that larger elements are treated deterministically while variability is applied for smaller elements. That approach admits nonphysical jump discontinuities in the standard deviation, whereas Figure 6(a) shows that seeding all elements with a Weibull distribution results in a smooth transition toward a deterministic response (zero standard deviation) as the element size increases. Both higher strength and greater variability are evident in the time-zero initialization of a statistically homogeneous material in the computational domain of Figure 6(b), which was generated using concepts developed in the next section to accommodate smooth, trajectory-dependent, and generally non-Weibull seeding of strength for 3D stress states.

Failure probability functions are valuable in engineering if one seeks a probability of failure initiation for the purpose of, say, component reliability and safety assessment [39, 40]. Theories are also available for expected values of fragment sizes in impact and shock loading after failure initiation [41, 42]. However, if one is interested in detailed analysis of the dynamic effects of the failure process, then realizations of strength, consistent with the  $P^{safe}(\mathbf{T}, V)$  function, are required to generate realistic crack spacings and fragment sizes [43]. Simulating realistic damage and fragmentation events is especially important for assessing degradation of strength that could lead to particular vulnerabilities to subsequent impact events. The next section of this paper recognizes and accounts for the existence of a  $P^{safe}(\mathbf{T}, V)$  function without requiring that it ever be explicitly constructed in simulations. Accordingly, we will now focus on the goal of converting failure probabilities into strength realizations on a computational domain.

## 3. “FUZZY FAILURE” THEORY AND STRENGTH REALIZATIONS

Classical deterministic plasticity models presume existence of a scalar-valued yield function,  $f(\mathbf{T}, V)$ , such that a stress state  $\mathbf{T}$  applied to a specimen of volume  $V$  is elastic if  $f(\mathbf{T}, V) < 0$ . The yield surface is the set of  $\mathbf{T}$  for which  $f(\mathbf{T}, V) = 0$ . Scale-insensitive models omit dependence on  $V$ . The yield function also implicitly depends on internal variables, such as dislocation density and crack texture parameters, which can evolve to allow the yield surface to expand (harden) or contract (soften) in response to plastic loading. The yield surface bounds the domain of *elastically* attainable states; engineering damage theories also use a so-called limit function  $F(\mathbf{T}, V)$  that defines the boundary of stress states attainable by *any* quasistatic means [44]. A quasi-brittle model employs distinct yield and limit functions ( $f \neq F$ ), whereas a purely brittle model uses coincident functions ( $f = F$ ). Damage corresponds to collapse of the limit surface and concomitant degradation of elastic stiffness.

These conventional sharp-threshold plasticity models have required considerable time and resources to develop. Many (such as [45] and [46]) are also equipped with large material property databases, compiled over many years. By presuming that existing discrete threshold functions are actually fitted only to median experimental data, so-called “fuzzy failure theory” incorporates spatial variability and scale effects in these models. Moreover, because of the size effect, these fits apply only to the specimen size used in the calibration experiments, where specimen size is defined using “data-delocalization” methods as explained in Section 6.

## 3.1. Aleatory uncertainty theory

Conceptually, spatial variability is incorporated into these conventional plasticity models by replacing a limit function  $F(\mathbf{T}, V)$  by a scale-dependent probability field  $P^{\text{safe}}(\mathbf{T}, V)$  that ranges continuously from a value of 1 at stress states that are safe from failure to a value of 0 at stress states that are certain to induce failure. The conventional limit function could, for example, be identified with the 50% quantile isosurface:  $F(\mathbf{T}, V) = P^{\text{safe}}(\mathbf{T}, V) - 1/2$ . Of course, similar revision of the yield function applies to quasi-brittle models. Strength is statistically variable because of aleatory uncertainty in the micromorphology (such as crack densities and textures). Recalling that  $F(\mathbf{T}, V)$  depends implicitly on internal variables that directly or indirectly account for micromorphology, it follows that a natural way to indirectly incorporate the  $P^{\text{safe}}(\mathbf{T}, V)$  in a simulation is to retain an existing code infrastructure for a scale-insensitive  $F(\mathbf{T})$ , but to initialize each finite element with different  $F(\mathbf{T})$  functions by statistically perturbing the parameters defining  $F(\mathbf{T})$  in a way that recovers observed strength probabilities captured in  $P^{\text{safe}}(\mathbf{T}, V)$ , where  $V$  is taken as the finite-element volume. As this initialization phase of the simulation is the only point at which  $P^{\text{safe}}(\mathbf{T}, V)$  is needed, there is no algorithmic burden to otherwise incorporate this function into run-time calculations, making the implementation of aleatory uncertainty quite manageable.

The  $P^{\text{safe}}(\mathbf{T}, V)$  function allows the effects of micromorphology to be treated *in ensemble* through empirical strength CDFs. With this engineering approach, the median (which has been presumably already fit to laboratory data) is retained but scaled appropriately with specimen (finite-element) size. Statistically perturbed realizations of strength surfaces in stress space are centered about the scale-adjusted median (i.e., 50% quantile) limit surface according to the statistics evident in laboratory data (and consistent with limit surface admissibility requirements such as convexity).

“Fuzziness” in the limit surfaces of conventional plasticity models has been formalized by Brannon and Gowen [31], who introduce the concept of quantile (isoprobability) surfaces in stress space, replacing the single failure surface used in deterministic models by the family of surfaces in Figure 7(a). The red in Figure 7(a) depicts the median (50% quantile) isoprobability surface viewed down the hydrostat, with the 5% and 95% ‘quantile surfaces depicted as blue. The images in Figure 7(b) are similar except that quantile surface colors are replaced with statistical dot densities centered about the median quantile surface. The “weak tail” of a strength distribution refers to the smaller quantile surfaces, which (being closer to the origin in stress space) represent instances of low-stress failure. For the idealized micromorphology used to generate Figure 7(a), the median 50% quantile surface (colored red) is a rounded triangle (modeled using the Willam–Warnke third-

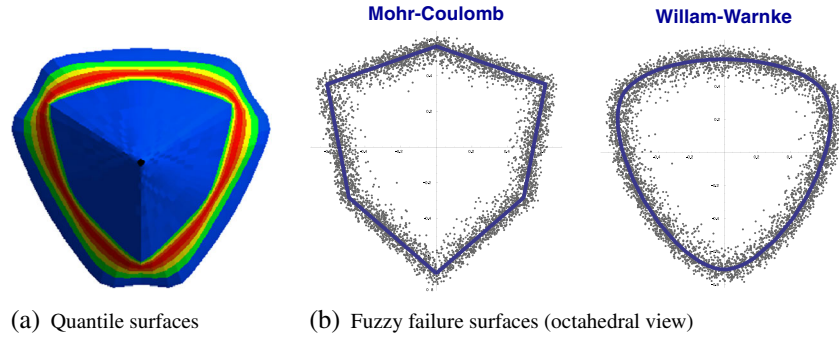


Figure 7. Accounting for variability in flaw size affects the location of the failure threshold itself, producing quantile surfaces in stress space depicted here looking down the hydrostat and visualized according to initial failure probability density by (a) colors showing theoretical shapes of each quantile surface [31] or (b) dot densities that were generated using self-similar quantile shapes (more like the current implementation).

invariant function common in geomechanics; cf. [6]), whereas the weak 5% quantile is the blue distorted hexagon that is easily modeled with Mohr–Coulomb third-invariant dependence. The quantile functions in Figure 7(b) did not include this recent finding of Lode-angle sensitivity.

In our fuzzy failure theory, natural subscale heterogeneities are accommodated in a macroscale simulation by using different realizations of strength in each finite element, such that the median strength of a finite volume is preserved, with a key goal being that failure-initiation statistics ought to be unaffected by mesh refinement. In numerical implementations, each finite element is regarded as a finite-sized specimen requiring assignment of strength consistent with scale effects observed in the laboratory. For illustration purposes, suppose that the stress at failure  $\sigma$  for a specimen size  $\bar{V}$  is known to be Weibull distributed with median  $\bar{\sigma}$  for a given stress trajectory  $\mathbf{N}$  (keep in mind that non-Weibull strengths are still possible for other trajectories). To assign strength properties to a finite element, one must generate a uniformly distributed random number  $R$  on the interval  $0 \leq R < 1$ . The realization of failure stress can then be found by replacing  $P^{\text{safe}}$  with  $R$  in Equation (2) and solving for  $\Sigma$ . The result is

$$\Sigma = \left[ \frac{\ln R}{\alpha \ln(1/2)} \right]^{1/m} \quad \text{i.e.,} \quad \sigma = \bar{\sigma} \left[ \frac{\bar{V} \ln R}{V \ln(1/2)} \right]^{1/m}. \quad (7)$$

For a different sample volume ( $V \neq \bar{V}$ ), this equation predicts a different strength. In our case, the strength  $\sigma$  assigned to a finite element depends on the volume  $V$  of the element relative to the sample volume used in laboratory testing. Equal-sized regions always have the same probability of failure initiation under homogeneous loading regardless of whether or not those regions are subdivided into many or few elements. As a special case, Equation (7) ensures that a collection of small elements having total volume  $\bar{V}$  will have the same failure initiation statistics (under uniform loading) as observed in the original laboratory sample of volume  $\bar{V}$  (non-uniform loading is addressed in Section 6). Equation (7) therefore contains an implicit length scale that ensures mesh-insensitive predictions for the *onset* of failure under homogeneous loading. This is verified numerically for a spherical stress trajectory,  $\mathbf{N} = \text{DIAG}[1, 1, 1]/\sqrt{3}$ , in Section 5 (Figure 14). As discussed in Section 7, mesh insensitivity for the subsequent *progression* of failure is an active area of research in which nonlocal effects and energy release considerations are important for mathematical, numerical, and physical reasons.

Plots of the material strength associated with each realization parameter  $R$  in Equation (7), and strength distribution curves associated with multiple realizations, are shown in Figure 8.

### 3.2. Algorithm for implementing aleatory uncertainty

A computational procedure to support aleatory uncertainty (fuzzy thresholds) in material properties can be designed such that an existing code infrastructure is impacted minimally and so that existing

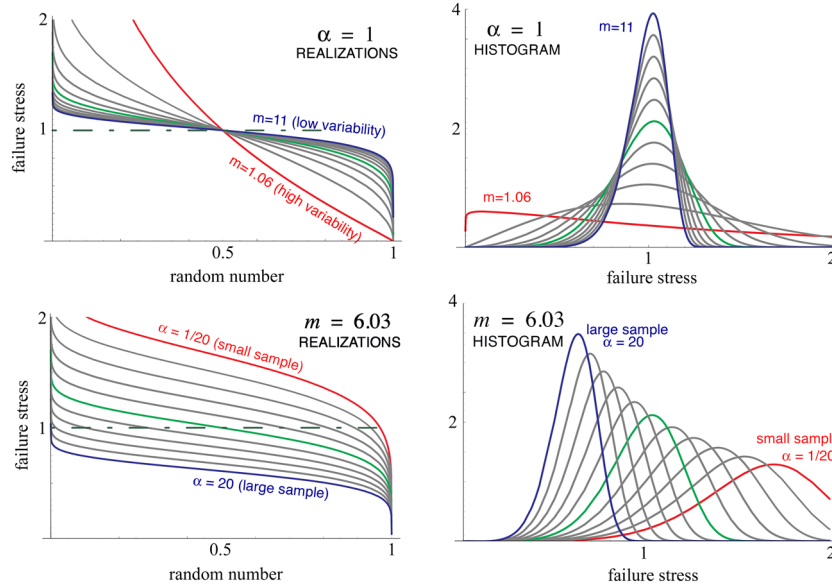


Figure 8. Strength realization functions for various Weibull moduli, given by Equation (7), and probability density functions for failure stress, given by Equation (4).

conventional deterministic material models require no internal revisions of their subroutines. Our approach is outlined below

- (1) Modify the user input parser so that material parameters that were formerly deterministic and specified (for example) by a syntax

$$\text{PROP} = \text{value}$$

are treated statistically whenever the user employs a syntax (for example) of the form

$$\text{PROP} = \text{Weibull}(\text{median}, V_{\text{ref}}, W_{\text{mod}})$$

where *median* is the median measured in the laboratory using a sample of volume  $V_{\text{ref}}$  and  $W_{\text{mod}}$  is the Weibull modulus. Because *median* usually equals *value* from the deterministic model, this approach to aleatory uncertainty requires two new user inputs ( $V_{\text{ref}}$  and  $W_{\text{mod}}$ ), both of which will be available if the original strength measurements were accompanied by specimen size and statistically significant standard deviations (as should be the case if testing followed ASTM and similar standards).

- (2) For each statistical user input, allocate storage for one field variable, giving it a name such as PROPSR (where “PROP” is the property name and “SR” stands for statistical realization).
- (3) Within an initialization loop over each finite element, compute a uniformly distributed random number  $R$  on  $[0, 1)$ , and use the finite-element volume  $V$  in Equation (7) (or any other distribution of interest) to assign a realization value “ $\sigma$ ” for the statistical property. Save the result in the PROPSR field variable array.
- (4) During each time step, before calling the constitutive model, construct a copy of the material property array in which PROP is replaced by PROPSR, and send this revised property array to the material model.

Virtually all of the CPU overhead resides in step 3, which is a once-only initialization setup task typically requiring only a tiny fraction of overall CPU for the calculation. Because step 4 requires nothing more than a small amount of additional memory storage and an array copy, this approach to aleatory uncertainty entails almost no change in overall CPU costs relative to the deterministic model. Because the constitutive model is treated as a “black box”, the necessary code architecture may be written in a way that supports all constitutive models in the code and not just the damage

model of interest. Users must be cautioned, however, that carelessly or arbitrarily perturbing parameters might violate the model's admissibility conditions (such as the convexity requirements). As discussed in the next section, achieving physically reasonable distributions for material parameters might be facilitated by introducing a change of variables.

#### 4. DIRECTIONALLY-WEIBULL THREE-INVARIANT THEORY

The aleatory quantile function,  $P^{\text{safe}}(\mathbf{T}, V)$ , is the probability that a sample of size  $V$  is safe from failure at an applied stress  $\mathbf{T}$ . This function is used indirectly in implementations of aleatory uncertainty. Specifically, the material parameters implicitly contained in the limit function  $F(\mathbf{T}, V)$  are statistically perturbed, thus giving each finite element its own perturbed limit surface distinct from other elements. If  $F(\mathbf{T}, V)$  depends on only a *single* material parameter (as for von Mises or Tresca theory), then model calibration requires measured strength statistics along only a *single* loading trajectory  $\mathbf{N}$  (such as simple shear). In contrast to calibration, *validation* requires additional data along other trajectories (such as uniaxial compression, biaxial tension, etc.). Such efforts might imply the need for a *three*-parameter limit function, for which calibration requires strength data along *three* non-crossing trajectories, as later discussed in Section 5.2.

Measured strengths are interpreted in the context of aleatory quantile (i.e., constant failure initiation probability) surfaces embedded in full 6D tensorial stress space. These multi-dimensional surfaces are typically visualized as 2D cross-sections of 6D stress space. The “octahedral diagram” in Figure 7, for example, is a visualization as seen looking down the hydrostat in principal stress space, where the median (50%) quantile surface is surrounded by other quantile surfaces to define a “cloud” that replaces the discrete limit surface in deterministic models. The “meridional diagram” in Figure 9 depicts a different 2D slice of 6D stress space to show that equivalent shear strength  $\tau \equiv \sqrt{J_2}$  increases with pressure  $p \equiv -I_1/3$ , where  $I_1$  and  $J_2$  are the first and second invariants of the stress tensor  $\mathbf{T}$ , defined  $I_1 = \text{tr}\mathbf{T}$  and  $J_2 = \frac{1}{2}\text{tr}(\text{dev}\mathbf{T})^2$ . Factors of  $\sqrt{2}$  and  $\sqrt{3}$  in the axes labels make the plot isomorphic to stress space (ensuring that the 6D unit tensor  $\mathbf{N}$  maps to a 2D unit vector that also forms the correct angle with the hydrostat). In upcoming plots, isomorphic scaling factors are omitted whenever geometric accuracy is not needed. Quantile surfaces in Figure 9 are depicted as a gray-scale “cloud”. The red line is the median (50% quantile), taken from the limit surface in a deterministic damage model. The 5% and 95% quantiles are lighter-gray contours.

The loading increment direction is typically fixed in model calibration experiments, meaning that the stress tensor is of the form  $\mathbf{T} = \sigma\mathbf{M} + \mathbf{C}$ , where  $\sigma$  is a time varying load intensity parameter,  $\mathbf{M}$  is a unit tensor in the loading increment direction, and  $\mathbf{C}$  is a constant tensor that is usually an elastic prestress reference state. If, for example, the experiment involves a hydrostatic compression to a pressure  $p$  followed by axisymmetric compression with lateral stress held constant at  $p$ , then

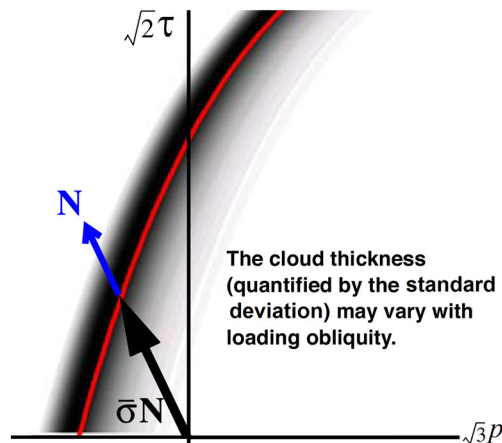


Figure 9. Loading trajectory through a fuzzy failure “cloud”.

$\mathbf{C} = -\text{DIAG}[p, p, p]$  and  $\mathbf{M} = \text{DIAG}[-1, 0, 0]$ . If  $\mathbf{C} = \mathbf{0}$ ,  $\mathbf{M}$  is the same as  $\mathbf{N}$  defined in Section 2.2 and sketched in Figure 9. For these experiments, strength distributions are actually distributions for the scalar loading parameter  $\sigma$ , and the quantiles for  $\sigma$  are then simply mapped to full 6D stress space to generate points for fitting the corresponding quantile surfaces. If  $\bar{\sigma}$  is the median, for example, then  $\mathbf{T} = \bar{\sigma}\mathbf{M} + \mathbf{C}$  must be a point on the median quantile surface. Points on other quantile surfaces in 6D stress space are similarly identified. The process of finding stress tensors on quantile surfaces must be repeated for at least as many distinct loading paths as there are distinct material parameters in the limit function  $F(\mathbf{T}, V)$ , thus providing a means of setting all material parameters corresponding to any quantile realization of strength. The strength  $\sigma$  in Equation (7) does not necessarily represent a single component of stress or even any particular stress invariant. It is merely a scalar parameter in the range from 0 to  $\infty$ , whose distribution  $p(\sigma, V)$  indirectly quantifies variation of the tensor-based CDF  $P^{\text{fail}}(\mathbf{T}, V)$  along a fixed loading trajectory  $\mathbf{M}$ . Specifically, the generally meaningless scalar  $\sigma$  is mapped to the physically significant stress tensor,  $\mathbf{T} = \sigma\mathbf{M} + \mathbf{C}$ . The scalar-based PDF  $p(\sigma, V)$  then provides information about the tensor-based CDF  $P^{\text{fail}}(\mathbf{T}, V)$  through the directed derivative relationship,  $p(\sigma, V) = (\partial P^{\text{fail}}/\partial \mathbf{T}) : \mathbf{M}$ .

It should be clear from examination of Figure 9 that the absolute spread in failure strength data (quantified by variance of  $\sigma$ ) will be affected by the angle of attack in stress space. More broadly, each different loading direction  $\mathbf{M}$  (i.e., each test, such as uniaxial tension, compression, shear, *etc.*) is associated with a different distribution of strength. They are connected into a unified theory by fitting stress *tensors* (not  $\sigma$ ) on each quantile surface to the limit function  $F(\mathbf{T}, V)$  of a “conventional” deterministic damage model.

For our own baseline “conventional” deterministic damage constitutive theory, we created a new model, called Kayenta [6], which revised an existing geomaterial model [47] to incorporate loss of strength in a manner similar to the Johnson–Holmquist class of ceramic models [45, 48–50]. Kayenta is a generalized isotropic damage model (as opposed to an anisotropic model, e.g., [28], or a combined plasticity/linear-fracture-mechanics model, e.g., [51]). Similar to the strain-based Johnson–Cook fracture [46] and Johnson–Holmquist models, Kayenta allows equivalent shear strength  $\tau \equiv \sqrt{J_2}$  in axisymmetric compression to depend on pressure  $p \equiv -I_1/3$  according to

$$\tau = a_1 - a_3 e^{-a_2 p} \quad (8)$$

where  $\{a_1, a_2, a_3\}$  are empirical fitting parameters (an additional term,  $+a_4 p$ , is also available for linear Drucker–Prager (and similar) functions but is omitted here for simplicity).

The Johnson–Holmquist and Johnson–Cook models fall into the category of nonlinear Drucker–Prager generalized plasticity (i.e., their octahedral profile is always a circle). Kayenta incorporates the third invariant (the Lode angle  $\theta$ ) by multiplying  $\tau$  by a function of  $\theta$  and  $p$  that equals unity in axisymmetric compression but is larger than unity at other Lode angles (giving non-circular octahedral profiles such as those in Figure 7). In Kayenta, the Lode angle multiplier function also includes pressure dependence that allows the octahedral profile to vary with pressure, smoothly (and automatically) transitioning from a maximum principal stress (or strain) criterion at low pressures to a von Mises criterion at high pressures. A cap for porous media is also available as an option but is not used in the sample calculations herein.

The limit surface defined by Equation (8) *could* be perturbed by changing one or more of the  $\{a_1, a_2, a_3\}$  parameters. However, these parameters ought not to be varied independently. The user must ensure that PDFs used to perturb material properties will not violate any fundamental physical constraints. Moreover, the user must recognize that achieving an observed Weibull (or other) distribution in strength might require complicated non-Weibull joint distribution functions for the model parameters. For example, the constraint  $a_1 > a_3$  must be satisfied to ensure positive shear strength at zero pressure. Other considerations, such as convexity of the limit surface, place additional coupled constraints on the  $\{a_1, a_2, a_3\}$  parameters. None of these strength-defining parameters are candidates for independent Weibull (or similar) perturbation. To address this inconvenience, we re-parameterized Equation (8) into the form

$$\tau = Y \left[ 1 - e^{-s(p+h)/Y} \right], \quad (9)$$

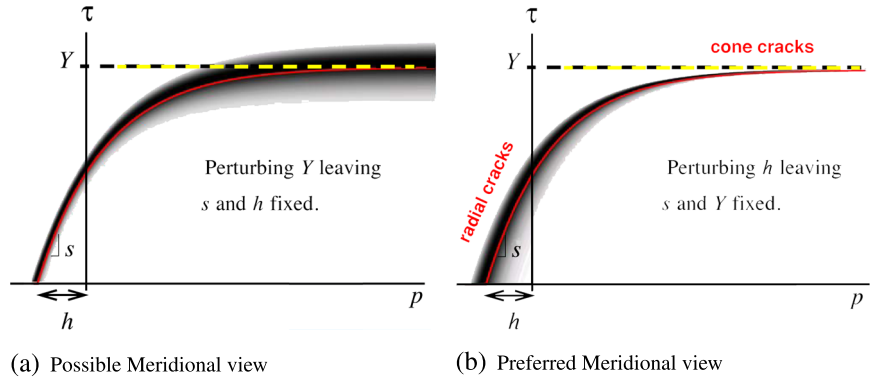


Figure 10. Meridional profiles (including probability clouds depicting variability) corresponding to (a) shear dominant variability obtained by perturbing  $Y$ , which failed to match observed damage patterns in ceramic fragmentation simulations, and (b) pressure-dominant variability obtained by perturbing  $h$ , which is consistent with both spall and Hugoniot elastic limit data and gives more realistic damage patterns in simulations.

where  $Y$  is the shear strength limit,  $h$  is the spherical tensile strength, and  $s$  is the initial slope. The meanings of these alternative parameters are labeled in Figure 10, which also shows probability density clouds resulting from statistically perturbing  $Y$  or  $h$  individually.

The Weibull modulus for SiC was reported to be 17 by the manufacturer Coorstek-Vista [52], but spall data [53] for SiC-N seem to have a Weibull modulus of only 4 (i.e., much more variable). Brazilian tensile strengths (presented later in Figure 17) appear to have a Weibull modulus of about 8. Hugoniot data for  $B_4C$  [54] has a Weibull modulus of 60, with similar numbers for SiC. Moreover, it is not even clear that some of these data sets are very well fitted to a Weibull distribution. To match the general character of observed strength data (i.e., lessening of variability with pressure and non-Weibull strength distributions for most loading trajectories), it is proposed that the limit surface be perturbed as sketched in Figure 10(b).

Unlike the original  $\{a_1, a_2, a_3\}$  fitting parameters, there is no coupling of the constraints on Kayenta's alternative limit-surface parameters  $\{Y, h, s\}$ , which only must be positive. Hence, any of them may be Weibull perturbed independently according to Equation (7) (where  $\sigma$  in that equation is replaced with  $Y, h, \text{ or } s$ ). Consistent with a similar statistical uncertainty assumption used to analyze the nearly spherical tension states that exist during spall [55], and also consistent with the tendency (evident in Figure 13) of strength measurements to be more variable at low pressure, we take the  $h$  parameter to be Weibull distributed, whereas the other limit surface parameters ( $Y$  and  $s$ ) are deterministic. This choice leads to non-Weibull statistics and non-Weibull scale effects for all non-spherical loading paths. This approach, which truly accounts for the 6D tensorial nature of stress, differs significantly from related efforts [56] that apply a 1D Weibull distribution to a single stress invariant such as equivalent shear stress, maximum principal stress, and so on. We conjecture that variability decreases with pressure because the additional confinement provides friction that can lock crack faces and hence suppress brittle failure modes in shear. Of course, high-pressure data often comes from high-rate experiments, but we have tentatively ruled out the possibility that variability decreases with increasing loading rate; this choice is motivated by the fact that quasistatic indentation experiments also seem to exhibit pressure-based variability because they tend not to bifurcate into symmetry-breaking radial cracks in regions of high confinement such as immediately beneath an indenter [57]. Further insight into the influence of pressure and loading rate is provided in a recent investigation of rate effects in brittle media [58], which applies micromorphological dimensional analysis to obtain a universal curve which closely fits rate-dependent and statistically variable strength for several different cracked and porous brittle materials (including various rocks, ceramics, and concrete).

In Figure 11(a), the value of the spherical tension limit  $h$  is exaggerated to illustrate the character of perturbation effects. The actual value of  $h$  for brittle ceramics is typically so small in comparison with the high pressure strength  $Y$  that the actual median limit curve might appear to pass through

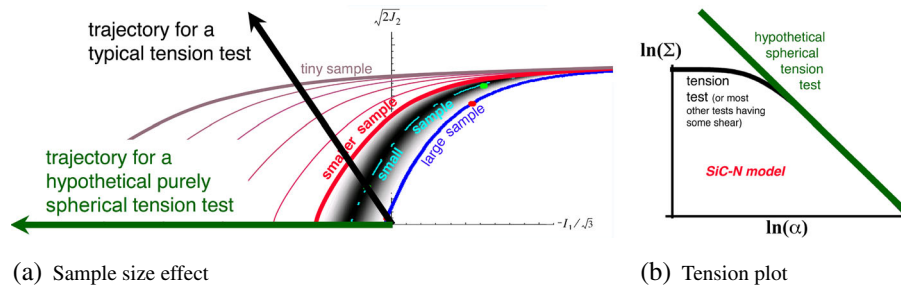


Figure 11. The strength size effect in “directed-Weibull” theory varies with stress space trajectory. Here, strength can increase without bound for progressively smaller samples in *spherical tension* (horizontal green loading line in (a) and straight green line in (b)), but it naturally saturates to a classical plasticity limit in uniaxial tension (black angled loading line in (a) and black curved line in (b)), with clear deviation from Equation (3) for non-spherical loading. This general framework naturally accommodates any other experimentally observed distributions of quantile surfaces, including non-monotonic scaling of strength.

the origin for lab scale ( $\text{cm}^3$ ) samples, as illustrated in Figure 13(a) for a large ( $> 1 \text{ cm}^3$ ) sample of SiC-N. Using Figure 7 to generate strength envelopes means that realizations of the  $\{Y, h, s\}$  parameters are scale dependent. Smaller samples have, on average, higher strength. To illustrate, Figure 13(b) shows the median limit surface for a grain-scale ( $\mu\text{m}^3$ ) sample.

When only the  $h$  parameter is Weibull-perturbed, the median spherical tensile strength approaches infinity as the element size approaches zero. However (as seen in Figure 11), the median uniaxial tension and shear strengths have finite limits, and the strength realizations in these non-spherical directions deviate realistically from a Weibull distribution. This “directionally-Weibull” approach gives a scaling effect similar to that reported by Bažant [23] on the basis of experimental data for concrete.

Bažant has amassed direct experimental data for concrete supporting deviation from the Weibull size dependence similar to what is depicted in Figure 11(b). No similar direct evidence of the deviation from the Weibull size effect has (to our knowledge) been collected for sufficiently small specimens of structural ceramics (whose “aggregate” size is orders of magnitude smaller than that of concrete). We arrived at this effect indirectly by recognizing that macroscale tensile strengths are far more variable than compressive strengths (Figure 13), suggesting that limit surface realizations require perturbations of the type in Figure 10(b).

The nature of the limit surface “variability cloud” and its median master surface strongly influences the appearance of cone or radial cracks in dynamic indentation. In indentation experiments (discussed in Section 8), cone cracks tend to emerge around and under the area of contact where the pressure is extremely high. They are also symmetric [59], indicating a low degree of spatial variability in the high-pressure realm. Radial cracks, on the other hand, tend to appear in relatively low-pressure regions, and they significantly disrupt symmetry, indicating a high degree of spatial variability in strength akin to Figure 10(b). As seen in Figure 12, similar observations have been made for brittle coatings [57, 60], glass [61], and boron carbide plates [62]. The order of the appearance of radial cracks and cone cracks can be altered by changes in materials and indenter geometries.

Mapping out aleatory quantile surfaces is clearly labor intensive. Given that it is often impossible or impractical to obtain statistically significant samples of laboratory-measured strengths, judicious use of virtual data (e.g., [64, 65]) might help to complete the aleatory model parameterization and/or to identify appropriate functional forms of the limit function  $F(\mathbf{T}, V)$  about which perturbations must be applied. Virtual data (if validated against available data and verified against classical idealized exact solutions) provides insight into the microphysical sources of aleatory uncertainty, which likewise aids in the formulation of path-dependent macroscale models for evolution of internal state variables that control the loss of stiffness and collapse of the limit surface from failure.

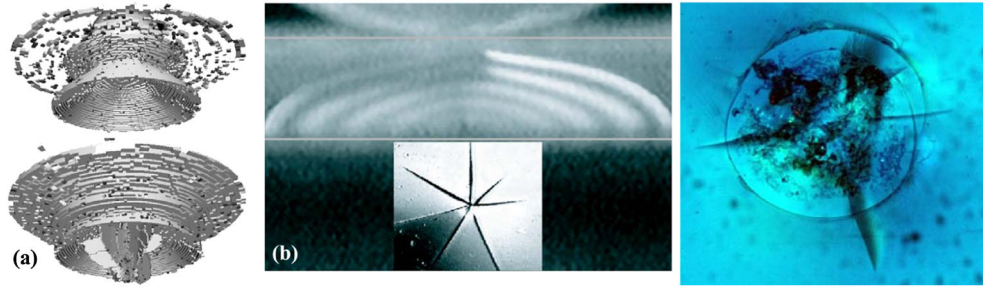


Figure 12. (a) Simulated dynamic indentation: transition from conical damage to radial cracking beneath the cone, similar in character to observations in glass indentation [57, 60]. (b) Quasi-static indentation: symmetry-breaking radial cracks beneath symmetry-preserving ring cracks [61, 63].

## 5. MODEL CALIBRATION OF VARIABILITY AND SCALE EFFECTS

Methods to account for strength data from differently sized experimental samples and methods to fit repeated strength measurements to a Weibull (or similar size-dependent) distribution are discussed in this section. These techniques provide median strength values and distribution properties (such as the mean and standard deviation) for a fixed experimental configuration.

### 5.1. Normalizing experimental data to a standard reference volume

Ideally, the same size sample would be used in a variety of strength tests for various loading modes such as triaxial compression, simple shear, buttonhead tension, and so on. If median strengths are measured in these different loading modes using the same sample size, then the results directly map to a strength limit surface in stress space. However, if different loading modes are measured using samples of different sizes, then all of the data must be scaled to a common reference volume before mapping the median strength surface.

Weibull scaling in Equation (2), including the reformulated versions in Equation (7), and Equation (11), requires the volume to be expressed relative to a reference volume  $\bar{V}$ . Similarly, the median stress  $\bar{\sigma}$  corresponds to measurements using a sample of size  $\bar{V}$ . Recognizing that different experiments might be performed using samples of different volumes, this section briefly explains how all data from homogeneous loading experiments can be scaled to a common reference volume. The far more subtle problem of scaling data from heterogeneous loading experiments (such as Brazilian or bending tests) remains unsolved, but the relevant issues are introduced in Section 6.

Consider a set of experiments that measure a median strength  $\sigma^*$  and Weibull modulus  $m^*$  for a sample volume  $V^*$ . According to Equation (2), the same set of experiments conducted using a sample with the reference volume  $\bar{V}$  would produce the same Weibull modulus ( $m = m^*$ ) and a volume-scaled median strength given by

$$\bar{\sigma} = \sigma^* \left[ \frac{\bar{V}}{V^*} \right]^{-1/m}. \quad (10)$$

By using this formula, experiments that use different specimen sizes may all be standardized to a single reference volume  $\bar{V}$ . The choice for the reference volume is arbitrary, but a reference volume is essential to generate failure threshold realizations by comparing the finite-element size to the reference volume size used in laboratory calibration experiments. In geomechanics, it is common to employ two different sets of material properties, one for the so-called laboratory scale and another for “field scale” (cf. [66]). The outlined implementation of Weibull theory in Equation (10) automates scale effects so that only one material data set spans a large range of size scales.

The directionally-Weibull theory developed in Section 4 produces a different apparent (and approximate) Weibull modulus  $m$  for different loading trajectories, which should be accounted for when scaling sample sizes. The Weibull modulus for a selected model and loading direction can be determined by numerically (or analytically) generating a series of realizations of the limit surface,

collecting the corresponding CDF (failure probability) values along the fixed loading trajectory, and then applying the analysis outlined below.

### 5.2. Measuring strength parameters

The following procedure for measuring the strength parameters is performed under two idealized conditions. One is that the loading is statistically homogeneous with no significant macroscale stress gradients prior to failure (e.g., quasi-static triaxial compression and extension and buttonhead tension tests). The second condition presumes that the same reference specimen size  $V$  is used in all experiments. Even though these conditions can be difficult to meet, understanding the method used to calibrate data under these idealizations is a prerequisite to subsequent discussions in Section 6 that release these assumptions by accounting for stress gradients within the specimen.

Although expanded testing and analysis are likely to be ongoing for many years, preliminary scale-normalized strength data for SiC-N are shown in Figure 13 for a variety of pressures, with median *scaled* data fitted to Kayenta's limit surface function. Without scaling all data to a common specimen size, no fitting exercise would be meaningful. The raw data measured using a sample of volume  $V$  were scaled to a common reference volume  $\bar{V}$  via Equation (10). For experiments having a nonuniform stress field, the specimen volume  $V$  is understood to be the effective volume from a data delocalization scheme (Section 6).

A suite of experiments using different loading directions must be run to identify the median failure envelope, as it is loading path-dependent. This conventional part of the calibration process identifies the dependence of  $\bar{\sigma}$  on stress triaxiality (or, more generally, on  $\mathbf{N}$ ). Once the loading path for a particular experiment is selected, the experiment must be conducted multiple times to obtain a statistically significant collection of strength realizations  $\{\sigma_1, \sigma_2, \dots, \sigma_n\}$ . Deciding if the data fit a Weibull distribution requires assigning values to the reference volume  $\bar{V}$ , the median strength  $\bar{\sigma}$  at the reference volume  $\bar{V}$ , and the Weibull modulus  $m$ . Specifically, the data must be fit to the "linearized" version of the Weibull equation

$$y = mx + b \quad \text{where} \quad y = \ln \ln \left( \frac{1}{P_{\text{safe}}} \right) - \ln \ln(2), \quad x = \ln \frac{\sigma}{\bar{\sigma}}, \quad \text{and} \quad b = \ln \frac{V}{\bar{V}}, \quad (11)$$

which is obtained by taking the natural logarithm of Equation (2) twice and rearranging. Any data, whether Weibull distributed or not, may be depicted in a plot of  $y$  versus  $x$ . Such data are proved to be Weibull distributed if the data fall on a straight line with respect to these Weibull axes, though we caution that even *exactly* Weibull-distributed data will exhibit stair-stepped deviations from the line as a result of finite-sampling errors (Figures 14 and 16). As seen in Figure 16, finite sampling errors in *exactly* Weibull-distributed data tend to produce a "dip" away from the lower-left portion of the theoretical Weibull line. On the other hand, simple microphysically derived strength CDFs (cf. [31]) also have a dip that is not the result of finite sampling error. Resolving whether a dip is a genuine

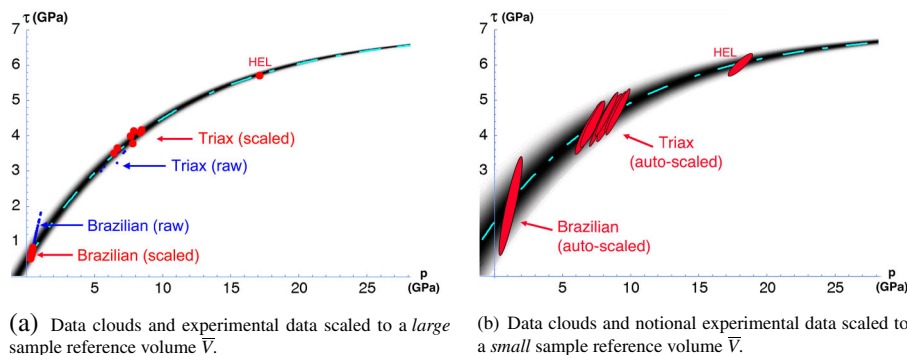


Figure 13. Examples of strength scaling and data clouds for comparison with experiments of silicon carbide. a) Data clouds and experimental data scaled to a large sample reference volume  $\bar{V}$ ; (b) data clouds and notional experimental data scaled to a small sample reference volume  $\bar{V}$ .

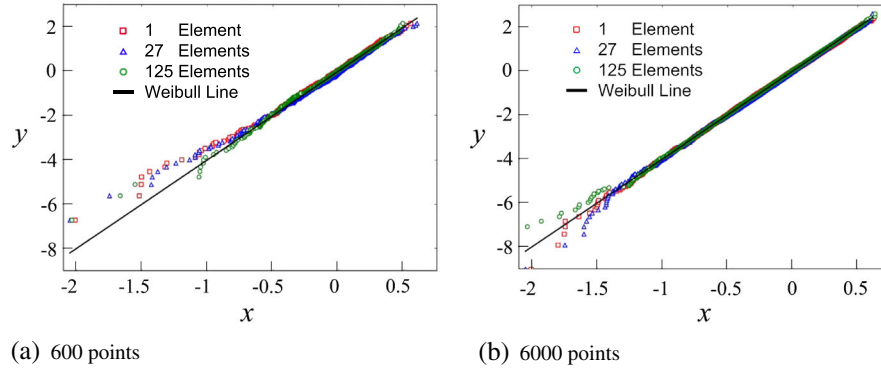


Figure 14. Verification of mesh independence for the onset of failure. The figures show simulated  $(x_k, y_k)$  data from Equation (13) for three mesh resolutions to demonstrate that failure statistics for the onset of failure are mesh insensitive and approach the Weibull distribution as the number of realizations is increased. (a) 600 points and (b) 6000 points.

attribute of the failure statistics or merely a byproduct of finite sampling is a strong motivation to increase the number of strength tests in laboratory testing.

As discussed in Sections 2.2 and 4, “strength” is regarded as the value of the loading parameter  $\sigma$  along a loading trajectory at failure. Here, “failure” is the initiation of stiffness and strength reduction upon attainment of a peak value of  $\sigma$ , but the definition could be altered to mean phase transformation, or any other “event” marking a distinct irreversible change in material behavior. Repeating the experiment  $n$  times will give a set of strength realizations that can be put into canonical order,  $\sigma_1 \leq \sigma_2 \leq \dots \leq \sigma_k \leq \dots \leq \sigma_n$ . Associated cumulative probabilities for sample number  $k$  are given by

$$P_k^{\text{fail}} = \frac{k - 1/2}{n} \quad \text{and} \quad P_k^{\text{safe}} = 1 - P_k^{\text{fail}}. \quad (12)$$

This estimator is advocated by Sullivan and Lauzon [67], who also recommend using no fewer than 20 measurements and preferably more than 50. An even larger number of measurements is required to determine if the Weibull distribution is, in fact, the actual distribution. For example, more than 1000 tests would be required to distinguish between the Weibull distribution and the log-normal distribution [68]. Other possibilities such as the Gumbel distribution [69] undoubtedly require an equally large sample to validate. Consequently, no part of this paper presupposes that a Weibull distribution applies to measured strength distributions. A Weibull distribution is used primarily as a familiar example to illustrate considerations in data fitting. As explained in Section 4, the empirical framework naturally predicts strengths fitting no standard distribution, Weibull or otherwise. Minimally, distributions must be used that will generate physically admissible realizations.

For fitting data to a Weibull distribution, the volumes  $V$  and  $\bar{V}$  are both set equal to the specimen volume used in the experiment, making  $b = 0$  in Equation (11). From the collection of  $(\sigma_k, P_k^{\text{safe}})$  data pairs,  $\bar{\sigma}$  is identified to be the median (i.e., the value corresponding to  $P^{\text{safe}} = 1/2$ ). Thus, only the Weibull modulus  $m$  in Equation (11) remains yet to be inferred from the measured data. To obtain  $m$ , the  $(\sigma_k, P_k^{\text{safe}})$  data pairs are converted to  $(x_k, y_k)$  data pairs by the following transformations:

$$x_k = \ln \frac{\sigma_k}{\bar{\sigma}} \quad \text{and} \quad y_k = \ln \ln \frac{1}{P_k^{\text{safe}}} - \ln \ln 2. \quad (13)$$

The Weibull modulus  $m$  is then determined via *true linear* (not affine) regression fitting of these data pairs to  $y = mx$ . Namely,

$$m = \frac{\sum_{k=1}^n y_n x_n}{\sum_{k=1}^n x_n^2} \quad (14)$$

Naturally, the regression error is a direct quantitative indicator of the extent to which a Weibull distribution is appropriate. Sullivan and Lauzon [67] offer an alternative to simple linear regression. However, by analyzing synthetically generated exactly Weibull-distributed data, we find their method to be superior to linear regression only when the total number of strength measurements *far* exceeds their recommended minimum of 20 data points. For data sets of fewer than 14 points, we similarly find that Equation (14) more accurately predicts the Weibull modulus than a scheme that sets the two (shape and scale) Weibull parameters by matching the sample mean and standard deviation to Equations 5 and 6. Even more than 14 points are needed for this alternative fitting scheme to be superior whenever data beyond 2 or 3 standard deviations from the mean are misidentified as “outliers”. For a typical engineering sample, which rarely exceeds 20 data points and is often far below what is needed for statistical significance, these finite-sampling errors and inadvertent removal of outliers seem to be best managed using the standard CDF estimator in Equation (12) and the simple linear regression in Equation (14). Overall, the median is the most convenient measure of central tendency because our approach aims to identify generally non-Weibull quantile surfaces (for which the median is the 50% quantile).

Figure 14 verifies mesh independence (and recovery of the user-prescribed Weibull strength distribution) for the onset of failure under idealized spherical tension. The effect of finite sampling on a Weibull plot (i.e., a plot of  $y$  vs.  $x$ ) is apparent in the numerical data; finite sampling tends to cause dips away from the Weibull line near the tail of the distribution (see also Figure 16).

In Figure 15(a), a ceramic flyer (the darker color on right side of both images) impacts a ceramic target twice the thickness of the flyer. The damaged (spalled) material has been hidden from the visualization of the simulation on the left, and the velocities are shown on the right (blue to red color scale indicates low to high velocities, respectively). Spatial variability induces non-uniform velocity on the surface of the target that would not occur in a deterministic simulation. Spall is a dynamic tensile failure produced by the interaction of two rarefaction release waves, making it a truly internal failure mode that is unaffected by surface smoothness at the specimen boundary. Because of the strongly non-uniform stress fields occurring in spall tests, the data delocalization methods discussed in Section 6 must be applied to iterate on the specimen volume associated with measured spall strengths.

In Figure 16, Weibull plots and plots of the complementary CDF  $P^{\text{safe}}$  are shown for two different experiments: the Hugoniot elastic limit strengths for  $B_4C$  ceramic [54] and previously unpublished tensile strengths for SiC-N for the Brazilian diametrical compression test. Both appear to be Weibull-distributed, but the  $B_4C$  shock compression data clearly exhibit much lower variability than the Brazilian data. These data deviate from the best Weibull fit but only by amounts well within the range of deviations of Weibull data pulled from the *exact* analytical Weibull distribution for the same modulus  $m$  (also shown in the figure). As seen, finite-sampling deviations in the exactly Weibull-distributed data are similar in magnitude and character to deviations of the measured data from the Weibull fit. Thus, measured data may be closer to a Weibull distribution than might seem evident from a Weibull plot.

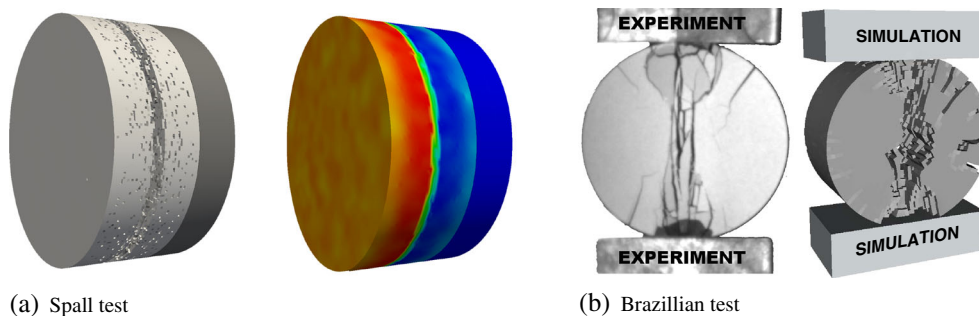


Figure 15. Simulations with variability demonstrating bifurcation from nominal symmetries. (a) Spall test and (b) Brazillian test.

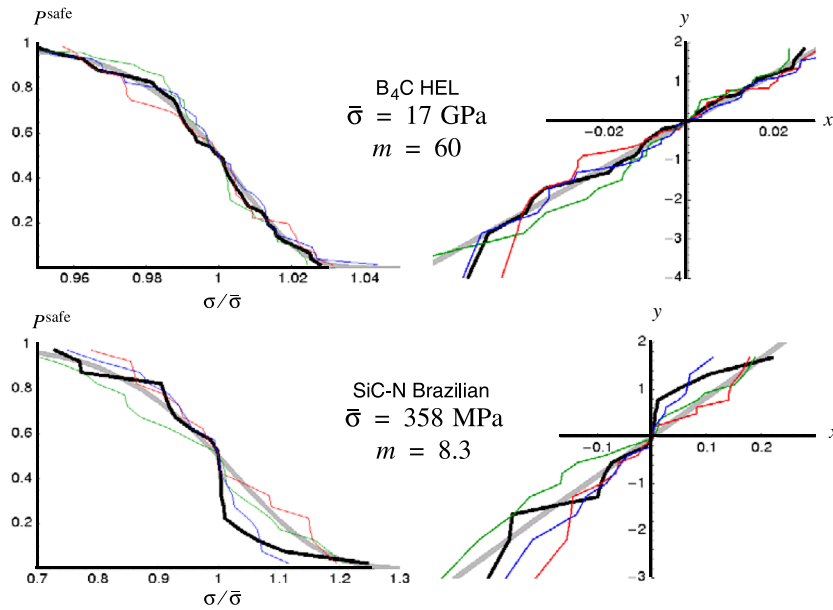


Figure 16. Measured strength data (thick black line) and best Weibull fit (thick straight gray line), along with the same number of *exactly* Weibull distributed simulated data points (thin colored lines).

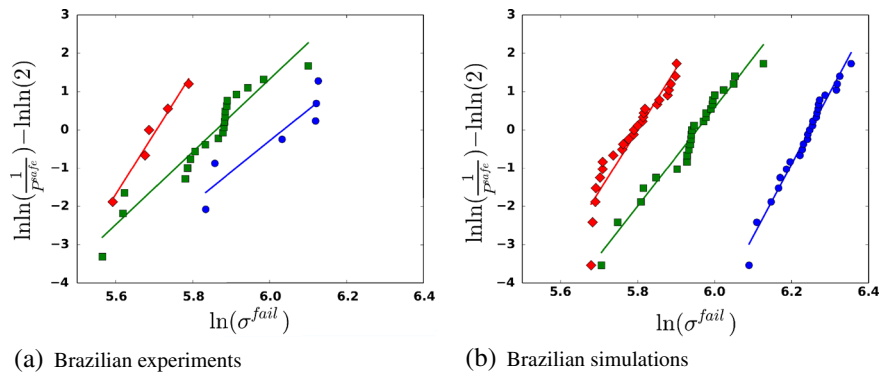


Figure 17. (a) Laboratory strength data show effects of three specimen sizes (five strengths measured for the largest specimen are near the left-most linear regression fit, twenty strengths for the medium-sized specimen are in the center, and six for the smallest specimen are on the right). (b) Simulations for the same sample sizes, with initial spherical strength perturbed using a Weibull modulus of 6, update those in [70].

Finite-sampling effects (the “dipping” deviation from the Weibull line) are especially apparent in the Brazilian strength data plotted in Figure 17, because those experiments acquired fewer data points than the experiments in the left side Hugoniot elastic limit data of Figure 16. Examples of Brazilian experiments and comparable simulations are in Figure 15(b).

The strength size scaling of the Brazilian material samples is noted in both the experiments and simulations in Figure 17. These plots omit the typical normalization of the abscissa by the median stress at failure  $\bar{\sigma}$  so that the relative strengths and slopes can more readily be seen for the different sample sizes. An interesting aspect of Figure 17 is that the experimental data appear to show not only an increase in strength with decreasing specimen size but also a decreasing Weibull modulus (i.e., decreasing slope and hence greater relative variability in strength) with decreasing specimen size. However, because of the small number of experimental samples and the impact of finite-sampling errors, this trend is inconclusive. The numerical simulations in Figure 17 do not show the same

trend, and further work is warranted to understand the scaling of the apparent Weibull modulus as a function of sample size for the Brazilian test (both experimentally and numerically).

Although it would be possible to tune the parameters of the model to better match the data, doing so would be grossly misleading, because the simulations failed to exhibit the same mitigation of mesh sensitivity that was seen in the dynamic indentation simulations of Figure 1 (i.e., changing the mesh would require inadmissible re-tuning of material properties). In the Brazilian simulations, the methods of Section 6 have been confirmed to provide a high rate of convergence for the *onset* of failure in the Brazilian test. However, further research breakthroughs are required to likewise identify regularizations for the subsequent *progression* of failure (both at the element and structural level) while also predicting realistic strength distributions and scale effects.

## 6. NONLOCAL CORRECTIONS OF DISCRETIZATION ERRORS

It has long been asserted that finite-element analysis of distributed damage requires some type of nonlocal theory [71, 72]. Nonlocality is attributed to many physical sources such as crack interactions [73], which might be important on fine discretizations where aleatory uncertainty in crack locations implies the need for correlated strength distributions. However, recent work [32] suggests purely numerical motivations for nonlocality.

When attempting to calibrate scale-dependent strength from laboratory experiments, so-called data delocalization techniques are recommended to assign the corresponding specimen volume. If the stress field is uniform, then the specimen volume is the actual volume. However, for bending, Brazilian indirect tension, spall, and other experiments having a nonuniform stress field, data-delocalization theory [38, 74–76] sets the specimen volume using a stress-weighted average. For a uniformly loaded and statistically homogeneous specimen, we adopt the axiom that the probability of a local failure must scale with volume, not area (i.e., with  $V$ , not  $V^{2/3}$ ). However, our own experimental data for the Brazilian indirect tension tests in Figure 17 scale with area, not volume. This apparent area-based scaling is a natural result when the data are analyzed using data-delocalization theory. Specifically, whenever stress is nonuniform (so that only a portion of the actual volume  $V$  is loaded to an appreciable stress level), the “effective volume” is defined by

$$V_{\text{eff}} = \int_V (\sigma/\sigma_{\text{peak}})^m dV \quad (15)$$

where  $\sigma$  is a scalar measure of stress (such as the maximum principal stress) and  $\sigma_{\text{peak}}$  is the maximum of  $\sigma$  in the domain  $V$ . In the Brazilian test, for example, higher stresses are located over a narrow region at the center of the domain, making this region have higher weight in Equation (15), thus predicting that  $V_{\text{eff}}$  scales with area as observed in the experiments.

Of course, evaluating  $V_{\text{eff}}$  requires a description of the non-uniform stress field prior to failure. Iterations are also required to find  $V_{\text{eff}}$  because it depends on one of the properties being measured (the Weibull modulus,  $m$ ). The explicit presence of the Weibull modulus corresponds to an implicit assumption that the strength is Weibull distributed. Furthermore, because Equation (15) depends on a single scalar stress invariant, it fails to account for the tensorial nature of stress, and it implicitly introduces a failure criterion into a situation where the failure criterion is part of the information being sought in the experiment itself. To alleviate some of these shortcomings, we expect that a practical generalization of Equation (15) might be

$$V_{\text{eff}} = \int_V \frac{X(\mathbf{T})}{X(\mathbf{T})_{\text{peak}}} dV \quad (16)$$

where the form and function of  $X(\mathbf{T})$  should be regarded as the primary information sought in laboratory testing. As explained in an idealized context, namely Equation (11) of [31], the generalized function “ $X$ ” is expected to be a function of the loading mode and a functional of underlying micromorphology (e.g., evolving flaw size and orientation statistics).

Data-delocalization is required to interpret data from experiments having a nonuniform stress field to assign a volume-dependent strength under uniform stress. Kamojjala and Brannon [32] have developed a “data-relocalization” technique to reduce errors due to under-resolution of stress fields. They explain that, to determine the failure initiation probability, the effective *uniform* stress acting on a finite element should be set as an  $m$ -norm of the actual pre-failure stress field:

$$P^{\text{safe}} = 2^{-\frac{V}{V}} \left( \frac{\sigma_{\text{eff}}}{\sigma} \right)^m \quad \text{where} \quad \sigma_{\text{eff}} \equiv \left( \frac{\iiint \sigma^m dV}{V} \right)^{1/m} \quad (17)$$

Here,  $m$  is the Weibull modulus, implying that this theory would require revision for non-Weibull distributions, perhaps replacing  $\sigma^m$  with a micro-physically based  $X(\mathbf{T})$  function, as in the delocalization proposal of Equation (16). Also,  $\sigma$  is the pre-failure *analytical* principal stress invariant field over the finite element, which is known for the Brazilian test and the other simple cases studied by Kamojjala and Brannon. To adapt this relocalization formula to cases for which the analytical stress field is unknown, a nonlocality algorithm is required to estimate the stress concentrations within an element on the basis of stresses in surrounding elements. This recommendation is a purely *numerical* motivation for nonlocality, and it might be accommodated in a “semi-local” way (i.e., without needing to query fields in neighboring elements) by applying an  $m$ -norm to the non-constant stress field available within higher-order elements (Kamojjala and Brannon’s exploration of higher-order elements essentially missed this opportunity by treating the stress field as constant in the neighborhood of the Gauss points).

The previous sections have focused on statistics for the *onset* of failure, which controls the statistically variable strength assigned at initialization and data-relocalization methods to adjust the strength of an element in response to stress gradients. In a rigorous analysis of quasistatic mode-I opening of a sharp crack, Bažant [23] addresses additional considerations associated with failure *progression* that may or may not lead to cascading structural failure after failure initiation. Specifically, Bažant pointed out that stress singularities near a crack tip increase more rapidly with mesh refinement than element strengthening associated with the size effect, thus encouraging crack growth. The related topic of damage progression is the subject of the next section.

## 7. STRENGTH AND STIFFNESS DEGRADATION

The preceding sections focused almost exclusively on the need to assign different values of material strength to each element in a computational domain. The strength assigned to different elements is statistically variable and size-scale dependent, consistent with laboratory observations. This variability in strength provides perturbations in the stress field necessary to stimulate inherent instabilities such as bifurcation of axisymmetric loading into radial cracking. However, once cascading failure has begun, additional scale effects are well known to be necessary to mitigate mesh dependence during the progression of failure [77].

Seminal cohesive zone formulations [78, 79] and more recent models (cf. [30, 80, 81]) introduce a scale effect by requiring a certain amount of energy be required to advance a crack. Until recently, cohesive zone formulations have been developed on the basis of an implicit assumption that a finite domain contains exactly one localized zone. However, high-rate loading can produce a network of actively growing cracks within a domain, therefore calling for an increase in required fracture energy [58, 77]. As explained in this section, this overall effect is accommodated in our model by requiring a finite time to elapse before an element can lose strength.

To simulate crack growth, conventional isotropic damage models usually employ a scalar damage parameter,  $D$ , that varies from 0 for undamaged material to 1 for fully failed material [44]. These models lead to a drop in both stiffness and strength with increasing strain (i.e., softening) and accordingly induce localization of material response. Without such localization, the governing equations can change type, potentially requiring solution of a diffusion equation [82]. On the other hand, traditional viscous rate dependence in the governing equations can have a regularizing effect [83]. Similarly, the extreme localization of brittle and quasi-brittle materials (i.e., formation of macrocracks) allows the governing equations to remain hyperbolic as long as boundary conditions

at the new surfaces are handled properly. In a smeared damage model, these boundary conditions are approximated by a boundary layer of failed material that offers resistance to compression but progressively reduced resistance to tension and ultimately only frictional resistance to shear. Further discourse about the character of conventional damage models is beyond the scope of this paper, where our focus here is on seeding these models realistically with uncertainty and scale effects. The conventional damage model used for our case study (Kayenta) is well documented [6], so we merely call attention to some aspects that differ from similar models available in commercial codes.

The Kayenta damage model employs a distinctive approach to degrading strength. Taking a cue from experimental observations [84] that cracks in high-stress environments grow at approximately constant speed (the effect of which is seen in Figure 18), the model smoothly decreases strength and elastic stiffness to failed values over a scale-dependent time interval (conceptually similar to the method applied in [85]). Consistent with the notion that element failure corresponds to traversal of a macroscale “train” of coalescing microcracks across the element, a small element fails sooner than a large element, and hence, the model keys the “time-to-failure” to the element size. For simplicity, elements with large aspect ratios are avoided so that the time-to-failure has a length scale of approximately  $V^{1/3}$ , where  $V$  is the volume of the finite element that is softening. Not only will larger elements be initialized with lower median strength, this straightforward (and computationally efficient) method for scaling the time-to-failure ensures that larger elements might still require greater fracture energy, depending on how the material parameters are set. Implementation details, including formulas for degrading strength and stiffness with damage, are provided in [6].

The time-to-failure scaling is insensitive to the loading rate. Accordingly, a specimen loaded slowly will seem to lose strength suddenly because the time-to-failure will be very small relative to the total loading time. If the same specimen is loaded very rapidly to the same strain level, it will lose strength over a larger fraction of the total loading interval (because a fixed amount of time must pass before the stiffness and strength degradations fully take effect). Thus, when loaded rapidly, our model will appear to have a finite plastic strain-to-failure (which is similar in character to the Johnson–Holmquist class of ceramic models [45, 48–50] whenever those models are compared with ours *at a given loading rate*, but a key distinction is that our model’s use of time-to-failure provides different *apparent strain-to-failure* at different loading rates).

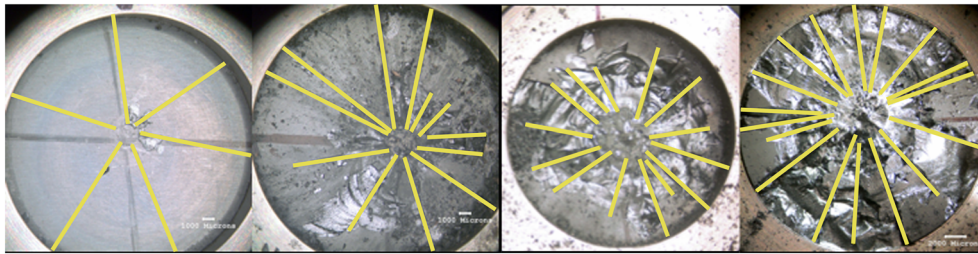


Figure 18. Laboratory observations of the effect of impact speed on the number of radial cracks [70]. Shown are B<sub>4</sub>C samples impacted at speeds ranging from 100 to 400 m/s from left to right.

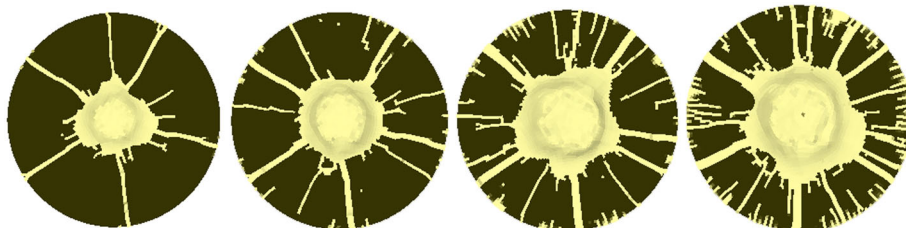


Figure 19. Realistic trends in damage patterns when including spatial variability, scale effects, and time-to-failure. Changing the WC impactor speed (from 200 to 500 m/s from left to right, using the same mesh and property seeding) produces a corresponding increased number of radial cracks in simulations of SiC targets.

Because our model requires a fixed amount of time to pass before loss of strength occurs, the higher strain rate caused by higher impact speed allows more material to reach the failure threshold before release waves (or Mott waves—see also the related “event horizon” [76]) can arrive from the radial cracks that formed slightly sooner at weaker points. The well-known experimental observation that the number of radial cracks in axisymmetric loading increases with loading rate [86, 87] (see, for example, observations for  $B_4C$  in Figure 18) is therefore easily predicted with this time-to-failure softening algorithm because it essentially allows a larger number of flaws to become activated at high rates (see similar trends in Figure 19 for simulations of SiC). Consistent with [58], this higher probability of flaw activation at higher loading rates was also mentioned by Brannon and Gowen [31] as being a simple case of a larger portion of the 3D stress state (as depicted in our Figure 4) falling transiently farther above each crack’s size-dependent failure line, hence naturally activating larger populations of both flaw sizes and orientations.

## 8. NECESSITY OF BOTH STRENGTH VARIABILITY AND TIME-TO-FAILURE SCALING

As seen in Figure 20, imposing both strength variability and time-to-failure scaling significantly reduces mesh dependencies of damaged zones, crack patterns, and rate of crack propagation for the dynamic indentation of ceramics. The images in Figure 20 depict simulations of a 6.35 mm diameter tungsten-carbide sphere launched at 375 m/s into a 25.4 mm diameter SiC-N ceramic cylinder, described in [7]. Each set of images in Figure 20(a)–(d) contains a low resolution simulation next to successively higher resolutions. All simulation parameters were kept the same, except for the mesh resolution and the feature being explored in the model (which was changed for each subset).

Figure 20(a) contains results for a conventional model (Kayenta) calibrated to the median of the experimental data and ignoring size effects. Here, the high-pressure limit surface parameter  $Y$  was calibrated to the data used in [49], the tensile strength parameter  $h$  was calibrated to buttonhead tension results in [88], and the shear strength parameter  $s$  was calibrated to triaxial compression results in [11]. To create an objective comparison between different features of the model, the time-to-failure was computed for the average element size of the mesh in the medium resolution simulation for the fully enhanced model in Figure 20(d) (the middle image) and applied as a constant time-to-failure for the conventional model. In Figure 20(a), the number of radial cracks and especially the depth of penetration depend on the element size in the mesh. For the simulations in Figure 20(b), the time-to-failure was scaled according to element size as discussed in Section 7, while other parameters were the same as in the conventional model in Figure 20(a). Note the reduction in the number

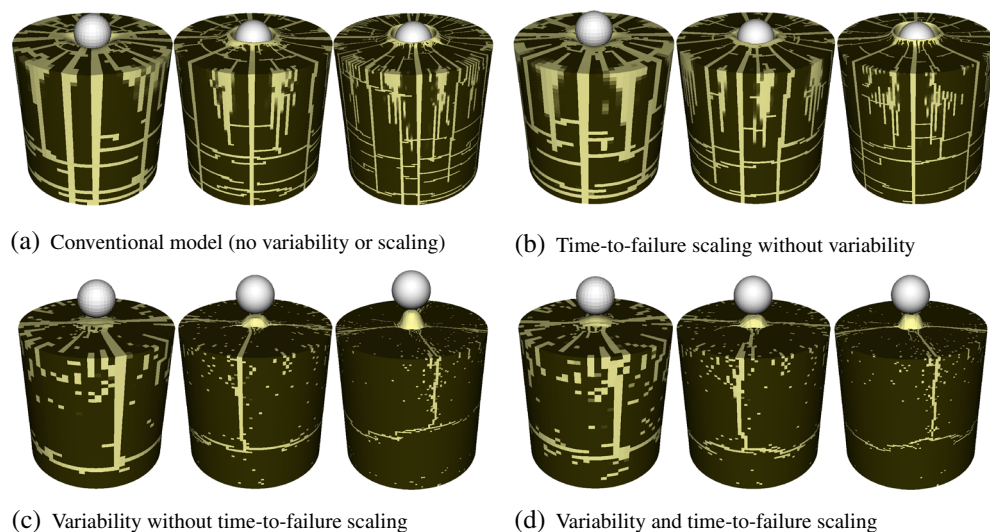


Figure 20. Influence of strength variability and time-to-failure scaling in dynamic indentation. Subfigures each show three results at the same moment in time on progressively finer mesh resolutions.

of radial cracks in the high-resolution simulation on the right due to the lower time-to-failure for the smaller elements in that mesh. For the simulations in Figure 20(c), a Weibull variability of strength, including the volume scaling term in Equation (7), was applied to the  $h$  parameter in the Kayenta model, as discussed in Section 4. The other parameters (including constant time-to-failure) were the same as in the conventional model in Figure 20(a). Note the significant effect of including spatial variability. The size scaling effects have resulted in a significantly higher median element strength (and hence more realistic ballistic performance) while the variability has, in principle (modulo issues with the progression of damage discussed in Section 7 that can be encountered in simulations of calibration experiments), preserved the match of the model to the suite of calibration experiments discussed in Section 5. Also, using a size-independent time-to-failure causes the horizontal lateral cracking (on the perimeter toward the bottom of the target) to be located at successively higher elevations as the mesh resolution increases (contrast this to the location of these cracks in Figure 20(d)).

The model in Figure 20(d) includes both the spatial variability of strength discussed in Figure 20(c) and the time-to-failure scaling discussed for Figure 20(b). The two effects work together, resulting in a model with substantially reduced mesh dependence, both in the amount of damage induced, and in the rate of growth of damage. The approximate number of radial cracks and location of lateral cracking are similar between the meshes. Mesh dependence has been reduced, but not eliminated. Significant work is still required for truly predictive ballistic simulations of ceramics, but spatial variability, size scaling, and time-to-failure scaling represent technical advances in comparison with predecessor theories. Moreover, these features can be added without significant increases in computational (or implementation/development) overhead.

## 9. CONCLUSIONS

Enhancing an existing conventional deterministic damage model to support spatial variability requires (in principle) no revisions of the existing constitutive model's source code. Code changes are applied at the host-code level, where fields are allocated (during initialization) for spatially varying material parameters. For example, each element in a simulation might have its own randomly assigned scale-dependent strength. These perturbed parameters are assigned values by applying whatever statistical distribution is evident in laboratory testing of finite-sized specimens but with care given to anticipate parameter correlations required to produce admissible material realizations (e.g., to avoid inadvertently generating non-convex strength limit surfaces).

If a material parameter  $\eta$  happens to be observed in the laboratory to be Weibull distributed and if it also has corresponding Weibull scale effects, then the spatially variable and size-dependent value assigned to a finite element of volume  $V$  is found by computing a random number  $R$  and then saving

$$\eta = \bar{\eta} \left[ \frac{\bar{V} \ln R}{V \ln(1/2)} \right]^{1/m} \quad (18)$$

into a field array, which then provides each finite element with its own perturbed material property  $\eta$ . Here,  $\bar{\eta}$  is the median value of  $\eta$  observed in the laboratory using a specimen of volume  $\bar{V}$ , and  $m$  is the Weibull modulus quantifying observed variability in the property. Appropriate revisions naturally apply if the parameter is not Weibull distributed.

Given the overarching goal to introduce variability into an already available deterministic damage model, each material property  $\eta$  that is already available in existing property databases is merely re-interpreted as the median about which uncertainty and scale effects are imposed. Thus, the only new inputs are  $\bar{V}$  and (if Weibull-distributed)  $m$ , both of which can be determined by re-inspecting the laboratory data originally used to determine  $\eta$ . "Data delocalization" schemes [38, 74–76] are needed to set the effective specimen volume  $V$  for any strength experiments involving non-uniform stress fields, whereas "relocalization" [32] essentially re-introduces stress variation that is not resolved on low-order finite elements.

If the value of a property  $\eta$  was determined from only a single experiment, then the experiment would need to be repeated enough times to quantify the scale-dependent central tendency (e.g.,

median) and the aleatory uncertainty (e.g., standard deviation). At least 20 (and preferably more than 50) such tests are recommended for each loading mode (shear, tension, compression, etc.), but indirect methods, such as “line VISAR” in spall testing seem promising to quantify variability in a *single* experiment [53, 89].

Overall, parameterizing material models from observed aleatory uncertainty and associated scale effects shows promise for enhancing the predictive capability of numerical methods of fracture and fragmentation while still keeping computational and development costs manageable. An ordinary deterministic limit surface is effectively replaced with a family of level sets (called aleatory quantile surfaces) associated with a notional failure probability function. Various scalar-based strengths (e.g., buttonhead tension, triaxial compression, spall, simple shear, Brazilian indirect tension, etc.) each have different distributions, yet they are all simultaneously modeled by reinterpreting them as directed derivatives of the single, unified tensor-based quantile function. The tensorial nature of stress is therefore directly and conveniently accounted for, giving non-Weibull, pressure-dependent, and generally loading-mode-dependent strength distributions based on laboratory evidence. A scale-dependent time-to-failure algorithm for degradation of stiffness (and strength) is shown to not only alleviate mesh sensitivity in dynamic indentation but it also naturally ensures activation of larger populations of flaws during higher-rate loading, automatically giving more fragments with increased loading rate.

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